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Thanasis Bouzidis and Giannis Karagiannis

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Thanasis Bouzidis ${ }^{a}$ and Giannis Karagiannis

Department of Economics, University of Macedonia, 156 Egnatia Str., 540 06, Thessaloniki, Greece
a thabou@uom.edu.gr; +30 2310891 781; ORCID: 0000-0002-8255-997X

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# A Note on the Zero-Sum Gains Data Envelopment Analysis Model 


#### Abstract

In the case of the proportional output reduction strategy with a single output, the Variable-Returns-to-Scale (VRS) Zero-Sum Gains Data Envelopment Analysis (ZSGDEA) efficiency scores can be obtained from the VRS conventional DEA efficiency scores by means of the Target's Assessment Theorem (TAT). Using TAT as a departure point, two relations for computing the ZSG-DEA efficiency scores appear in the literature. Our objective in this note is to compare, contrast and challenge them on both theoretical and empirical grounds. For the latter, three different data sets are used.


## Keywords

Output Interdependency; ZSG-DEA Efficiency; Conventional DEA Efficiency

## 1 Introduction

Conventional Data Envelopment Analysis (DEA) assumes that the output of each Decision Making Unit (DMU) is independent of that of any other DMU. This implies that each DMU may expand its output as far as it is needed to improve its efficiency independently of the other DMUs. However, this does not hold in the presence of output interdependency that occurs, for example, if (i) outputs are ranks in a contest where the higher is a participant's ranking, the lower is the ranking of another participant; (ii) the aggregate output of DMUs is a priori fixed as when market share or the total number of wins during a league season are considered as outputs; (iii) DMUs use inputs to both expand their output and shrink that of their rivals as in head-to-head competition; or (iv) the aggregate desirable (undesirable) output is regulated by quotas (permits).

To deal with such cases, Lins et al. (2003) proposed the Zero-Sum Gains DEA (ZSG-DEA) model that provides efficiencies adjusted for output interdependency. ${ }^{1}$ This model is operationalized by means of the equal, the proportional or the minimal output reduction strategy; see Lins et al. (2003), Collier et al. (2011), and Yang et al. (2011). In the ZSG-DEA model, the extra output that each DMU under evaluation may require to become efficient is taken from all other DMUs in such a way that the sum of output gain and output losses across DMUs equals zero and the aggregate resultant output equals the aggregate actual output.

One difficulty with this model is that it is non-linear and only under certain circumstances, it can be simplified. For example, Lins et al. (2003) have shown that, in the case of the proportional output reduction strategy with a single output, the Variable-Returns-to-Scale (VRS) ZSG-DEA efficiency scores are related to the VRS conventional DEA efficiency scores by means of the Target's Assessment Theorem (TAT), which states that the potential output of each inefficient DMU in the ZSG-DEA model is a fraction of its potential output in the conventional DEA model. This implies that the distance of each inefficient DMU from the ZSG-DEA efficient frontier is always shorter than its distance from the conventional DEA efficient frontier or, in other

[^0]words, that the ZSG-DEA efficiency score of each inefficient DMU is always greater than its conventional DEA efficiency score. In addition, Lins et al. (2003) have shown that the ZSG-DEA and the conventional DEA models result in the same set of intensity variables; this is known as the Benchmarks' Contribution Equality Theorem (BCET) ${ }^{2}$ and it implies that the efficient frontiers in both models are formed by the same DMUs or, in other words, that the ZSG-DEA (in)efficient DMUs are also DEA (in)efficient and vice versa.

The relation for computing the ZSG-DEA efficiency scores by means of the TAT under the above circumstances is not however directly operational. In an attempt to simplify things, two alternative relations appear in the literature for computing the ZSG-DEA efficiency scores; see Hu and Fang (2010) and Bi et al. (2014). The main objective of this note is to compare, contrast and challenge them on both theoretical and empirical grounds. In theoretical terms, we examine whether they are consistent with the postulates of the ZSG-DEA model, namely the TAT and the BCET. Specifically, we expect the resulting ZSG-DEA efficiency scores to be (i) between zero and one, (ii) greater-than-or-equal-to their conventional DEA efficiency scores, and (iii) less than (equal to) one if their conventional DEA efficiency scores are less than (equal to) one. In empirical terms, we use three different data sets to examine their behavior and relationship.

## 2 Theoretical Framework

The envelopment form of the output-oriented VRS ZSG-DEA model is given as (Lins et al., 2003):

$$
\begin{array}{cll}
\begin{array}{cl}
\text { Max } & \hat{h}^{k} \\
\hat{h}^{k}, \lambda_{k}^{i}, \tilde{y}_{r}^{i} & \\
\text { s.t. } & \sum_{i=1}^{I} \lambda_{k}^{i} x_{m}^{i} \leq x_{m}^{k},
\end{array} & \\
& \sum_{i=1}^{I} \lambda_{k}^{i} \tilde{y}_{r}^{i} \geq \hat{h}^{k} y_{r}^{k}, & r=1, \ldots, M ; \\
& \sum_{i=1}^{I} \lambda_{k}^{i}=1, & \\
& \lambda_{k}^{i} \geq 0, & i=1, \ldots, k, \ldots, I ; \tag{1}
\end{array}
$$

[^1]where $\hat{h}^{k}$ refers to the expansion factor $\left(1 \leq \hat{h}^{k}<\infty\right), x$ to input quantities, $y$ to output quantities, $\tilde{y}$ to output quantities after accounting for gain and losses among DMUs, $\lambda_{k}^{i}$ to the intensity variables estimated in the $k^{\text {th }}$ "run" of (1), $m$ is used to index inputs, $r$ to index outputs, and $i$ to index DMUs. The ZSG-DEA efficiency scores are given as $\hat{\theta}^{k}=\frac{1}{\hat{h}^{k}}$. The only difference between the above and the corresponding conventional DEA model is in the left-hand side of the second constraint in (1) where we have $\tilde{y}$ instead of $y$. In (1), $\tilde{y}$ is a choice variable defined as $\tilde{y}^{k}=y^{k}+z^{k}=\hat{h}^{k} y^{k}$ for the $k^{t h}$ DMU and as $\tilde{y}^{i}=y^{i}-l_{k}^{i}$ for all other DMUs, where $z^{k}$ and $l_{k}^{i}$ refer respectively to output gain and output losses. Then, (1) may be rewritten as:
\[

$$
\begin{array}{cll}
\begin{array}{c}
\text { Max }
\end{array} & \hat{h}^{k} & \\
\hat{h}^{k}, \lambda_{k}^{i}, l_{k r}^{i}, z_{r}^{k} & & \\
\text { s.t. } & \sum_{i \neq k} \lambda_{k}^{i} x_{m}^{i}+\lambda_{k}^{k} x_{m}^{k} \leq x_{m}^{k}, & m=1, \ldots, M ; \\
& \sum_{i \neq k} \lambda_{k}^{i}\left(y_{r}^{i}-l_{k r}^{i}\right)+\lambda_{k}^{k}\left(y_{r}^{k}+z_{r}^{k}\right) \geq \hat{h}^{k} y_{r}^{k}, & r=1, \ldots, R ; \\
& \sum_{i \neq k} \lambda_{k}^{i}+\lambda_{k}^{k}=1, &  \tag{2}\\
& \lambda_{k}^{i} \geq 0, & i=1, \ldots, k, \ldots, I ;
\end{array}
$$
\]

If $\hat{h}^{k}=1$, then (2) is identical to the VRS conventional DEA model since in this case the $k^{\text {th }}$ DMU requires no output gain to become ZSG-DEA efficient (i.e., $z_{r}^{k}=0 \forall r=$ $1, \ldots, R$ ) and thus, no other DMU is forced to lose output from it (i.e., $l_{k r}^{i}=0 \forall i \neq$ $k, r=1, \ldots, R$ ). This implies that the same DMUs are on both the conventional DEA and the ZSG-DEA efficient frontiers, which in turn implies that the ZSG-DEA (in)efficient DMUs are also DEA (in)efficient and vice versa. On the other hand, if $\hat{h}^{k}>1$, then (2) seems to resemble the super-efficiency DEA model (Andersen and Petersen, 1993) because in this case $\lambda_{k}^{k}=0$ and thus, there is no second term in the lefthand side of the first, second and third constraint in (2). ${ }^{3}$ Although, a difference may be that (2) also estimates both the output gain of the $k^{t h}$ DMU and the output loss of each DMU $i \neq k$.

Noticeably, (2) is a non-linear model that can be simplified only under certain circumstances. One such a case considered by Lins et al. (2003) is for the proportional

[^2]output reduction strategy with a single output (i.e., $r=1$ ), where $z^{k}=y^{k}\left(\hat{h}^{k}-1\right) \geq$ 0 and $l_{k}^{i}=\frac{y^{i} z^{k}}{Y-y^{k}} \geq 0 \forall i \neq k$ with $Y=\sum_{i=1}^{I} y^{i}$. In this case, (2) is written as:
\[

$$
\begin{array}{cll}
\text { Max } & \hat{h}^{k} & \\
\hat{h}^{k}, \lambda_{k}^{i} & & \\
\text { s.t. } & \sum_{i \neq k} \lambda_{k}^{i} x_{m}^{i}+\lambda_{k}^{k} x_{m}^{k} \leq x_{m}^{k}, & m=1, \ldots, M ; \\
& \sum_{i \neq k} \lambda_{k}^{i} y^{i} R C^{k}+\lambda_{k}^{k} \hat{h}^{k} y^{k} \geq \hat{h}^{k} y^{k}, &  \tag{3}\\
& \sum_{i \neq k} \lambda_{k}^{i}+\lambda_{k}^{k}=1, & \\
& \lambda_{k}^{i} \geq 0, & i=1, \ldots, k, \ldots, I ;
\end{array}
$$
\]

where $0<R C^{k}=1-\frac{y^{k}\left(\hat{h}^{k}-1\right)}{Y-y^{k}} \leq 1$ is the reduction coefficient estimated in the $k^{t h}$ "run" of (3). ${ }^{4}$ Lins et al. (2003) have shown that $\hat{\theta}^{k}$ can be obtained, without using an optimizer, from the conventional DEA efficiency scores $\left(\theta^{k}\right)$ by means of the TAT, which states that the potential output of the $k^{t h}$ DMU evaluated by means of (3) is equal to its potential output evaluated by using the conventional DEA model multiplied by its reduction coefficient, namely:

$$
\begin{equation*}
\frac{y^{k}}{\hat{\theta}^{k}}=\left(\frac{y^{k}}{\theta^{k}}\right) R C^{k} \tag{4}
\end{equation*}
$$

The above relation is not however directly operational since $R C^{k}$ contains $z^{k}$, which in turn contains $\hat{\theta}^{k}$, the variable that we want to estimate.

Using (4) as a point of departure, two relations appear in the literature for computing the VRS ZSG-DEA efficiency scores under the proportional output reduction strategy with a single output. One, due to Hu and Fang (2010), is given as:

$$
\begin{equation*}
\breve{\theta}^{k}=\frac{\theta^{k} y^{k}\left(Y-y^{k}\right)+y^{k^{2}}}{y^{k}\left(Y-y^{k}+1\right)}=\frac{\theta^{k}\left(Y-y^{k}\right)+y^{k}}{Y-y^{k}+1} \tag{5}
\end{equation*}
$$

and the other, due to Bi et al. (2014), is given in terms of the expansion factor, namely $\hat{h}^{k}=\frac{h^{k} Y}{Y-y^{k}+h^{k} y^{k}}$, which can be converted into ZSG-DEA efficiency score terms as:

[^3]\[

$$
\begin{equation*}
\hat{\theta}^{k}=\frac{Y-y^{k}+\frac{y^{k}}{\theta^{k}}}{\frac{Y}{\theta^{k}}}=\frac{\theta^{k}\left(Y-y^{k}\right)+y^{k}}{Y} \tag{6}
\end{equation*}
$$

\]

in order to be directly comparable with (5). One can verify that (6) is implied by (4) but we were unable to show that the same is true for (5). To prove the former, substitute $z^{k}$ into $R C^{k}$ and then rewrite (4) as follows:

$$
\begin{equation*}
\frac{1}{\hat{\theta}^{k}}=\frac{1-\frac{y^{k}\left(\frac{1}{\hat{\theta}^{k}}-1\right)}{Y-y^{k}}}{\theta^{k}}=\frac{\hat{\theta}^{k}\left(Y-y^{k}\right)-y^{k}\left(1-\hat{\theta}^{k}\right)}{\theta^{k} \hat{\theta}^{k}\left(Y-y^{k}\right)} \tag{7}
\end{equation*}
$$

which implies that:

$$
\begin{equation*}
\hat{\theta}^{k} Y-y^{k}=\theta^{k}\left(Y-y^{k}\right) \tag{8}
\end{equation*}
$$

Then, by solving (8) for $\hat{\theta}^{k}$, we can obtain (6). Notice that (6) may also be written as:

$$
\begin{equation*}
\hat{\theta}^{k}=\theta^{k}+s^{k}-\theta^{k} S^{k} \tag{9}
\end{equation*}
$$

which states that the VRS ZSG-DEA efficiency score of the $k^{t h}$ DMU is equal to its VRS conventional DEA efficiency score plus its output share $\left(s^{k}=\frac{y^{k}}{Y}\right)$ minus their product.

If one takes (5) at face value, it may at first glance be seen that the only difference with (6) is in their denominators, unless $y^{k}=1$. In particular, if $y^{k}<(>) 1$ then $\breve{\theta}^{k}<(>) \hat{\theta}^{k}$. Apart from this, (5) may imply (depending on the values of $y^{k}$ ) ZSG-DEA efficiency scores that are: (i) less than their conventional DEA efficiency scores; (ii) greater-than-one; and (iii) less than (equal to) one despite that their conventional DEA efficiency scores are equal to (less than) one. As these results are inconsistent with the postulates of the ZSG-DEA model, namely the TAT and the BCET, doubts are raised about the use of (5).

To see that, consider first the case that (5) may imply ZSG-DEA efficiency scores that are less than their conventional DEA efficiency scores, i.e., $\breve{\theta}^{k}<\theta^{k}$ for $\theta^{k}<1$. This is inconsistent with the TAT, which implies exactly the opposite since in
this case the reduction coefficient, which measures the vertical distance between the conventional DEA and the ZSG-DEA efficient frontier at the evaluated DMU's input level, is positive but less than one (see Figure 1). ${ }^{5}$ To prove that the difference between the ZSG-DEA efficiency scores implied by (5) and their conventional DEA efficiency scores can be negative, substitute (5) into $\breve{\theta}^{k}-\theta^{k}$ to get:

$$
\begin{equation*}
\breve{\theta}^{k}-\theta^{k}=\frac{y^{k}-\theta^{k}}{Y-y^{k}+1} \tag{10}
\end{equation*}
$$

which is non-negative only if $y^{k} \geq \theta^{k}$. If however $y^{k}<\theta^{k}$ then $\breve{\theta}^{k}<\theta^{k}$. On the contrary, the ZSG-DEA efficiency scores implied by (6) are never less than their conventional DEA efficiency scores. This can be verified by substituting (6) into $\hat{\theta}^{k}-$ $\theta^{k}$ to get:

$$
\begin{equation*}
\hat{\theta}^{k}-\theta^{k}=\frac{y^{k}\left(1-\theta^{k}\right)}{Y} \tag{11}
\end{equation*}
$$

that is always non-negative since by definition $0<\theta^{k} \leq 1$.
Second, consider the case that (5) may imply ZSG-DEA efficiency scores that are greater than one, i.e., $\breve{\theta}^{k}>1$, which may occur if either $\theta^{k}=1$ or $\theta^{k}<1$. This is inconsistent with the definition of efficiency. Nevertheless, for $\theta^{k}=1$, (5) implies:

$$
\begin{equation*}
\breve{\theta}^{k}=\frac{Y}{Y-y^{k}+1} \tag{12}
\end{equation*}
$$

which differs from one unless $y^{k}=1$. If however $y^{k}>1$ then $\breve{\theta}^{k}>1$. On the other hand, in the case of $\theta^{k}<1$, for (5) to imply $\breve{\theta}^{k}<1$ it is necessary that:

$$
\begin{equation*}
\left(Y-y^{k}\right)\left(\theta^{k}-1\right)+\left(y^{k}-1\right)<0 \tag{13}
\end{equation*}
$$

As the first term in (13) is negative for $\theta^{k}<1$, a sufficient condition for the above inequality to hold is that $y^{k} \leq 1$. If however $y^{k}>1$ then it is possible for the second term in (13) to be greater than the absolute value of the first term and thus, for (5) to

[^4]imply $\breve{\theta}^{k}>1$. On the contrary, the ZSG-DEA efficiency scores implied by (6) are never greater-than-one. Indeed, in terms of (6), $\hat{\theta}^{k}>1$ requires that:
\[

$$
\begin{equation*}
\theta^{k}\left(Y-y^{k}\right)-\left(Y-y^{k}\right)>0 \tag{14}
\end{equation*}
$$

\]

which is impossible since by definition $0<\theta^{k} \leq 1$.
Third, consider the case that (5) may imply $\breve{\theta}^{k}<1$ even though $\theta^{k}=1$. From (12), we can see that this occurs if $y^{k}<1$. On the other hand, if $y^{k}>1$, then it is possible for the second term in (13) to be equal to the absolute value of the first term and thus, for (5) to imply $\breve{\theta}^{k}=1$ even though $\theta^{k}<1$. As we have seen, both of these are inconsistent with the BCET, which postulates that $\hat{\theta}^{k}=1$ as long as $\theta^{k}=1$ and $\hat{\theta}^{k}<1$ as long as $\theta^{k}<1 .{ }^{6}$ On the contrary, if $\theta^{k}=1$ then (6) implies that $\hat{\theta}^{k}=$ $\frac{Y-y^{k}+y^{k}}{Y}=1$. On the other hand, if $\theta^{k}<1$ then (6) implies that $\hat{\theta}^{k}<1$. This can be verified by considering that, in terms of (6), $\hat{\theta}^{k}<1$ requires that:

$$
\begin{equation*}
\left(Y-y^{k}\right)\left(\theta^{k}-1\right)<0 \tag{15}
\end{equation*}
$$

which clearly holds for $\theta^{k}<1$.

## 3 Empirical Results

To further demonstrate that (5) may provide results that are inconsistent with the main postulates of the ZSG-DEA model, we provide some empirical evidence using three different data sets. First, we closely examine the conventional DEA and the ZSG-DEA efficiency scores reported in Table B1 of Hu and Fang (2010), who evaluated the performance of a sample of securities firms operated in Taiwan from 2001 to 2005 considering three inputs (fixed assets, financial capital and expenses) and a single output (market share). Descriptive statistics of these data are given in Table 1. Nevertheless, as Hu and Fang (2010) do not report the raw data, we cannot compute $\hat{\theta}^{k}$ directly from $\theta^{k}$. For this reason, in Table 2 we report only $\theta^{k}$ and $\breve{\theta}^{k}$ implied by (5).

As it follows from Table 3, 31.5 to $46 \%$ (depending on the year under consideration) of the results based on (5) are counterintuitive. This means that 16 to 23

[^5]firms have inappropriate ZSG-DEA efficiency scores. Specifically, 89.5 to $100 \%$ (depending on the year under consideration) of the counterintuitive results (or in other words 29.4 to $44 \%$ of all ZSG-DEA efficiency scores) belong to the case where $\breve{\theta}^{k}<$ $\theta^{k}$ for $\theta^{k}<1$ and $y^{k}<\theta^{k}$. The reason that most of the counterintuitive results belong to this case is that the average level of firms' market share ranged from 1.64 to $2.00 \%$ through years and its minimum level from 0.01 to $0.05 \%$; see Table 1 in Hu and Fang (2010). Consequently, there were several firms whose actual market share was smaller than their conventional DEA efficiency score. For 17 firms, in particular, this was the case for all their yearly observations (see firms \#2, \#3, \#6, \#15, \#16, \#18, \#22, \#24, \#34, \#37, \#38, \#42, \#49, \#55, \#60, \#64 and \#65 in Table 2). On the other hand, there are no counterintuitive results belonging to either the case where $\breve{\theta}^{k}>1$ for $\theta^{k}=1$ and $y^{k}>$ 1 or the case where $\breve{\theta}^{k}<1$ even though $\theta^{k}=1$ because $y^{k}<1$. Therefore, all firms deemed efficient by the conventional DEA model had a ZSG-DEA efficiency score that is equal to one. This implies in turn that their market share was very close to $1 \%$ since, in any other case, their $\breve{\theta}$ 's would differ from one (see (12)). Similarly, there are no counterintuitive results belonging to the case where $\breve{\theta}^{k}>1$ for $\theta^{k}<1$ and $y^{k}>1$, while 0 to $10.5 \%$ (depending on the year under consideration) of the counterintuitive results (or in other words 0 to $4 \%$ of all ZSG-DEA efficiency scores) belong to the case where $\breve{\theta}^{k}=1$ even though $\theta^{k}<1$ because $y^{k}>1$.

The second data set refers to Sydney 2000 Olympic Games and it is used to estimate efficiency with countries' population and Gross Domestic Product (GDP) as inputs and their medal index as a single output based on an output-oriented VRS conventional DEA model. The medal index is a weighted average of each country's gold, silver and bronze medals won computed for robustness purposes by means of five alternative weighting schemes, the first of which was proposed by Lins et al. (2003) and the other four by Churilov and Flitman (2006). The resulted model variables are reported in Table 4.

From Table 5, where for each alternative medal index we report the conventional DEA efficiency scores and the ZSG-DEA efficiency scores implied by (5) and (6), we can see that the ZSG-DEA efficiency scores implied by (5) are greater (less) than those implied by (6) for values of medal indices greater (less) than one, as required by the models given in section 2 . In addition, we can see that for all values of medal indices, (6) implies ZSG-DEA efficiency scores that are: (i) greater-than-or-
equal-to their conventional DEA efficiency scores; (ii) between zero and one; and (iii) equal to (less than) one for countries deemed efficient (inefficient) by the conventional DEA model.

On the contrary, as it follows from Table 6, 15 to $18 \%$ (depending on the medal index considered) of the results implied by (5) are counterintuitive. This means that 12 to 14 countries have inappropriate ZSG-DEA efficiency scores. Specifically, 23.1 to $30.8 \%$ (depending on the medal index considered) of the counterintuitive results (or in other words 3.8 to $5.1 \%$ of all ZSG-DEA efficiency scores implied by (5)) belong to the case where $\breve{\theta}^{k}<\theta^{k}$ for $\theta^{k}<1$ and $y^{k}<\theta^{k}$. Another 41.7 to $50 \%$ (depending on the medal index considered) of the counterintuitive results (or in other words 6.3 to $8.9 \%$ of all ZSG-DEA efficiency scores implied by (5)) belong to the cases where $\breve{\theta}^{k}>$ 1 for either $\theta^{k}=1$ or $\theta^{k}<1$ and $y^{k}>1$. Finally, an additional 21.4 to $33.3 \%$ (depending on the medal index considered) of the counterintuitive results (or in other words 3.8 to $5.1 \%$ of all ZSG-DEA efficiency scores implied by (5)) belong to the case where $\breve{\theta}^{k}<1$ even though $\theta^{k}=1$ because $y^{k}<1$. Notice that we found no results belonging to the case where $\breve{\theta}^{k}=1$ even though $\theta^{k}<1$ because $y^{k}>1$. This is not surprising as it is rather rare for the second term in (13) to be exactly equal to the absolute value of the first term.

The last data set refers to the 32 teams participated in the regular season of the 2009 National Football League (NFL) and it was taken from Collier et al. (2011), where an output-oriented VRS conventional DEA model is used to estimate teams' efficiency scores with the number of their total wins as the single output (reported in the second column of Table 7) and three indices capturing teams' skills in offense and defense as inputs (i.e., offensive yards per play to defensive yards per play, offensive third-down conversion success to defensive third-down conversion success and defensive penalty yards to offensive penalty yards). These conventional DEA efficiency scores along with their ZSG-DEA efficiency scores implied by (5) and (6) are reported in the last three columns of Table 7. From this table, we can see that (with the exception of team \#29 whose actual output is equal to one) the ZSG-DEA efficiency scores implied by (5) are greater than those implied by (6) as each team won more than one games in the league season under consideration. In addition, we can see that the efficiency scores implied by (6) satisfy the postulates of the ZSG-DEA model. On the contrary, as it follows from Table 8, 40.5\% of the results implied by (5) are counterintuitive and 13
teams have inappropriate ZSG-DEA efficiency scores (see the efficiency scores of teams \#2, \#3, \#4, \#7, \#8, \#14, \#16, \#17, \#18, \#20, \#23, \#24 and \#26 in Table 7). In particular, all counterintuitive results belong to the cases where $\breve{\theta}^{k}>1$ for either $\theta^{k}=$ 1 or $\theta^{k}<1$ and $y^{k}>1$.

## 4 Concluding Remarks

In this note, we have provided both theoretical and empirical evidence for choosing between the two alternative relations used to compute VRS ZSG-DEA efficiency scores under the proportional output reduction strategy with a single output. Both types of evidence are in favor of (6) rather than (5) as the latter fails in several occasions to fulfill either the postulates of the ZSG-DEA model (the TAT and the BCET) or the very definition of efficiency. We also provided an alternative to (6) by means of (9) where the VRS ZSG-DEA efficiency score of each DMU under evaluation is equal to its VRS conventional DEA efficiency score plus its output share minus their product.

Empirical results from three different data sets indicate that, for the data at hand, most of the counterintuitive ZSG-DEA efficiency scores implied by (5) fall into the cases where $($ i $) \breve{\theta}^{k}<\theta^{k}$ for $\theta^{k}<1$ and $y^{k}<\theta^{k}$, (ii) $\breve{\theta}^{k}>1$ for $\theta^{k}=1$ and $y^{k}>1$, and (iii) $\breve{\theta}^{k}<1$ even though $\theta^{k}=1$ because $y^{k}<1$. This does not mean that the other two cases, where either $\breve{\theta}^{k}>1$ for $\theta^{k}<1$ and $y^{k}>1$ or $\breve{\theta}^{k}=1$ even though $\theta^{k}<1$ because $y^{k}>1$, are less important as they may account for the majority of the counterintuitive results in some other data sets. However, despite the fact that more empirical analysis is always welcome, it is our belief that the empirical and theoretical evidence presented in this note is sufficient to warn researchers working with the ZSGDEA model.

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Figure 1: The conventional DEA and the ZSG-DEA frontiers for the proportional output reduction strategy


Table 1: Descriptive Statistics of Model's Variables, Securities Firms in Taiwan

| Fixed Assets NT\$ | Financial Capital NT\$ <br> $(000,000,000 \mathrm{~s})$ | Expenses NT\$ <br> $(000,000,000 \mathrm{~s})$ | Market Share <br> $(000,000,000 \mathrm{~s})$ |  |
| :---: | ---: | ---: | ---: | ---: |
| Max | $4,455.02$ | $7,815.06$ | $22,315.20$ | 8.66 |
| Min | 5.74 | 48.70 | 150.00 | 0.05 |
| Average | 932.13 | $1,637.57$ | $4,413.48$ | 1.64 |
| Standard Deviation | $1,074.12$ | $1,733.19$ | $4,730.23$ | 1.89 |
| M002 | $4,306.02$ |  |  |  |
| Max | 1.84 | $24,689.52$ | $10,560.52$ | 10.48 |
| Min | $1,006.18$ | 151.29 | 10.64 | 0.02 |
| Average | $1,132.26$ | $5,112.88$ | $1,935.13$ | 1.85 |
| Standard Deviation | $5,439.46$ | $2,323.28$ | 2.42 |  |
| 2003 | $4,413.71$ |  |  |  |
| Max | 0.00 | $25,382.95$ | $8,587.78$ | 11.01 |
| Min | $1,067.66$ | 154.58 | 11.12 | 0.01 |
| Average | $1,259.69$ | $5,594.04$ | $2,126.09$ | 2.02 |
| Standard Deviation | $5,958.09$ | $2,555.90$ | 2.53 |  |
| 2004 | $6,203.25$ |  |  |  |
| Max | 0.00 | $31,988.93$ | $14,008.10$ | 9.39 |
| Min | $1,135.49$ | 156.84 | 21.18 | 0.02 |
| Average | $6,163.88$ | $3,010.77$ | 2.00 |  |
| Standard Deviation | $1,439.44$ | $6,822.10$ | $3,757.83$ | 2.51 |
| 2005 |  |  |  |  |
| Max | $6,692.11$ | $33,559.95$ | $12,772.28$ | 7.63 |
| Min | 0.00 | 157.81 | 26.80 | 0.02 |
| Average | $6,284.13$ | $3,213.83$ | 2.00 |  |
| Standard Deviation | $1,141.64$ | $7,069.14$ | $3,682.73$ | 2.31 |

Source: Table 1 in Hu and Fang (2010).

Table 2: Estimated Efficiency Scores, Securities Firms in Taiwan

| Securities <br> Firm | 2001 |  | 2002 |  | 2003 |  | 2004 |  | 2005 |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | $\theta^{k}$ | $\breve{\theta}^{k}$ | $\theta^{k}$ | $\breve{\theta}^{k}$ | $\theta^{k}$ | $\breve{\theta}^{k}$ | $\theta^{k}$ | $\breve{\theta}^{k}$ | $\theta^{k}$ | $\breve{\theta}^{k}$ |
| 1. Jih Sun | 0.951 | 0.989 | 1.000 | 1.000 | 0.763 | 0.810 | 0.884 | 0.930 | 1.000 | 1.000 |
|  | 0 | 3 | 0 | 0 | 0 | 5 | 0 | 7 | 0 | 0 |
| 2. Jen Hsin | 0.597 | 0.595 |  |  |  |  |  |  |  |  |
|  | 0 0.596 | 8 0.595 |  |  |  |  |  |  |  |  |
| 3. First | 0.596 | 0.595 | 0.854 | 0.851 | 0.626 | 0.624 | 0.642 | 0.638 | 0.430 | 0.427 |
|  | 0 | 4 | 0 | 2 | 0 | 7 | 0 | 7 | 0 | 5 |
| 4. Asia | 0.748 | 0.752 | 0.584 | 0.587 | 0.564 | 0.566 |  |  |  |  |
|  | 0 | 4 | 0 | 2 | 0 | 4 |  |  |  |  |
| 5. Tingkong | 0.653 | 0.655 |  |  |  |  |  |  |  |  |
|  | 0 | 5 |  |  |  |  |  |  |  |  |
| 6. Entrust | 0.552 | 0.550 |  |  |  |  |  |  |  |  |
|  | 0 | 6 |  |  |  |  |  |  |  |  |
| 7. Horizon | 0.452 | 0.460 | 0.376 | 0.378 | 0.478 | 0.481 | 0.385 | 0.386 | 0.467 | 0.466 |
|  | 0 | 1 | 0 | 9 | 0 | 8 |  | 5 | 0 | 6 |
| 8. Macquarie | 0.501 | 0.498 | 0.623 | 0.619 | 1.000 | 1.000 | 1.000 | 1.000 | 0.898 | 0.893 |
|  | 0 | 6 | 0 | 8 | 0 | 0 | 0 | 0 | 0 | 6 |
| 9. ABN | 1.000 | 1.000 | 1.000 | 1.000 | 1.000 | 1.000 | 1.000 | 1.000 | 1.000 | 1.000 |
| Amro | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 10. Merrill | 1.000 | 1.000 | 1.000 | 1.000 | 1.000 | 1.000 | 1.000 | 1.000 | 1.000 | 1.000 |
| Lynch | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 11. Nomura (HK) | 1.000 | 1.000 | 1.000 | 1.000 | 1.000 | 1.000 | 1.000 | 1.000 | 0.671 | 0.665 |
|  | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 7 |
| 12. Societe | 1.000 | 1.000 | 1.000 | 1.000 | 1.000 | 1.000 | 0.672 | 0.665 |  |  |
| Generale <br> (HK) | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 7 | 0 | 0 |
| 13. Goldman | 1.000 | 1.000 | 1.000 | 1.000 | 0.499 | 0.494 |  | 1.000 |  |  |
| Sachs <br> (Asia) | 0 | 0 | 0 | 0 | 0 | 4 | 0 | 0 | 0 | 0 |
| 14. Oriental | 0.980 | 0.986 | 0.490 | 0.492 | 1.000 | 1.000 | 1.000 | 1.000 | 1.000 | 1.000 |
|  | 0 | 5 | 0 | 4 | 0 | 0 | 0 | 0 | 0 | 0 |
| 15. First | 0.480 | 0.478 | 0.442 | 0.440 |  |  |  |  |  |  |
| Taiwan | 0 | 7 | 0 | 2 |  |  |  |  |  |  |
| 16. Tachan | 0.735 | 0.733 | 0.781 | 0.778 | 0.534 | 0.533 | 0.993 | 0.988 | 0.765 | 0.763 |
|  | 0 | 7 | 0 | 1 | 0 | 3 | 0 | 1 | 0 | 7 |
| 17. Hua Nan | 0.809 | 0.827 | 0.957 | 0.979 | 0.786 | 0.804 | 0.950 | 0.965 | 0.860 | 0.879 |
|  | 0 | 2 | 0 | 7 | 0 | 9 | 0 | 4 | 0 | 6 |
| 18. Full Long | 0.625 | 0.623 | 0.648 | 0.646 | 0.314 | 0.313 | 0.547 | 0.543 | 0.186 | 0.185 |
|  | 0 | 5 | 0 | 3 | 0 | 0 | 0 | 3 | 0 | 5 |
| 19. Pacific | 0.827 | 0.831 | 0.614 | 0.616 | 0.516 | 0.517 | 0.460 | 0.460 | 0.465 | 0.464 |
|  | 0 | 0 | 0 | 6 | 0 | 7 | 0 | 2 | 0 | 8 |
| 20. Ta Ching | 0.844 | 0.843 | 0.794 | 0.793 | 0.691 | 0.691 | 0.777 | 0.774 | 0.686 | 0.683 |
|  | 0 | 9 | 0 | 6 | 0 | 2 | 0 | 3 | 0 | 6 |
| 21. Capital | 0.887 | 0.927 | 0.812 | 0.858 | 0.728 | 0.770 | 1.000 | 1.000 | 0.920 | 0.966 |
|  | 0 | 3 | 0 | 8 | 0 | 7 | 0 | 0 | 0 | 6 |
| 22. Chung | 0.623 | 0.622 | 0.558 | 0.557 |  |  |  |  |  |  |
| Hsing | 0 | 4 | 0 | 1 |  |  |  |  |  |  |
| 23. First | 0.908 | 0.908 | 0.798 | 0.799 | 0.813 | 0.817 | 0.751 | 0.758 | 0.925 | 0.942 |
| Taisec | 0 | 8 | 0 | 4 | 0 | 6 | 0 | 9 | 0 | 8 |
| 24. Forwin | 0.446 | 0.445 | 0.353 | 0.351 | 0.384 | 0.382 | 0.492 | 0.488 | 0.173 | 0.171 |
|  | 0 | 1 | 0 | 8 | 0 | 6 | 0 | 4 | 0 | 6 |
| 25. Sinopac | 0.899 | 0.932 | 0.854 | 0.907 | 0.876 | 0.938 | 0.876 | 0.922 | 0.965 | 1.000 |
|  | 0 | 9 | 0 | 3 | 0 | 7 | 0 | 5 | 0 | 0 |
| 26. Taiwan | 1.000 | 1.000 | 0.899 | 0.945 | 0.846 | 0.900 | 1.000 | 1.000 | 1.000 | 1.000 |
|  | 0 | 0 | 0 | 3 | 0 | 0 | 0 | 0 | 0 | 0 |
| 27. Taiyu | 0.659 | 0.661 | 0.669 | 0.673 |  |  |  |  |  |  |
|  | 0 | 7 | 0 | 9 |  |  |  |  |  |  |
| 28. KGI | 1.000 | 1.000 | 0.833 | 0.874 | 0.759 | 0.804 | 0.995 | 1.000 | 1.000 | 1.000 |
|  | 0 | 0 | 0 | 0 | 0 | 4 | 0 | 0 | 0 | 0 |
| 29. IBT | 0.977 | 0.976 | 0.672 | 0.673 | 1.000 | 1.000 | 1.000 | 1.000 | 0.921 | 0.931 |
|  | 0 | 5 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 8 |


| 30. Grand Cathay | 0.831 | 0.869 | 0.697 | 0.733 | 0.780 | 0.818 | 0.948 | 1.000 | 1.000 | 1.000 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | 0 | 1 | 0 | 2 | 0 | 3 | 0 | 0 | 0 | 0 |
| 31. Taiwan | 0.881 | 0.904 | 0.850 | 0.877 | 0.687 | 0.711 | 0.714 | 0.736 | 0.785 | 0.811 |
| Intl. | 0 | 0 | 0 | 7 | 0 | 3 | 0 | 0 | 0 | 7 |
| 32. President | 0.928 | 0.966 | 0.958 | 1.000 | 0.892 | 0.935 | 0.945 | 0.981 | 0.946 | 0.981 |
|  | 0 | 7 | 0 | 0 | 0 | 8 | 0 | 4 | 0 | 1 |
| 33. Masterlin k | 0.858 | 0.890 | 0.756 | 0.791 | 0.809 | 0.852 | 0.913 | 0.950 | 0.898 | 0.929 |
|  | 0 | 2 | 0 | 0 | 0 | 0 | 0 | 6 | 0 | 6 |
| 34. Primasia | 0.482 | 0.481 | 0.651 | 0.648 | 0.607 | 0.605 | 0.799 | 0.797 | 0.487 | 0.484 |
|  | 0 | 6 | 0 | 2 | 0 | 0 | 0 | 0 | 0 | 5 |
| 35. Chinatrust | 1.000 | 1.000 | 0.511 | 0.515 | 0.777 | 0.785 | 0.909 | 0.918 | 1.000 | 1.000 |
|  | 0 | 0 | 0 | 9 | 0 | 9 | 0 | 8 | 0 | 0 |
| 36. Barits | 1.000 | 1.000 | 0.712 | 0.718 |  |  |  |  |  |  |
|  | 0 | 0 | 0 | $2$ |  |  |  |  |  |  |
| 37. Grand | 0.583 | 0.582 | 0.530 | 0.528 | 0.601 | 0.598 | 0.412 | 0.408 | 0.563 | 0.557 |
| Fortune | 0 | 4 | 0 | 2 | 0 | 9 | 0 | 7 | 0 | 9 |
| 38. Ta Chong | 0.937 | 0.935 | 0.632 | 0.631 | 0.532 | 0.531 | 0.776 | 0.772 | 0.646 | 0.644 |
|  | 0 | 3 | 0 | 2 | 0 | 4 | 0 | 9 | 0 | 5 |
| 39. Reliance | 0.772 | 0.769 | 1.000 | 1.000 | 0.406 | 0.408 | 0.589 | 0.585 | 0.581 | 0.577 |
|  | 0 | 9 | 0 | 0 | 0 | 7 | 0 | 6 | 0 | 0 |
| 40. Mega | 0.621 | 0.627 | 0.569 | 0.574 | 0.863 | 0.904 | 0.757 | 0.786 | 0.884 | 0.912 |
|  | 0 | 7 | 0 | 3 | 0 | 4 | 0 | 4 | 0 | 1 |
| 41. Concord Intl. | 0.653 | 0.653 | 0.905 | 0.902 | 0.578 | 0.576 | 1.000 | 1.000 | 0.477 | 0.476 |
|  | 0 | 2 | 0 | 4 | 0 | 7 | 0 | 0 | 0 | 0 |
| 42. Jinhwa | 0.605 | 0.601 |  |  |  |  |  |  |  |  |
|  | 0 | 3 |  |  |  |  |  |  |  |  |
| 43. Waterland | 0.623 | 0.627 | 0.607 | 0.616 | 0.529 | 0.538 | 0.696 | 0.702 | 0.556 | 0.562 |
|  | 0 | 7 | 0 | 0 | 0 | 7 | 0 | 6 | 0 | 0 |
| 44. Hsinbao | 0.995 | 0.999 |  |  |  |  |  |  |  |  |
|  | 0 | 1 |  |  |  |  |  |  |  |  |
| 45. J.P. | 0.657 | 0.657 | 0.688 | 0.687 | 0.619 | 0.617 | 0.601 | 0.599 | 0.796 | 0.794 |
| Morgan | 0 | 5 | 0 | 1 | 0 | 8 | 0 | 7 | 0 | 0 |
| 46. Concord | 0.832 | 0.840 | 0.648 | 0.655 | 0.628 | 0.636 | 0.797 | 0.802 | 0.724 | 0.732 |
|  | 0 | 3 | 0 | 8 | 0 | 1 | 0 | 1 | 0 | 8 |
| Continue |  |  |  |  |  |  |  |  |  |  |
| 47. Concourse | 0.678 | 0.678 |  |  |  |  |  |  |  |  |
|  | 0 | 1 |  |  |  |  |  |  |  |  |
| 48. Sinopac (Old) | 0.787 | 0.792 |  |  |  |  |  |  |  |  |
|  | 0 | 8 |  |  |  |  |  |  |  |  |
| 49. Grand | 0.611 | 0.610 |  |  |  |  |  |  |  |  |
| Orient | 0 | 5 |  |  |  |  |  |  |  |  |
| 50. Shinkong | 0.614 | 0.614 | 0.259 | 0.259 | 0.768 | 0.765 | 0.678 | 0.674 | 0.728 | 0.727 |
|  | 0 | 4 | 0 | 8 | 0 | 6 | 0 | 5 | 0 | 5 |
| 51. Citibank | 1.000 | 1.000 | 1.000 | 1.000 | 1.000 | 1.000 | 1.000 | 1.000 | 1.000 | 1.000 |
|  | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 52. Fu Hwa | 0.831 | 0.849 | 0.879 | 0.905 | 0.828 | 0.854 | 0.880 | 0.903 | 0.703 | 0.725 |
|  | 0 | 8 | 0 | 2 | 0 | 0 | 0 | 4 | 0 | 9 |
| 53. Sun-Fund | 1.000 | 1.000 | 0.498 | 0.495 | 0.374 | 0.372 | 0.409 | 0.405 | 0.321 | 0.318 |
|  | 0 | 0 | 0 | 5 | 0 | 1 | 0 | 7 | 0 | 2 |
| 54. Ho Tung | 1.000 | 1.000 | 0.503 | 0.499 | 0.823 | 0.816 | 0.633 | 0.627 |  |  |
|  | 0 | 0 | 0 | 0 | 0 | 5 | 0 | 9 |  |  |
| 55. E. Sun | 0.807 | 0.801 | 0.705 | 0.700 | 0.442 | 0.441 | 0.618 | 0.616 | 0.655 | 0.653 |
|  | 0 | 3 | 0 | 4 | 0 | 4 | 0 | 0 | 0 | 9 |
| 56. Daiwa | 1.000 | 1.000 | 0.943 | 0.936 | 1.000 | 1.000 | 1.000 | 1.000 | 1.000 | 1.000 |
|  | 0 | 0 | 0 | 3 | 0 | 0 | 0 | 0 | 0 | 0 |
| 57. Fubon | 1.000 | 1.000 | 1.000 | 1.000 | 1.000 | 1.000 | 1.000 | 1.000 | 1.000 | 1.000 |
|  | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 58. Polaris | 0.953 | 0.994 | 0.975 | 1.000 | 0.995 | 1.000 | 1.000 | 1.000 | 1.000 | 1.000 |
|  | 0 | 4 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 59. Yuanta | 1.000 | 1.000 | 1.000 | 1.000 | 1.000 | 1.000 | 1.000 | 1.000 | 1.000 | 1.000 |
|  | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 60. Far | 0.660 | 0.655 | 0.609 | 0.604 | 0.443 | 0.439 | 0.554 | 0.549 | 0.252 | 0.250 |
| Eastern | 0 | 9 | 0 | 2 | 0 | 8 | 0 | 3 | 0 | 3 |
| 61. Yuan Li | 0.916 | 0.913 | 0.743 | 0.740 | 1.000 | 1.000 |  |  |  |  |
|  | 0 | 3 | 0 | 1 | 0 | 0 |  |  |  |  |


| 62. Deutsche (Asia) |  |  | 1.000 | 1.000 | 1.000 | 1.000 | 1.000 | 1.000 | 1.000 | 1.000 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 63. Lehman |  |  |  |  | 1.000 | 1.000 | 1.000 | 1.000 | 1.000 | 1.000 |
| Brothers |  |  |  |  | 0 | 0 | 0 | 0 | 0 | 0 |
| 64. HSBC <br> (HK) |  |  |  |  |  |  |  |  | 0.467 | 0.463 |
|  |  |  |  |  |  |  |  |  | 0 | 0 |
| 65. Cathay |  |  |  |  |  |  | 0.384 | 0.380 | 0.691 | 0.686 |
|  |  |  |  |  |  |  | 0 | 7 | 0 | 6 |
| Average | 0.801 | 0.806 | 0.748 | 0.756 | 0.748 | 0.758 | 0.808 | 0.814 | 0.767 | 0.772 |
|  | 0 | 8 | 9 | 4 | 2 | 1 | 7 | 4 | 3 | 1 |

Source: Table B1 in Hu and Fang (2010).

Table 3: Counterintuitive Results Implied by (5), Securities Firms in Taiwan

| Cases | 2001 | 2002 | 2003 | 2004 | 2005 |
| :--- | :---: | :---: | :---: | :---: | :---: |
| $\breve{\theta}^{k}<\theta^{k}$ for $\theta^{k}<1$ and $y^{k}<\theta^{k}$ | 20 | 19 | 15 | 17 | 22 |
| $\breve{\theta}^{k}>1$ for $\theta^{k}=1$ and $y^{k}>1$ | - | - | - | - | - |
| $\breve{\theta}^{k}>1$ for $\theta^{k}<1$ and $y^{k}>1$ | - | - | - | - | - |
| $\breve{\theta}^{k}<1$ even though $\theta^{k}=1$ because $y^{k}<1$ | - | - | - | - | - |
| $\breve{\theta}^{k}=1$ even though $\theta^{k}<1$ because $y^{k}>1$ | - | 2 | 1 | 2 | 1 |
| Percentage (\%) of the total | 32.8 | 38.9 | 31.4 | 38.0 | 46.0 |

Table 4: Model's Variables, Sydney 2000 Olympic Games

|  |  |  | GDP US\$ | Medal | Medal | Medal | Medal |
| :--- | :--- | ---: | ---: | ---: | ---: | ---: | ---: | | Medal |
| ---: |
|  |
|  |
|  |
|  |
|  |
| Country |

$\left.\left.\begin{array}{lrrrrrrr}\hline \text { Continue } & & & & & & \\ \hline \text { 61. } & \text { Romania } & 22,327 & 37,911 & 9.4317 & 8.9474 & 9.5000 & 9.3125\end{array}\right) 8.6658\right)$

Note: Medal index A was computed with the use of a weighting scheme that employs the following values $(0.5814,0.2437,0.1749)$ as weights respectively assigned to the number of gold, silver and bronze medals. On the other hand, medal indices B, C, D and E were computed with the use of weighting schemes that respectively use the following values ( 0.5263 , $0.3684,0.1053)$, $(0.6250,0.2500,0.1250),(0.6250,0.3125,0.0625)$ and $(0.3333,0.3333$, 0.3333 ) as weights.

Source: The data in the first three columns were taken from Churilov and Flitman (2006).
The data in the last five columns come from authors' calculations.

Table 5: Estimated Efficiency Scores, Sydney 2000 Olympic Games

|  |  | Medal Index A |  |  | Medal Index B |  |  | Medal Index C |  |  | Medal Index D |  |  | Medal Index E |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Country | $\theta^{k}$ | $\breve{\theta}^{k}$ | $\hat{\theta}^{k}$ | $\theta^{k}$ | $\breve{\theta}^{k}$ | $\hat{\theta}^{k}$ | $\theta^{k}$ | $\breve{\theta}^{k}$ | $\hat{\theta}^{k}$ | $\theta^{k}$ | $\breve{\theta}^{k}$ | $\hat{\theta}^{k}$ | $\theta^{k}$ | $\breve{\theta}^{k}$ | $\hat{\theta}^{k}$ |
| 1. Algeria | 0.1080 | 0.1121 | 0.1120 | 0.0950 | 0.0987 | 0.0986 | 0.0980 | 0.1018 | 0.1017 | 0.0870 | 0.0904 | 0.0904 | 0.1410 | 0.1460 | 0.1457 |
| 2. Argentina | 0.0440 | 0.0466 | 0.0466 | 0.0480 | 0.0510 | 0.0510 | 0.0400 | 0.0423 | 0.0424 | 0.0390 | 0.0414 | 0.0414 | 0.0690 | 0.0731 | 0.0730 |
| 3. Armenia | 0.4190 | 0.4182 | 0.4193 | 0.2300 | 0.2296 | 0.2303 | 0.3290 | 0.3283 | 0.3293 | 0.1670 | 0.1667 | 0.1672 | 0.5270 | 0.5264 | 0.5275 |
| 4. Australia | 1.0000 | 1.0605 | 1.0000 | 1.0000 | 1.0649 | 1.0000 | 1.0000 | 1.0608 | 1.0000 | 1.0000 | 1.0630 | 1.0000 | 1.0000 | 1.0634 | 1.0000 |
| 5. Austria | 0.1700 | 0.1741 | 0.1738 | 0.1620 | 0.1662 | 0.1659 | 0.1800 | 0.1844 | 0.1841 | 0.1820 | 0.1866 | 0.1862 | 0.1160 | 0.1189 | 0.1189 |
| 6. Azerbaijan | 0.8770 | 0.8785 | 0.8775 | 0.7300 | 0.7314 | 0.7310 | 0.9080 | 0.9095 | 0.9084 | 0.8500 | 0.8515 | 0.8507 | 0.5830 | 0.5844 | 0.5844 |
| 7. Bahamas | 1.0000 | 0.9994 | 1.0000 | 1.0000 | 0.9997 | 1.0000 | 1.0000 | 0.9996 | 1.0000 | 1.0000 | 0.9998 | 1.0000 | 1.0000 | 0.9989 | 1.0000 |
| 8. Barbados | 1.0000 | 0.9973 | 1.0000 | 1.0000 | 0.9970 | 1.0000 | 1.0000 | 0.9971 | 1.0000 | 1.0000 | 0.9969 | 1.0000 | 1.0000 | 0.9978 | 1.0000 |
| 9. Belarus | 0.7680 | 0.7801 | 0.7714 | 0.6520 | 0.6627 | 0.6564 | 0.6870 | 0.6980 | 0.6911 | 0.5870 | 0.5967 | 0.5918 | 1.0000 | 1.0154 | 1.0000 |
| 10. Belgium | 0.1000 | 0.1030 | 0.1030 | 0.0980 | 0.1012 | 0.1011 | 0.0860 | 0.0886 | 0.0886 | 0.0780 | 0.0804 | 0.0805 | 0.1580 | 0.1629 | 0.1626 |
| 11. Brazil | 0.0820 | 0.0900 | 0.0896 | 0.0940 | 0.1031 | 0.1025 | 0.0730 | 0.0802 | 0.0799 | 0.0730 | 0.0803 | 0.0799 | 0.1360 | 0.1487 | 0.1472 |
| 12. Britain | 0.4550 | 0.4875 | 0.4730 | 0.4490 | 0.4823 | 0.4676 | 0.4630 | 0.4963 | 0.4812 | 0.4640 | 0.4981 | 0.4826 | 0.4160 | 0.4458 | 0.4337 |
| 13. Bulgaria | 0.9580 | 0.9705 | 0.9587 | 0.9950 | 1.0086 | 0.9951 | 0.9750 | 0.9880 | 0.9754 | 1.0000 | 1.0139 | 1.0000 | 0.8890 | 0.9003 | 0.8906 |
| 14. Cameroon | 0.1340 | 0.1355 | 0.1357 | 0.1190 | 0.1203 | 0.1205 | 0.1410 | 0.1426 | 0.1428 | 0.1380 | 0.1396 | 0.1398 | 0.0770 | 0.0778 | 0.0780 |
| 15. Canada | 0.1990 | 0.2112 | 0.2092 | 0.1730 | 0.1842 | 0.1826 | 0.1860 | 0.1974 | 0.1957 | 0.1660 | 0.1765 | 0.1752 | 0.2300 | 0.2446 | 0.2417 |
| 16. Chile | 0.0150 | 0.0155 | 0.0156 | 0.0090 | 0.0093 | 0.0093 | 0.0100 | 0.0104 | 0.0104 | 0.0050 | 0.0052 | 0.0052 | 0.0290 | 0.0300 | 0.0301 |
| 17. China | 0.7430 | 0.8211 | 0.7623 | 0.7320 | 0.8084 | 0.7517 | 0.7580 | 0.8386 | 0.7767 | 0.7610 | 0.8423 | 0.7796 | 0.6650 | 0.7307 | 0.6864 |
| 18. Colombia | 0.0380 | 0.0398 | 0.0398 | 0.0340 | 0.0356 | 0.0357 | 0.0400 | 0.0419 | 0.0420 | 0.0400 | 0.0419 | 0.0420 | 0.0230 | 0.0240 | 0.0241 |
| 19. Costa Rica | 0.0860 | 0.0869 | 0.0871 | 0.0500 | 0.0505 | 0.0507 | 0.0600 | 0.0606 | 0.0608 | 0.0290 | 0.0293 | 0.0294 | 0.1780 | 0.1796 | 0.1798 |
| 20. Croatia | 0.1680 | 0.1699 | 0.1701 | 0.1360 | 0.1376 | 0.1378 | 0.1640 | 0.1659 | 0.1661 | 0.1460 | 0.1478 | 0.1479 | 0.1590 | 0.1606 | 0.1608 |
| 21. Cuba | 1.0000 | 1.0315 | 1.0000 | 1.0000 | 1.0327 | 1.0000 | 1.0000 | 1.0324 | 1.0000 | 1.0000 | 1.0334 | 1.0000 | 1.0000 | 1.0290 | 1.0000 |
| 22. Czech Rep. | 0.2520 | 0.2592 | 0.2579 | 0.2490 | 0.2564 | 0.2551 | 0.2430 | 0.2501 | 0.2489 | 0.2370 | 0.2441 | 0.2430 | 0.2900 | 0.2978 | 0.2962 |
| 23. Denmark | 0.3740 | 0.3796 | 0.3783 | 0.3860 | 0.3922 | 0.3906 | 0.3810 | 0.3868 | 0.3853 | 0.3910 | 0.3972 | 0.3955 | 0.3520 | 0.3574 | 0.3562 |
| 24. Estonia | 0.5880 | 0.5891 | 0.5893 | 0.4410 | 0.4420 | 0.4424 | 0.5330 | 0.5341 | 0.5343 | 0.4370 | 0.4380 | 0.4384 | 0.6870 | 0.6880 | 0.6880 |
| 25. Ethiopia | 1.0000 | 1.0069 | 1.0000 | 1.0000 | 1.0060 | 1.0000 | 1.0000 | 1.0071 | 1.0000 | 1.0000 | 1.0067 | 1.0000 | 0.9700 | 0.9755 | 0.9703 |
| 26. Finland | 0.2920 | 0.2962 | 0.2957 | 0.2650 | 0.2692 | 0.2687 | 0.2980 | 0.3024 | 0.3018 | 0.2880 | 0.2924 | 0.2918 | 0.2400 | 0.2436 | 0.2433 |
| 27. France | 0.5830 | 0.6251 | 0.6007 | 0.5780 | 0.6214 | 0.5964 | 0.5860 | 0.6286 | 0.6038 | 0.5860 | 0.6296 | 0.6041 | 0.5640 | 0.6049 | 0.5820 |
| 28. FYROM | 0.1760 | 0.1760 | 0.1765 | 0.0990 | 0.0990 | 0.0993 | 0.1230 | 0.1230 | 0.1234 | 0.0590 | 0.0590 | 0.0592 | 0.3180 | 0.3180 | 0.3187 |
| 29. Georgia | 0.5940 | 0.5955 | 0.5954 | 0.3410 | 0.3420 | 0.3424 | 0.4260 | 0.4271 | 0.4274 | 0.2080 | 0.2086 | 0.2090 | 1.0000 | 1.0033 | 1.0000 |
| 30. Germany | 0.6930 | 0.7489 | 0.7100 | 0.6630 | 0.7177 | 0.6812 | 0.6670 | 0.7211 | 0.6849 | 0.6370 | 0.6895 | 0.6559 | 0.7830 | 0.8459 | 0.7964 |
| 31. Greece | 0.4220 | 0.4349 | 0.4302 | 0.4340 | 0.4481 | 0.4427 | 0.4240 | 0.4372 | 0.4323 | 0.4310 | 0.4449 | 0.4396 | 0.4260 | 0.4388 | 0.4341 |


| Continue |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 32. Hungary | 0.7070 | 0.7269 | 0.7134 | 0.6950 | 0.7154 | 0.7018 | 0.7200 | 0.7407 | 0.7264 | 0.7220 | 0.7435 | 0.7285 | 0.6340 | 0.6506 | 0.6407 |
| 33. Iceland | 0.4750 | 0.4740 | 0.4753 | 0.3090 | 0.3083 | 0.3092 | 0.3590 | 0.3582 | 0.3593 | 0.1950 | 0.1946 | 0.1952 | 0.7700 | 0.7686 | 0.7702 |
| 34. India | 0.0060 | 0.0066 | 0.0066 | 0.0030 | 0.0033 | 0.0033 | 0.0040 | 0.0044 | 0.0044 | 0.0020 | 0.0022 | 0.0022 | 0.0110 | 0.0120 | 0.0121 |
| 35. Indonesia | 0.1020 | 0.1071 | 0.1069 | 0.1130 | 0.1187 | 0.1184 | 0.0990 | 0.1040 | 0.1038 | 0.1020 | 0.1073 | 0.1070 | 0.1290 | 0.1351 | 0.1347 |
| 36. Iran | 0.0970 | 0.1030 | 0.1027 | 0.0840 | 0.0893 | 0.0891 | 0.1000 | 0.1063 | 0.1059 | 0.0960 | 0.1021 | 0.1018 | 0.0690 | 0.0731 | 0.0730 |
| 37. Ireland | 0.0600 | 0.0606 | 0.0608 | 0.0850 | 0.0859 | 0.0861 | 0.0610 | 0.0616 | 0.0618 | 0.0740 | 0.0748 | 0.0750 | 0.0810 | 0.0818 | 0.0820 |
| 38. Israel | 0.0280 | 0.0285 | 0.0286 | 0.0160 | 0.0163 | 0.0163 | 0.0200 | 0.0203 | 0.0204 | 0.0100 | 0.0102 | 0.0102 | 0.0520 | 0.0529 | 0.0530 |
| 39. Italy | 0.5370 | 0.5753 | 0.5549 | 0.4930 | 0.5295 | 0.5117 | 0.5340 | 0.5724 | 0.5521 | 0.5110 | 0.5485 | 0.5296 | 0.5070 | 0.5434 | 0.5252 |
| 40. Jamaica | 0.6160 | 0.6189 | 0.6179 | 0.7040 | 0.7076 | 0.7058 | 0.5490 | 0.5517 | 0.5510 | 0.5550 | 0.5579 | 0.5571 | 1.0000 | 1.0044 | 1.0000 |
| 41. Japan | 0.2020 | 0.2205 | 0.2170 | 0.2150 | 0.2348 | 0.2309 | 0.2010 | 0.2196 | 0.2162 | 0.2070 | 0.2263 | 0.2226 | 0.2160 | 0.2351 | 0.2313 |
| 42. Kazakhstan | 0.2920 | 0.3000 | 0.2983 | 0.3200 | 0.3291 | 0.3269 | 0.3030 | 0.3115 | 0.3096 | 0.3220 | 0.3314 | 0.3290 | 0.2660 | 0.2727 | 0.2716 |
| 43. Kenya | 0.4730 | 0.4788 | 0.4769 | 0.5010 | 0.5072 | 0.5049 | 0.4660 | 0.4719 | 0.4700 | 0.4750 | 0.4811 | 0.4790 | 0.5100 | 0.5160 | 0.5137 |
| 44. Korea People's | 0.1780 | 0.1799 | 0.1801 | 0.1580 | 0.1597 | 0.1599 | 0.1420 | 0.1436 | 0.1438 | 0.1120 | 0.1133 | 0.1135 | 0.3160 | 0.3193 | 0.3190 |
| 45. Korea Rep. | 0.4320 | 0.4601 | 0.4484 | 0.4140 | 0.4421 | 0.4308 | 0.4240 | 0.4518 | 0.4404 | 0.4100 | 0.4375 | 0.4266 | 0.4510 | 0.4807 | 0.4677 |
| 46. Kuwait | 0.0740 | 0.0743 | 0.0745 | 0.0420 | 0.0422 | 0.0423 | 0.0520 | 0.0522 | 0.0524 | 0.0250 | 0.0251 | 0.0252 | 0.1480 | 0.1486 | 0.1489 |
| 47. Kyrgyzstan | 0.3830 | 0.3823 | 0.3834 | 0.2060 | 0.2057 | 0.2063 | 0.3020 | 0.3014 | 0.3023 | 0.1520 | 0.1517 | 0.1522 | 0.4750 | 0.4745 | 0.4756 |
| 48. Latvia | 0.4760 | 0.4777 | 0.4777 | 0.4550 | 0.4568 | 0.4568 | 0.4630 | 0.4648 | 0.4648 | 0.4460 | 0.4478 | 0.4478 | 0.4930 | 0.4946 | 0.4946 |
| 49. Lithuania | 0.4500 | 0.4541 | 0.4531 | 0.3520 | 0.3554 | 0.3549 | 0.4220 | 0.4260 | 0.4251 | 0.3620 | 0.3656 | 0.3650 | 0.4830 | 0.4869 | 0.4858 |
| 50. Mexico | 0.0620 | 0.0670 | 0.0669 | 0.0610 | 0.0660 | 0.0659 | 0.0580 | 0.0628 | 0.0627 | 0.0550 | 0.0596 | 0.0595 | 0.0780 | 0.0843 | 0.0840 |
| 51. Moldova | 1.0000 | 0.9981 | 1.0000 | 1.0000 | 0.9983 | 1.0000 | 1.0000 | 0.9979 | 1.0000 | 1.0000 | 0.9979 | 1.0000 | 1.0000 | 0.9989 | 1.0000 |
| 52. Morocco | 0.0840 | 0.0868 | 0.0868 | 0.0690 | 0.0714 | 0.0714 | 0.0660 | 0.0683 | 0.0683 | 0.0480 | 0.0497 | 0.0498 | 0.1570 | 0.1619 | 0.1616 |
| 53. Mozambique | 1.0000 | 0.9986 | 1.0000 | 0.9510 | 0.9496 | 0.9511 | 1.0000 | 0.9988 | 1.0000 | 1.0000 | 0.9988 | 1.0000 | 0.4510 | 0.4506 | 0.4516 |
| 54. Netherlands | 0.6390 | 0.6702 | 0.6507 | 0.6160 | 0.6482 | 0.6288 | 0.6630 | 0.6956 | 0.6744 | 0.6650 | 0.6989 | 0.6767 | 0.5140 | 0.5400 | 0.5272 |
| 55. New Zealand | 0.2700 | 0.2727 | 0.2727 | 0.1950 | 0.1971 | 0.1972 | 0.2410 | 0.2435 | 0.2435 | 0.1890 | 0.1911 | 0.1912 | 0.3310 | 0.3343 | 0.3339 |
| 56. Nigeria | 0.0400 | 0.0423 | 0.0423 | 0.0600 | 0.0635 | 0.0634 | 0.0410 | 0.0433 | 0.0434 | 0.0500 | 0.0529 | 0.0530 | 0.0570 | 0.0601 | 0.0601 |
| 57. Norway | 0.7540 | 0.7634 | 0.7569 | 0.6990 | 0.7084 | 0.7025 | 0.7570 | 0.7665 | 0.7599 | 0.7320 | 0.7417 | 0.7352 | 0.6880 | 0.6967 | 0.6914 |
| 58. Poland | 0.3270 | 0.3433 | 0.3386 | 0.3240 | 0.3408 | 0.3359 | 0.3320 | 0.3489 | 0.3438 | 0.3350 | 0.3524 | 0.3471 | 0.2970 | 0.3114 | 0.3077 |
| 59. Portugal | 0.0370 | 0.0380 | 0.0381 | 0.0210 | 0.0216 | 0.0217 | 0.0260 | 0.0267 | 0.0268 | 0.0130 | 0.0134 | 0.0134 | 0.0710 | 0.0729 | 0.0730 |
| 60. Qatar | 0.1590 | 0.1591 | 0.1595 | 0.0890 | 0.0891 | 0.0893 | 0.1090 | 0.1091 | 0.1094 | 0.0520 | 0.0520 | 0.0522 | 0.3470 | 0.3470 | 0.3477 |
| 61. Romania | 0.8330 | 0.8621 | 0.8382 | 0.7730 | 0.8008 | 0.7797 | 0.8240 | 0.8534 | 0.8295 | 0.7920 | 0.8211 | 0.7984 | 0.8120 | 0.8382 | 0.8173 |
| 62. Russia | 1.0000 | 1.1067 | 1.0000 | 1.0000 | 1.1065 | 1.0000 | 1.0000 | 1.1078 | 1.0000 | 1.0000 | 1.1085 | 1.0000 | 1.0000 | 1.1014 | 1.0000 |
| 63. Saudi Arabia | 0.0300 | 0.0313 | 0.0313 | 0.0320 | 0.0335 | 0.0335 | 0.0260 | 0.0271 | 0.0272 | 0.0260 | 0.0272 | 0.0272 | 0.0480 | 0.0500 | 0.0501 |
| 64. Slovakia | 0.2820 | 0.2860 | 0.2855 | 0.3190 | 0.3237 | 0.3229 | 0.2790 | 0.2830 | 0.2826 | 0.2930 | 0.2974 | 0.2968 | 0.3380 | 0.3423 | 0.3416 |
| 65. Slovenia | 0.4970 | 0.4992 | 0.4989 | 0.4270 | 0.4291 | 0.4290 | 0.5200 | 0.5224 | 0.5220 | 0.5000 | 0.5025 | 0.5021 | 0.3060 | 0.3072 | 0.3075 |
| 66. South Africa | 0.0640 | 0.0671 | 0.0671 | 0.0660 | 0.0693 | 0.0693 | 0.0550 | 0.0577 | 0.0577 | 0.0500 | 0.0525 | 0.0526 | 0.1090 | 0.1141 | 0.1138 |
| 67. Spain | 0.1650 | 0.1755 | 0.1742 | 0.1520 | 0.1622 | 0.1610 | 0.1600 | 0.1703 | 0.1690 | 0.1510 | 0.1609 | 0.1598 | 0.1750 | 0.1864 | 0.1848 |
| 68. Sri Lanka | 0.0260 | 0.0265 | 0.0266 | 0.0150 | 0.0153 | 0.0153 | 0.0180 | 0.0184 | 0.0184 | 0.0090 | 0.0092 | 0.0092 | 0.0520 | 0.0529 | 0.0530 |
| 69. Sweden | 0.4550 | 0.4670 | 0.4623 | 0.4500 | 0.4628 | 0.4578 | 0.4590 | 0.4712 | 0.4664 | 0.4600 | 0.4727 | 0.4676 | 0.4300 | 0.4417 | 0.4374 |


| Continue |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 70. Switzerland | 0.3190 | 0.3258 | 0.3244 | 0.3710 | 0.3796 | 0.3771 | 0.3150 | 0.3218 | 0.3204 | 0.3380 | 0.3456 | 0.3438 | 0.3860 | 0.3946 | 0.3920 |
| 71. Thailand | 0.0540 | 0.0569 | 0.0569 | 0.0420 | 0.0443 | 0.0443 | 0.0500 | 0.0527 | 0.0527 | 0.0420 | 0.0443 | 0.0444 | 0.0600 | 0.0631 | 0.0631 |
| 72. Trinidad-Tobago | 0.2480 | 0.2486 | 0.2490 | 0.2670 | 0.2677 | 0.2681 | 0.2140 | 0.2145 | 0.2150 | 0.2050 | 0.2056 | 0.2060 | 0.4480 | 0.4487 | 0.4492 |
| 73. Turkey | 0.0960 | 0.1020 | 0.1017 | 0.0830 | 0.0883 | 0.0881 | 0.0990 | 0.1053 | 0.1049 | 0.0950 | 0.1011 | 0.1008 | 0.0680 | 0.0721 | 0.0720 |
| 74. Ukraine | 0.5090 | 0.5271 | 0.5186 | 0.5310 | 0.5505 | 0.5408 | 0.4750 | 0.4923 | 0.4847 | 0.4650 | 0.4824 | 0.4750 | 0.6970 | 0.7202 | 0.7046 |
| 75. United States | 1.0000 | 1.1239 | 1.0000 | 1.0000 | 1.1193 | 1.0000 | 1.0000 | 1.1253 | 1.0000 | 1.0000 | 1.1240 | 1.0000 | 1.0000 | 1.1134 | 1.0000 |
| 76. Uruguay | 0.0700 | 0.0706 | 0.0707 | 0.1010 | 0.1019 | 0.1021 | 0.0700 | 0.0706 | 0.0708 | 0.0840 | 0.0848 | 0.0849 | 0.1010 | 0.1018 | 0.1020 |
| 77. Uzbekistan | 0.2450 | 0.2481 | 0.2479 | 0.2300 | 0.2329 | 0.2328 | 0.2310 | 0.2340 | 0.2339 | 0.2150 | 0.2178 | 0.2178 | 0.2860 | 0.2894 | 0.2891 |
| 78. Vietnam | 0.0230 | 0.0237 | 0.0238 | 0.0340 | 0.0351 | 0.0352 | 0.0230 | 0.0237 | 0.0238 | 0.0290 | 0.0299 | 0.0300 | 0.0340 | 0.0350 | 0.0350 |
| 79. Yugoslavia | 0.2040 | 0.2066 | 0.2066 | 0.1990 | 0.2016 | 0.2016 | 0.2010 | 0.2036 | 0.2036 | 0.1970 | 0.1997 | 0.1997 | 0.2050 | 0.2076 | 0.2076 |
| Max | 1.0000 | 1.1239 | 1.0000 | 1.0000 | 1.1193 | 1.0000 | 1.0000 | 1.1253 | 1.0000 | 1.0000 | 1.1240 | 1.0000 | 1.0000 | 1.1134 | 1.0000 |
| Min | 0.0060 | 0.0066 | 0.0066 | 0.0030 | 0.0033 | 0.0033 | 0.0040 | 0.0044 | 0.0044 | 0.0020 | 0.0022 | 0.0022 | 0.0110 | 0.0120 | 0.0121 |
| Average | 0.3850 | 0.3970 | 0.3893 | 0.3621 | 0.3741 | 0.3664 | 0.3723 | 0.3843 | 0.3765 | 0.3538 | 0.3659 | 0.3581 | 0.4074 | 0.4192 | 0.4117 |
| Standard Deviation | 0.3249 | 0.3364 | 0.3251 | 0.3210 | 0.3331 | 0.3214 | 0.3267 | 0.3387 | 0.3270 | 0.3293 | 0.3419 | 0.3299 | 0.3216 | 0.3318 | 0.3212 |

Table 6: Counterintuitive Results Implied by (5), Sydney 2000 Olympic Games

| Cases | Medal <br> Index | Medal <br> Index | Medal <br> Index | Medal <br> Index | Medal <br> Index |
| :---: | :---: | :---: | :---: | :---: | :---: |
| $\breve{\theta}^{k}<\theta^{k}$ for $\theta^{k}<1$ and $y^{k}<\theta^{k}$ | A | B | C | D | E |
| $\breve{\theta}^{k}>1$ for $\theta^{k}=1$ and $y^{k}>1$ | 3 | 4 | 3 | 3 | 4 |
| $\breve{\theta}^{k}>1$ for $\theta^{k}<1$ and $y^{k}>1$ | 5 | 5 | 5 | 6 | 7 |
| $\breve{\theta}^{k}<1$ even though $\theta^{k}=1$ because $y^{k}<1$ | 4 | 3 | 4 | 4 | 3 |
| $\breve{\theta}^{k}=1$ even though $\theta^{k}<1$ because $y^{k}>1$ | - | - | - | - | - |
| Percentage $(\%)$ of the total | - | 1 | - | - | - |

Table 7: Model's Variables and Estimated Efficiency Scores, 2009 NFL

| Team | Wins | Yards | Third-Down | Penalty | $\theta^{k}$ | $\breve{\theta}^{k}$ | $\hat{\theta}^{k}$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1. Arizona Cardinals | 10.000 | 1.048 | 1.027 | 0.948 | 0.950 | 0.987 | 0.952 |
| 2. Atlanta Falcons | 9.000 | 0.928 | 0.929 | 1.343 | 1.000 | 1.032 | 1.000 |
| 3. Baltimore Ravens | 9.000 | 1.142 | 1.133 | 0.677 | 1.000 | 1.032 | 1.000 |
| 4. Buffalo Bills | 6.000 | 0.959 | 0.637 | 1.075 | 1.000 | 1.020 | 1.000 |
| 5. Carolina Panthers | 8.000 | 1.000 | 1.049 | 0.905 | 0.795 | 0.824 | 0.801 |
| 6. Chicago Bears | 7.000 | 0.977 | 0.907 | 0.890 | 0.853 | 0.878 | 0.857 |
| 7. Cincinnati Bengals | 10.000 | 0.994 | 1.053 | 0.889 | 1.000 | 1.036 | 1.000 |
| 8. Cleveland Browns | 5.000 | 0.738 | 0.838 | 1.198 | 1.000 | 1.016 | 1.000 |
| 9. Dallas Cowboys | 11.000 | 1.214 | 1.160 | 0.861 | 0.911 | 0.952 | 0.915 |
| 10. Denver Broncos | 8.000 | 1.058 | 0.976 | 0.908 | 0.828 | 0.857 | 0.833 |
| 11. Detroit Lions | 2.000 | 0.757 | 0.887 | 1.229 | 0.370 | 0.376 | 0.375 |
| 12. Green Bay Packers | 11.000 | 1.213 | 1.306 | 0.865 | 0.901 | 0.942 | 0.905 |
| 13. Houston Texans | 9.000 | 1.101 | 1.023 | 0.874 | 0.889 | 0.922 | 0.893 |
| 14. Indianapolis Colts | 14.000 | 1.184 | 1.093 | 1.628 | 1.000 | 1.053 | 1.000 |
| 15. Jacksonville Jaguars | 7.000 | 0.929 | 1.003 | 0.919 | 0.815 | 0.840 | 0.820 |
| 16. Kansas City Chiefs | 4.000 | 0.814 | 0.717 | 1.231 | 1.000 | 1.012 | 1.000 |
| 17. Miami Dolphins | 7.000 | 0.860 | 1.406 | 0.920 | 1.000 | 1.024 | 1.000 |
| 18. Minnesota Vikings | 12.000 | 1.108 | 1.300 | 1.192 | 0.957 | 1.002 | 0.959 |
| 19. New England Patriots | 10.000 | 1.085 | 1.177 | 1.050 | 0.836 | 0.873 | 0.842 |
| 20. New Orleans Saints | 13.000 | 1.142 | 1.177 | 0.911 | 1.000 | 1.049 | 1.000 |
| 21. New York Giants | 8.000 | 1.056 | 1.108 | 0.845 | 0.755 | 0.784 | 0.763 |
| 22. New York Jets | 9.000 | 1.177 | 1.177 | 1.004 | 0.685 | 0.719 | 0.696 |
| 23. Oakland Raiders | 5.000 | 0.797 | 0.830 | 0.743 | 1.000 | 1.016 | 1.000 |
| 24. Philadelphia Eagles | 11.000 | 1.185 | 1.097 | 0.830 | 1.000 | 1.041 | 1.000 |
| 25. Pittsburgh Steelers | 9.000 | 1.160 | 0.932 | 1.174 | 0.866 | 0.899 | 0.871 |
| 26. San Diego Chargers | 13.000 | 1.123 | 1.099 | 1.395 | 1.000 | 1.049 | 1.000 |
| 27. San Francisco 49ers | 8.000 | 0.996 | 0.813 | 1.278 | 0.943 | 0.971 | 0.945 |
| 28. Seattle Seahawks | 5.000 | 0.871 | 0.854 | 1.156 | 0.708 | 0.725 | 0.714 |
| 29. St. Louis Rams | 1.000 | 0.763 | 0.742 | 0.735 | 1.000 | 1.000 | 1.000 |
| 30. Tampa Bay Buccaneers | 3.000 | 0.850 | 0.810 | 1.050 | 0.491 | 0.501 | 0.497 |
| 31. Tennessee Titans | 8.000 | 1.008 | 1.019 | 0.882 | 0.818 | 0.847 | 0.824 |
| 32. Washington Redskins | 4.000 | 1.008 | 1.002 | 1.124 | 0.393 | 0.407 | 0.402 |
| Max | 14.000 | 1.214 | 1.406 | 1.628 | 1.000 | 1.053 | 1.000 |
| Min | 1.000 | 0.738 | 0.637 | 0.677 | 0.370 | 0.376 | 0.375 |
| Average | 8.000 | 1.008 | 1.009 | 1.023 | 0.868 | 0.896 | 0.871 |
| Standard Deviation | 3.223 | 0.144 | 0.178 | 0.214 | 0.175 | 0.181 | 0.173 |

Source: The data in the first six columns were taken from Collier et al. (2011).
The data in the last two columns come from authors' calculations.

Table 8: Counterintuitive Results Implied by (5), 2009 NFL

| Cases | Frequency |
| :--- | :---: |
| $\breve{\theta}^{k}<\theta^{k}$ for $\theta^{k}<1$ and $y^{k}<\theta^{k}$ | - |
| $\breve{\theta}^{k}>1$ for $\theta^{k}=1$ and $y^{k}>1$ | 12 |
| $\breve{\theta}^{k}>1$ for $\theta^{k}<1$ and $y^{k}>1$ | 1 |
| $\breve{\theta}^{k}<1$ even though $\theta^{k}=1$ because $y^{k}<1$ | - |
| $\breve{\theta}^{k}=1$ even though $\theta^{k}<1$ because $y^{k}>1$ | - |
| Percentage (\%) of the total | 40.5 |


[^0]:    ${ }^{1}$ Lins et al. (2003) estimated the ZSG-DEA efficiency scores of countries in the Olympic Games by using as a single output the a priori fixed number of their total (gold, silver and bronze) medals won.

[^1]:    ${ }^{2}$ Gomes and Lins (2008) coined the names TAT and BCET since these are respectively referred to as Theorem and Corollary in Lins et al. (2003).

[^2]:    ${ }^{3}$ For this reason, Bi et al. (2014) wrote the ZSG-DEA model by excluding the DMU under evaluation from the reference set.

[^3]:    ${ }^{4}$ If $\hat{h}^{k}>1$, then $\lambda_{k}^{k}=0$ and thus (3) is the same as the model in Bi et al.'s (2014) equation (3).

[^4]:    ${ }^{5}$ The left-hand side term in (4) is equal to $x^{k} c^{\prime}$ in Figure 1, the first right-hand side term in (4) is equal to $x^{k} c$, and thus $R C^{k}=\frac{x^{k} c^{\prime}}{x^{k} c}$ corresponds to the vertical distance between $\mathrm{T}_{\mathrm{DEA}}$ and $\mathrm{T}_{\text {ZSG-DEA }}$ at $x^{k}$.

[^5]:    ${ }^{6}$ In terms of Figure 1, this means that DMUs $a$ and $b$ are on both the conventional DEA and the ZSGDEA frontiers while DMU $k$ is inefficient with respect to both frontiers.

