Empirical modeling of the impact factor distribution

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Abstract

The distribution of impact factors has been modeled in the recent informetric literature using two-exponent law proposed by Mansilla et al. (2007). This paper shows that two distributions widely-used in economics, namely the Dagum and Singh-Maddala models, possess several advantages over the two-exponent model. Compared to the latter, the former give as good as or slightly better fit to data on impact factors in eight important scientific fields. In contrast to the two-exponent model, both proposed distributions have closed-from probability density functions and cumulative distribution functions, which facilitates fitting these distributions to data and deriving their statistical properties.

Keywords: impact factor, two-exponent law, Dagum model, Singh-Maddala model, maximum likelihood estimation, model selection

1. Introduction

The distribution of journal impact factors has been recently studied in the informetric literature from both theoretical and empirical perspectives. Mansilla et al. (2007) proposed the two-exponent law to model rank-frequency distributions of impact factors. This model was used by Campanario (2010) to study empirically changes in the distribution of impact factors over time. A theoretical derivation of the rank-frequency distribution of impact factors was derived by Egghe (2009); see also Egghe (2011) and Egghe & Waltman

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(2011).¹ Mishra (2010) has fitted several well-established statistical distributions to data on impact factors for journals from several scientific disciplines. The two-exponent law introduced by Mansilla et al. (2007) has been recently studied by Sarabia et al. (2012). The authors have obtained the probabilistic quantile function corresponding to the two-exponent law as well as derived several statistical measures and tools associated with this law like the moments, Lorenz and Leimkuhler curves and the Gini index of inequality. Moreover, they fitted the two-exponent law to data on impact factors for eight science categories and found that the fit of the model was satisfactory.

The present paper contributes to the literature on empirical modeling of the journal impact factor distribution by verifying in a statistically rigorous way if the two-exponent law is consistent with data on impact factors. In particular, the paper fits the two-exponent law do data using the maximum likelihood approach, which is more efficient than the least squares approach used in previous studies (Mansilla et al., 2007; Sarabia et al., 2012). The fit of the model to data is evaluated using an appropriate goodness-of-fit test. Finally, the two-exponent law is compared to alternative models using a likelihood ratio test. As alternatives we have chosen the Singh-Maddala and Dagum statistical models (see, e.g., Kleiber & Kotz, 2003), which are widelyused in economics to model the distribution of income and other variables.² The analysis is performed for impact factors from eight science categories studied previously by Sarabia et al. (2012).

The reminder of the paper is structured as follows. Section 2 introduces definition and basic properties of the two-exponent law and the compared alternative models. Statistical tests used to assess goodness of fit to data and tests used for model selection are presented as well. Section 3 describes our data on impact factors, while Section 4 presents and discusses empirical findings. Finally, Section 5 concludes.

¹Waltman & Van Eck (2009) have criticized the approach of Egghe (2009) from an empirical point of view.

 $^{^{2}}$ These distributions, among others, were previously used by Mishra (2010) to study the distribution of impact factors. However, they were not compared with the two-exponent model.

2. Methods

2.1. The two-exponent law versus Dagum and Singh-Maddala models

In order to model the distribution of impact factors, Mansilla et al. (2007) introduced the two-exponent law in terms of rank-frequency distribution taking the form:

$$f(r) = K \frac{(N+1-r)^b}{r^a},$$
 (1)

where a > 0, b > 0 are shape parameters, K > 0 is a scale parameter, N is the total number of sources (journals in the case of modeling impact factors), r = 1, 2, ..., N is the ranking number and f(r) it the impact factor. If b = 0, then (1) reduces to Zipf's law (Egghe, 2005); if a = b it becomes the Lavalette law (Lavalette, 1996), and when a = 0 it becomes a power law. Sarabia et al. (2012) derive the quantile function that corresponds to the two-exponent law (1), which takes the form:

$$F^{-1}(u) = K \frac{u^b}{(1-u)^a}, \text{ for } 0 < u < 1.$$
 (2)

The distribution defined by (2) was introduced in statistical literature by Hankin & Lee (2006), who called it the Davies distribution. Unfortunately, neither probability density function (pdf) nor cumulative distribution function (cdf) for (2) is available in closed form except for some special cases like the Zipf's law, the power law or the Lavalette law. However, they can be calculated numerically by inverting the quantile function (2).

The Singh-Maddala and Dagum models were introduced in economics in the context of modeling income distribution by, respectively, Singh & Maddala (1976) and Dagum (1977). In statistics, these distributions appeared first in the system of distributions of Burr (1942) and are known as Burr XII distribution (Singh-Maddala) and Burr III distribution (Dagum). For the Singh-Maddala and Dagum distributions, the pdfs, cdfs and quantile functions are available in closed forms. Specifically, for a sample of positive impact factors in a given scientific filed, $x_1, ..., x_N$, the cdf for the Singh-Maddala distribution is given by:

$$F(x) = 1 - \left[1 + \left(\frac{x}{b}\right)^a\right]^{-q},\tag{3}$$

where a > 0, q > 0 are shape parameters and b > 0 is a scale parameter.

The cdf for the Dagum distribution is

$$F(x) = \left[1 + \left(\frac{x}{b}\right)^{-a}\right]^{-p},\tag{4}$$

where a > 0, p > 0 are shape parameters and b > 0 is a scale parameter. The Singh-Maddala and Dagum distributions are closely related in the following way: $X \sim \text{Dagum}(a, b, p) \iff \frac{1}{X} \text{SM}(a, \frac{1}{b}, p))$, where ~ means "is distributed as". The upper tail of the Singh-Maddala distribution is governed by two parameters (a and q), while the lower tail by a only (Kleiber, 1996). The opposite holds for the Dagum model (two parameters govern the behaviour of the lower tail and only one shapes the upper tail) and for this reason this model is more flexible in the lower tail. Therefore, the two models can be considered complementary as they have advantages in modeling different parts of the data. Theoretical properties of the Singh-Maddala and Dagum distributions are very well known; see Kleiber (1996, 2008) and Kleiber & Kotz (2003) for a detailed discussion.³ In particular, while Sarabia et al. (2012)offer expressions for the Lorenz curve and the Gini index of inequality for the two-exponent distribution, in case of the Singh-Maddala and Dagum models expressions for a wide variety of inequality measures exist (Kleiber & Kotz, 2003; Jenkins, 2009). Moreover, the conditions that allow for testing Lorenz dominance (i.e. inference on inequality robust to the choice of an inequality measure) are available (Wilfling & Krämer, 1993; Kleiber, 1996).

When a = b the two-exponent law becomes the Lavalette law (Lavalette, 1996). As noticed by Sarabia et al. (2012), the Lavalette law is known in the economic literature as the Fisk (or log-logistic) distribution. The Fisk distribution is a special case of both the Singh-Maddala model (with q set to 1) and the Dagum model (with p set to 1). Kleiber & Kotz (2003) discuss the properties of the Fisk distribution in more detail.

The Singh-Maddala and Dagum distributions were shown to fit income distributions well. McDonald (1984) studied the fit of several three- and fourparameter statistical models to grouped income distribution data and found that the Singh-Maddala model performed best among the three-parameter distributions and even better than one four-parameter distribution. Dastrup et al.

³The Singh-Maddala and Dagum distributions are nested within a four-parameter Generalized Beta of the Second Kind (GB2) model introduced by McDonald (1984). We have experimented with fitting this model to data on impact factors, but the gains from the additional complexity were small.

(2007) found that for the disposable income variable the Singh-Maddala model was the best-fitting distribution as often as the Dagum distribution in the group of three-parameter distributions. On the other hand, Bandourian et al. (2003) compared several distributional models for 82 data sets on gross income data and found that the Dagum model is the best three-parameter model in 84% of cases.

2.2. Estimation methods, goodness-of-fit and model selection tests

We estimate all models using maximum likelihood (ML) estimation. The log-likelihood functions for the Singh-Maddala and Dagum models are available in Kleiber & Kotz (2003). The log-likelihood for the two-exponent model can be calculated by numerical inversion of the quantile function (2) (Hankin & Lee, 2006). Parameter estimates are obtained by numerical maximization of the log-likelihood functions. The variance matrix is calculated as the negative inverse of the Hessian evaluated at the parameter estimates.

Sarabia et al. (2012) estimate the parameters of the two-exponent law using ordinary least squares (OLS) approach, which was earlier proposed by Hankin & Lee (2006). However, Hankin & Lee (2006) show in a simulation study that the OLS approach should be used only when the exponents a and b in (2) are equal. Since there is no theoretical reason to assume that, this paper uses rather the ML approach to estimate the parameters and their variances of the two-exponent law.⁴

In order to assess whether our samples of impact factors are consistent with a given statistical model, we follow Clauset et al. (2009) in using a goodness-of-fit test based on parametric bootstrap approach. In particular, we use the well-known Kolmogorov-Smirnov test defined by:

$$KS = \max_{x} \left| F(x) - F(x;\hat{\theta}) \right|, \tag{5}$$

where F(x) is the cdf of the data, $F(x; \hat{\theta})$ is the cdf of a tested model and $\hat{\theta}$ is the vector of the model's parameter estimates obtained, for example, using

⁴Notice also that Sarabia et al. (2012) estimate standard errors for parameter estimates of the two-exponent model by directly taking the standard errors from the OLS regression of logged sample order statistics (or observed quantiles) on their expected values. However, this approach does not take into account the covariance between order statistics and produces erroneous results. Hankin & Lee (2006) provide a correct method of calculating standard errors for parameters of (2) estimated using the OLS.

ML estimation. The distribution of the KS statistic is known for data sets drawn from a given model. However, when the underlying model is not known or when its parameters are estimated from the data, which is the case studied in this paper, the distribution of the KS must be obtained by simulation. The appropriate simulation is implemented in the following way. First, we fit a given model, $F(x; \theta)$, to our data set obtaining a vector of estimated parameters $\hat{\theta}$. Next, we calculate a KS statistic for the fitted model and denote it by KS_{org} . Third, we draw a large number, B, of synthetic samples of the original size from a fitted model $F(x; \hat{\theta})$. For each simulated sample, we fit the model $F(x; \theta)$ and calculate its KS statistic denoted by KS_b . Finally, the *p*-value for the test is obtained as the proportion of KS_b greater than KS_{org} . The hypothesis that our data set follows $F(x; \theta)$ is rejected if the *p*-value is smaller than the chosen threshold (set to 0.1 in this paper).

We also compare formally whether the two-exponent law gives a better fit to the impact factor data than the the Singh-Maddala and Dagum distributions. To this aim, we use the likelihood ratio test, which tests if the compared models are equally close to the true model against the alternative that one is closer. The test computes the logarithm of the ratio of the likelihoods of the data under two competing distributions, LR, which is negative or positive depending on which model fits data better. Vuong (1989) showed that in the case of non-nested models the normalized log-likelihood ratio $NLR = n^{-1/2}LR/\sigma$, where σ is the estimated standard deviation of LR, has a limit standard normal distribution. This result can be used to compute a p-value for the test discriminating between the competing models.

3. Data

We use data on impact factors from the latest available (2012) edition of Thompson Reuters Journal Citation Reports (JCR). Following Sarabia et al. (2012), we use impact factors for scientific journals belonging to the following scientific fields: Chemistry, Economics, Education, Information Science and Library Science (abbreviated further as Information SLS), Mathematics, Neurosciences, Psychology and Physics.⁵ We have removed journals with

⁵In some cases, we have grouped several JCR subject categories into one science category. For example, our category "Mathematics" consists of the following JCR subject categories: "Mathematical & Computational Biology", "Mathematics", "Mathematics, Applied" and "Mathematics, Interdisciplinary Applications".

zero impact factors from our samples. Descriptive statistics for our data sets are presented in Table 1.

Table 1: Descriptive statistics for the impact factors in eight scientific fields (N denotes the number of journals within a given field).

Field	N	Mean	Median	Std. Dev.	Gini
Chemistry	513	2.703	1.684	3.669	0.513
Economics	333	1.062	0.795	0.922	0.440
Education	254	0.868	0.679	0.686	0.404
Information SLS	84	1.001	0.755	0.928	0.465
Mathematics	573	0.950	0.717	0.769	0.379
Neurosciences	252	3.574	2.872	3.486	0.419
Psychology	556	1.766	1.361	1.852	0.435
Physics	381	2.535	1.400	4.514	0.568

Comparing our data with data used by Sarabia et al. (2012), we observe that the number of journals indexed in JCR has increased between JCR 2010 Edition and JCR 2012 edition for each science category analyzed; the increases lie in the range between 5% and 20%. Other descriptive statistics are close to those reported by Sarabia et al. (2012).

4. Empirical results and discussion

Table 2 presents results of fitting the two-exponent law to the data. Beside parameter estimates and their standard errors, we give also the value of log-likelihood and the *p*-value from the goodness-of-fit test described in Section 2.2.

Table 2: Maximum likelihood fits of the two-exponent law to data on impact factors in selected fields of science. Standard errors in parentheses.

Field	K	b	a	Log-likelihood	<i>p</i> -value
Chemistry	1.8972(0.1371)	0.6331(0.0459)	0.4997(0.0372)	-982.63	0.650
Economics	1.1403(0.0892)	0.7827(0.0618)	0.3197(0.0361)	-343.19	0.615
Education	0.8453(0.0791)	0.6155(0.0615)	0.3334(0.0448)	-198.02	0.013
Information SLS	1.0459(0.1832)	0.8308(0.1246)	0.3466(0.0841)	-82.99	0.006
Mathematics	0.6730(0.0377)	0.3334(0.0290)	0.4386(0.0337)	-427.29	0.105
Neurosciences	3.4131(0.2438)	0.6386(0.0558)	0.3213(0.0330)	-549.75	0.002
Psychology	1.5840(0.0845)	0.6096(0.0384)	0.3577(0.0278)	-819.86	0.314
Physics	1.2452(0.0827)	0.4722(0.0388)	0.5965(0.0444)	-657.62	0.008

Our estimates of the parameters are in general close to the OLS estimates obtained by Sarabia et al. (2012). However, our estimates of standard errors are roughly one order higher than those of Sarabia et al. (2012). As pointed out before, this is due to the fact that Sarabia et al. (2012) do not account in their approach for the dependence between order statistics. For this reason, their standard errors are severely underestimated.

The goodness-of-fit test used suggests that the two-exponent model is a plausible hypothesis for impact factors of journals in Chemistry, Economics, Mathematics and Psychology.

We can use our estimates to test for the Lavalette law, which holds that a = b. Using Wald test (see Sarabia et al., 2012), we cannot reject the hypothesis that a = b at the 5% level for Mathematics and Physics (with *p*-values equal to 0.063 and 0.075, respectively). Also, the significance level of the test is close to 5% for Chemistry (*p*-value = 0.048). For other science categories, the hypothesis tested is rejected even at the 1% level. Our results with respect to the Lavalette law are inconsistent with those of Sarabia et al. (2012), who found that the law was rejected for each of the science category studied. This inconsistency is due to the previously mentioned fact that Sarabia et al. (2012) substantially underestimated the variance of the parameter estimates of the two-exponent law.

Tables 3-4 presents our results of fitting the Singh-Maddala and Dagum models. In most of the cases, the parameters are estimated with sufficient precision. There are some exceptions, especially in case of fitting the Singh-Maddala distribution to impact factors for Economics, Education, and Information SLS. However, even in these cases all three parameters are jointly different from zero in a statistically significant way.

Field	a	b	q	Log-likelihood	<i>p</i> -value
Chemistry	1.6208(0.1030)	2.2040(0.3789)	1.3613(0.2477)	-982.60	0.759
Economics	1.3534(0.1004)	3.6112(2.1859)	5.5194(3.4980)	-341.64	0.916
Education	1.6679(0.1536)	1.4947(0.5821)	2.8249(1.2977)	-197.07	0.061
Information SLS	1.2895(0.1962)	3.1271(3.8256)	4.7948(5.7605)	-82.44	0.043
Mathematics	2.8287(0.2140)	0.6471(0.0602)	0.7957(0.1303)	-428.15	0.037
Neurosciences	1.7098(0.1354)	4.8304(1.1260)	2.1769(0.5982)	-552.07	0.001
Psychology	1.7595(0.0980)	2.0466(0.3225)	1.8321(0.3432)	-821.01	0.085
Physics	2.1123 (0.1623)	1.0978 (0.1274)	0.7382(0.1133)	-657.02	0.038

Table 3: Maximum likelihood fits of the Singh-Maddala model to data on impact factors in selected fields of science. Standard errors in parentheses.

The goodness-of-fit test suggests that both models are plausible for impact factors of journals in Chemistry and Economics. The Dagum model

Table 4: Maximum likelihood fits of the Dagum model to data on impact factors in selected fields of science. Standard errors in parentheses.

Field	a	b	p	Log-likelihood	<i>p</i> -value
Chemistry	1.9727(0.1454)	2.0778(0.2616)	0.7701(0.1191)	-982.87	0.528
Economics	2.7042(0.3041)	1.3517(0.1585)	0.4420(0.0835)	-344.56	0.648
Education	2.7038(0.3170)	0.9520(0.1276)	0.5702(0.1248)	-198.64	0.016
Information SLS	2.4773(0.7232)	1.2474(0.4136)	0.4605(0.2290)	-83.73	0.005
Mathematics	2.3283(0.1400)	0.5924(0.0704)	1.4299(0.2592)	-426.87	0.146
Neurosciences	3.0261(0.3013)	4.1831(0.3626)	0.4615(0.0741)	-547.63	0.014
Psychology	2.6482(0.1986)	1.8506(0.1547)	0.5701(0.0775)	-819.85	0.293
Physics	1.7539(0.1255)	1.1489(0.1890)	1.2473(0.2334)	-658.09	0.020

seems also to be consistent with data for Mathematics and Psychology. The Singh-Maddala model has a higher probability of being consistent with data for Education, Information SLS, and Physics than both the two-exponent law and the Dagum model. Neither model seems to be consistent with data for Neurosciences.

We can also test for the Lavalette law (Fisk distribution) using fitted Singh-Maddala and Dagum models. For the Singh-Maddala model this amounts to testing the hypothesis that q = 1, while for the Dagum model the relevant hypothesis is that p = 1. Using Wald tests, we have found that at the 5% level both the Singh-Maddala and Dagum models can be reduced to the Fisk distribution in case of Chemistry and Mathematics. Moreover, the Singh-Maddala model can be reduced to the Fisk distribution in case of Economics, Education, and Information SLS, while the Dagum model reduces to the Fisk model in case of Physics.⁶ These results suggest that in modeling the distribution of impact factors, the Lavalette law should be treated as a serious alternative to the three-parameter distributions, especially for formal and natural sciences.

The results of model selection tests based on the approach of Vuong (1989) are reported in Table 5. The compared models are empirically indistinguishable for impact factors of journals in Chemistry, Mathematics, Psychology and Physics. For Economics, Education, and Information SLS, the Singh-Maddala model is preferred to both the two-exponent model and the Dagum model. Finally, the Dagum distribution is preferred for data in

⁶However, these results for the Singh-Maddala model may be due to the previously discussed insufficient precision of the estimates for the parameter q.

Neurosciences.

Table 5: Model selection tests for impact factor distributions, selected science fields. "A \succ B" denotes that distribution A gives a better fit than distribution B. Positive values of LR indicate that the likelihood for the two-exponent model is higher than the likelihood for the alternative.

Science category	Two-exp.	vs. Dagum	Two-exp	p. vs. SM	Conclusion
	LR	p-value	LR	p-value	
Chemistry	0.237	0.483	-0.027	0.939	Distributions are indistinguishable
Economics	1.365	0.071	-1.554	0.094	Two-exp. \succ Dagum, SM \succ Two-exp.
Education	0.623	0.184	-0.949	0.058	Two-exp. \succ Dagum, SM \succ Two-exp.
Information SLS	0.746	0.022	-0.545	0.089	Two-exp. \succ Dagum, SM \succ Two-exp.
Mathematics	-0.416	0.336	0.864	0.117	Distributions are indistinguishable
Neurosciences	-2.115	0.005	2.322	0.020	Dagum \succ Two-exp., Two-exp. \succ SM
Psychology	-0.017	0.981	1.148	0.214	Distributions are indistinguishable
Physics	0.472	0.234	-0.599	0.137	Distributions are indistinguishable

The models can be also compared visually using quantile-quantile (q-q) plots. Figure 1 plots empirical quantiles versus theoretical quantiles for the two-exponent law and the better of the two alternatives (Singh-Maddala or Dagum) for these cases, when results from Table 5 suggest that the two-exponent law is worse fit to the data. If the estimated model fitted the data perfectly, the q-q plot would coincide with the 45-degree line. Figure 1 shows that the differences between compared models are in general small. However, it is clearly visible that the two-exponent law gives a slightly worse fit in these cases, especially for the highest quantiles.

Overall, both our formal tests and visual inspection from the q-q plots suggest that compared to the two-exponent law, the Singh-Maddala and Dagum models offer as good as or slightly better fit to data on impact factors.

5. Conclusions

The distribution of impact factors has been modeled in the recent informetric literature using the two-exponent law proposed by Mansilla et al. (2007). This paper shows that two distributions widely-used in economics, namely the Singh-Maddala and Dagum models, possess several advantages over the two-exponent model. They are either empirically indistinguishable from the two-exponent law or give slightly better fit to data on impact factors in eight important scientific fields. Contrary to Sarabia et al. (2012), the present paper also found that the Lavalette law (Lavalette, 1996), which

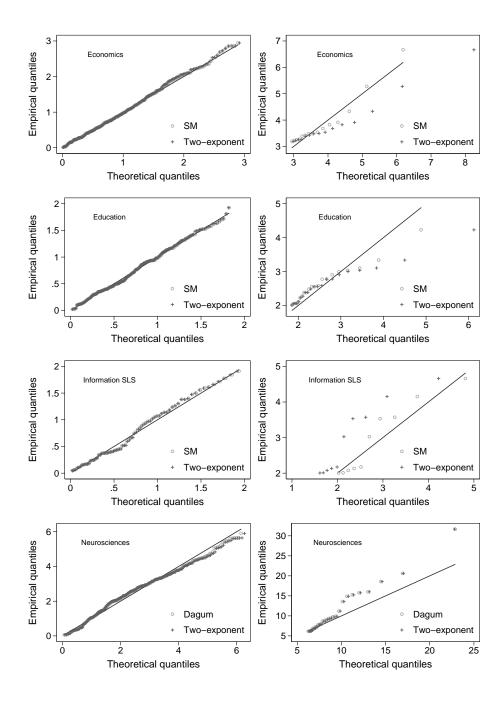


Figure 1: Quantile-quantile plots for the two-exponent law vs. the better of the Singh-Maddala and Dagum models. Left panels show bottom and middle quantiles, right panels show upper quantiles.

is a special case of both the two-exponent law and the Singh-Maddala and Dagum models, often cannot be rejected as an appropriate distribution for impact factors, especially for formal and natural sciences.

In contrast to the two-exponent model, both models proposed in this paper have closed-from probability density functions and cumulative distribution functions. This advantage allows one to estimate model parameters in a more straightforward way without the need for numerical estimation of the model's basic functions. There exists also a well-developed literature (see, e.g., Kleiber & Kotz, 2003; Kleiber, 2008), that explores various theoretical properties of the Singh-Maddala and Dagum distributions, including ready-made expressions for several inequality measures as well as conditions required for Lorenz ordering or stochastic dominance of various degrees. These properties may be exploited in modeling the distribution of impact factors and other informetric variables. For these reasons, this paper suggests that the Singh-Maddala and Dagum models are more useful for empirical modeling of the distribution of journal impact factors than the two-exponent law.

References

- Bandourian, R., McDonald, J. B., & Turley, R. S. (2003). A comparison of parametric models of income distribution across countries and over time. *Estadística*, 55, 135–152.
- Burr, I. W. (1942). Cumulative frequency functions. The Annals of Mathematical Statistics, 13, 215–232.
- Campanario, J. M. (2010). Distribution of changes in impact factors over time. Scientometrics, 84, 35–42.
- Clauset, A., Shalizi, C. R., & Newman, M. E. (2009). Power-law distributions in empirical data. SIAM review, 51, 661–703.
- Dagum, C. (1977). A new model of personal income distribution: Specification and estimation. *Economie Appliquée*, 30, 413–437.
- Dastrup, S. R., Hartshorn, R., & McDonald, J. B. (2007). The impact of taxes and transfer payments on the distribution of income: A parametric comparison. *The Journal of Economic Inequality*, 5, 353–369.

- Egghe, L. (2005). Power laws in the information production process: Lotkaian informetrics. Oxford: Elsevier.
- Egghe, L. (2009). Mathematical derivation of the impact factor distribution. Journal of Informetrics, 3, 290–295.
- Egghe, L. (2011). The impact factor rank-order distribution revisited. *Scientometrics*, 87, 683–685.
- Egghe, L., & Waltman, L. (2011). Relations between the shape of a sizefrequency distribution and the shape of a rank-frequency distribution. *In*formation Processing & Management, 47, 238–245.
- Hankin, R. K., & Lee, A. (2006). A new family of non-negative distributions. Australian & New Zealand Journal of Statistics, 48, 67–78.
- Jenkins, S. P. (2009). Distributionally-sensitive inequality indices and the GB2 income distribution. *Review of Income and Wealth*, 55, 392–398.
- Kleiber, C. (1996). Dagum vs. Singh-Maddala income distributions. *Economics Letters*, 53, 265–268.
- Kleiber, C. (2008). A Guide to the Dagum Distribution. In D. Chotikapanich (Ed.), Modelling Income Distributions and Lorenz Curves: Essays in Memory of Camilo Dagum (pp. 97–268). Berlin, New York: Springer.
- Kleiber, C., & Kotz, S. (2003). Statistical size distributions in economics and actuarial sciences. Hoboken, NJ: John Wiley.
- Lavalette, D. (1996). Facteur dimpact: impartialité ou impuissance. Report, INSERM U350, 91405, Orsay (France): Institut CurieRecherche, Bât. 112.
- Mansilla, R., Köppen, E., Cocho, G., & Miramontes, P. (2007). On the behavior of journal impact factor rank-order distribution. *Journal of Informetrics*, 1, 155–160.
- McDonald, J. B. (1984). Some generalized functions for the size distribution of income. *Econometrica*, 52, 647–663.
- Mishra, S. K. (2010). A note on empirical sample distribution of journal impact factors in major discipline groups. Available at: http://works.bepress.com/sk_mishra/8.

- Sarabia, J. M., Prieto, F., & Trueba, C. (2012). Modeling the probabilistic distribution of the impact factor. *Journal of Informetrics*, 6, 66–79.
- Singh, S., & Maddala, G. (1976). A function for size distribution of incomes. *Econometrica*, 44, 963–970.
- Vuong, Q. H. (1989). Likelihood ratio tests for model selection and nonnested hypotheses. *Econometrica*, 57, 307–333.
- Waltman, L., & Van Eck, N. J. (2009). Some comments on Egghe's derivation of the impact factor distribution. *Journal of Informetrics*, *3*, 363–366.
- Wilfling, B., & Krämer, W. (1993). The Lorenz-ordering of Singh-Maddala income distributions. *Economics Letters*, 43, 53–57.