Absent-minded passengers

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Abstract

Passengers board a fully booked airplane in order. The first passenger picks one of the seats at random. Each subsequent passenger takes his or her assigned seat if available, otherwise takes one of the remaining seats at random. It is well known that the last passenger obtains her own seat with probability 1/2. We study the distribution of the number of incorrectly seated passengers, and we also discuss the case of several absent-minded passengers.

Keywords: Absent-minded passengers, combinatorial probability, Stirling numbers of the first kind, random permutation

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1 Introduction.

The following problem is posed and solved in [3, p. 177], and [10, p. 35]. An airplane with $n \ge 2$ passengers is fully booked. Passengers are boarding in chronological order, according to the numbers on their boarding passes. The first passenger loses his boarding pass and picks one of the seats at random. Each subsequent passenger takes his or her seat if available, otherwise takes one of the remaining seats at random. The following intuitive argument shows that the last passenger will sit in her own seat number n with probability 1/2. The first absent-minded passenger chooses either his assigned seat number 1, or seat number n, or one of the seats numbered $2, \ldots, n-1$. In the first case, the last passenger will sit in her seat, and in the second case she has to sit down in seat number 1. In each of the remaining cases, the role of the absent-minded passenger will be taken over by the first passenger who finds his seat occupied. In the end, everything boils down to a toss of a fair coin that decides whether one of the passenger will sit in his assigned seat with probability 2/3. Likewise, the probability is 3/4 for the antepenultimate passenger to sit in his assigned seat, etc. In this article we study the distribution of the number of incorrectly seated passengers in the more general case where the first k passengers are absent-minded.

We find it both useful and informative to first focus on the case k = 1. Even though the results are then basically known (see [3, 8, 9, 10]), we take a partly new approach based on probabilistic arguments. As long as the first seat is not occupied, there is, at each stage of the boarding procedure, exactly one incorrectly seated passenger in the remaining seats. The (suitably defined) conditional distribution of the number of this seat is uniform; see Lemma 2.2. This fact easily implies a certain independence property (Proposition 2.3) and then in turn an explicit formula for the distribution of the number of incorrectly seated passengers and a central limit theorem. In Section 3 we extend these arguments to the case where the first k passengers are absent-minded. While there is still an independence property (Theorem 3.2), the full distribution of the number of incorrectly seated passengers seems to be difficult to obtain. Still we have been able to derive explicit formulae for mean and variance along with a central limit theorem. Except for the mean, these results are new.

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2 One absent-minded passenger.

First we describe our model in a neutral, more abstract language. Consider n balls numbered $1, \ldots, n$, to be placed in chronological order in a box with n numbered pits. The first ball is placed at random. The second ball is placed in pit 2, provided this pit is empty. Otherwise it is placed at random in one of the remaining n - 1 pits. This procedure continues until the *n*th ball is placed.

In this section, we derive the distribution of the random number W_n (say) of balls that do not meet their assigned pit. Moreover, we show that the limit distribution of W_n , when suitably normalized, is standard normal as $n \to \infty$. Most parts of this article are based on elementary probability theory as can be found in numerous textbooks (see, e.g., [4] or [6]).

In what follows, let X_1 denote the number of the pit taken by the first ball. By definition, X_1 is uniformly distributed on $\{1, \ldots, n\}$. For each $i \in \{2, \ldots, n-1\}$, consider the state of the box after the placement of the *i*th ball. If pit 1 is occupied, then each subsequently placed ball matches the number of its assigned pit. In this case we define $X_i := 1$. If pit 1 is empty, then each of the pits $2, \ldots, i$ is occupied, as well as exactly one further pit having number j (say), where j > i. In this case we put $X_i := j$. Finally, we define $X_n := 1$.

To illustrate the notation, we consider the case n = 9. If the first ball takes pit 4, the fourth ball takes pit 8, and ball 8 takes pit 1, then the balls numbered 2,3,5,6,7, and 9 find their assigned pits, and we have $X_1 = X_2 = X_3 = 4$, $X_4 = X_5 = X_6 = X_7 = 8$, and $X_8 = X_9 = 1$.

If we denote by N_j the number of the pit taken by the *j*th ball, where j = 1, ..., n, then $(N_1, ..., N_n)$ is a random permutation of (1, ..., n), and we have $X_i = 1$ if $\min(N_1, ..., N_i) = 1$ and $X_i = \max(N_1, ..., N_i)$ otherwise. In other words, $X_i = 1$ if and only if pit 1 is occupied after the placement of the *i*th ball. Otherwise, X_i is the largest number of the occupied pits at that stage. Thus, for example, the permutation (2,1,3,4) of (1,2,3,4) arises if the first ball takes pit 2 and the second ball (that finds its pit occupied) chooses pit 1. Clearly, balls 3 and 4 then find their assigned pits.

Table 1 shows the feasible permutations (N_1, \ldots, N_n) for the case n = 4, together with their respective probabilities. For each permutation, the last column displays the number of "incorrectly placed balls," i.e., balls that do not meet their assigned pit.

N_1	N_2	N_3	N_4	prob	W_4
1	2	3	4	1/4	0
2	1	3	4	1/12	2
2	3	1	4	1/24	3
2	3	4	1	1/24	4
2	4	3	1	1/12	3
3	2	1	4	1/8	2
3	2	4	1	1/8	3
4	2	3	1	1/4	2

Table 1: The feasible permutations (N_1, N_2, N_3, N_4) and their probabilities.

The model has the following simple but crucial symmetry property.

Lemma 2.1. Let $i \in \{1, \ldots, n\}$. Then X_i is uniformly distributed on $\{1, i+1, \ldots, n\}$.

Proof. We use induction on i. For i = 1 and i = n the assertion is true. Assume that it holds for some $i \in \{2, ..., n-1\}$ and let $j \in \{1, i+2, ..., n\}$. Conditioning on X_i and noting that $X_{i+1} = j$ implies either $X_i = j$ or $X_i = i + 1$, we obtain

$$\mathbb{P}(X_{i+1} = j) = \mathbb{P}(X_i = j)\mathbb{P}(X_{i+1} = j|X_i = j) \\ + \mathbb{P}(X_i = i+1)\mathbb{P}(X_{i+1} = j|X_i = i+1) \\ = \frac{1}{n-i+1} \cdot 1 + \frac{1}{n-i+1} \cdot \frac{1}{n-i} \\ = \frac{1}{n-i},$$

and the proof is finished.

For $i \in \{1, ..., n\}$, let A_i denote the event that the *i*th ball is placed in its assigned pit. Since $A_i = \{X_{i-1} \neq i\}$ for $i \ge 2$, Lemma 2.1 yields

$$\mathbb{P}(A_i) = \frac{n-i+1}{n-i+2}, \quad i = 2, \dots, n;$$
(1)

see also [9]. Thus, in particular, $\mathbb{P}(A_n) = 1/2$, which is the solution to the initial problem given in [3] and [10].

Incidentally, it follows by induction that there are 2^{n-1} feasible permutations. In fact, the last ball can only meet either its assigned pit or pit 1, and each of these cases gives rise to the same number of feasible permutations (N_1, \ldots, N_n) . Indeed, if $N_n = n$, then the admissible assignments of balls $1, \ldots, n-1$ to pits $1, \ldots, n-1$ coincide with the feasible permutations in the case of n-1pits. On the other hand, if $N_n = 1$, then we are in the situation of n-1 balls and n-1 pits numbered $n, 2, \ldots, n-1$. Here, pit n is assigned to the first ball to be placed, and this ball is "absent-minded"; see Table 1 for the case n = 4. Hence, the case $N_n = 1$ also gives rise to 2^{n-1} feasible permutations.

Writing $\mathbf{1}_A$ for the indicator function of an event A, let

$$C_n := \sum_{i=1}^n \mathbf{1}_{A_i}$$

denote the number of correctly placed balls. Since $\mathbb{P}(A_1) = 1/n$, (1) gives

$$\mathbb{E}(C_n) = \frac{1}{n} + \sum_{i=2}^n \left(1 - \frac{1}{n-i+2}\right) = \frac{1}{n} + \sum_{j=0}^{n-2} \left(1 - \frac{1}{j+2}\right),$$

so that

$$\mathbb{E}(C_n) = n - H_{n-1},\tag{2}$$

where $H_m := 1 + 1/2 + \cdots + 1/m$, $m \in \mathbb{N}$, is the *m*th harmonic number.

Below we will show that the events A_2, \ldots, A_n are independent, a result that might come as a surprise. To do so we need the following lemma.

Lemma 2.2. Suppose that $r \in \{1, \ldots, n-2\}$ and $2 \leq i_1 < \cdots < i_r \leq n-1$. Let $i \in \{i_r+1, \ldots, n\}$. Then, under the condition $A_{i_1} \cap \cdots \cap A_{i_r}$, the random variable X_i has a uniform distribution on $\{1, i+1, \ldots, n\}$.

Proof. For ease of notation we prove the result in the case r = 1, the general case being analogous. Let $m := i_1$. We need to show that $\mathbb{P}(X_i = j | A_m) = 1/(n - i + 1)$ for each $i \in \{m + 1, \dots, n\}$ and each $j \in \{1, i + 1, \dots, n\}$. We proceed as in the proof of Lemma 2.1 and use induction on i.

Let $j \in \{1, m + 2, ..., n\}$. By the definition of our model and Lemma 2.1,

$$\mathbb{P}(X_{m+1} = j, X_{m-1} \neq m) = \mathbb{P}(X_{m-1} = j) + \mathbb{P}(X_{m-1} = m+1, X_{m+1} = j)$$
$$= \frac{1}{n-m+2} + \mathbb{P}(X_{m-1} = m+1) \cdot \frac{1}{n-m}$$
$$= \frac{1}{n-m+2} \cdot \frac{n-m+1}{n-m}.$$

Using Lemma 2.1 again, we hence obtain

$$\mathbb{P}(X_{m+1} = j | A_m) = \frac{\mathbb{P}(X_{m+1} = j, X_{m-1} \neq m)}{\mathbb{P}(X_{m-1} \neq m)} = \frac{1}{n-m}$$

Assume now the assertion is true for each $i \in \{m+1, \ldots, n-1\}$ and take some $j \in \{1, i+2, \ldots, n\}$. Then

$$\mathbb{P}(X_{i+1} = j | A_m) = \mathbb{P}(X_i = j | A_m) + \mathbb{P}(X_{i+1} = j, X_i = i+1 | A_m)$$

= $\frac{1}{n-i+1} + \mathbb{P}(X_{i+1} = j | \{X_i = i+1\} \cap A_m) \mathbb{P}(X_i = i | A_m)$
= $\frac{1}{n-i+1} + \frac{1}{n-i} \cdot \frac{1}{n-i+1} = \frac{1}{n-i}.$

This finishes the induction and hence the proof.

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The following independence property can be found (with a different proof) as Theorem 3.5 in [7].

Theorem 2.3. The events A_2, \ldots, A_n are independent.

Proof. Let $r \in \{1, ..., n-1\}$ and $2 \le i_1 < \cdots < i_r \le n-1$. By Lemma 2.2,

$$\mathbb{P}(A_{i_1} \cap \dots \cap A_{i_r}) = \mathbb{P}(A_{i_r} | A_{i_1} \cap \dots \cap A_{i_{r-1}}) \mathbb{P}(A_{i_1} \cap \dots \cap A_{i_{r-1}})$$
$$= \frac{n - i_r + 1}{n - i_r + 2} \cdot \mathbb{P}(A_{i_1} \cap \dots \cap A_{i_{r-1}}).$$

Repeating this reasoning r-1 times and using Lemma 2.1 in the final step, we obtain

$$\mathbb{P}(A_{i_1} \cap \dots \cap A_{i_r}) = \frac{n - i_r + 1}{n - i_r + 2} \cdots \frac{n - i_1 + 1}{n - i_1 + 2},$$

as required.

Writing A^c for the complement of an event A, let

$$W_n := n - C_n = \sum_{i=1}^n \mathbf{1}_{A_i^c}$$

denote the number of balls that do not meet their assigned pit. From Table 1, the probability distribution of W_4 is given by

$$\mathbb{P}(W_4 = 0) = \frac{6}{24}, \ \mathbb{P}(W_4 = 2) = \frac{11}{24}, \ \mathbb{P}(W_4 = 3) = \frac{6}{24}, \ \mathbb{P}(W_4 = 4) = \frac{1}{24}.$$

Here, the numbers $6 = {4 \choose 1}$, $11 = {4 \choose 2}$, $6 = {4 \choose 3}$, and $1 = {4 \choose 4}$ figuring in the respective numerators are *Stirling numbers of the first kind*. For general $n \in \mathbb{N}$ and $j \in \{1, \ldots, n\}$, the Stirling number ${n \choose j}$ is the number of permutations of $\{1, \ldots, n\}$ having exactly j cycles; see, e.g., [5, Section 6.1]. Notice that ${n \choose 1} = (n-1)!$. The next result shows that the occurrence of the Stirling numbers in the special case n = 4 is not a lucky coincidence; see [7, Theorem 3.6].

Theorem 2.4. The distribution of W_n is given by

$$\mathbb{P}(W_n = 0) = \frac{\binom{n}{1}}{n!}, \quad \mathbb{P}(W_n = j) = \frac{\binom{n}{j}}{n!}, \quad j = 2, \dots, n.$$

Proof. For $i \in \{2, \ldots, n\}$ define $B_i := A_{n-i+2}$. Then

$$W_n = R_n - \mathbf{1}_{A_1},\tag{3}$$

where $R_n := 1 + \sum_{i=2}^n \mathbf{1}_{B_i}$. By Theorem 2.3, the events B_2, \ldots, B_n are independent, and we have $\mathbb{P}(B_i) = 1/i, i \in \{2, \ldots, n\}$. We thus have, for each $n \ge 1$ and each $j \in \{2, \ldots, n\}$,

$$\mathbb{P}(R_{n+1} = j) = \mathbb{P}(\{R_{n+1} = j\} \cap B_{n+1}^c) + \mathbb{P}(\{R_{n+1} = j\} \cap B_{n+1})$$
$$= \frac{n}{n+1} \mathbb{P}(R_n = j) + \frac{1}{n+1} \mathbb{P}(R_n = j-1).$$
(4)

Hence, the generating function $f_n(z) := \mathbb{E}(z^{R_n}), z \in [0, 1]$, of R_n satisfies the recursion

$$f_{n+1}(z) = \frac{n+z}{n+1} f_n(z).$$

Since $f_1(z) = z$, it follows that

$$f_n(z) = \frac{z(z+1)\cdots(z+n-1)}{n!}, \quad z \in [0,1].$$

Up to the factor 1/n! this expression is the generating function of the Stirling numbers of the first kind (see [5, p. 252]), so that

$$\mathbb{P}(R_n = j) = \frac{\binom{n}{j}}{n!}, \quad j = 1, \dots, n.$$
(5)

Since, by (3), $R_n = W_n$ on the event A_1^c and $W_n = 0$ on the event A_1 (since $A_1 = A_1 \cap A_2 \cap \cdots \cap A_n$), the result follows.

Formula (5) gives the distribution of the number of cycles in a purely random permutation of $\{1, 2, \ldots, n\}$ (see, e.g., [2, Section 6.2]), as well as the distribution of the number of records in a sequence of independent and identically distributed continuous random variables (see, e.g., [2, Section 9.5]). In both cases, the reason is a recursion of the type (4).

Figure 1 shows a bar chart of the distribution of W_{100} . The most probable value (= 0.2112) for the number of incorrectly seated passengers is attained at $\ell = 5$. Since $|W_n - R_n| \leq 1$, the Lindeberg–Feller central limit theorem and Sluzki's lemma show that the limit distribution of $(W_n - \log n)/\sqrt{\log n}$ as $n \to \infty$ is standard normal; see, e.g., [1, p. 383], and, in particular, Example 27.3.

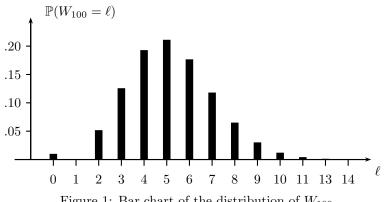


Figure 1: Bar chart of the distribution of W_{100} .

3 Several absent-minded passengers.

Generalizing the problem discussed in [3], [9], and [10], we now consider the case when the first k passengers are absent-minded, where $k \in \{1, \ldots, n-1\}$. If these passengers take their seats completely at random and the subsequent passengers board the airplane according to the rule stated in the introduction, what is the probability that the last passenger finds his or her seat available? We will see that the answer is 1/(k+1). Moreover, we will compute the expectation and variance of and a central limit theorem for the number of incorrectly seated passengers.

In a more neutral formulation, suppose again that we have n balls to be allocated to n pits in a box. Let $k \in \{1, \ldots, n-1\}$, and assume that the first k balls are distributed at random, each subset of size k of all n pits having the same probability of being chosen. The (k + 1)st ball takes pit k + 1 provided that pit is available. Otherwise it chooses its position at random. This procedure continues until the last ball is placed.

Given $i \in \{k, \ldots, n\}$, we define a random subset Z_i of $\{1, \ldots, k\} \cup \{i+1, \ldots, n\}$ as follows. Consider the state of the box after the placement of the *i*th ball. If i = k, let Z_i be the set of pits occupied by the "absent-minded balls." If i > k, then all pits from $\{k + 1, \ldots, i\}$ are occupied, and we write Z_i for the set of all *other* occupied pits. Then Z_i has k elements, i.e., we have $|Z_i| = k$.

To illustrate the new notation, consider the case n = 10, k = 3 and the placement

Thus, ball 1 takes pit 7, ball 2 chooses pit 3, ball 3 meets pit 9, etc. In this case, we have $Z_3 = Z_4 = Z_5 = Z_6 = \{3, 7, 9\}, Z_7 = Z_8 = \{1, 3, 9\}, Z_9 = \{1, 3, 10\}, \text{ and } Z_{10} = \{1, 2, 3\}.$

Let \mathcal{P}_i denote the system of all sets $A \subset \{1, \ldots, k\} \cup \{i+1, \ldots, n\}$ with |A| = k. The next result generalizes Lemma 2.1.

Lemma 3.1. Let $i \in \{k, \ldots, n\}$. Then the distribution of Z_i is uniform on \mathcal{P}_i .

Proof. We proceed by induction on *i*. For i = k the assertion holds. Suppose it is true for some $i \in \{k, \ldots, n-1\}$ and let $A \in \mathcal{P}_{i+1}$. Since $\{Z_i = A\} \subset \{Z_{i+1} = A\}$ and $Z_i \neq Z_{i+1}$ if and only if $i+1 \in Z_i$, we have

$$\mathbb{P}(Z_{i+1} = A) = \mathbb{P}(Z_{i+1} = A = Z_i) + \mathbb{P}(Z_{i+1} = A, Z_i \neq A)$$

= $\mathbb{P}(Z_i = A) + \sum_{j \in A} \mathbb{P}(Z_i = (A \setminus \{j\}) \cup \{i+1\}) \cdot \frac{1}{n-i}$
= $\binom{n-i+k}{k}^{-1} \binom{1+\frac{k}{n-i}}{l} = \binom{n-i+k}{k}^{-1} \frac{n-i+k}{n-i}$
= $\binom{n-(i+1)+k}{k}^{-1}$.

This finishes the induction and hence the proof.

For $i \in \{1, \ldots, n\}$ again let A_i denote the event that the *i*th ball is placed in its pit. Clearly,

$$\mathbb{P}(A_i) = \frac{1}{n}, \quad i = 1, \dots, k.$$
(6)

Since $A_i = \{i \notin Z_{i-1}\}$ for $i \ge k+1$, Lemma 3.1 entails

$$\mathbb{P}(A_i) = \frac{n-i+1}{n-i+k+1}, \quad i = k+1, \dots, n,$$
(7)

and in particular $\mathbb{P}(A_n) = 1/(k+1)$.

We note in passing that there are $k!(k+1)^{n-k}$ feasible permutations; see [7, Theorem 3.9]. The proof uses the so-called canonical cycle representation of a permutation. An alternative argument can be based on induction, similarly to the case k = 1 discussed in the paragraph after equation (1).

A conceptual proof of Lemma 3.1 uses the fact that, after the placement of the *i*th ball, where i > k, each of the places $k + 1, \ldots, i$ is occupied. Therefore, the *other* k occupied places are among the n - i + k places numbered $1, \ldots, k$ and $i + 1, \ldots, n$. By symmetry, each choice of k (occupied) places from these n - i + k places has the same probability. In particular, the probability that the (i + 1)st ball finds its assigned pit empty equals (n - i)/(n - i + k).

Writing $W_{n,k} = \sum_{i=1}^{n} \mathbf{1}_{A_i^c}$ for the number of balls that do not meet their assigned pit, we now have

$$\mathbb{E}(W_{n,k}) = \sum_{i=1}^{n} \mathbb{P}(A_i^c) = k\left(1 - \frac{1}{n}\right) + \sum_{i=k+1}^{n} \frac{k}{n - i + k + 1}$$
$$= k\left(1 + H_{n-1} - H_k\right).$$

The following result can be proved with the help of an analog of Lemma 2.2. We leave this to the reader.

Theorem 3.2. The events A_{k+1}, \ldots, A_n are independent.

We now derive a formula for the variance of $W_{n,k}$. Since $W_{n,k} + C_{n,k} = n$, where $C_{n,k} = \sum_{i=1}^{n} \mathbf{1}_{A_i}$ is the number of balls that meet their assigned pit, we have $\mathbb{V}(W_{n,k}) = \mathbb{V}(C_{n,k})$. Now, $C_{n,k}$ is a sum of indicator random variables, whence

$$\mathbb{V}(C_{n,k}) = \sum_{i=1}^{n} \mathbb{P}(A_i) \left(1 - \mathbb{P}(A_i)\right) + 2 \sum_{1 \le i < j \le n} \left(\mathbb{P}(A_i \cap A_j) - \mathbb{P}(A_i)\mathbb{P}(A_j)\right).$$
(8)

Using (6) and (7), the single sum equals

1

$$\frac{k(n-1)}{n^2} + k(H_n - H_k) - k^2 \sum_{\ell=k+1}^n \frac{1}{\ell^2}.$$

In view of Theorem 3.2, only pairs (i, j) satisfying either $1 \leq i < j \leq k$ or $1 \leq i \leq k < j \leq n$ make a nonzero contribution to the double sum figuring in (8). Since $\mathbb{P}(A_i \cap A_j) = 1/(n(n-1))$ if $1 \leq i < j \leq k$, it follows that

$$\sum_{\leq i < j \leq k} \left(\mathbb{P}(A_i \cap A_j) - \mathbb{P}(A_i)\mathbb{P}(A_j) \right) = \binom{k}{2} \frac{1}{n^2(n-1)}.$$

If $i \in \{1, ..., k\}$ and $j \in \{k+1, ..., n\}$, we write $\mathbb{P}(A_i \cap A_j) = \mathbb{P}(A_i)\mathbb{P}(A_j|A_i)$. Under the condition A_i , the situation is that of a box containing n-1 pits numbered 1, ..., i-1, i+1, ..., n, and the balls numbered 1, ..., i-1, i+1, ..., k are distributed at random. By relabelling each ball j, where j > i, with j - 1, we can use formula (7) with n, k, and i replaced with n - 1, k - 1, and j - 1, respectively, and obtain

$$\mathbb{P}(A_i \cap A_j) = \frac{1}{n} \cdot \frac{n-j+1}{n-j+k}, \quad j = k+1, \dots, n.$$

Consequently,

$$\sum_{i=1}^{k} \sum_{j=k+1}^{n} (\mathbb{P}(A_i \cap A_j) - \mathbb{P}(A_i)\mathbb{P}(A_j)) = \sum_{i=1}^{k} \sum_{j=k+1}^{n} \frac{1}{n} \left(\frac{n-j+1}{n-j+k} - \frac{n-j+1}{n-j+k+1}\right),$$

and some algebra shows that this expression equals $\frac{k}{n}(H_{n-1} - H_{k-1} - 1 + k/n)$. Putting everything together, straightforward calculations give

$$\mathbb{V}(W_{n,k}) = k \left[\frac{2(1-n+kn)-n^2-k}{n^2(n-1)} + \frac{2}{nk} + \left(1+\frac{2}{n}\right)(H_n - H_k) - k \sum_{\ell=k+1}^n \frac{1}{\ell^2} \right]$$

We have not been able to find a closed-form expression for the distribution of $W_{n,k}$ if $k \geq 2$. The asymptotic distribution of $W_{n,k}$ as $n \to \infty$, however, is available. To this end, writing $a_n \sim b_n$ if $a_n/b_n \to 1$ as $n \to \infty$, and using $H_n \sim \log n$, it follows that $\mathbb{E}(W_{n,k}) \sim k \log n$ and $\mathbb{V}(W_{n,k}) \sim k \log n$. By the Lindeberg–Feller central limit theorem, the random variable $\sum_{i=k+1}^{n} \mathbf{1}_{A_i^c}$, after standardization, has a standard normal limit as $n \to \infty$ (see [1, p. 383]. Since $|W_{n,k} - \sum_{i=k+1}^{n} \mathbf{1}_{A_i^c}| \leq k$, Sluzki's lemma shows that the limit distribution of $(W_{n,k} - k \log n)/\sqrt{k \log n}$ as $n \to \infty$ is standard normal.

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References

- [1] Billingsley, P. (2012). Probability and Measure, Anniversary Edition. New York, NY: Wiley.
- [2] Blom, G., Holst, L., Sandell, D. (1994). Snapshots from the World of Probability. New York, NY: Springer.
- [3] Bolobás, B. (2006). The Art of Mathematics: Coffee Time in Memphis. Cambridge, UK: Cambridge Univ. Press.
- [4] Chung, K. L., AitSahlia, F. (2003). Elementary Probability Theory. With Stochastic Processes and an Introduction to Mathematical Finance. 4th ed. Undergraduate Texts in Mathematics. New York, NY: Springer.

- [5] Graham, R. L., Knuth, D. E., Patashnik, O. (1994). Concrete Mathematics. 2nd ed. Reading, MA: Addison-Wesley.
- [6] Grimmet, G., Welsh, D. (2014). Probability Theory: An Introduction. 2nd ed. Oxford, UK: Oxford Univ. Press.
- [7] Lengyel, T. (2010). The distribution of the size of the union of cycles of two types of random permutations. Int. J. Comb., Article ID 751861, 10 pages. doi.org/10.1155/2010/751861.
- [8] Misra, N. (2008). The missing boarding pass. Resonance 13(7): 662-679.
- [9] Nigussie, Y. (2014). Finding your seat versus tossing a coin. Amer. Math. Monthly 121(6): 545-546.
- [10] Winkler, P. (2004). Mathematical Puzzles: A Connoisseur's Collection. Natick, MA: A.K. Peters.