# Breaching Euclidean Distance-Preserving Data Perturbation Using Few Known Inputs 

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#### Abstract

We examine Euclidean distance-preserving data perturbation as a tool for privacy-preserving data mining. Such perturbations allow many important data mining algorithms (e.g. hierarchical and k -means clustering), with only minor modification, to be applied to the perturbed data and produce exactly the same results as if applied to the original data. However, the issue of how well the privacy of the original data is preserved needs careful study. We engage in this study by assuming the role of an attacker armed with a small set of known original data tuples (inputs). Little work has been done examining this kind of attack when the number of known original tuples is less than the number of data dimensions. We focus on this important case, develop and rigorously analyze an attack that utilizes any number of known original tuples. The approach allows the attacker to estimate the original data tuple associated with each perturbed tuple and calculate the probability that the estimation results in a privacy breach. On a real 16-dimensional dataset, we show that the attacker, with 4 known original tuples, can estimate an original unknown tuple with less than $7 \%$ error with probability exceeding 0.8 .


Keywords: Euclidean distance, privacy, data mining, data perturbation

## 1. Introduction

Owners of sensitive information face a dilemma in many situations. On the one hand, making this data available for statistical analysis can violate the privacy of the individuals represented in

[^0]the data or reveal sensitive information about the data owner. On the other hand, making the data available can lead to discoveries that provide societal benefits. For example, mining health-care data for security/fraud issues may require analyzing clinical records and pharmacy transaction data of many individuals over a certain area. While the release of such data may violate privacy laws, mining it can improve the overall quality of the health-care system. Privacy-Preserving Data Mining (PPDM) strives to provide a solution to this dilemma. It aims to allow useful data patterns to be extracted without compromising privacy.

Data perturbation represents one common approach in PPDM. Here, the original private dataset $X$ is perturbed and the resulting dataset $Y$ is released for analysis. Perturbation approaches typically face a "privacy/accuracy" trade-off. On the one hand, perturbation must not allow the original data records to be adequately recovered. On the other hand, it must allow "patterns" that hold in the original data to be recovered. In many cases, increased privacy comes at the cost of reduced accuracy and vice versa. For example, Agrawal and Srikant [1] proposed adding randomly generated i.i.d. noise to the dataset. They showed how the distribution from which the original data arose can be estimated using only the perturbed data and the distribution of the noise. However, Kargupta et al. [2] and Huang et al. [3] pointed out how, in many cases, the noise can be filtered off leaving a reasonably good estimation of the original data (further investigated by Guo et al. [4]). These results point to the fact that unless the variance of the additive noise is sufficiently large, original data records can be recovered unacceptably well. However, this increase in variance reduces the accuracy with which the original data distribution can be estimated. This privacy/accuracy trade-off is not limited to additive noise; some other data transformation techniques suffer from a similar problem, e.g. k-anonymity [5].

Recently, Euclidean distance-preserving data perturbation for the census model ${ }^{1}$ has gained attention $([7,8,9,10,11,12,13,14])$ because it mitigates the privacy/accuracy trade-off by guaranteeing perfect accuracy. The census model using Euclidean distance-preserving data perturbation can be illustrated as follows. An organization has a private, real-valued dataset $X$ (represented as a matrix where each column is a data record) and wishes to make it publicly available for data analysis while keeping the individual records (columns) private. To accomplish this, $Y=T(X)$ is released to the public where $T($.) is a function, known only to the data owner that preserves Euclidean distances between columns. With this nice property, many useful data mining algorithms, with only minor modification, can be applied to $Y$ and produce exactly the same patterns that would be extracted if the algorithm was applied directly to $X$. For example, assume single-link, agglomerative hierarchical clustering (using Euclidean distance) is applied directly to $Y$ [15]. The cluster memberships in the resulting dendrogram will be identical to those in the dendrogram produced if the same algorithm is applied to $X$.

However, the issue of how well the private data is hidden after Euclidean distance-preserving data perturbation needs careful study. Without any prior knowledge, the attacker can do very little (if anything) to accurately recover the private data. However, no prior knowledge seems an unreasonable assumption in many situations. Consideration of prior knowledge-based attack techniques against Euclidean distance-preserving transformations is an important avenue of study. In this paper, we engage in this study by considering known input prior knowledge wherein the attacker knows a small set of original data tuples (inputs), but does not know their associated perturbed data tuples. As pointed out in [13, 14], this knowledge could be obtained through insider information. For example, consider a dataset where each record corresponds to

[^1]information about an individual (e.g. medical data, census data). It is reasonable to assume that the individuals know (1) that a record for themselves appears in the dataset, and (2) the attributes of the dataset. As such, each individual knows one record in the original dataset. A small group of malicious individuals could then combine their insider information to produce a larger set of known original data tuples.

Summary of our contributions: The goal of the attacker is to use the perturbed data tuples and known original data tuples to produce good estimates of unknown original data tuples along with links to their perturbed counterparts. To achieve this, we develop an attack technique called the known input attack which proceeds in three steps.

1. The attacker links as many of the known original data tuples (columns in $X$ ) to their corresponding perturbed counterparts (columns in $Y$ ).
2. For each unlinked perturbed data tuple, the attacker computes the breach probability of the associated unknown original data tuple. This is the probability that the following stochastic procedure will result in an accurate enough estimate of the associated unknown original data tuple to be considered a privacy breach (the probability calculation is done by applying a closed-form expression we derive later).
(a) A Euclidean distance-preserving transformation is uniformly chosen from the space of such transformations that satisfy the original-perturbed (input-output) constraints from step 1.
(b) The inverse of the chosen transformation is used to estimate original data tuples from their perturbed counterparts.
3. The attacker chooses the perturbed data tuples which are most vulnerable to breach based their probabilities from step 2 , e.g. chooses the one with the maximum probability or chooses all whose probability exceeds a threshold, and generates estimates of their associated known original data tuples.

When the number of linked, linearly independent known original data tuples exceeds the number of data dimensions, the privacy breach probability, for all unknown original data tuples, equals one as the estimates are guaranteed to be error-free. However, to our knowledge, little work has been done for the case where the number of known original data tuples is less than the number of data dimensions. This is an important case, since obtaining original data tuples is likely difficult. The attacker ought to be able to utilize however as many as she can get. Our results demonstrate how the attacker can do this and with increasing probability of success with respect to the number original data tuples obtained. Experiments on real and synthetic data show that even with the number of known original data tuples significantly smaller than the number of data dimensions, privacy can be breached with high probability. For example, on a real 16dimensional dataset, we show that the attacker can use 4 known original data tuples to estimate an unknown original tuple with less than $7 \%$ error with probability exceeding 0.8.

Paper organization: Section 2 describes related work in data perturbation for privacypreserving data analysis. Section 3 discusses some background material - the definition of $T$, a Euclidean distance-preserving data perturbation, and the definition of a privacy breach. Section 4 describes the main contribution of the paper - the known input attack outlined above. Section 5 discusses the results of experiments on real and synthetic data to evaluate the behavior of the attack. Section 6 provides a brief summary of the paper and a pointer to an idea for future work. Proofs and some detailed derivations are included in an appendix.

## 2. Related Work

In this section, we give a brief overview of a wide variety of data-perturbation techniques. We first introduce methods that do not preserve Euclidean distance between data tuples. Then we focus on research most relevant to this paper, a majority of which aim to preserve Euclidean distance by projecting private data to a new space.

### 2.1. General Data Perturbation and Transformation Methods

Additive perturbation: Adding i.i.d. white noise to protect data privacy is one common approach for statistical disclosure control [6]. The perturbed data allows the retrieval of aggregate statistics of the original data (e.g. sample mean and variance) without disclosing values of individual records. Moreover, additive white noise perturbation has received attention in the data mining literature [1, 2, 3, 4]. Clearly, additive noise does not preserve Euclidean distance and, therefore, is fundamentally different than the data perturbation we consider. An interesting example along these lines is given by Mukherjee et al. [16]. They considered additive noise to the most dominate principal components of the dataset along with a modification of k-nearestneighbor classification [17] on the perturbed data to improve accuracy. Moreover, they nicely extend to additive noise the $\rho_{1}$-to- $\rho_{2}$ privacy breach measure originally introduced for categorical data in [18]. Another example is Liu et al. [19]. They argued that the level of additive noise ought to be flexible per record. They developed a modified addative noise approach allowing the level of noise to be varied per record based on data owner preference.
Multiplicative perturbation: Two traditional multiplicative data perturbation schemes were studied in the statistics community [20]. One scheme multiplies each data element by a random number that has a truncated Gaussian distribution with mean one and small variance. The other takes a logarithmic transformation of the data first, adds multivariate Gaussian noise, then takes the exponential function $\exp ($.$) of the noise-added data. These perturbations allow summary$ statistics (e.g., mean, variance) of the attributes to be estimated, but do not preserve Euclidean distances among records.

To assess the security of traditional multiplicative perturbation together with additive perturbation, Trottini et al. [21] proposed a Bayesian intruder model that considers both prior and posterior knowledge of the data. Their overall strategy of attacking the privacy of perturbed data using prior knowledge is the same as ours. However, they particularly focused on linkage privacy breaches, where an intruder tries to identify the identity (of a person) linked to a specific record; while we are primarily interested in data record recovery. Moreover, they did not consider Euclidean distance-preserving perturbation as we do.
k-anonymization: Samarati and Sweeney [5, 22] originally developed the $k$-anonymity model to transform person-specific data. Their work shows that an attacker can link a subset of data attributes (called quasi-identifiers) with third-party information to uniquely identify a person even when his personally identifiable information is not present in the original data. To mitigate the risk, the authors proposed the suppression or generalization of values of these quasi-identifiers so that any records in the database, when projected onto the quasi-identifiers, cannot be distinguished from at least $k-l$ others. This model has drawn much of attention because of its simple privacy definition. Since its initial appearance, a variety of extensions have been developed to anonymize transactional data [23], sequential data [24], and mobility data [25]. We refer interested readers to the survey book [26] for more details. It should be noted that none of these approaches consider Euclidean distance-preserving perturbation as we do.

Data micro-aggregation: Two multivariate micro-aggregation approaches have been proposed by researchers in the data mining area. The technique presented by Aggarwal and Yu [27] partitions the original data into multiple groups of predefined size. For each group, a certain level of statistical information (e.g., mean and covariance) is maintained. This statistical information is used to create anonymized data that has similar statistical characteristics to the original dataset. Li et al. [28] proposed a kd-tree based perturbation method, which recursively partitions a dataset into subsets which are progressively more homogeneous after each partition. The private data in each subset is then perturbed using the subset average. The relationships between attributes are argued to be preserved reasonably well. However, neither of these two approaches preserves Euclidean distance between the original data tuples.
Data swapping and shuffling: Data swapping transforms a database by exchanging values of sensitive attributes among individual records. Records are exchanged in such a way that the lower-order frequency counts or marginals are maintained. A variety of refinements and applications of data swapping have been addressed since its initial appearance. We refer readers to [29] for a thorough treatment. Data shuffling [30] is similar to swapping, but is argued to improve on many of the shortcomings of swapping for numeric data. However, neither swapping or shuffling preserves Euclidean distance, which is the focus of this paper.
Other techniques: Evfimievski et al. [18], Rizvi and Haritza [31] considered the use of categorical data perturbation in the context of association rule mining. Their algorithms delete real items and add bogus items to the original records. Association rules present in the original data can be estimated from the perturbed data. Along a related line, Verykios et al. [32] considered perturbation techniques which allow the discovery of some association rules while hiding others considered to be sensitive. We refer interested readers to Chapter 11 of the survey book [26] for a nice overview of association rule hiding methods.

Oliveira and Zaiane [9] consider the application of a rotation, additive noise, and multiplicative noise, separately to each original data attribute. As such, their transformation is not guaranteed to preserve Euclidean distance between data tuples. However, Oliveira and Zaiane argue, through experiments, that their overall data perturbation technique preserves the accuracy of two clustering algorithms.

Similar to [9], Ting et al. [33] considered perturbation of the data attributes by an orthogonal transformation. More precisely, Ting et al. considered left-multiplication of the original data matrix by a randomly generated orthogonal matrix. However, they assume the original data tuples are rows rather than columns, as we do. As a result, Euclidean distance between original data tuples is not preserved, but, sample mean and covariance is. If the original data arose as independent samples from multi-variate Gaussian distribution, then the perturbed data allows inferences to be drawn about this underlying distribution just as well as the original data. For all but small or very high-dimensional datasets, their approach is more resistant to prior knowledge attacks than Euclidean distance-preserving perturbations. Their perturbation matrix is $m \times m$ ( $m$ is the number of original data tuples), much bigger than Euclidean distance-preserving perturbation matrices, $n \times n$ ( $n$ is the number of data dimensions).
A survey: Fung et al. [34] provided a detailed survey of work related to this paper (using the descriptive term "privacy-preserving data publishing"). They discussed a wide range of data perturbation and transformation techniques, as well as, approaches to breach privacy. They also discussed scenarios other than the census model, e.g. multiple release data publishing and statistical database querying.

### 2.2. Euclidean Distance-Preserving Data Perturbation

In this part, we describe research most related to this paper. The majority of the work focuses on Euclidean distance-preserving data perturbation.

Chen and Liu [7] observe that some classifiers are invariant with respect to Euclidean distance between the training tuples. The authors quantify the privacy offered by a Euclidean distance preserving perturbation in terms of the empirical covariance matrix with respect to the difference between the original and perturbed data attributes. The authors' privacy quantification does not take into account prior knowledge, hence, the attack based on prior knowledge presented in our paper applies directly to the Euclidean distance preserving data perturbation method of Chen and Liu. An important issue not discussed by Chen and Liu is how the classifier learned from perturbed data will be used to classify new tuples. Perturbing the new tuples and applying the classifier would produce the same result as if a classifier built from the unperturbed training data was applied to the unperturbed new tuples. But, the process of perturbing the new tuples and applying the classifier need be done with great care to not leak information that could be used to recover the original training tuples.

Oliveira and Zaiane [8] observe that some clustering algorithms are invariant with respect to Euclidean distance between data tuples. The authors quantify privacy using an approach related to that in Chen and Liu. Like Chen and Liu, Oliveira and Zaiane do not consider prior knowledge, hence, the attack based on it presented in our paper applies directly to the Euclidean distance preserving data perturbation method of Oliveira and Zaiane.

Liu et al. [10] developed two types of attacks to breach the privacy of distance-preserving data perturbation.

1. Liu developed the known-sample attack which assumes that the attacker has a moderatesized collection of independent samples chosen i.i.d. from the same distribution as the private data. By mapping the principal components of the perturbed data to the principle components of the original data (estimated from the sample), the attacker can reconstruct the perturbation matrix and consequently recover the private data. The prior knowledge assumption made by this attack is different than the assumption in our manuscript of a very small set of known original tuples. For example, the known sample prior knowledge of Liu requires the original dataset and known sample be drawn i.i.d. (from the same distribution), while the assumption in our manuscript requires no i.i.d. or any other distribution assumptions. If, in our manuscript, we make the additional assumption that the original data is drawn i.i.d., then the known sample attack of Liu can, in theory, be applied. But, the attack's accuracy will be very low as the attack requires a much larger sample than the size of the known tuples we are considering.
2. Liu developed the known input-output attack which assumes that the attacker knows a very small subset of the original (private) data tuples and their correspondences to perturbed tuples (i.e. for each known original tuple, the attacker is assumed to know which is its corresponding perturbed tuple). Their attack technique is the same as a part of our attack - choose an orthogonal matrix randomly from the set of those that satisfy the input-output constraints. Then use a closed-form expression for the breach probability for each private tuple to choose the best one to re-estimate. However, we significantly weaken and make more realistic (providing an explicit scenario) the prior knowledge assumption. We assume only that the attacker knows a very small subset of original (private) data tuples, but does not know their correspondences to perturbed tuples. We extend the attack algorithm of Liu to first infer the correspondences between the known original tuples and the perturbed
tuples. Also, we provide a complete and rigorous mathematical analysis of the attack (Liu did not do this). We also correct a mistake in $\rho\left(x_{j}, \epsilon\right)$, the probability closed-form expression. Finally, we provide experimental results (run-time and accuracy) for the attack (Liu did not do this).

Chen et al. [12] also discussed a known input attack technique. Unlike ours, they considered a combination of distance-preserving data perturbation followed by additive noise. They also assumed a stronger form of known input prior knowledge: the attacker knows a subset of private data records and knows to which perturbed tuples they correspond. Finally, they assume that the number of linearly independent known input data records is no smaller than the number of data dimensions. They pointed out that linear regression can be used to re-estimate private data tuples.

Mukherjee et al. [11] considered the use of discrete Fourier transformation (DFT) and discrete cosine transformation (DCT) to perturb the data. Only the high energy DFT/DCT coefficients are used, and the transformed data in the new domain approximately preserves Euclidean distance. The DFT/DCT coefficients were further permuted to enhance the privacy protection level. Note that DFT and DCT are (complex) orthogonal transforms. Hence, their perturbation technique can be expressed as left multiplication by a (complex) orthogonal matrix (corresponding to the DFT/DCT followed by a perturbation of the resulting coefficients), then a left multiplication by an identity matrix with some zeros on the diagonal (corresponding to dropping all but the high-energy coefficients). They did not consider attacks based on prior knowledge. For future work, it would be interesting to do so.

Turgay et al. [13] extended some of the results in [10]. They assume that the similarity matrix of the original data is made public rather than, $Y$, the perturbed data itself. They describe how an attacker, given at least $n+1$ linearly independent original data tuples and their corresponding entries in the similarity matrix, can recover the private data ( $n$ is the number of data dimensions). Like Chen et al., this differs from our known input attack in two main ways: (i) we do not require prior knowledge beyond the known input tuples; (ii) our attack analysis smoothly encompasses the case where the number of linearly independent known input tuples is greater than $n$ as well as less.

Wong et al. [14] considered data perturbation as a solution to privacy problems introduced by data outsourcing wherein an un-trusted party holds the perturbed data and computes k-nearestneighbor queries against it on behalf of other parties. Among other things, they examined the vulnerabilities of the perturbed data against an attacker armed with known input prior knowledge (their "level 2" prior knowledge). Independently of us, they briefly discussed a basic idea for linking the known inputs to their perturbed counterparts that is similar to our linking technique (although they provide only a cursory description omitting many details). $\sqrt{2}$ They point out how a distance-preserving data perturbation can be undone if the number of linearly independent known inputs that can be linked to perturbed tuples exceeds the number of data dimensions. Their work differs from ours in that it says nothing about the case where the number of linearly independent, linked known tuples is less than the number of data dimensions.

Kaplan et al. [35] considered the estimation of private trajectories (vectors of real numbers) given various kinds of prior knowledge like Euclidean distances from the private trajectories to

[^2]a known one. They develop an innovative algorithm that can incorporate a wide variety of types of prior knowledge and produce estimates. The primary differences between Kaplan's and our work are as follows. Our work applies to the more general problem where the attacker has only a collection of known inputs and does not know their perturbed counterparts. We develop a novel technique for linking the known inputs to their perturbed counterparts. Once this is done, Kaplan's algorithm can be applied to estimate unknown private tuples from the known input-output pairs. However, unlike Kaplan's, our approach provides precise estimation error guarantees, namely, the precise value of the estimation error probability. Thus, with our approach, the attacker can know (in probability) how good each of the estimates is, and, for example, pick the best one. Through experiments, we found our approach to be significantly more accurate than Kaplan's. On the other hand, Kaplan's approach has the advantage over ours of being more general in the sense that it can incorporate a larger variety of prior knowledge into the attack. Our approach is tailored to known input-output prior knowledge.

Before we briefly describe another two attacks based on independent component analysis (ICA) [36], it is necessary to give a brief ICA overview.

### 2.2.1. ICA Overview

Given an $n^{\prime}$-variate random vector $\mathcal{V}$, one common ICA model posits that this random vector was generated by a linear combination of independent random variables, i.e., $\mathcal{V}=A \mathcal{S}$ with $\mathcal{S}$ an $n$-variate random vector with independent components. Typically, $\mathcal{S}$ is further assumed to satisfy the following additional assumptions: (i) at most one component is distributed as a Gaussian; (ii) $n^{\prime} \geq n$; and (iii) $A$ has rank $n$ (full rank).

One common scenario in practice: there is a set of unobserved samples (the columns of $n \times q$ matrix $S$ ) that arose from $\mathcal{S}$ which satisfies (i) - (iii) and whose components are independent. But observed is $n^{\prime} \times q$ matrix $V$ whose columns arose as linear combination of the rows of $S$. The columns of $V$ can be thought of as samples that arose from a random vector $\mathcal{V}$ which satisfies the above generative model. There are ICA algorithms whose goal is to recover $S$ and $A$ from $V$ up to a row permutation and constant multiple. This ambiguity is inevitable due to the fact that for any diagonal matrix (with all non-zeros on the diagonal) $D$, and permutation matrix $P$, if $A, S$ is a solution, then so is $(A D P),\left(P^{-1} D^{-1} \mathcal{S}\right)$.

### 2.2.2. ICA Based Attacks

Liu et al. [37] considered matrix multiplicative data perturbation, $Y=M X$, where $M$ is an $n^{\prime} \times n$ matrix with each entry generated independently from the same distribution with mean zero and variance $\sigma^{2}$. They discussed the application of the above ICA approach to estimate $X$ directly from $Y: \mathcal{S}=\mathcal{X}, \mathcal{V}=\mathcal{Y}, S=X, V=Y$, and $A=M$. They argued the approach to be problematic because the ICA generative model imposes assumptions not likely to hold in many practical situations: the components of $\mathcal{X}$ are independent with at most one such being Gaussian distributed. Moreover, they pointed out that the row permutation and constant multiple ambiguity further hampers accurate recovery of $X$. A similar observation is made later by Chen et al. [12].

Guo and Wu [38] considered matrix multiplicative perturbation assuming only that $M$ is an $n \times n$ matrix (orthogonal or otherwise). They assumed the attacker has known input prior knowledge, i.e. she knows, $\widetilde{X}$, a collection of original data columns from $X$. They develop an ICA-based attack technique for estimating the remaining columns in $X$. To avoid the ICA problems described in the previous paragraph, they instead applied ICA separately to $\widetilde{X}$ and $Y$ producing representations $\left(A_{\widetilde{X}}, S_{\widetilde{X}}\right)$ and $\left(A_{Y}, S_{Y}\right)$. They argued that these representations are
related in a natural way allowing $X$ to be estimated. Their approach, however, will be quite inaccurate for extremely small numbers of known inputs. Moreover, their approach does not provide the attacker with any sort of error information and she will thus not know which (if any) of her original data tuple estimates are accurate.

## 3. Euclidean Distance-Preserving Perturbation and Privacy Breaches

This section provides: some common notation used throughout the article, the definition of $T$ a Euclidean distance-preserving data perturbation, the definition of a privacy breach, and a small example illustrating a Euclidean distance-preserving perturbation.

### 3.1. Notation and Conventions

In the rest of this paper, unless otherwise stated, the following notations and conventions are used. "Euclidean distance-preserving" and "distance-preserving" are used interchangeably. All matrices and vectors discussed are assumed to have real entries (unless otherwise stated). All vectors are assumed to be column vectors and $M^{\prime}$ denotes the transpose of any matrix $M$. Given a vector $x,\|x\|$ denotes its Euclidean norm. An $m \times n$ matrix $M$ is said to be orthogonal if $M^{\prime} M=I_{n}$, the $n \times n$ identity matrix ${ }^{3}$ The set of all $n \times n$, orthogonal matrices is denoted by $\mathbb{O}_{n}$.

Given $n \times p$ and $n \times q$ matrices $A$ and $B$, let $[A \mid B]$ denote the $n \times(p+q)$ matrix whose first $p$ columns are $A$ and last $q$ are $B$. Likewise, given $p \times n$ and $q \times n$ matrices $A$ and $B$, let $\left[\begin{array}{c}A \\ B\end{array}\right]$ denote the $(p+q) \times n$ matrix whose first $p$ rows are $A$ and last $q$ are $B$.

The data owner's private dataset is represented as an $n \times m$ matrix $X$, with each column a record and each row an attribute (each record is assumed to be non-zero). The data owner applies a Euclidean distance-preserving perturbation to $X$ to produce an $n \times m$ data matrix $Y$, which is then released to the public or another party for analysis. That $Y$ was produced from $X$ by a Euclidean distance-preserving data perturbation (but not which one) is also make public.

### 3.2. Euclidean Distance-Preserving Perturbation

A function $H: \mathfrak{R}^{n} \rightarrow \mathfrak{R}^{n}$ is Euclidean distance-preserving if for all $x, y \in \mathfrak{R}^{n},\|x-y\|$ $=\|H(x)-H(y)\|$. Here $H$ is also called a rigid motion. It has been shown that any distancepreserving function is equivalent to an orthogonal transformation followed by a translation [39, pg. 128]. In other words, $H$ may be specified by a pair $(M, v) \in \mathbb{O}_{n} \times \mathfrak{R}^{n}$, in that, for all $x \in \mathfrak{R}^{n}$, $H(x)=M x+v$. If $v=0, H$ preserve Euclidean length: $\|x\|=\|H(x)\|$, as such, it moves $x$ along the surface of the hyper-sphere with radius $\|x\|$ and centered at the origin.

Recall that columns of $X$ (denoted $x_{1}, \ldots, x_{m}$ ) refer to private data records. And, columns of $T(X)=Y$ (denoted $\left.y_{1}, \ldots, y_{m}\right)$ refer to perturbed data records. The correspondence between the private and perturbed data records is not assumed known, e.g. the perturbed version of $x_{i}$ is not necessarily $y_{i}$. Instead, the columns of $X$ are transformed using a Euclidean distance-preserving function, then are permuted to produce the columns of the perturbed dataset $Y$. Formally, the perturbed dataset $Y$, is produced as follows. The private data owner chooses $\left(M_{T}, v_{T}\right)$, a secret Euclidean distance-preserving function, and $\pi$, a secret permutation of $\{1, \ldots, m\}$. Then, for $1 \leq i \leq m$, the data owner produces $y_{\pi(i)}=M_{T} x_{i}+v_{T}$.

Euclidean distance between the private data tuples is preserved in the perturbed dataset: for all $1 \leq i, j \leq m,\left\|x_{i}-x_{j}\right\|=\left\|y_{\pi(i)}-y_{\pi(j)}\right\|$. Moreover, if $v_{T}=0$, then length of the private data tuples is also preserved: for all $1 \leq i \leq m,\left\|x_{i}\right\|=\left\|y_{\pi(i)}\right\|$.

[^3]
### 3.3. Privacy Breach

For simplicity, we assume the attacker produces an estimate for a single unknown original data tuple 4 Formally, the attacker will employ a stochastic procedure and produce $1 \leq j \leq m$ and non-zero, $\hat{x} \in \mathfrak{R}^{n}$. Here, $\hat{x}$ is an estimate of $x_{\hat{j}}$ (with $\hat{j}$ denoting $\pi^{-1}(j)$ ), the private original data tuple that was perturbed to produce $y_{j}{ }^{5}$ Given $\epsilon>0$, we define a privacy breach as follows.

Definition 3.1. An $\epsilon$-privacy breach occurs if $\left\|\hat{x}-x_{\dot{j}}\right\| \leq\left\|x_{\hat{j}}\right\| \epsilon$, i.e. if the attacker's estimate is wrong with Euclidean relative error no more than $\epsilon$.

In the next section, we describe and analyze the known input attack. The main focus of analysis concerns, $\rho(\epsilon)$, the probability that an $\epsilon$-privacy breach occurred.

### 3.4. Example

Figure 1 illustrates a small private dataset (left) and the result of applying a simple Euclidean distance-preserving perturbation (a 90-degree clockwise rotation and identity $\pi$ ). In general, Euclidean distance-preserving perturbations can be much more complex than the one illustrated here.


Figure 1: A four record, original private dataset (left) and the result of applying a simple Euclidean distance-preserving perturbation: 90-degree clockwise rotation (the perturbed records permutation, $\pi$, is the identity).

## 4. Known Input Attack

For $1 \leq a \leq m-1$, let $X_{a}$ denote the first $a$ columns of $X$. The attacker is assumed to know $X_{a}$ and her attack proceeds in three steps. For the remainder of the paper we use interchangeably "known inputs" and "known original data tuples".

1. Infer as many as possible of the input-output mappings in $\pi_{a}$ (the restriction of $\pi$ to $\{1, \ldots, a\})$, that is, find as many as possible perturbed counterparts of $X_{a}$ in $Y$.

[^4]2. For each perturbed tuple $y_{j}$ in $Y$ which is not mapped onto by $\pi_{a}$, compute the probability that the following stochastic procedure will result in an $\epsilon$-privacy breach when estimating the original tuple associated with $y_{j}$ (the probability calculation is done using a closedform expression derived later).
(a) Estimate $M_{T}$ by choosing a matrix, $\hat{M}$, uniformly from the space of all orthogonal matrices that map the tuples in $X_{a}$ to their $\pi_{a}$ counterparts in $Y$ (as computed in step 1).
(b) Estimate the original tuple associated with $y_{j}$ as $\hat{M}^{\prime} y_{j}$.
3. Choose the $y_{j}$ with the highest probability from step 2 and produce $\hat{x}=\hat{M}^{\prime} y_{j}$.

The bulk of our work involves the development and analysis of an attack technique in the case where the data perturbation is assumed to be orthogonal (does not involve a fixed translation, $v_{T}=0$ ). The majority of this section is dedicated to developing and analyzing an attack in this case. Then, in Subsection 4.5, we briefly describe how the attack and analysis can be extended to arbitrary Euclidean distance-preserving perturbation $\left(v_{T} \neq 0\right)$.

### 4.1. Inferring $\pi_{a}$

The attacker may not have enough information to infer $\pi_{a}$, so, her goal is to infer $\pi_{I}$ (the restriction of $\pi$ to $I \subseteq\{1, \ldots, a\}$ ), for as large an $I$ as possible. Next, we describe how this goal can be precisely stated as an algorithmic problem that the attacker can address given her available information.

Given $I \subseteq\{1, \ldots, a\}$, an assignment on $I$ is a 1-1 function $\beta: I \rightarrow\{1, \ldots, m\}$. An assignment $\beta$ on $I$ is valid if it satisfies both of the following conditions for all $i, j \in I$, (1) $\left\|x_{i}\right\|=\left\|y_{\beta(i)}\right\|$ and (2) $\left\|x_{i}-x_{j}\right\|=\left\|y_{\beta(i)}-y_{\beta(j)}\right\|$. Importantly, if $\beta$ is not valid, it cannot be a correct linkage between tuples in $X_{a}$ and $Y$, i.e. $\beta \neq \pi_{I}$. As such, there is at least one valid assignment on $I$, namely $\pi_{I}$, but, there may be more. If $\beta$ is the only valid assignment on $I$, then it must equal $\pi_{I}$.

For notational convenience, we say that $I$ is uniquely valid if there is only one valid assignment on $I$. The attacker's goal is to find a maximal uniquely valid $I$, i.e. a uniquely valid $I$ such that there does not exist uniquely valid $J$ with $|J|>|I|$. It can be shown that there exists only one maximal uniquely valid subset of $\{1, \ldots, a\}$. Thus, the attacker's goal is to find the maximal uniquely valid subset of $\{1, \ldots, a\}$ along with its corresponding assignment.

The following straight-forward algorithm will meet the attacker's goal by employing a topdown, level-wise search of the subset space of $\{1, \ldots, a\}$. The inner for-loop uses an implicit linear ordering to enumerate the size $\ell$ subsets without repeats and requiring $O(1)$ space.

```
Algorithm 1 Overall Algorithm For Finding the Maximal Uniquely Valid Subset
    For \(\ell=a, \ldots, 1\), do
        For all \(I \subseteq\{1, \ldots, a\}\) and \(|I|=\ell\), do
            If \(I\) is uniquely valid, then output \(I\) along with its corresponding assignment and terminate the algorithm.
    Otherwise output \(\emptyset\).
```

Example revisited-part 1: consider the dataset and its perturbed version illustrated in Figure 1 and assume that $X_{3}=\left[x_{1} x_{2} x_{3}\right]$ are the known original data tuples $(a=3)$. Algorithm 1 proceeds as follows.

- Check if $I=\{1,2,3\}$ is uniquely valid. Since the distances of $y_{3}$ to $y_{2}$ and $y_{1}$ are the same as those of $y_{4}$ to $y_{2}$ and $y_{1}$, then the assignment $\beta: 1 \mapsto 1,2 \mapsto 2,3 \mapsto 4$ is valid. The identity assignment on $I$ is also valid because, in this example, $\pi$ is the identity permutation. Thus, $I$ has more than one valid assignment (is not uniquely valid).
- Check if $I=\{1,2\}$ is uniquely valid. To see that $I$ is uniquely valid note that any valid assignment, $\beta$, must assign 2 to itself or else $\left\|x_{2}\right\| \neq\left\|y_{\beta(2)}\right\|$. And, it can be checked that $\beta(1) \neq 3,4$ in order to satisfy $1=\left\|x_{1}-x_{2}\right\|=\left\|y_{\beta(1)}-y_{\beta(2)}\right\|$. Therefore, $\beta$ must be the identity assignment on $I$.
- The algorithm terminates and outputs $I=\{1,2\}$ as the maximal uniquely valid subset of $\{1,2,3\}$ with assignment $1 \mapsto 1,2 \mapsto 2$. Note: any ordering on the subsets of size two may be considered. We chose lexicographic order for simplicity.

Now we develop an algorithm that, given $I \subseteq\{1, \ldots, a\}$, determines if $I$ is uniquely valid, and, if so, also computes the corresponding assignment. The idea is to search the space of all assignments on $I$ for valid ones. Once more than one valid assignment is identified, the search is cut-off and the algorithm outputs that $I$ is not uniquely valid. Otherwise, exactly one valid assignment, $\pi_{I}$, will be found. In this case, the algorithm outputs that $I$ is uniquely valid and returns the corresponding assignment. The algorithm performs a depth-first search with each node, $\mathcal{N}_{1}$, in the search tree representing $\emptyset \subseteq I_{1} \subseteq I$ and $\beta_{1}$ a valid assignment on $I_{1}$. The search proceeds by considering all $\hat{I}_{1}=I_{1} \cup\left\{\hat{i}_{1}\right\}$ where $\hat{i} \in\left(I \backslash I_{1}\right)$ and all possible ways of extending $\beta_{1}$ to be a valid assignment, $\hat{\beta_{1}}$ on $\hat{I_{1}}$. In turn, $\hat{I_{1}}$ and $\hat{\beta_{1}}$ represent a node, $\hat{\mathcal{N}_{1}}$, in the search tree immediately below $\mathcal{N}_{1}$. If $\hat{I}_{1}=I$, then a NumValidAssignFound counter is incremented. If the counter exceeds one, then $I$ has more than one valid assignment, and the search is terminated.

To make this search efficient, we employ a simple, but effective, pruning rule to quickly eliminate possible extensions of $\beta_{1}$ that are not valid assignments on $\hat{I_{1}}=I_{1} \cup\left\{\hat{i_{1}}\right\}$. Let $C\left(I_{1}, \beta_{1}, \hat{i}\right)$ denote the set of all extensions of $\beta_{1}, \hat{i} \mapsto j$, which appropriately preserve Euclidean distances; formally put, all $j \in\left(\{1, \ldots, m\} \backslash \beta_{1}\left(I_{1}\right)\right)$ which satisfies both of the following conditions: (1) $\left\|x_{i}\right\|=\left\|y_{j}\right\|$, and (2) for all $i_{1} \in I_{1},\left\|x_{i_{1}}-x_{i}\right\|=\left\|y_{\beta_{1}\left(i_{1}\right)}-y_{j}\right\|$. It can be shown that $j \notin C\left(I_{1}, \beta_{1}, \hat{i}\right)$ does not represent a valid assignment. Therefore, to enumerate all possible, valid, extensions of $\beta_{1}$ on $\hat{I}_{1}$, it suffices to consider those assignments $\hat{\beta_{1}}$ on $\hat{I}_{1}$ which are of the following form: (i) for all $\ell \in I_{1}, \hat{\beta_{1}}(\ell)=\beta_{1}(\ell)$ and (ii) $\hat{\beta_{1}}\left(\hat{i_{1}}\right)=j$ for some $j$ in $C\left(I_{1}, \beta_{1}, \hat{i}\right)$.

Algorithms 2 and 3 describe the precise details of the determination whether $I$ is uniquely valid (namely, the details of the search discussed in the previous two paragraphs).

```
Algorithm 2 Determining Unique Validity Main
Inputs: \(I \subseteq\{1, \ldots, a\}\).
    : Set global variable NumValidAssignFound \(=0\).
    Call Algorithm 3 on inputs \(\emptyset\) and \(\beta_{\emptyset}\) ( \(\beta_{\emptyset}\) denotes the unique valid assignment on \(\emptyset\) ).
    If NumValidAssignFound \(>1\), then return "I IS NOT UNIQUELY VALID". Else, return "I IS UNIQUELY VALID
    WITH ASSIGNMENT" \(\beta_{I}\).
```

Comment: The order by which the elements of $\left(I \backslash I_{1}\right)$ and $C\left(I_{1}, \beta_{1}, \hat{i}\right)$ are chosen in iterating through the for loops in Algorithm 3 does not affect the correctness of the algorithm. However, it may affect efficiency. For simplicity, the loops order the elements in these sets from smallest to largest index number.

Algorithm 1 has worst-case computational complexity $O\left(m^{a}\right)$. While this is no better than a simple brute-force approach, in our experiments, quite reasonable running times are observed because few original data tuples will have the same length and/or few pairs of original data tuples will have the same Euclidean distance.

```
Algorithm 3 Determining Unique Validity Recursive
Inputs: \(I_{1} \subseteq I\) and \(\beta_{1}\) a valid assignment on \(I_{1}\).
    : If \(I_{1}=I\), then
        NumValidAssignFound \(=\) NumValidAssignFound +1
        If NumValidAssignFound \(==1\), then set \(\beta_{I}\) to \(\beta_{1}\).
        End If.
    Else, do
        For \(\hat{i} \in\left(I \backslash I_{1}\right)\) and as long as NumValidAssignFound \(\leq 1\), do
            For \(j \in C\left(I_{1}, \beta_{1}, i\right)\) and as long as NumValidAssignFound \(\leq 1\), do
                Extend \(\beta_{1}\) to \(\hat{\beta_{1}}\) s.t. \(\hat{\beta_{1}}(\hat{i})=j\). Let \(\hat{I_{1}}=I_{1} \cup \hat{i}\).
                Call algorithm 3 on inputs \(\hat{I}_{1}\) and \(\hat{\beta_{1}}\).
```


### 4.2. Known Input-Output Attack

Assume, without loss of generality, that the attacker applies Algorithm 1 and learns $\pi_{q}(0 \leq$ $q \leq a)$, i.e. $\{1, \ldots, q\}$ is the maximal uniquely valid subset of $\{1, \ldots, a\}$. Further, to simplify notation, we may also assume that $\pi_{q}(i)=i{ }^{6}$ Let $Y_{q}$ denote the first $q$ columns of $Y$. As such, the attacker is assumed to know $X_{q}$ and the fact that $Y_{q}=M_{T} X_{q}$ where $M_{T}$ is an unknown orthogonal matrix. Based on this, she will apply an attack, called the known input-output attack, to produce $q<j \leq m$, and $\hat{x}$, which is an estimate of $x_{\hat{j}}$, the private tuple that was perturbed to produce $y_{j}$. The known input-output attack was described in the last two steps in the algorithm at the beginning of Section 4 More formally, the known input-output attack is as follows. Let $\mathbb{M}\left(X_{q}, Y_{q}\right)$ denote the set of all $M \in \mathbb{O}_{n}$ such that $M X_{q}=Y_{q}$.

1. For each $q<j \leq m$, compute the probability that the following stochastic procedure will result in an $\epsilon$-privacy breech when estimating $x_{\hat{j}}$.
(a) Estimate $M_{T}$ by choosing a matrix, $\hat{M}$, uniformly from $\mathbb{M}\left(X_{q}, Y_{q}\right)$.
(b) Estimate $x_{\hat{j}}$ as $\hat{M}^{\prime} y_{j} \nabla^{7}$
2. Choose the $y_{j}$ with the highest probability from step 2 and produce $\hat{x}=\hat{M}^{\prime} y_{j}$.

A key component of the known input-output attack is the computation of $\rho\left(x_{\hat{j}}, \epsilon\right)=\operatorname{Pr}\left(\| \hat{M}^{\prime} y_{j-}\right.$ $x_{\hat{j}}\|\leq\| x_{\hat{j}} \| \epsilon$, the probability that an $\epsilon$-privacy breach will result from the attacker estimating $x_{\hat{j}}$ as $\hat{M}^{\prime} y_{j}$. In Section4.4, we will develop a closed-form expression for $\rho\left(x_{\hat{j}}, \epsilon\right)$. This expression will only involve information known to the attacker; therefore, she can choose $q<j \leq m$ so as to maximize $\rho\left(x_{\hat{j}}, \epsilon\right)$. Another key component of the known input-output algorithm is in choosing $\hat{M}$ uniformly from $\mathbb{M}\left(X_{q}, Y_{q}\right)$. In most cases, $\mathbb{M}\left(X_{q}, Y_{q}\right)$ is uncountable and it is not obvious how to choose $\hat{M}$. We will develop and algorithm for doing so in Section 4.4 Before getting to Section 4.4 we discuss some important linear algebra background.

### 4.3. Linear Algebra background

Let $\operatorname{Col}\left(X_{q}\right)$ denote the column space of $X_{q}$ and $\operatorname{Col}_{\perp}\left(X_{q}\right)$ denote its orthogonal complement, i.e., $\left\{z \in \mathbb{R}^{n}: z^{\prime} w=0, \forall w \in \operatorname{Col}\left(X_{q}\right)\right\}$. Likewise, let $\operatorname{Col}\left(Y_{q}\right)$ denote the column space of $Y_{q}$ and $\operatorname{Col}_{\perp}\left(Y_{q}\right)$ denote its orthogonal compliment. Let $k$ denote the dimension of $\operatorname{Col}\left(X_{q}\right)$. The "Fundamental Theorem of Linear Algebra" [40, pg. 95] implies that the dimension of $\operatorname{Col}_{\perp}\left(X_{q}\right)$

[^5]is $n-k$. Since $Y_{q}=M_{T} X_{q}$ and $M_{T}$ is orthogonal, then it can be shown that $\operatorname{Col}\left(Y_{q}\right)$ has dimension $k$. Thus, $\operatorname{Col}_{\perp}\left(Y_{q}\right)$ has dimension $n-k$.

Let $U_{k}$ and $V_{k}$ denote $n \times k$ matrices whose columns form an orthonormal basis for $\operatorname{Col}\left(X_{q}\right)$ and $\operatorname{Col}\left(Y_{q}\right)$, respectively. It can easily be shown that $\operatorname{Col}\left(M_{T} U_{k}\right)=\operatorname{Col}\left(Y_{q}\right)=\operatorname{Col}\left(V_{k}\right)$. Let $U_{n-k}$ and $V_{n-k}$ denote $n \times(n-k)$ matrices whose columns form an orthonormal basis for $\operatorname{Col}_{\perp}\left(X_{q}\right)$ and $\operatorname{Col}_{\perp}\left(Y_{q}\right)$, respectively. It can easily be shown that $\operatorname{Col}\left(M_{T} U_{n-k}\right)=\operatorname{Col}_{\perp}\left(Y_{q}\right)=\operatorname{Col}\left(V_{n-k}\right)$.

### 4.4. A Closed-Form Expression for $\rho\left(x_{\hat{j}}, \epsilon\right)$

Now we return to the issue of how to choose $\hat{M}$ uniformly from $\mathbb{M}\left(X_{q}, Y_{q}\right)$ and how to compute $\rho\left(x_{\hat{j}}, \epsilon\right)=\operatorname{Pr}\left(\left\|\hat{M}^{\prime} y_{j}-x_{\hat{j}}\right\| \leq\left\|x_{\hat{j}}\right\| \epsilon\right)=\operatorname{Pr}\left(\left\|\hat{M}^{\prime} M_{T} x_{\hat{j}}-x_{\hat{j}}\right\|\right)$.

To choose $\hat{M}$ uniformly from $\mathbb{M}\left(X_{q}, Y_{q}\right)$, the basic idea is to utilize standard algorithms for choosing a matrix $P$ uniformly from $\mathbb{O}_{n-k}$, the set of all $(n-k) \times(n-k)$ orthogonal matrices, then apply an appropriately designed transformation to $P$. The transformation will be an affine, bijection from $\mathbb{O}_{n-k}$ to $\mathbb{M}\left(X_{q}, Y_{q}\right)$ The following technical result, proven in Appendix A provides this transformation 9

Theorem 4.1. Let $L$ be the mapping $P \in \mathbb{O}_{n-k} \mapsto M_{T} U_{k} U_{k}^{\prime}+V_{n-k} P U_{n-k}^{\prime}$. Then, $L$ is an affine bijection from $\mathbb{O}_{n-k}$ to $\mathbb{M}\left(X_{q}, Y_{q}\right)$. And, $L^{-1}$ is the mapping $M \in \mathbb{M}\left(X_{q}, Y_{q}\right) \mapsto V_{n-k}^{\prime} M U_{n-k}$.

```
Algorithm 4 Uniform Choice From \(\mathbb{M}\left(X_{q}, Y_{q}\right)\)
Inputs: \(U_{k}\), an \(n \times k\) matrix whose columns form an orthonormal basis of \(\operatorname{Col}\left(X_{q}\right)\), and \(M_{T} U_{k}\) ( \(M_{T}\) is unknown); \(U_{n-k}\)
    and \(V_{n-k}, n \times(n-k)\) matrices whose columns form an orthonormal basis of \(\operatorname{Col}_{\perp}\left(X_{q}\right)\) and \(\operatorname{Col}_{\perp}\left(Y_{q}\right)\), respectively.
Outputs: \(\hat{M}\) a uniformly chosen matrix from \(\mathbb{M}\left(X_{q}, Y_{q}\right)\).
    Choose \(P\) uniformly from \(\mathbb{O}_{n-k}\) using algorithm [41].
    Set \(\hat{M}=L(P)\), i.e., \(M_{T} U_{k} U_{k}^{\prime}+V_{n-k} P U_{n-k}^{\prime}\).
```

Two comments are in order regarding Algorithm4 First, some special cases are interesting to highlight: when $k=n, \hat{M}$ is chosen as $M_{T}$; when $k=n-1, \hat{M}$ is one of two choices (one of which equals $M_{T}$ ); otherwise, $\hat{M}$ is, in theory, chosen from an uncountable set (containing $M_{T}$ ). Second, it is not obvious how the attacker can compute the inputs to the algorithm, e.g. $M_{T} U_{k}$. This issue will be discussed later when spelling out the details of the Known Input Attack Algorithm, Algorithm 5.

Now we develop a closed-form expression for $\rho\left(x_{\hat{j}}, \epsilon\right)$. The key points are outlined, while a more rigorous justification is provided in Appendix A First of all, from Algorithm $4 \hat{M}=$ $M_{T} U_{k} U_{k}^{\prime}+V_{n-k} P U_{n-k}^{\prime}$ where $P$ is chosen uniformly from $\mathbb{O}_{n-k}$. Therefore,

$$
\begin{aligned}
\rho\left(x_{\hat{j}}, \epsilon\right) & =\operatorname{Pr}\left(\left\|\hat{M}^{\prime} M_{T} x_{\hat{j}}-x_{\hat{j}}\right\| \leq\left\|x_{j}\right\| \epsilon\right) \\
& =\operatorname{Pr}\left(\left\|U_{k} U_{k}^{\prime} x_{\hat{j}}+U_{n-k} P^{\prime} V_{n-k}^{\prime} M_{T} x_{\hat{j}}-x_{\hat{j}}\right\| \leq\left\|x_{j}\right\| \epsilon\right) .
\end{aligned}
$$

Since $\left[\begin{array}{c}U_{k}^{\prime} \\ U_{n-k}^{\prime}\end{array}\right] \in \mathbb{O}_{n}$, then it can left-multiply each term in the left $\|\ldots\|$ of the second probability without changing the equality. As a result, the derivation continues

[^6]\[

$$
\begin{aligned}
\cdots & =\operatorname{Pr}\left(\left\|\left[\begin{array}{c}
U_{k}^{\prime} x_{\hat{i}} \\
0
\end{array}\right]+\left[\begin{array}{c}
0 \\
P^{\prime} V_{n-k}^{\prime} M_{T} x_{\hat{j}}
\end{array}\right]-\left[\begin{array}{c}
U_{k}^{\prime} x_{\hat{j}} \\
U_{n-k}^{\prime} x_{\hat{j}}
\end{array}\right]\right\| \leq\left\|x_{\hat{j}}\right\| \epsilon\right) \\
& =\operatorname{Pr}\left(\left\|P^{\prime} V_{n-k}^{\prime} M_{T} x_{\hat{j}}-U_{n-k}^{\prime} x_{\hat{j}}\right\| \leq\left\|x_{\hat{j}}\right\| \epsilon\right) .
\end{aligned}
$$
\]

Since $\operatorname{Col}\left(M_{T} U_{n-k}\right)=\operatorname{Col}\left(V_{n-k}\right)$, then there exists $(n-k) \times(n-k)$ matrix $B$ such that $M_{T} U_{n-k} B=$ $V_{n-k}$. It follows that (i) $V_{n-k}^{\prime}=B^{\prime} U_{n-k}^{\prime} M_{T}^{\prime}$, (ii) $B=U_{n-k}^{\prime} M_{T}^{\prime} V_{n-k}$. Thus, $B$ is orthogonal ${ }^{10}$ Using (i), the derivation continues

$$
\begin{align*}
\cdots & =\operatorname{Pr}\left(\left\|P^{\prime} B^{\prime}\left(U_{n-k}^{\prime} x_{\hat{j}}\right)-\left(U_{n-k}^{\prime} x_{\hat{j}}\right)\right\| \leq\left\|x_{j}\right\| \epsilon\right)  \tag{1}\\
& =\operatorname{Pr}\left(\left\|P^{\prime}\left(U_{n-k}^{\prime} x_{\hat{j}}\right)-\left(U_{n-k}^{\prime} x_{\hat{j}}\right)\right\| \leq\left\|x_{\hat{j}}\right\| \epsilon\right) \tag{2}
\end{align*}
$$

where the second equality is due to the fact that $B^{\prime} \in \mathbb{O}_{n-k}$, and thus $\left(P^{\prime} B^{\prime}\right)$ can be regarded as having been uniformly chosen from $\mathbb{O}_{n-k}$ just like $P^{\prime}$ (a rigorous proof of the second equality is provided in Appendix A. Putting the whole derivation together,

$$
\begin{equation*}
\rho\left(x_{\hat{j}}, \epsilon\right)=\operatorname{Pr}\left(P \text { uniformly chosen from } \mathbb{O}_{n-k} \text { satisfies }\left\|P^{\prime}\left(U_{n-k}^{\prime} x_{\hat{j}}\right)-\left(U_{n-k}^{\prime} x_{\hat{j}}\right)\right\| \leq\left\|x_{\dot{j}}\right\| \epsilon\right) . \tag{3}
\end{equation*}
$$

Let $S_{n-k}\left(\left\|U_{n-k}^{\prime} x_{\hat{j}}\right\|\right)$ denote the hyper-sphere in $\mathfrak{R}^{n-k}$ with radius $\left\|U_{n-k}^{\prime} x_{j}\right\|$ and centered at the origin. Since $P$ is chosen uniformly from $\mathbb{O}_{n-k}$, then any point on the surface of $S_{n-k}\left(\left\|U_{n-k}^{\prime} x_{j}\right\|\right)$ is equally likely to be $P^{\prime}\left(U_{n-k}^{\prime} x_{\dot{j}}\right)$. Let $S_{n-k}\left(U_{n-k}^{\prime} x_{\hat{j}},\left\|x_{j}\right\| \epsilon\right)$ denote the "hyper-sphere cap" consisting of all points in $S_{n-k}\left(\left\|U_{n-k}^{\prime} x_{j}\right\|\right)$ with distance from $U_{n-k}^{\prime} x_{\hat{j}}$ no greater than $\left\|x_{\hat{j}}\right\| \epsilon$. Therefore, (3) becomes

$$
\begin{align*}
\rho\left(x_{\hat{j}}, \epsilon\right) & =\operatorname{Pr}\left(\text { a uniformly chosen point on } S_{n-k}\left(\left\|U_{n-k}^{\prime} x_{j}\right\|\right) \text { is also in } S_{n-k}\left(U_{n-k}^{\prime} x_{\hat{j}},\left\|x_{\hat{j}}\right\| \epsilon\right)\right) \\
& =\frac{S A\left(S_{n-k}\left(U_{n-k}^{\prime} x_{\hat{j}},\left\|x_{\hat{j}}\right\| \epsilon\right)\right)}{S A\left(S_{n-k}\left(\left\|U_{n-k}^{\prime} x_{\hat{j}}\right\|\right)\right)} \tag{4}
\end{align*}
$$

where $S A\left(\right.$.) denotes the surface area of a subset of a hyper-sphere ${ }^{11}$ Based on equations (4), we prove, in Appendix A, the following closed form expression, for $\rho\left(x_{\hat{j}}, \epsilon\right)$, where, $\Gamma$ (.) denotes the standard gamma function, $a c_{[]-1}(x)$ denotes $\arccos \left(\left[\frac{\left\|x_{j}\right\| \epsilon}{\left\|U_{n-k}^{\prime} x_{j}\right\| \sqrt{2}}\right]^{2}-1\right)$, and $a c_{1-[]}(x)$ denotes $\arccos \left(1-\left[\frac{\left\|x_{j}\right\| \epsilon}{\left\|U_{n-k}^{\prime} x_{j}\right\| \sqrt{2}}\right]^{2}\right)$.

[^7]\[

$$
\begin{align*}
& \text { if } n-k=0 \text {; } \\
& \text { if }\left\|x_{j}\right\| \epsilon \geq\left\|U_{n-k}^{\prime} x_{j}\right\| 2 \text { and } n-k \geq 1 \text {; } \\
& \text { if }\left\|x_{\hat{j}}\right\| \epsilon<\left\|U_{n-k}^{\prime} x_{\hat{j}}\right\| 2 \text { and } n-k=1 \text {; } \\
& \text { if }\left\|U_{n-k}^{\prime} x_{\hat{j}}\right\| \sqrt{2}<\left\|x_{\hat{j}}\right\| \epsilon<\left\|U_{n-k}^{\prime} x_{\hat{j}}\right\| 2 \text { and } n-k=2 \text {; } \\
& \text { if }\left\|U_{n-k}^{\prime} x_{\hat{j}}\right\| \sqrt{2}<\left\|x_{\hat{j}}\right\| \epsilon<\left\|U_{n-k}^{\prime} x_{\hat{j}}\right\| 2 \text { and } n-k \geq 3 \text {; } \\
& \text { if }\left\|x_{\hat{j}}\right\| \epsilon \leq\left\|U_{n-k}^{\prime} x_{\hat{j}}\right\| \sqrt{2} \text { and } n-k=2 \text {; } \\
& \text { if }\left\|x_{\hat{j}}\right\| \epsilon \leq\left\|U_{n-k}^{\prime} x_{\hat{j}}\right\| \sqrt{2} \text { and } n-k \geq 3 \text {. } \tag{5}
\end{align*}
$$
\]

Comment: it can be shown that $\left\|U_{n-k}^{\prime} x_{\hat{j}}\right\|$ is the distance from $x_{\hat{j}}$ to its closest point in $\operatorname{Col}\left(X_{q}\right)$ (the column space of $X_{q}$ ). Thus, the sensitivity of a tuple to breach is dependent upon its length relative to its distance to the column space of $X_{q}$. In particular, if the distance from $x_{\hat{j}}$ to the column space of $X_{q}$ is sufficiently small, less than $\left(\left\|x_{j}\right\| \epsilon\right) / 2$, then the breach probability is one from the second case in equation (5).

Recall that the attacker seeks to use the closed-form expressions for $\rho\left(x_{\hat{j}}, \epsilon\right)$ to decide for which $q<j \leq m$ does $\hat{x}=\hat{M}^{\prime} y_{j}$ produce the best estimation of $x_{\hat{j}}$. This is naturally done by choosing $j$ to maximize $\rho\left(x_{\hat{j}}, \epsilon\right)$. To allow for this, observe that $\left\|x_{\hat{j}}\right\| \epsilon$ and $\left\|U_{n-k}^{\prime} x_{j}\right\|$ equa ${ }^{12}\left\|y_{j}\right\| \epsilon$ and $\left\|V_{n-k}^{\prime} y_{j}\right\|$, respectively, which are known to the attacker. Therefore, (5) can be rewritten as follows, where $a c_{[-1}(y)$ denotes $\arccos \left(\left[\frac{\left\|y_{j}\right\| \epsilon}{\left\|V_{n-k}^{\prime} y_{j j}\right\| \sqrt{2}}\right]^{2}-1\right)$, and $\operatorname{ac}_{1-[]}(y)$ denotes $\arccos \left(1-\left[\frac{\left\|y_{j}\right\| \epsilon}{\left\|V_{n-k}^{\prime} y_{j}\right\| \sqrt{2}}\right]^{2}\right)$.

Now we put together all the parts and provide the pseudo-code of the full known input attack algorithm (Algorithm[5). Before doing so, first note that $U_{k}, U_{n-k}, V_{k}$, and $V_{n-k}$ can be computed from $X_{q}$ and $Y_{q}$ using standard procedures [40]. Second, $M_{T} U_{k}=Y_{q} A$ where $A$ is an $q \times k$ matrix that can be computed ${ }^{13}$ from $U_{k}$ and $X_{q}$. Third, a recursive procedure for computing (6) is described in Appendix A

Comment: The $\epsilon$-privacy breach probability $\rho(\epsilon)$ equals $\max _{q<j \leq m} \rho\left(x_{\hat{j}}, \epsilon\right)$.
Example revisited -part 2: consider the dataset and its perturbed version illustrated in Figure 1 with known original tuples $x_{1}, x_{2}, x_{3}$. Part 1 of this example showed how Algorithm 1 inferred the following mappings to perturbed tuples $x_{1} \mapsto y_{1}$ and $x_{2} \mapsto y_{2}$, hence $X_{q}=\left[x_{1} x_{2}\right]$ and $Y_{q}=\left[y_{1} y_{2}\right]$. Consider perturbed tuple $y_{4}$. The Known Input Attack will compute $\hat{x}$ an estimate of the original tuple $x_{\hat{4}}$ associated with $y_{4}\left(x_{\hat{4}}=x_{4}\right.$ in this case) and $\rho\left(x_{\hat{4}}, \epsilon\right)$, the $\epsilon$-privacy breach probability. Since $x_{1}$ and $x_{2}$ are linearly dependent, $k=1$, so, $n-k=1$ and the second or third

[^8]```
Algorithm 5 Known Input Attack Algorithm
Inputs: \(Y, \epsilon \geq 0\), and \(X_{a}\).
Outputs: \(a<j \leq m\) and \(\hat{x} \in \mathfrak{R}^{n}\) the corresponding estimate of \(x_{\hat{j}}\).
    Compute \(Y_{q}=M_{T} X_{q}\) (where \(1 \leq q \leq a\) ) using Algorithm 1
    Compute \(U_{k}, V_{k}, U_{n-k}, V_{n-k}\), and \(M_{T} U_{k}\) as described earlier.
    For each \(q<j \leq m\) do
            Compute \(\rho\left(x_{\hat{j}}, \epsilon\right)\) using (6) as described in Appendix A
    End For.
    Choose the \(j\) from the previous loop producing the largest \(\rho\left(x_{\hat{j}}, \epsilon\right)\).
    Choose \(\hat{M}\) uniformly from \(\mathbb{M}\left(X_{q}, Y_{q}\right)\) by applying Algorithm 4
    Set \(\hat{x} \leftarrow \hat{M}^{\prime} y_{j}\)
```

cases of Equation6apply. It can be shown that $V_{n-k}^{\prime}=V_{1}^{\prime}=[0,1]$. So, $\left\|y_{4}\right\| \epsilon=2 \epsilon$ and $\left\|V_{n-k}^{\prime} y_{4}\right\| 2$ $=4$. Therefore, if $\epsilon \geq 2$, the second case applies and $\rho\left(x_{\hat{4}}, \epsilon\right)=1$, else, the third case applies and $\rho\left(x_{\hat{4}}, \epsilon\right)=0.5$.

There are only two Euclidean distance preserving transformations fixing the origin that satisfy the input-output constraints $x_{1} \mapsto y_{1}$ and $x_{2} \mapsto y_{2}$ : the 90-degree clockwise rotation (the actual perturbation applied) and the 90 -degree counter-clockwise rotation, these are the elements of $\mathbb{M}\left(X_{2}, Y_{2}\right)$. So $\hat{x}$ is chosen randomly between the inverse of these transformations applied to $y_{4}$ resulting in $\hat{x}=[2,0]^{\prime}$ or $[-2,0]^{\prime}$. If $\epsilon \geq 2$, then either of these choices represent an $\epsilon$-privacy breach so $\rho\left(x_{\hat{4}}, \boldsymbol{\epsilon}\right)=1$. If $\epsilon<2$, then only one of these choices, $[2,0]^{\prime}$, represent a breach so $\rho\left(x_{\hat{4}}, \epsilon\right)=0.5$.

### 4.5. Known Input Attack on General Distance-Preserving Data Perturbation

Previously, we considered the case where the data perturbation is assumed to be orthogonal (does not involve a fixed translation, $v_{T}=0$ ). Now we briefly discuss how the attack technique and its analysis can be extended to arbitrary Euclidean distance-preserving perturbation ( $v_{T} \neq 0$ ). Extending the algorithms for inferring $\pi_{a}$ : Since the length of the private data tuples may not be preserved, then the definition of validity in Section 4.1 must be changed: $\beta$ on $I$ is valid if $\forall i, j \in I,\left\|x_{i}-x_{j}\right\|=\left\|y_{\beta(i)}-y_{\beta(j)}\right\|$. As well, the definition of $C\left(I_{1}, \beta_{1}, \hat{i}\right)$ (given $I_{1} \subseteq I, \beta_{1}$ a valid assignment on $I_{1}$, and $\left.\hat{i} \in\left(I \backslash I_{1}\right)\right)$, must change: the set of all $j \in\left(\{1, \ldots, m\} \backslash \beta_{1}\left(I_{1}\right)\right)$ such that for all $i_{1} \in I_{1},\left\|x_{i_{1}}-x_{i}\right\|=\left\|y_{\beta_{1}\left(i_{1}\right)}-y_{j}\right\|$. With these changes, Algorithms 1, 2, and 3)work correctly as stated.
Extending the known input attack: The basic idea is simple and relies on the fact that the same $v_{T}$ is added to all tuples in the perturbation of $X_{q}$. Fix one tuple, say $x_{1}$ and $y_{1}$, and consider the following differences $x_{1}^{-}=\left(x_{q}-x_{1}\right), \ldots, x_{q-1}^{-}=\left(x_{q}-x_{q-1}\right)$ and $y_{1}^{-}=\left(y_{q}-y_{1}\right), \ldots$, $y_{q-1}^{-}=\left(y_{q}-y_{q-1}\right)$. Let $X_{q-1}^{-}$denote the matrix with columns $x_{1}^{-}, \ldots, x_{q-1}^{-}$and $Y_{q-1}^{-}$denote the matrix with columns $y_{1}^{-}, \ldots, y_{q-1}^{-}$. Observe that $Y_{q-1}^{-}=M_{T} X_{q-1}^{-}$, hence, the attack and its analysis from the orthogonal data perturbation case can be applied. The details are straight-forward and are omitted for brevity. However, a caveat is in order. The attack depends on the choice of the tuple to fix. Therefore, the attacker examines them all and chooses the highest privacy breach probability.

## 5. Experiments and Discussion

The experiments are designed to assess the computational efficiency of the overall known input attack and its effectiveness at breaching privacy. We performed two sets of experiments: (a)
those involving only the known input attack, and (b) those comparing the known input attack with the attack of Kaplan et al. [35]. In both sets of experiments, we used two datasets as the original, private data tuples $X: 1$ ) a 100,000 tuple synthetic dataset generated from a 100 -variate Gaussian distribution ${ }^{[14}$; 2) the Letter Recognition dataset, 20,000 tuples and 16 numeric attributes, from UCI machine learning repository [42] - we removed tuples which were duplicated over the numeric attributes yielding a final dataset of 18,668 tuples. The attacks were implemented in Matlab 7 (R14) and all experiments were carried out on a Thinkpad laptop with 1.83 GHz Intel Core 2 CPU, 1.99GB RAM, and WindowsXP system. We did not compare our attack technique against the ICA-based attack in [38] and the known sample attack in [10] because the extremely small size of the known inputs will render these attacks ineffective. In all Figures, the error bars show one standard deviation above and below the average.

### 5.1. Experiments Only Involving the Known Input Attack

The first experiment fixes $X$ and its perturbed version $Y$, but changes the number of known input tuples, $a$. It proceeds by carrying out ten trials as follows. Select $a$ linearly independent tuples randomly from $X$ (these become the know inputs). Use Algorithm 1 to compute $I$, the maximal uniquely valid assignment. Use steps 2-5 in Algorithm 5 to compute the $\rho(\epsilon)$, the $\epsilon$-privacy breach probability (a closed-form was given immediately above Algorithm 5).

To measure the accuracy of the attack, we report the average of $\rho(\epsilon)$ and $|I|$ over all ten trials. To measure the efficiency, we report the average time taken to compute $I$ (the rest is ignored as the overall attack computation time is dominated by Algorithm (1). In Figures 2and 3 results are shown with $\epsilon=0.15$. In Figure 4 accuracy results are shown with varying $\epsilon$ and $a$ fixed at four.

The second experiment fixes the number of known input tuples (and $\epsilon$ at 0.15 ) but changes the size of the original data $X$ in order to assess the computational efficiency of the attack. For the Gaussian data, it uses the first $k$ tuples as $X$ where $k$ takes a value in $\{10000,20000, \ldots, 100000\}$. Then, the attack proceeds by carrying out the following operations ten times. Select $a=50$ linearly independent tuples randomly from $X$ and use Algorithm 1 to compute the maximal uniquely valid assignment $I$. The average time taken to compute $I$ is given in Figure 5 top. For the Letter Recognition data, $k$ takes a value in $\{2000,4000, \ldots, 18000\}$ and the attack randomly select $a=10$ linearly independent tuples as the known inputs. The average time taken to find $I$ is given in Figure 5 bottom.

Regarding the known input attack accuracy, the linking phase of the attack (Algorithm (1), exhibits excellent performance. For synthetic data, its performance is perfect in that all known input tuples have their corresponding perturbed tuple inferred (see Figure 2 top). For real data, its performance is nearly perfect - see Figure 3 top. As expected, $\rho(\epsilon)$ approaches one as $a$ increases see Figures 2 and 3 bottom. Interestingly on the synthetic dataset, the transition from $\rho(\epsilon)=0 \rightarrow 1$ occurs very sharply around $a=60$. Moreover, on the real dataset, $\rho(\epsilon)=1$ with $a$ as small as 4 (and we also observe in Figure4 that the probability remains fairly high for $\epsilon$ as small as 0.07 ).

Regarding computational efficiency, the algorithm appears to require quite reasonable time in all cases observed, e.g. less that 450 seconds on the synthetic dataset with 100 known tuples (see Figure 2 middle) and less than 45 seconds on the real dataset with 16 known inputs (see Figure 3

[^9]

Figure 2: Known input attack on Gaussian data with different number of known inputs and $\epsilon=0.15$.


Figure 3: Known input attack on Letter Recognition data with different number of known inputs and $\epsilon=0.15$.


Figure 4: Known input attack on Letter Recognition data with fixed data size, fixed number of known inputs $a=4$, but varying $\epsilon$.
middle). With respect to known input set size (a), the average computation time exhibits a linear (synthetic data) or slower (real data) trend (see Figure 2 and Figure 3 middle). With respect to dataset size (number of private data tuples), the average computation time exhibits a clear linear trend for both synthetic and real data (see Figure 5). These results demonstrate that, despite the high worst-case computational complexity, the computation times on both real and synthetic data are quite reasonable.

The experimental results support the conclusion that the attack can breach privacy in plausible situations. For example, on the 16-dimensional, 18688 tuple real dataset, the known input attack achieves a privacy breach with probability one using four known inputs and less than 30 seconds of run-time.

### 5.2. Experiments Comparing the Known Input-Output Attack with Kaplan's Attack

We compare the accuracy of the attacks with respect to an attacker's goal of producing a single perturbed tuple and an estimate of its unperturbed counterpart. Since Kaplan's attack does not provide the attacker with any means to know how good the estimate is, the attacker has no reason to choose one perturbed tuple over another, hence, we assume the attacker picks randomly. On the other hand, our attack provides the attacker with breech proababilities, so, the attacker chooses the perturbed tuple to maximize the breech probability (as done in Algorithm [5].

The experiments proceed as follows. $X$ and $Y$ are fixed and $a$, the number of known input tuples, is varied. 100 trials are carried out as follows. Select $a$ linearly independent tuples randomly from $X$ (these become the know inputs) and do both of the following. (i) Choose a tuple $y_{\ell}$ randomly from $Y$ whose unperturbed counterpart $x_{\hat{\ell}}$ in $X$ is not among the $a$ known inputs. Use Kaplan's attack ${ }^{15}$ to produce an estimate, $\hat{x}$, of $x_{\hat{\ell}}$. Record the Euclidean relative error of the estimate, $\left\|x_{\hat{\ell}}-\hat{x}\right\| /\left\|x_{\hat{\imath}}\right\|$. (ii) Use our Algorithm ${ }^{[6]}$ to choose the tuple $y_{j}$ from $Y$ with

[^10]

Figure 5: Known input attack on Gaussian (left) and Letter Recognition data (right) with varying size, but fixed number of known inputs $a=50,10$ (respectively) and fixed $\epsilon=0.15$.
maximum $\epsilon$-privacy breech probability and whose unperturbed counterpart $x_{\hat{j}}$ does not appear among the $a$ known inputs, then produce an estimate $\hat{x}$ of $x_{\hat{j}}$. Record the Euclidean relative error of the estimate: $\left\|x_{\hat{j}}-\hat{x}\right\| /\left\|x_{j}\right\|$.

Figure 6 shows the average relative error of the attacks on both datasets. From the Figure it is clear that our approach allows the attacker to produce a significantly more accurate estimate. We do not provide a figure comparing the run-times of the attacks because Kaplan's computes an estimate from only one perturbed tuple while ours, in effect, computes an estimate from all perturbed tuples. However, the time required by our algorithm to produce an estimate from a single, randomly chosen, perturbed tuple is 100 to 1000 times faster than Kaplan's.

## 6. Conclusion

We examined the vulnerability of Euclidean distance-preserving data perturbation when a small set of original data tuples are known to the attacker. We developed a stochastic technique


Figure 6: A comparison of Kaplan's attack and our Algorithm 5] in terms of average relative error over 100 trials. The top and bottom charts show the average error on the Gaussian and Letter recognition datasets, respectively.
allowing the attacker to estimate, for each perturbed tuple, the original unknown data tuple and calculate the probability that the estimation results in a privacy breach. For perturbations which fix the origin, this probability is dependent on the length of the original tuple relative to its distance from the column space of the known inputs. Therefore, the probability increases as the number of known original tuples does, reaching one when the number of linearly independent known original tuples reaches the number of data dimensions. The assumption of fixing the origin can be dropped, resulting in a slightly more complicated breach probability calculation. Our experiments on real and synthetic data showed that even with the number of known original tuples significantly smaller than the number of data dimensions, privacy is breached with high probability. For example, on a real 16 -dimensional dataset, 4 known original tuples is enough for the attacker to estimate an unknown original tuple with less than $7 \%$ error with probability exceeding 0.8 .

We conclude the paper by pointing to an interesting direction for future work, extending techniques in this paper to apply to random projection data perturbation: $Y=\ell^{-1 / 2} \hat{R} X$ where $\hat{R}$ is an $\ell \times n$ matrix with each entry generated independently and from a standard normal distribution (this type of data perturbation for $\ell \leq n$ was discussed in [37]). It can be shown that matrix $R$ is orthogonal on expectation and the probability of orthogonality approaches one exponentially fast with $\ell$. By increasing $\ell$, the data owner can guarantee that distances are preserved with arbitrarily high probability. However, such an increase intuitively would seem to increase the vulnerability with respect to a known input attack. Some preliminary results along these lines can be found in [43].

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## Appendix A. Supplementary Material

## Appendix A.1. Known Input Attack: Proof of Theorem 4.1

Theorem 4.1 Let $L$ be the mapping $P \in \mathbb{O}_{n-k} \mapsto M_{T} U_{k} U_{k}^{\prime}+V_{n-k} P U_{n-k}^{\prime}$. Then, $L$ is an affine bijection from $\mathbb{O}_{n-k}$ to $\mathbb{M}\left(X_{q}, Y_{q}\right)$. And, $L^{-1}$ is the mapping $M \in \mathbb{M}\left(X_{q}, Y_{q}\right) \mapsto V_{n-k}^{\prime} M U_{n-k}$.

To prove this theorem we rely upon the following key technical result.
Lemma Appendix A.1. Let $\mathbb{P}$ denote the set $\left\{M_{T} U_{k} U_{k}^{\prime}+V_{n-k} P U_{n-k}^{\prime}: P \in \mathbb{O}_{n-k}\right\}$. Then $\mathbb{M}\left(X_{q}, Y_{q}\right)=\mathbb{P}$.

Proof: Let $\mathbb{M}\left(U_{k}, M_{T} U_{k}\right)$ denote the set of all $M \in \mathbb{O}_{n}$ such that $M U_{k}=M_{T} U_{k}$. First we show that $\mathbb{M}\left(X_{q}, Y_{q}\right)=\mathbb{M}\left(U_{k}, M_{T} U_{k}\right)$. Since $\operatorname{Col}\left(X_{q}\right)=\operatorname{Col}\left(U_{k}\right)$, then there exists $k \times p$ matrix $A$ such that $U_{k} A=X_{q}$. Since $A$ has $k$ columns, then $\operatorname{rank}(A) \leq k$. Furthermore, [40, pg. 201] implies that $k=\operatorname{rank}\left(U_{k} A\right) \leq \min \{k, \operatorname{rank}(A)\}$, thus, $\operatorname{rank}(A)=k$. Therefore, from [40, pg. 90], $A$ has a right inverse.

For any $M \in \mathbb{O}_{n}$, we have

$$
\begin{aligned}
M \in \mathbb{M}\left(X_{q}, Y_{q}\right) & \Leftrightarrow \quad M U_{k} A=M_{T} U_{k} A \\
& \Leftrightarrow M U_{k}=M_{T} U_{k} .
\end{aligned}
$$

The last $\Leftrightarrow$ follows from the fact that $A$ has a right inverse. We conclude that $\mathbb{M}\left(X_{q}, Y_{q}\right)=$ $\mathbb{M}\left(U_{k}, M_{T} U_{k}\right)$. Now we complete the proof by showing that $\mathbb{M}\left(U_{k}, M_{T} U_{k}\right)=\mathbb{P}$.
(1) For any $M \in \mathbb{P}$, there exists $P \in \mathbb{O}_{n-k}$ such that $M=\left\{M_{T} U_{k} U_{k}^{\prime}+V_{n-k} P U_{n-k}^{\prime}\right\}$. We have then

$$
\begin{aligned}
M U_{k} & =M_{T} U_{k} U_{k}^{\prime} U_{k}+V_{n-k} P U_{n-k}^{\prime} U_{k} \\
& =M_{T} U_{k} .
\end{aligned}
$$

If we can show that $M$ is orthogonal, then $M \in \mathbb{M}\left(U_{k}, M_{T} U_{k}\right)$, so, $\mathbb{P} \subseteq \mathbb{M}\left(U_{k}, M_{T} U_{k}\right)$, as desired. Let $U$ denote $\left[U_{k} \mid U_{n-k}\right]$ (clearly $U \in \mathbb{O}_{n}$ ). Observe

$$
\begin{aligned}
M^{\prime} M & =U_{k} U_{k}^{\prime} M_{T}^{\prime} M_{T} U_{k} U_{k}^{\prime}+U_{k} U_{k}^{\prime} M_{T}^{\prime} V_{n-k} P U_{n-k}^{\prime} \\
& +U_{n-k} P^{\prime} V_{n-k}^{\prime} M_{T} U_{k} U_{k}^{\prime}+U_{n-k} P^{\prime} V_{n-k}^{\prime} M_{T} U_{n-k} P U_{n-k}^{\prime} \\
& =U_{k} U_{k}^{\prime}+0+0+U_{n-k} U_{n-k}^{\prime} \\
& =U U^{\prime}=I_{n} .
\end{aligned}
$$

where the first zero in the second equality is due to the fact that $\operatorname{Col}\left(M_{T} U_{k}\right)=\operatorname{Col}\left(Y_{q}\right)$, so, $V_{n-k}^{\prime} M_{T} U_{k}=0$.
(2) Now consider $M \in \mathbb{M}\left(U_{k}, M_{T} U_{k}\right)$. It can be shown that $\operatorname{Col}\left(V_{n-k}\right)=\operatorname{Col}\left(M U_{n-k}\right){ }^{17}$ Thus, there exists $(n-k) \times(n-k)$ matrix $P$ with $V_{n-k} P=M U_{n-k}$. Observe that

$$
\begin{aligned}
P^{\prime} P & =P^{\prime}\left(V_{n-k}^{\prime} V_{n-k}\right) P \\
& =\left(V_{n-k} P\right)^{\prime}\left(V_{n-k} P\right) \\
& =\left(M U_{n-k}\right)^{\prime}\left(M U_{n-k}\right)=I_{n-k}
\end{aligned}
$$

[^11]Thus, $P \in \mathbb{O}_{n-k}$. Moreover,

$$
\begin{aligned}
M U & =M\left[U_{k} \mid U_{n-k}\right] \\
& =\left[M_{T} U_{k} \mid M U_{n-k}\right] \\
& =\left[M_{T} U_{k} \mid V_{n-k} P\right] .
\end{aligned}
$$

Thus,

$$
\begin{aligned}
M & =\left[M_{T} U_{k} \mid V_{n-k} P\right]\left[\begin{array}{c}
U_{k}^{\prime} \\
U_{n-k}^{\prime}
\end{array}\right] \\
& =M_{T} U_{k} U_{k}^{\prime}+V_{n-k} P U_{n-k}^{\prime}
\end{aligned}
$$

Therefore, $M \in \mathbb{P}$, so, $\mathbb{M}\left(U_{k}, M_{T} U_{k}\right) \subseteq \mathbb{P}$, as desired.
Now we prove Theorem 4.1
Proof: Clearly $L$ is an affine map. Moreover, Lemma Appendix A. 1 directly implies that $L$ maps $\mathbb{O}_{n-k}$ onto $\mathbb{M}\left(X_{q}, Y_{q}\right)$. To see that $L$ is one-to-one, consider $P_{1}, P_{2} \in \mathbb{O}_{n-k}$ such that $L\left(P_{1}\right)=$ $L\left(P_{2}\right)$. By definition, $M_{T} U_{k} U_{k}^{\prime}+V_{n-k} P_{1} U_{n-k}^{\prime}=M_{T} U_{k} U_{k}^{\prime}+V_{n-k} P_{2} U_{n-k}^{\prime}$, thus, $V_{n-k} P_{1} U_{n-k}^{\prime}=$ $V_{n-k} P_{2} U_{n-k}^{\prime}$. Therefore $P_{1}=V_{n-k}^{\prime} V_{n-k} P_{1} U_{n-k}^{\prime} U_{n-k}=V_{n-k}^{\prime} V_{n-k} P_{2} U_{n-k}^{\prime} U_{n-k}=P_{2}$.

To complete the proof, consider $P \in \mathbb{O}_{n-k}$. We have, $V_{n-k}^{\prime} L(P) U_{n-k}=V_{n-k}^{\prime} M_{T} U_{k} U_{k}^{\prime} U_{n-k}+$ $V_{n-k}^{\prime} V_{n-k} P U_{n-k}^{\prime} U_{n-k}=0+P$. Moreover, consider $M \in \mathbb{M}\left(X_{q}, Y_{q}\right)$. By Lemma Appendix A.1, there exists $P_{M} \in \mathbb{O}_{n-k}$ such that $M=M_{T} U_{k} U_{k}^{\prime}+V_{n-k} P_{M} U_{n-k}^{\prime}$. We have $L\left(V_{n-k}^{\prime} M U_{n-k}\right)=L\left(P_{M}\right)$ $=M$. Therefore, the inverse of $L$ is $M \in \mathbb{M}\left(X_{q}, Y_{q}\right) \mapsto V_{n-k}^{\prime} M U_{n-k}$.

Appendix A.2. Known Input Attack: A Rigorous Development of the Closed-Form Expression for $\rho\left(x_{\hat{j}}, \epsilon\right)$
Up to (1), we had derived the following result (for $P$ chosen uniformly from $\mathbb{O}_{n-k}$ ):

$$
\begin{equation*}
\rho\left(x_{\hat{j}}, \epsilon\right)=\operatorname{Pr}\left(\left\|P^{\prime} B^{\prime}\left(U_{n-k}^{\prime} x_{\hat{j}}\right)-\left(U_{n-k}^{\prime} x_{\hat{j}}\right)\right\| \leq\left\|x_{j}\right\| \epsilon\right) \tag{A.1}
\end{equation*}
$$

where $B \in \mathbb{O}_{n-k}$ and satisfies $M_{T} U_{n-k} B=V_{n-k}$. Now we provide a rigorous proof of (2), i.e. the r.h.s. above equals $\operatorname{Pr}\left(\left\|P^{\prime}\left(U_{n-k}^{\prime} x_{\hat{j}}\right)-\left(U_{n-k}^{\prime} x_{\hat{j}}\right)\right\| \leq\left\|x_{\hat{j}}\right\| \epsilon\right)$. To do so, we need some material from measure theory.

Because $\mathbb{0}_{n-k}$ is a locally compact topological group [39, pg. 293], it has a Haar probability measure, denoted by $\mu$, over $\mathbb{B}$, the Borel algebra on $\mathbb{O}_{n-k}$. This is commonly regarded as the standard uniform probability measure over $\mathbb{O}_{n-k}$. Its key property is left-invariance: for all $\mathcal{B} \in \mathbb{B}$ and all $M \in \mathbb{O}_{n-k}, \mu(\mathcal{B})=\mu(M \mathcal{B})$, i.e., shifting $\mathcal{B}$ by a rigid motion does not change its probability assignment.

Let $\mathbb{O}_{n-k}\left(U_{n-k}^{\prime} x_{\hat{j}},\left\|x_{j}\right\| \epsilon\right)$ denote the set of all $P \in \mathbb{O}_{n-k}$ such that $\left\|P^{\prime}\left(U_{n-k}^{\prime} x_{\hat{j}}\right)-\left(U_{n-k}^{\prime} x_{\hat{j}}\right)\right\| \leq$ $\left\|x_{j}\right\| \epsilon$. Let $\mathbb{O}_{n-k}^{B^{\prime}}\left(U_{n-k}^{\prime} x_{\hat{j}},\left\|x_{\hat{j}}\right\| \epsilon\right)$ denote the set of all $P \in \mathbb{O}_{n-k}$ such that $\left\|P^{\prime} B^{\prime}\left(U_{n-k}^{\prime} x_{\hat{j}}\right)-\left(U_{n-k}^{\prime} x_{\hat{j}}\right)\right\| \leq$ $\left\|x_{\hat{j}}\right\| \epsilon \underbrace{18}$ By definition of $\mu$ we have,

$$
\begin{aligned}
\mu\left(\mathbb{O}_{n-k}\left(U_{n-k}^{\prime} x_{\hat{j}},\left\|x_{\hat{j}}\right\| \epsilon\right)\right) & =\operatorname{Pr}\left(P \text { uniformly chosen from } \mathbb{O}_{n-k} \text { lies in } \mathbb{O}_{n-k}\left(U_{n-k}^{\prime} x_{\hat{j}},\left\|x_{\dot{j}}\right\| \epsilon\right)\right) \\
& =\operatorname{Pr}\left(\left\|P^{\prime}\left(U_{n-k}^{\prime} x_{\hat{j}}\right)-\left(U_{n-k}^{\prime} x_{\hat{j}}\right)\right\| \leq\left\|x_{\hat{j}}\right\| \epsilon\right),
\end{aligned}
$$

[^12]and,
\[

$$
\begin{aligned}
\mu\left(\mathbb{O}_{n-k}^{B^{\prime}}\left(U_{n-k}^{\prime} x_{j},\left\|x_{j}\right\| \epsilon\right)\right) & =\operatorname{Pr}\left(P \text { uniformly chosen from } \mathbb{O}_{n-k} \text { lies in } \mathbb{O}_{n-k}^{B^{\prime}}\left(U_{n-k}^{\prime} x_{\hat{j}},\left\|x_{\hat{j}}\right\| \epsilon\right)\right) \\
& =\operatorname{Pr}\left(\left\|P^{\prime} B^{\prime}\left(U_{n-k}^{\prime} x_{\hat{j}}\right)-\left(U_{n-k}^{\prime} x_{\hat{j}}\right)\right\| \leq\left\|x_{\hat{j}}\right\| \epsilon\right),
\end{aligned}
$$
\]

Therefore,

$$
\begin{align*}
\operatorname{Pr}\left(\left\|P^{\prime} B^{\prime}\left(U_{n-k}^{\prime} x_{\hat{j}}\right)-\left(U_{n-k}^{\prime} x_{\dot{j}}\right)\right\| \leq\left\|x_{\hat{j}}\right\| \epsilon\right) & =\mu\left(\mathbb{O}_{n-k}^{B^{\prime}}\left(U_{n-k}^{\prime} x_{\hat{j}},\left\|x_{\dot{j}}\right\| \epsilon\right)\right) \\
& =\mu\left(B \mathbb{O}_{n-k}^{B^{\prime}}\left(U_{n-k}^{\prime} x_{\hat{j}},\left\|x_{j}\right\| \epsilon\right)\right) \\
& =\mu\left(\mathbb{O}_{n-k}\left(U_{n-k}^{\prime} x_{\hat{j}},\left\|x_{j}\right\| \epsilon\right)\right)  \tag{A.2}\\
& =\operatorname{Pr}\left(\left\|P^{\prime}\left(U_{n-k}^{\prime} x_{\hat{j}}\right)-\left(U_{n-k}^{\prime} x_{\hat{j}}\right)\right\| \leq\left\|x_{j}\right\| \epsilon\right)
\end{align*}
$$

where the second equality is due to the left-invariance of $\mu$ and the third equality is due to the fact that $B \mathbb{O}_{n-k}^{B^{\prime}}\left(U_{n-k}^{\prime} x_{\hat{j}},\left\|x_{j}\right\| \epsilon\right)$ can be shown to equal $\mathbb{O}_{n-k}\left(U_{n-k}^{\prime} x_{\hat{j}},\left\|x_{\hat{j}}\right\| \epsilon\right)$.

Since the last equality above was for intuitive purposes only, we will ignore it in completing the derivation of a closed form expression. (A.1) and (A.2) imply

$$
\rho\left(x_{\hat{j}}, \epsilon\right)=\mu\left(\mathbb{O}_{n-k}\left(U_{n-k}^{\prime} x_{\hat{j}},\left\|x_{\hat{j}}\right\| \epsilon\right)\right) .
$$

Recall that $S_{n-k}\left(\left\|U_{n-k}^{\prime} x_{j}\right\|\right)$ denotes the hyper-sphere in $\mathfrak{R}^{n-k}$ with radius $\left\|U_{n-k}^{\prime} x_{j}\right\|$ and centered at the origin and $S_{n-k}\left(U_{n-k}^{\prime} x_{\hat{j}},\left\|x_{j}\right\| \epsilon\right)$ denotes the points contained by $S_{n-k}\left(\left\|U_{n-k}^{\prime} x_{j}\right\|\right)$ whose distance from $U_{n-k}^{\prime} x_{\hat{j}}$ is no greater than $\left\|x_{j}\right\| \epsilon$. Using basic principles from measure theory, it can be shown tha ${ }^{19}$

$$
\mu\left(\mathbb{O}_{n-k}\left(U_{n-k}^{\prime} x_{\hat{j}},\left\|x_{j}\right\| \epsilon\right)\right)=\frac{S A\left(S_{n-k}\left(U_{n-k}^{\prime} x_{\hat{j}},\left\|x_{\hat{j}}\right\| \epsilon\right)\right)}{S A\left(S_{n-k}\left(\left\|U_{n-k}^{\prime} x_{j}\right\|\right)\right)}
$$

We have arrived at Equation (4) from Section 4.4 Next, we derive the desired closed-form expression (5). To simplify exposition, we prove the following result for $m \geq 0, z \in \mathfrak{R}^{m}$, and $c \geq 0$ (by plugging in $m=n-k, z=U_{n-k}^{\prime} x_{\hat{j}}$, and $c=\left\|x_{\dot{j}}\right\| \epsilon$, (5) follows).

Before proving (A.3) we establish:

[^13]

Figure A.7: The hyper-sphere $S_{m}(\|z\|)$ and two "north pole" caps $(c \leq\|z\| \sqrt{2})$.


Figure A.8: The hyper-sphere $S_{m}(\|z\|)$ and one "south pole" cap ( $\|z\| \sqrt{2}<c<\|z\| 2$ ).

$$
\begin{equation*}
\text { For } b \geq 2 \text { and } r>0, S A\left(S_{b}(r)\right)=\frac{b r^{b-1} \pi^{b / 2}}{\Gamma((b+2) / 2)} \tag{A.4}
\end{equation*}
$$

Indeed, with $\operatorname{Vol}($.$) denoting volume, it can be shown that S A\left(S_{b}(r)\right)=\frac{d \operatorname{Vol}\left(S_{b}(r)\right)}{d r}=\operatorname{Vol}\left(S_{b}(1)\right) \frac{d r_{b}}{d r}$ $=\frac{\pi^{b / 2} b b^{r-1}}{\Gamma((b+2) / 2)}$. The last equality follows from [44]. Now we return to proving (A.3).

If $m=0$, then the surface area ratio equals 1 by definition. If $c \geq\|z\| 2$ and $m \geq 1$, then the ratio equals 1 since $S_{m}(z, c)=S_{m}(\|z\|)$. If $c<\|z\| 2$ and $m=1$, then, the ratio equals 0.5 since $S_{1}(z, c)=\{z\}$ and $S_{1}(\|z\|)=\{z,-z\}$. For the remainder of the derivation, we assume that $m \geq 2$ and, without loss of generality, $z$ is at the "north pole" of the hyper-sphere $S_{m}(\|z\|)$, i.e. $z$ $=(1,0,0, \cdots, 0)$.

Case $c \leq\|z\| \sqrt{2}$ : The set of points on $S_{m}(\|z\|)$ whose distance from $z$ equals $c$ is the intersection of $S_{m}(\|z\|)$ with the hyper-plane whose perpendicular to $z$ is of length $h$ as seen in Figure A. 7 Thus, $S_{m}(z, c)$ are all those points on $S_{m}(\|z\|)$ not below that hyper-plane.

Sub-case $m=2$ : Since $S_{2}(\|z\|)$ is an ordinary circle, then the angle $\theta$ in Figure A.7determines the surface area ratio as follows $\frac{\left.(2 \theta / 2 \pi) S A\left(S_{2}\| \|\| \|\right)\right)}{S A\left(S_{2}(\| \| \|)\right)}=\theta / \pi$. Moreover, since $\theta$ is the top angle of an isosceles triangle with sides of length $\|z\|$ and base of length $c$, then $\sin (\theta / 2)=c /(2\|z\|)$. The half-angle formula implies that $\theta=\arccos \left(1-[c /(\|z\| \sqrt{2})]^{2}\right)$. Therefore, as desired,

$$
\begin{equation*}
\frac{S A\left(S_{2}(z, c)\right)}{S A\left(S_{2}(\|z\|)\right)}=(1 / \pi) \arccos \left(1-[c /(\|z\| \sqrt{2})]^{2}\right) \tag{A.5}
\end{equation*}
$$

Sub-case $m \geq 3$ : Here, computing the surface area ratio is more complicated and requires an appeal to the integral definition of the cap surface area. Consider the intersection of $S_{m}(\|z\|)$ with the hyper-plane whose perpendicular to $z$ is of length $0 \leq h_{1} \leq h$ as seen in Figure A. 7 The surface area of this intersection equals the surface area of $S_{m-1}\left(r\left(h_{1}\right)\right)$. Thus, (A.4) implies

$$
\begin{aligned}
S A\left(S_{m}(z, c)\right) & =\int_{h_{1}=0}^{h} S A\left(S_{m-1}\left(r\left(h_{1}\right)\right)\right) d h_{1} \\
& =\left(\frac{(m-1) \pi^{(m-1) / 2}}{\Gamma((m+1) / 2)}\right) \int_{h_{1}=0}^{h}\left(r\left(h_{1}\right)\right)^{m-2} d h_{1} .
\end{aligned}
$$

To evaluate the integral, we change coordinates with $h_{1}=\|z\|\left(1-\cos \left(\theta_{1}\right)\right)$. So, $h_{1}=0, h$ implies that $\theta_{1}=0, \arccos (1-h /\|z\|)$. And, $r\left(\|z\|\left(1-\cos \left(\theta_{1}\right)\right)=\|z\| \sin \left(\theta_{1}\right)=\frac{d h_{1}}{d \theta_{1}}\right.$. Therefore,

$$
\begin{aligned}
\int_{h_{1}=0}^{h}\left(r\left(h_{1}\right)\right)^{m-2} d h_{1} & =\int_{\theta_{1}=0}^{\arccos (1-h /\|z\|)} r\left(\|z\|\left(1-\cos \left(\theta_{1}\right)\right)^{m-2} \frac{d h_{1}}{d \theta_{1}} d \theta_{1}\right. \\
& =\int_{\theta_{1}=0}^{\arccos (1-h /\|z\|)}\|z\|^{m-2} \sin ^{m-2}\left(\theta_{1}\right)\|z\| \sin \left(\theta_{1}\right) d \theta_{1} \\
& =\|z\|^{m-1} \int_{\theta_{1}=0}^{\arccos (1-h /\|z\|)} \sin ^{m-1}\left(\theta_{1}\right) d \theta_{1} .
\end{aligned}
$$

Plugging this into the previous equations for $S A\left(S_{m}(z, c)\right)$ and using (A.4), we get

$$
\begin{aligned}
\frac{S A\left(S_{m}(z, c)\right)}{S A\left(S_{m}(\|z\|)\right)} & =\left(\frac{(m-1) \pi^{(m-1) / 2}\|z\|^{m-1} \Gamma((m+2) / 2)}{\Gamma((m+1) / 2) m\|z\|^{m-1} \pi^{m / 2}}\right) \int_{\theta_{1}=0}^{a r c c o s(1-h /\|z\|)} \sin ^{m-1}\left(\theta_{1}\right) d \theta_{1} \\
& =\left(\frac{(m-1) \Gamma((m+2) / 2)}{\Gamma((m+1) / 2) m \sqrt{\pi}}\right) \int_{\theta_{1}=0}^{\arccos (1-h /\|z\|)} \sin ^{m-1}\left(\theta_{1}\right) d \theta_{1}
\end{aligned}
$$

Since $h=\frac{c^{2}}{2\|z\|}$, then, as desired, we get

$$
\begin{equation*}
\frac{S A\left(S_{m}(z, c)\right)}{S A\left(S_{m}(\|z\|)\right)}=\left(\frac{(m-1) \Gamma((m+2) / 2)}{\Gamma((m+1) / 2) m \sqrt{\pi}}\right) \int_{\theta_{1}=0}^{\arccos \left(1-[c /\|z\| \sqrt{2}]^{2}\right)} \sin ^{m-1}\left(\theta_{1}\right) d \theta_{1} \tag{A.6}
\end{equation*}
$$

Case $\|z\| \sqrt{2}<c<\|z\| 2$ : As depicted in Figure A.8 $S_{m}(z, c)$ contains the entire northern hemisphere of $S_{m}(\|z\|)$. Let $S_{n}(-z, c)$ denote the "south pole" cap defined by $h^{\prime}$ (and $c^{\prime}$ ) in Figure A. 8 (clearly $c^{\prime} \leq\|z\| \sqrt{2}$ ). We have

$$
\begin{equation*}
\frac{S A\left(S_{m}(z, c)\right)}{S A\left(S_{m}(\|z\|)\right)}=1-\frac{S A\left(S_{m}\left(-z, c^{\prime}\right)\right.}{S A\left(S_{m}(\|z\|)\right)} \tag{A.7}
\end{equation*}
$$

By replacing " $c$ " with " $c$ '" in (A.5) and (A.6) then plugging the resulting expression into A.7) we get,

$$
\frac{S A\left(S_{m}(z, c)\right)}{S A\left(S_{m}(\|z\|)\right)}= \begin{cases}1-(1 / \pi) \arccos \left(1-\left[c^{\prime} /(\|z\| \sqrt{2})\right]^{2}\right) & \text { if } m=2  \tag{A.8}\\ 1-\frac{(m-1) \Gamma([m+2] / 2)}{m \sqrt{\pi \Gamma([m+1] / 2)}} \int_{\theta_{1}=0}^{\arccos \left(1-\left[c^{\prime} /(\|z\| \sqrt{2})\right]^{2}\right)} \sin ^{m-1}\left(\theta_{1}\right) d \theta_{1} & \text { if } m \geq 3\end{cases}
$$

From Figure A.8, it can be seen that $\theta$ is the top angle on an isosceles triangle with sides of length $\|z\|$ and base of length $c^{\prime}$. So, $\sin (\theta / 2)=\frac{c^{\prime}}{2\|z\|}$. The half-angle formula implies $\cos (\theta)=$ $1-\left[c^{\prime} /(\|z\| \sqrt{2})\right]^{2}$. Similar reasoning shows $\cos (\pi-\theta)=1-[c /(\|z\| \sqrt{2})]^{2}$. Since $0 \leq \theta \leq \pi / 2$, then $\cos (\pi-\theta)=-\cos (\theta)$. Thus, $[c /(\|z\| \sqrt{2})]^{2}-1=1-\left[c^{\prime} /(\|z\| \sqrt{2})\right]^{2}$. Plugging $2-\left[\frac{c}{\|z\| \sqrt{2}}\right]^{2}$ in for $\left[\frac{c^{\prime}}{\|z\| \sqrt{2}}\right]^{2}$ in A.8) yields the desired results.

## Appendix A.3. Known Input Attack: Computing the Closed-Form Expression for $\rho\left(x_{\hat{j}}, \epsilon\right)$

Next we develop recursive procedures for computing (6). This amounts to computing the following two functions: (i) $G R(m)=\Gamma([m+2] / 2) / \Gamma([m+1] / 2$ ) for $m \geq 1$; (ii) $S I(z, m)=$ $\int_{\theta_{1}=0}^{\arccos (z)} \sin ^{m-1}\left(\theta_{1}\right) d \theta_{1}$ for $1 \geq z \geq 0$ and $m \geq 1$. Indeed, (6) is equivalent to

$$
\rho\left(x_{j}, \epsilon\right)= \begin{cases}1 & \text { if } n-k=0 ;  \tag{A.9}\\ 1 & \text { if }\| \|\left\|_{j}\right\| \geq\left\|V_{n-k}^{\prime} y_{j}\right\| 2 \text { and } n-k \geq 1 ; \\ \text { if }\left\|y_{j}\right\| \epsilon<\left\|V_{n-k}^{\prime} y_{j}\right\| 2 \text { and } n-k=1 ; \\ 1-(1 / \pi) \arccos \left(\left[\frac{\left\|y_{j}\right\| \epsilon}{\left\|V_{n-k} j_{j}\right\| \sqrt{2}}\right]^{2}-1\right) & \text { if }\left\|V_{n-k}^{\prime} y_{j}\right\| \sqrt{2}<\left\|y_{j}\right\| \epsilon<\left\|V_{n-k}^{\prime} y_{j}\right\| 2 \text { and } n-k=2 ; \\ 1-\frac{(n-k-1) G R(n-k)}{(n-k) \sqrt{\pi}} S I\left(\left[\frac{\left\|y_{j}\right\| \epsilon}{\| V_{n-k}^{\prime}-k j l \sqrt{2}}\right]^{2}-1, n-k\right) & \text { if }\left\|V_{n-k}^{\prime} y_{j}\right\| \sqrt{2}<\left\|y_{j}\right\| \epsilon<\left\|V_{n-k}^{\prime} y_{j}\right\| 2 \text { and } n-k \geq 3 ; \\ (1 / \pi) \arccos \left(1-\left[\frac{\left\|y_{j}\right\| \epsilon}{\left\|V_{n-k}^{\prime} y_{j}\right\| \sqrt{2}}\right]^{2}\right) & \text { if }\left\|y_{j}\right\| \epsilon \leq\left\|V_{n-k}^{\prime} y_{j}\right\| \sqrt{2} \text { and } n-k=2 ; \\ \frac{(n-k-1) G R(n-k)}{(n-k) \sqrt{\pi}} S I\left(1-\left[\frac{\left\|y_{j}\right\| \epsilon}{\left\|V_{n-k}^{\prime} y_{j}\right\| \sqrt{2}}\right]^{2}, n-k\right) & \text { if }\left\|y_{j}\right\| \epsilon \leq\left\|V_{n-k}^{\prime} y_{j}\right\| \sqrt{2} \text { and } n-k \geq 3 .\end{cases}
$$

To compute $G R(m)$ for $m \geq 1$, we use the following facts: $\Gamma(z+1)=z \Gamma(z)$ for $z>0, \Gamma(1 / 2)$ $=\sqrt{\pi}$, and $\Gamma(1)=1$. Thus, we get a recursive procedure for computing $G R(m)$.

$$
G R(m)= \begin{cases}\frac{\sqrt{\pi}}{2} & \text { if } m=1  \tag{A.10}\\ \frac{2}{\sqrt{\pi}} & \text { if } m=2 \\ \left(\frac{m}{m-1}\right) G R(m-2) & \text { if } m \geq 3\end{cases}
$$

To compute $S I(z, m)$ for $1 \geq z \geq 0$ and $m \geq 1$, we use the following facts. $\sin ^{m-2}(\arccos (z))=$ $\left[1-z^{2}\right]^{(m-2) / 2}$ if $m \geq 3$. And, $S I(z, m)=\left[\int \sin ^{m-1}\left(\theta_{1}\right) d \theta_{1}\right](\arccos (z))-\left[\int \sin ^{m-1}\left(\theta_{1}\right) d \theta_{1}\right](0)$. And,

$$
\left[\int \sin ^{m-1}\left(\theta_{1}\right) d \theta_{1}\right](w)= \begin{cases}w & \text { if } m-1=0  \tag{A.11}\\ -\cos (w) & \text { if } m-1=1 \\ \frac{m-2}{m-1}\left[\int \sin ^{m-3}\left(\theta_{1}\right) d \theta_{1}\right](w)-\frac{\sin ^{m-2}(w) \cos (w)}{m-1} & \text { if } m-1 \geq 2\end{cases}
$$

Therefore,

$$
S I(z, m)= \begin{cases}\arccos (z) & \text { if } m=1  \tag{A.12}\\ 1-z & \text { if } m=2 \\ \frac{m-2}{m-1} S I(z, m-2)-\frac{z\left[1-z^{2}\right]^{(m-2) / 2}}{m-1} & \text { if } m \geq 3\end{cases}
$$



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[^1]:    ${ }^{1}$ The census model is widely studied in the field of security control for statistical databases [6].

[^2]:    ${ }^{2}$ We described our linking technique in an earlier, unpublished, technical report version of this paper (citation omitted because of the double-blind nature of this submission). This report appeared 3 months after Wong's paper, and, at the time we were unaware of Wong's work.

[^3]:    ${ }^{3}$ If $M$ is square, it is orthogonal if and only if $M^{\prime}=M^{-1}$ 39, pg. 17].

[^4]:    ${ }^{4}$ As described in Section 1 this can easily be extended to produce estimates for as many unknown original data tuples as desired.
    ${ }^{5}$ The attacker does not need to know $\hat{j}$; she is merely producing an estimate of the private data tuple that was perturbed to produce $y_{j}$.

[^5]:    ${ }^{6}$ This can be achieved by the attacker appropriately reordering the columns of $X_{a}$ and $Y$.
    ${ }^{7}$ This is equivalent to a maximum likelihood estimate of $x_{\hat{j}}$.

[^6]:    ${ }^{8}$ That the resulting $\hat{M}$ was chosen uniformly from $\mathbb{M}\left(X_{q}, Y_{q}\right)$ could be more rigorously justified using left-invariance of probability measures and the Haar probability measure over $\mathbb{O}_{n-k}$. But, such a discussion is not relevant to this paper and is omitted.
    ${ }^{9}$ We define $\mathbb{O}_{0}$ to contain a single, empty matrix. And, for $P \in \mathbb{O}_{0}$, we define $V_{n-k} P U_{n-k}^{\prime}$ to be the $n \times n$ zero matrix.

[^7]:    ${ }^{10} B^{\prime} B=B^{\prime} U_{n-k}^{\prime} M_{T}^{\prime} V_{n-k}=V_{n-k}^{\prime} V_{n-k}=I_{n-k}$.
    ${ }^{11} S_{1}\left(\left\|U_{1}^{\prime} x_{\hat{j}}\right\|\right)$ consists of two points. We define $\frac{S A\left(S_{1}\left(U_{1}^{\prime} x_{\hat{j}}\left\|x_{j}\right\| \epsilon\right)\right)}{S A\left(S_{1}\left(\left\|U_{1}^{\prime} x_{j}\right\| \|\right)\right.}$ as 0.5 if $S_{1}\left(U_{1}^{\prime} x_{\hat{j}},\left\|x_{\hat{j}}\right\| \epsilon\right)$ is one point, and as 1 otherwise. Moreover, we define $\frac{S A\left(S_{0}\left(U_{0}^{\prime} x_{j},\left\|x_{j}\right\| \epsilon\right)\right)}{S A\left(S_{0}\left(\left\|U_{0}^{\prime} x_{j}\right\|\right)\right)}$ as 1 .

[^8]:    ${ }^{12} M_{T} x_{\hat{j}}=y_{j}$, so, $\left\|x_{\hat{j}}\right\|=\left\|M_{T} x_{\hat{j}}\right\|=\left\|y_{j}\right\|$. Moreover, as shown earlier, there exists $B \in \mathbb{O}_{n-k}$ such that $V_{n-k}^{\prime}=B^{\prime} U_{n-k}^{\prime} M_{T}^{\prime}$. Thus, $\left\|U_{n-k}^{\prime} x_{j}\right\|=\left\|B^{\prime} U_{n-k}^{\prime} M_{T}^{\prime} M_{T} x_{j}\right\|=\left\|V_{n-k}^{\prime} y_{j}\right\|$.
    ${ }^{13}$ Since $\operatorname{Col}\left(U_{k}\right)=\operatorname{Col}\left(X_{q}\right)$, then by solving $k$ systems of linear equations (one for each column of $U_{k}$ ), a $q \times k$ matrix $A$ can be computed such that $X_{q} A=U_{k}$.

[^9]:    ${ }^{14}$ The mean vector is specified by independently generating 100 numbers from a univariate Gaussian with mean zero and variance one. The covariance matrix is specified by (i) independently generating 100 data tuples each with 100 independently generated entries a from a univariate Gaussian with mean zero and variance one, (ii) computing the empirical covariance of this 100 tuple dataset.

[^10]:    ${ }^{15}$ With a learning rate of 0.05 and 500 iterations, values observed empirically to produce the best results
    ${ }^{16}$ With $\epsilon=0.15$ and 0.1 for the Gaussian and Letter data, respectively.

[^11]:    ${ }^{17}$ Since $\left(M U_{n-k}\right)^{\prime} M U_{k}=0$, then $\operatorname{Col}\left(M U_{n-k}\right)=\operatorname{Col}_{\perp}\left(M U_{k}\right)$. Since $M U_{k}=M_{T} U_{k}$ and $\operatorname{Col}\left(M_{T} U_{k}\right)=\operatorname{Col}\left(Y_{q}\right)$, then it follows that $\operatorname{Col}_{\perp}\left(M U_{k}\right)=\operatorname{Col}_{\perp}\left(M_{T} U_{k}\right)=\operatorname{Col}_{\perp}\left(Y_{q}\right)=\operatorname{Col}\left(V_{n-k}\right)$.

[^12]:    ${ }^{18}$ Since $\mathbb{O}_{n-k}\left(U_{n-k}^{\prime} x_{j},\left\|x_{j}\right\| \epsilon\right)$ and $\mathbb{O}_{n-k}^{B^{\prime}}\left(U_{n-k}^{\prime} x_{\hat{j}},\left\|x_{j}\right\| \epsilon\right)$ are topologically closed sets, then they are Borel subsets of $\mathbb{O}_{n-k}$, therefore, $\mu$ is defined on each of these.

[^13]:    ${ }^{19} S_{1}\left(\left\|U_{1}^{\prime} x_{\hat{j}}\right\|\right)$ consists of two points. Recall that we define $\frac{S A\left(S_{1}\left(U_{1}^{\prime} x_{j}\left\|x_{j}\right\| \epsilon\right)\right.}{S A\left(S_{1}\left(\left\|U_{1}^{\prime} x_{j}\right\|\right)\right)}$ as 0.5 if $S_{1}\left(U_{1}^{\prime} x_{\hat{j}},\left\|x_{j}\right\| \epsilon\right)$ is one point, and as 1 otherwise. Moreover, we define $\frac{S A\left(S_{0}\left(U_{0}^{\prime} x_{j},\left\|x_{j}\right\| \epsilon\right)\right)}{S A\left(S_{0}\left(\left\|U_{0}^{\prime} x_{j}\right\| \|\right)\right.}$ as 1 .

