# Bounds on the Zero-Error List-Decoding Capacity of the $q /(q-1)$ Channel 

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#### Abstract

We consider the problem of determining the zero-error list-decoding capacity of the $q /(q-1)$ channel studied by Elias (1988). The $q /(q-1)$ channel has input and output alphabet consisting of $q$ symbols, say, $\mathcal{X}=\left\{x_{1}, x_{2}, \ldots, x_{q}\right\}$; when the channel receives an input $x \in \mathcal{X}$, it outputs a symbol other than $x$ itself. Let $n(m, q, \ell)$ be the smallest $n$ for which there is a code $C \subseteq \mathcal{X}^{n}$ of $m$ elements such that for every list $w_{1}, w_{2}, \ldots, w_{\ell+1}$ of distinct code-words from $C$, there is a coordinate $j \in[n]$ that satisfies $\left\{w_{1}[j], w_{2}[j], \ldots, w_{\ell+1}[j]\right\}=\mathcal{X}$. We show that for $\epsilon<1 / 6$, for all large $q$ and large enough $m, n(m, q, \epsilon q \ln q) \geq$ $\Omega\left(\exp \left(q^{1-6 \epsilon} / 8\right) \log _{2} m\right)$.


The lower bound obtained by Fredman and Komlós (1984) for perfect hashing implies that $n(m, q, q-1)=\exp (\Omega(q)) \log _{2} m$; similarly, the lower bound obtained by Körner (1986) for nearlyperfect hashing implies that $n(m, q, q)=\exp (\Omega(q)) \log _{2} m$. These results show that the zero-error list-decoding capacity of the $q /(q-1)$ channel with lists of size at most $q$ is exponentially small. Extending these bounds, Chakraborty et al. (2006) showed that the capacity remains exponentially small even if the list size is allowed to be as large as $1.58 q$. Our result implies that the zero-error list-decoding capacity of the $q /(q-1)$ channel with list size $\epsilon q$ for $\epsilon<1 / 6$ is $\exp \left(\Omega\left(q^{1-6 \epsilon}\right)\right)$. This resolves the conjecture raised by Chakraborty et al. (2006) about the zero-error listdecoding capcity of the $q /(q-1)$ channel at larger list sizes.

## I. Introduction

We study the zero-error-list-decoding capacity of the $q /(q-1)$ channel. The input and output alphabet of this channel are a set of $q$ symbols, namely $\mathcal{X}=\left\{x_{1}, x_{2}, \ldots, x_{q}\right\}$; when the symbol $x \in \mathcal{X}$ is input, the output symbol can be anything other than $x$ itself. We wish to design good error correcting codes for such a channel. For the $q /(q-1)$ channel it is impossible to recover the message without error if the code has at least two code-words: in fact, no matter how many letters are used for encoding, for every set of up to ( $q-1$ ) input code-words, one can construct an output word that is compatible with all of them. It is, however, possible to design codes where on receiving an output word from the channel, one can narrow down the input message to a set of size at most $(q-1)$-that is, we can list-decode with lists of size $(q-1)$. Such codes have rate exponentially small in $q$.

Definition 1.1 (Code, Rate). A code $C \subseteq\left\{x_{1}, \ldots, x_{q}\right\}^{n}$ is an $\ell$-list-decoding-code for the $q /(q-1)$ channel, if for every output word $\sigma^{\prime} \in \mathcal{X}^{n}$, we have $\mid\left\{\sigma \in \mathcal{X}^{n}\right.$ : the input word $\sigma$ is compatible with $\left.\sigma^{\prime}\right\} \mid \leq \ell$. Let $n(m, q, \ell)$ be the smallest $n$ such that there exists an $\ell$-list-decoding code for the $q /(q-1)$ channel with $m$ code-words. The zero-error-list-of- $\ell$-rate of $C,|C|=m$, is given by $\frac{1}{n} \log _{2}(m / \ell)$, and the list-of- $\ell$-capacity of the $q /(q-1)$ channel, denoted by $\operatorname{cap}(q, \ell)$, is the least upper bound on the attainable zero-error-list-of- $\ell$-rate across all $\ell$-list-decoding-codes.

The list-of-2-capacity of the $3 / 2$ channel was studied by Elias [1], who showed that $0.08 \approx \log _{2}(3)-1.5 \leq \operatorname{cap}(3,2) \leq$ $\log _{2}(3)-1 \approx 0.58$. For the $4 / 3$ channel, Dalai, Guruswami and Radhakrishnan [2] showed that $\operatorname{cap}(4,3) \leq 6 / 19 \approx$ 0.3158 , improving slightly on an earlier upper bound of 0.3512 shown by Arikan [3]; it was shown by Körner and Marton [4] that $\operatorname{cap}(4,3) \geq(1 / 3) \log _{2}(32 / 29) \approx 0.0473$. For general $q$, one can obtain the following upper bound using a routine probabilistic argument.
Proposition I.1. $n(m, q, q-1)=\exp (O(q)) \lg m$.
This implies that the $\operatorname{cap}(q, q-1)=\exp (-O(q))$. So for each fixed $q$ we do have codes with positive rate, but the rate promised by this construction goes to zero exponentially with $q$. Fredman and Komlós [5] showed that this exponential deterioration is inevitable; Körner showed that $\operatorname{cap}(q, q)=\exp (-\Omega(q))$. On the other hand, it can be shown that $\operatorname{cap}(q,\lceil q \ln q\rceil)=1 / q$, and that for all functions $\ell: \mathbb{Z} \rightarrow \mathbb{Z}$ we have $\operatorname{cap}(q, \ell(q)) \geq 1 / q$. Thus, the list-of- $\ell$ capacity of the $q /(q-1)$ channel cannot be better than $1 / q$ unless $\ell$ is allowed to grow with $m$.

We thus have the following situation. The list-of- $\ell$-rate of any code reaches the optimal value of $1 / q$ when the listsize is about $q \ln q$; however, the list-of- $(q-1)$ (as well as list-of- $q$ ) rate is exponentially small in $q$. It is interesting, therefore, to study the trade-off between the list size and the rate, and determine how the rate changes from inverse polynomial in $q$ to exponentially small in $q$. Chakraborty, Radhakrishnan, Raghunathan and Sasatte [6] addressed this question and showed the following.
Theorem I.2. For every $\epsilon>0$, there is $a \delta>0$ such that
for all large $q$ and large enough $m$, we have $n(m, q,(\eta-$ $\epsilon) q) \geq \exp (\delta q) \log _{2} m$, where $\eta=e /(e-1) \approx 1.58$. Thus, $\operatorname{cap}(q,(\eta-\epsilon) q)=\exp (-\Omega(q))$.

We show the following.
Theorem I. 3 (Result). For every $\epsilon<1 / 6$, for all large $q$ and large enough $m$, we have $n(m, q, \epsilon q \ln q) \geq \Omega\left(\exp \left(q^{1-6 \epsilon} / 8\right) \log _{2} m\right)$. Thus, for all $\epsilon<1 / 6, \operatorname{cap}(q, \epsilon q \ln q)=\exp \left(-\Omega\left(q^{1-6 \epsilon}\right)\right)$.

This establishes both parts of the conjecture of Chakraborty et al. which states the following.
Conjecture I.1. (a) For all constants $c>0$, there is a constant $\alpha$, such that for all large $m$, we have $n(m, q, c q) \geq$ $\exp (\alpha q) \log _{2} m$.
(b) For all functions $\ell(q)=o\left(q \log _{2} q\right)$ and all large $m$, we have $n(m, q, \ell(q)) \geq q^{\omega(1)} \log _{2} m$.

## A. Overview of our approach

We extend the approach of Chakraborty et al., which in turn was based on the approach used by Fredman and Komlós [5] to obtain lower bounds on the size of families of perfect hash functions. To describe our adaptation of this approach, it will be convenient to reformulate the problem using matrix terminology.

Consider $C \subseteq \mathcal{X}^{n}$ with $m$ code-words. We can build an $m \times n$ matrix $C=\left(c_{i j}: i=1, \ldots, m\right.$ and $\left.j=1, \ldots, n\right)$ (we use the name $C$ both for the code and the associated matrix) by writing the code-words as rows of the matrix (the order does not matter): so $c_{i j}=k$ iff the $j$-th component of the $i$-th codeword is $x_{k} \in \mathcal{X}$. Then, $C$ is an $\ell$-list-decoding code iff the matrix has the following property: for every choice $R$ of $\ell+1$ rows, there is a column $h$ such that $\left\{c_{r h}: r \in R\right\}=\mathcal{X}$. In this reformulation, $n(m, q, \ell)$ is the minimum $n$ so that there exists a matrix with this property. We refer to such a matrix as an $\ell$-list-decoding matrix. Furthermore, instead of writing $c_{r h}$ we write $h(r)$; indeed, in the setting of hash families (originally considered by Fredman and Komlós), the columns correspond to hash functions that assign a symbol in $\mathcal{X}$ to each row-index in $[m]$.

We can now describe the approach of Chakraborty et al. Fix a list-size $\ell=\alpha q$. Suppose there is an $\ell$-list decoding matrix $C$ with $n=\exp (\beta q) \log _{2} m$ columns. We wish to show that if $\beta$ is small then the matrix cannot have the required property; that is, we can find a set $R$ of $\ell+1$ rows for which $h(R)$ is a proper subset of $[q]$ for every column $h$. To exhibit such a set $R$ we will proceed in stages. In the first stage, we pick a subset $R_{1}$ of $q-2$ rows at random. Consider a column $h$. What can we expect? We expect to see a good number of collisions, where the same symbol appears in column $h$ at two different rows in $R_{1}$. In fact, we expect $h(R)$ to contain only about $q(1-1 / e)$ elements. By appealing to standard results (e.g., McDiarmid's inequality), we may conclude that with probability exponentially close to 1 (that is, of the form $1-\exp (-\gamma q)), h(R)$ is unlikely to have significantly more
elements. So we might settle on a choice of $R$, so that $h(R)$ deviates significantly (say by $\epsilon q$ for some small $\epsilon$ ) for at most $\exp (-\gamma q) \exp (\beta q) \log _{2} m$ columns. If the original $\beta$ is chosen to be much smaller than $\gamma$, this number is an exponentially small fraction of $\log _{2} m$.
The key idea now is to make these exceptional columns ineffective. We do this by focusing our attention on a reduced number of rows. For each exceptional column, we pick the symbol that appears most often in that column, and restrict attention to those rows that have this symbol in the exceptional column. This depletes the number of rows by a factor at $1 / q$ for each exceptional column; after we do this sequentially for all the $\exp (-(\gamma-\beta) q) \log _{2} m \ll \log _{2} m$ rows, we will be left with $m^{\prime}$ rows, where $\log _{2} m^{\prime}=\Omega\left(\log _{2} m\right)$. We may now add more rows to our existing list $R_{1}$. If we choose these from the set of $m^{\prime}$ rows, we are in no danger from the exceptional columns; in the other columns $R_{1}$ spans about $q(1-1 / e)$ symbols, so we can add to $R_{1}$ about $q / e$ rows $R_{2}$ (picked from the $m^{\prime}$-rows) and still ensure that in no column $h$, we are in danger of $h\left(R_{1} \cup R_{2}\right)$ becoming $\mathcal{X}$. It is clear that we can carry this approach further, e.g., by picking $R_{2}$ randomly, expecting a significant number of internal collisions, making the exceptional columns ineffective, focusing attention on a smaller but still significant number of rows, etc., then picking $R_{3}$ from the rows that survive, and so on. In fact, Chakraborty et al. derived Theorem I. 2 using precisely this approach.
In this paper, we follow the approach outlined above but implement the idea more precisely. Before we describe our contribution it will be useful to pin-point where the calculations in Chakraborty et al. were sub-optimal. We argued above that after $R_{1}$ is picked, we expect to span only about $q(1-1 / e)$ symbols in a given column $h$. What about after $R_{2}$ is picked? $R_{1} \cup R_{2}$ contains a total of $q+q / e$ rows: if all symbols in column $h$ appeared with the same frequency (and continued to do so in the $m^{\prime}$ rows after the exceptional columns were eliminated), then we should expect $h\left(R_{1} \cup R_{2}\right)$ to span about $(q+q / e)(1-\exp (1+1 / e))$ symbols. Notice that this is roughly the expected number of distinct coupons collected in the classical coupon collector problem after $q+q / e$ attempts. Unfortunately, there are technical difficulties that arise in claiming that this number will be reflected in our process because (i) $R_{1}$ and $R_{2}$ are not picked independently, and (ii) even if the symbols appeared with the same frequency initially, they may not do so after we focus on a depleted set of rows. Faced with these difficulties, Chakraborty et al. settled for less. Instead of matching the bound suggested by the coupon collector problem, when analysing the expected size of $h\left(R_{1} \cup R_{2}\right)$, they estimated $h\left(R_{2}\right)$ separately and bounded $\left|h\left(R_{1} \cup R_{2}\right)\right|$ by $\left|h\left(R_{1}\right)\right|+\left|h\left(R_{2}\right)\right|$, thereby ignoring $h\left(R_{1} \cap R_{2}\right)$. The loss in precision resulting from the use of this union bound increases as the number of phases increases. Indeed, when the coupon collector process is carried in phases by picking sets $R_{1}, R_{2}, \ldots, R_{t}$ for a large $t$, progress in collecting coupons is retarded more by collisions across sets
(because for some $i \neq j, h\left(R_{i}\right)$ and $h\left(R_{j}\right)$ have elements in common) than by collisions within some $h\left(R_{i}\right)$. By neglecting collisions across phases, and by failing to track the coupon collector process closely, the argument in Chakraborty et al. were unable to push the list size in Theorem I. 2 beyond $e /(e-1)$.

What is new?: We attempt to track the progress of the coupon collector faithfully. Instead of the set $R_{1}$ of size $q-2$ that was picked earlier, we pick an ensemble (a collection of sets) $\mathcal{R}^{1}$ of sets of size $q-2$. Similarly, in the later steps we will pick ensembles $\mathcal{R}^{2}, \mathcal{R}^{3}, \ldots$. However, in the end we pick one set $R_{i}$ from each of the ensembles $\mathcal{R}^{i}$ respectively, and assemble our list of rows: $R_{1} \cup R_{2} \cup \cdots \cup R_{t}$. That this process is more effective in bounding $\left|h\left(R_{1} \cup R_{2} \cup \ldots \cup R_{t}\right)\right|$ will be formally verified in later sections. For now, let us qualitatively see how it helps in bounding $\left|h\left(R_{1} \cup R_{2}\right)\right|$. We pick $\mathcal{R}^{1}$ at random: if the number of sets in the ensemble is large enough (we will set it to be $\exp (\Theta(q))$ ), then it should reflect a random set of rows that was obtained by picking rows independently $(q-2)$-times from the set of all rows. Fix a choice for $R_{2}$, the set to be picked at the second stage. Consider $\mathbf{X}=\mid h\left(R_{1} \cup\right.$ $\left.R_{2}\right) \mid$ where $R_{1}$ is picked uniformly from the ensemble $\mathcal{R}^{1}$; let $\mathbf{Y}=\left|h\left(\mathbf{R}_{1} \cup R_{2}\right)\right|$, where $\mathbf{R}_{1}$ is picked uniformly from the set of all rows. Then, we expect $X$ and $Y$ to have similar distribution. So, we proceed as follows. We pick an ensemble $\mathcal{R}^{1}$ at random. If for a certain column $h$, the ensemble $\mathcal{R}^{1}$ fails to deliver a good sample, we will need to make that column ineffective as before. Further, if some set in $\mathcal{R}^{1}$ spans a significantly larger number of symbols in some column, we will again make that column ineffective. After this, we pick $R_{2}$ from the remaining rows. We expect it to not only have a good number of internal collisions but also be such that $\mid h\left(\mathbf{R}_{1} \cup\right.$ $\left.R_{2}\right) \mid$ and $\left|h\left(\mathbf{R}_{1} \cup \mathbf{R}_{2}\right)\right|$ (where the set $\mathbf{R}_{2}$ is chosen uniformly from the available rows) are similar in expectation. Now, since we ensured that the ensemble $\mathcal{R}^{1}$ was good for column $h$, a random choice of $R_{1}$ from the ensemble will deliver a value of $\left|h\left(R_{1} \cup R_{2}\right)\right|$ that, with high probability, can be bounded by the number of distinct coupons picked up at the same stage by the coupon collector; in particular, it accounts for symbols common to $h\left(R_{1}\right)$ and $h\left(R_{2}\right)$. The outline above illustrates the advantages of picking an ensemble instead of committing to just one randomly chosen set. However, a large ensemble comes with its drawbacks. We need to ensure that no set in the ensemble spans too many elements in any column, or rather, we need to eliminate any column where some set spans many elements. This forces a more drastic reduction in the number of rows than before (that is, now $m^{\prime}$ when compared with $m$ is much smaller than in the calculation in [6]). Thus, it is important to keep the sizes of the ensembles small. The tradeoff between these opposing concerns needs to be handled with some care. The argument is presented in detail below.

## II. Proof of the Result

In what follows we assume that $q$ is a large natural number and $m \rightarrow \infty$.

We will need the following concentration result due to McDiarmid (1989).
Lemma II. 1 (McDiarmid). Let $X_{1}, X_{2}, \ldots, X_{n}$ be independent random variables where each $X_{k}$ takes values in a finite set $A$. Let $f: A^{n} \rightarrow \mathbb{R}$ be such that $|f(x)-f(y)| \leq c$ whenever $x$ and $y$ differ in only one coordinate. Let $Y=$ $f\left(X_{1}, X_{2}, \ldots, X_{n}\right)$; then, for all $t>0$,

$$
\operatorname{Pr}[\mathbb{E}[Y]-Y \geq t], \operatorname{Pr}[Y-\mathbb{E}[Y] \geq t] \leq \exp \left(\frac{-2 t^{2}}{n c^{2}}\right)
$$

Let $C$ be an $\ell$-list-decoding-code for the $q /(q-1)$ channel with $\ell<q \ln q / 6$. As mentioned in the introduction, we will view $C$ as an $m \times n$ matrix with entries from $[q]$. In other words, the rows are indexed by code-words and the columns are indexed by hash functions. Let wt be a function from $[q]$ to $\{0,1\}$; for $A \subseteq[q]$, let $\mathrm{wt}(A):=\sum_{a \in A} \mathrm{wt}(a)$. Let $\mathbf{R}$ be a random variable taking values in $\mathcal{P}([m])$. Sometimes we use $\mathbf{R}$ to also refer to the distribution of this random variable.

Following the idea mentioned in the introduction, we intend to keep an ensemble $\mathcal{R}$ of sets of rows such that when we pick a new set of rows $R_{2}$ from a depleted number of rows $m^{\prime}$, we not only observe the correct number of internal collisions within $R_{2}$ but also observe the correct number of collision between members of $\mathcal{R}$ and $R_{2}$. This motivates the following definition.
Definition II. 1 (Sampler). We say that an ensemble $\mathcal{R}=$ $\left(R_{1}, R_{2}, \ldots, R_{L}\right)$, where each $R_{i} \subseteq[m]$, is a $(\gamma, \delta)$ sampler for $\mathbf{R}$ wrt column $h$ if $\left(A_{1}, A_{2}, \ldots, A_{L}\right)$ := $\left(h\left(R_{1}\right), h\left(R_{2}\right), \ldots, h\left(R_{L}\right)\right)$ satisfies $\forall \mathrm{wt}:[q] \rightarrow\{0,1\}$

$$
\operatorname{Pr}_{j \in \mathbf{u}[L]}\left[\left|\mathrm{wt}\left(A_{j}\right)-\mathbb{E}[\mathrm{wt}(h(\mathbf{R}))]\right| \geq \gamma q\right] \leq \exp (-\delta q)
$$

The definition makes provision for all functions wt, because it tries to anticipate the appropriate internal collisions (see Lemma (II.2) with very little advance knowledge of what the distribution on $[q]$ looks like in column $h$ after a large number of rows have been discarded.

Let $\pi: S \rightarrow[0,1]$ be a probability mass function on a finite set $S$. Let $k \geq 1$, and let $X_{1}, X_{2}, \ldots, X_{k}$ be independent random variables each distributed according to $\pi$. Then, let $\pi^{\{k\}}$ denote the probability mass function of the set $\left\{X_{1}, X_{2}, \ldots, X_{k}\right\}$.
For distributions $\mathbf{A}$ and $\mathbf{B}$ on $\mathcal{P}([m])$, let $\mathbf{A} \vee \mathbf{B}$ be the distribution of $S \cup T$ where $S \sim \mathbf{A}$ and $T \sim \mathbf{B}$, with $S$ and $T$ chosen independently. The following lemma will be the main workhorse for our argument.
Lemma II. 2 (Ensemble Composition Lemma). Let $\mathbf{R}$ be a distribution on $\mathcal{P}([m])$ and let $D$ be a distribution on $[m]$. Let $\mathcal{R}$ be a $(\gamma, \delta)$-sampler for $\mathbf{R}$ wrt a column $h$; let
$\left(R_{1}, R_{2}, \ldots, R_{s}\right)$ be obtained by taking s independent samples from the ensemble $\mathcal{R}$. Similarly, let $\mathcal{R}^{\prime}=\left(R_{1}^{\prime}, R_{2}^{\prime}, \ldots, R_{s}^{\prime}\right)$ be obtained by taking s independent samples according to $\mathbf{R}^{\prime} \sim$ $D^{\{t q\}}$ where $t<1$. Let $\gamma^{\prime}, \delta^{\prime}>0$ be such that $\delta \leq 2\left(\gamma^{\prime}\right)^{2} / t$ and $\delta>\delta^{\prime}$ Let $s=\exp \left(\delta-\delta^{\prime}\right) q, \widetilde{\gamma}=\gamma+\gamma^{\prime}, \widetilde{\delta}=\delta-\delta^{\prime}$. Then, with probability $1-12 \exp \left(-\delta^{\prime} q\right)$ over the random choices, the composed ensemble

$$
\widetilde{\mathcal{R}}:=\left(R_{1} \cup R_{1}^{\prime}, R_{2} \cup R_{2}^{\prime}, \ldots, R_{s} \cup R_{s}^{\prime}\right)
$$

of cardinality s, is a $(\widetilde{\gamma}, \widetilde{\delta})$-sampler for $\mathbf{R} \vee \mathbf{R}^{\prime}$ wrt the column $h$, and furthermore $\forall i \in[s]$,

$$
\begin{equation*}
\left|\left|h\left(R_{i} \cup R_{i}^{\prime}\right)\right|-\mathbb{E}\left[\left|h\left(\mathbf{R} \vee \mathbf{R}^{\prime}\right)\right|\right]\right| \leq \widetilde{\gamma} q \tag{1}
\end{equation*}
$$

Note that this ensemble is generated according to the product distribution $\left(\mathcal{R} \vee \mathbf{R}^{\prime}\right)^{s}$.

Proof. Fix $f:[q] \rightarrow\{0,1\}$ and let $\mu_{f}:=\mathbb{E}\left[f\left(h\left(\mathbf{R} \cup \mathbf{R}^{\prime}\right)\right)\right]$; similarly, for $R^{\prime} \subseteq[m]$ let $\mu_{f}\left(R^{\prime}\right):=\mathbb{E}_{\mathbf{R}}\left[f\left(h\left(\mathbf{R} \cup R^{\prime}\right)\right)\right]$. First, we bound the probability that when $\mathcal{R}^{\prime}$ is chosen according to $\mathbf{R}^{\prime}$, it fails to have $\mu_{f}\left(R^{\prime}\right)$ close to $\mu_{f}$. Using McDiarmid's inequality over the $t q$ primitive choices for $R^{\prime}$, we have

$$
\begin{align*}
\operatorname{Pr}_{R^{\prime} \sim \mathbf{R}^{\prime}}\left[\left|\mu_{f}\left(R^{\prime}\right)-\mu_{f}\right| \geq \gamma^{\prime} q\right] & \leq 2 \exp \left(\frac{-2\left(\gamma^{\prime}\right)^{2} q^{2}}{t q}\right) \\
& =2 \exp \left(\frac{-2\left(\gamma^{\prime}\right)^{2} q}{t}\right) \tag{2}
\end{align*}
$$

Now, let wt : $[q] \rightarrow 0,1$ be defined by $\operatorname{wt}(x)=f(x)$ if $x \notin h\left(R^{\prime}\right)$ and $\mathrm{wt}(x)=0$ otherwise. Then, for $R \subseteq[m]$, we have, $f\left(h\left(R \cup R^{\prime}\right)\right)=f\left(h\left(R^{\prime}\right)\right)+\mathrm{wt}(h(R))$. Therefore (note here $R^{\prime}$ is fixed and $R$ varies randomly in $\mathcal{R}$ ),

$$
\begin{aligned}
& \operatorname{Pr}_{R \in \mathbf{u} \mathcal{R}}\left[\left|f\left(h\left(R \cup R^{\prime}\right)\right)-\mu_{f}\left(R^{\prime}\right)\right| \geq \gamma q\right] \\
& \quad=\operatorname{Pr}_{R \in \mathbf{u} \mathcal{R}}[|\operatorname{wt}(h(R))-\mathbb{E}[\operatorname{wt}(h(\mathbf{R}))]| \geq \gamma q]
\end{aligned}
$$

and since $\mathcal{R}$ is a $(\gamma, \delta)$-sampler wrt $\mathbf{R}$, we have

$$
\operatorname{Pr}_{R \in \mathfrak{u} \mathcal{R}}[|\operatorname{wt}(h(R))-\mathbb{E}[\operatorname{wt}(h(\mathbf{R}))]| \geq \gamma q] \leq \exp (-\delta q)
$$

Thus,

$$
\begin{align*}
& \operatorname{Pr}_{R \in \mathbf{u} \mathcal{R}, R^{\prime} \sim \mathbf{R}^{\prime}}\left[\left|f\left(h\left(R \cup R^{\prime}\right)\right)-\mu_{f}\right| \geq\left(\gamma+\gamma^{\prime}\right) q\right] \leq \\
& R \in \operatorname{Pr}_{\mathbf{u}} \operatorname{Pr}_{R^{\prime} \sim \mathbf{R}^{\prime}}\left[\left|\mu_{f}-\mu_{f}\left(R^{\prime}\right)\right| \geq \gamma^{\prime} q\right]+ \\
& \quad \operatorname{Pr}_{R \in \mathbf{u}} \operatorname{Pr}_{R^{\prime} \sim \mathbf{R}^{\prime}}\left[\left|\mu_{f}\left(R^{\prime}\right)-f\left(h\left(R \cup R^{\prime}\right)\right)\right| \geq \gamma q\right] \\
& \leq\left[2 \exp \left(\frac{-2\left(\gamma^{\prime}\right)^{2} q}{t}\right)+\exp (-\delta q)\right] \leq 3 \exp (-\delta q) . \tag{3}
\end{align*}
$$

(We used $\delta \leq 2\left(\gamma^{\prime}\right)^{2} / t$ to justify the last inequality.) Let $\Delta:=$ $3 \exp (-\delta q)$, the quantity on the right in (3). By taking $f$ to be the all-1's function, we conclude from (3) that for each $i$ with probability at least $1-\Delta,\left|\left|h\left(R_{i} \cup R_{i}^{\prime}\right)\right|-\mathbb{E}\left[\left|h\left(\mathbf{R} \vee \mathbf{R}^{\prime}\right)\right|\right]\right| \leq$ $\left(\gamma+\gamma^{\prime}\right) q$.

Now, $f\left(h\left(R_{i} \cup R_{i}^{\prime}\right)\right)=\left|h\left(R_{i}\right) \cup h\left(R_{i}^{\prime}\right)\right|$, and $\mu_{f}=$ $\mathbb{E}\left[\left|h\left(\mathbf{R} \vee \mathbf{R}^{\prime}\right)\right|\right]$. Now, by a union bound over the $s$ choices for $i$, we obtain

$$
\begin{gather*}
\operatorname{Pr}_{\tilde{\mathcal{R}}}\left[\exists i \in[s],\left|\left|h\left(R_{i} \cup R_{i}^{\prime}\right)\right|-\mathbb{E}\left[\left|h\left(\mathbf{R} \vee \mathbf{R}^{\prime}\right)\right|\right]\right| \geq\left(\gamma+\gamma^{\prime}\right) q\right] \\
\leq \Delta s \leq 3 \exp \left(-\delta^{\prime} q\right) \tag{4}
\end{gather*}
$$

This establishes (1).
It remains to establish our first claim that whp the ensemble picked according to $\left(\mathcal{R} \vee \mathbf{R}^{\prime}\right)^{s}$ is a $(\widetilde{\gamma}, \widetilde{\delta})$-sampler for $\mathbf{R} \vee \mathbf{R}^{\prime}$. Fix $f:[q] \rightarrow\{0,1\}$. Now, (3) implies that for each $i \in$ [s], the probability that $\left|f\left(h\left(R_{i} \cup R_{i}^{\prime}\right)\right)-\mu_{f}\right| \geq\left(\gamma+\gamma^{\prime}\right) q$ is exponentially small in $q$. Then, the tail probabilities for $\mathbb{Y}:=\sum_{i=1}^{s} \mathbb{I}\left[\left|f\left(h\left(R_{i} \cup R_{i}^{\prime}\right)\right)-\mu_{f}\right| \geq\left(\gamma+\gamma^{\prime}\right) q\right]$ can be bounded by considering $\operatorname{Bin}(s, \Delta)$. Therefore,

$$
\begin{align*}
& \underset{\widetilde{\mathcal{R}}}{\operatorname{Pr}}[\mathbb{Y}>\exp (-\widetilde{\delta} q) s] \\
& \leq\binom{ s}{\exp (-\widetilde{\delta} q) s}(\Delta)^{\exp (-\widetilde{\delta} q) s} \\
& \leq(e \exp (\widetilde{\delta} q) \Delta)^{\exp (-\widetilde{\delta} q) s} \\
& \leq 9 \exp \left(-\delta^{\prime} q\right) . \tag{5}
\end{align*}
$$

(We need to take a union bound against the $2^{q}$ possible functions $f:[q] \rightarrow\{0,1\}$ : by changing $s$ to $q s$ we may easily establsih this.) By (4) and (5), the probability that our ensemble fails to be a $(\widetilde{\gamma}, \widetilde{\delta})$-sampler, with $\widetilde{\gamma}=\gamma+\gamma^{\prime}$ and $\widetilde{\delta}=\delta-\delta^{\prime}$, or fails to satisfy (1) is at most $12 \exp \left(-\delta^{\prime} q\right)$.

Let us recall the template of our argument. At any stage we will have an ensemble of sets of rows, say $\mathcal{R}$, and a universe $U \subseteq[m]$ to choose sets of rows from to add to $\mathcal{R}$. We will add a specific number of randomly chosen sets of rows of a particular size from $U$ and then declare those columns bad where the modified $\mathcal{R}$ deviates from its expected behaviour. Consider a set $R \in \mathcal{R}$ : we want to say that the couponcollector process at $|R|$ probes into $[q]$ is the gold standard for good behaviour, i.e., no set in $\mathcal{R}$ will have expansion more than the coupon-collector at the same stage. The expected number of elements that the coupon-collector process picks up after $a$ i.i.d. uniform probes into $[q]$ is approximately $q(1-\exp (-a / q))$ : we will denote this as $\mu_{q}^{c c}(a)$. So, we need the following lemma, which is proved in the appendix.
Lemma II. 3 (Phased Coupon Collector). Let $a_{1}, a_{2}, \ldots, a_{k}$ be positive integers; let $a=a_{1}+a_{2}+\cdots+a_{k}$, and let $\pi_{1}, \pi_{2}, \ldots, \pi_{k}$ be probability mass functions. Let $\mathbf{A}_{1}, \mathbf{A}_{2}, \ldots, \mathbf{A}_{k}$ be independent random variables taking values in $\mathcal{P}([q])$, where $\mathbf{A}_{i} \sim \pi_{i}^{\left\{a_{i}\right\}}$. Suppose $a \leq \epsilon q \ln q$ and $k \leq e q^{\epsilon}$ for some $\epsilon<1 / 3$, then,

$$
\begin{array}{r}
\mathbb{E}\left[\left|\mathbf{A}_{1} \cup \mathbf{A}_{2} \cup \cdots \cup \mathbf{A}_{k}\right|\right] \leq q(1-\exp (-a / q))+o\left(q^{1-\epsilon}\right) \\
=\mu_{q}^{c c}(a)+o\left(q^{1-\epsilon}\right)
\end{array}
$$

Our next target is to understand the number of iterations we wish to perform, i.e., the number of times we need to enlarge
the sizes of the sets surviving the ensemble $\mathcal{R}$ so that the list size hits the target of $\epsilon q \ln q$, where $\epsilon<1 / 6$. At the first stage we will pick up sets of rows of size about $\ell_{1}=q$, and expect the image size to be close to $\mu_{q}^{c c}\left(\ell_{1}\right)$; we then prune out the exceptional columns. In the next stage, we pick sets of size about $\ell_{2}=q-\mu_{q}^{c c}\left(\ell_{1}\right)$ and expect the combined image size to be close to $\mu_{q}^{c c}\left(\ell_{1}+\ell_{2}\right)$. Hence, in the third iteration we pick sets of size close to $\ell_{3}=q-\mu_{q}^{c c}\left(\ell_{1}+\ell_{2}\right)$, and so on for the subsequent iterations. We are interested in the list size after $k$ iterations, i.e, $\ell_{\leq k}:=\sum_{i=1}^{k} \ell_{i}$. We have the following proposition, which is proved in the appendix.
Proposition II.4. Let $\ell_{1}=q$, and for $i \geq 1$ let $\ell_{i+1}=q-$ $\mu_{q}^{c c}\left(\sum_{j=1}^{i} \ell_{j}\right)$. Suppose $k=e q^{\epsilon}$ for some $\epsilon<1$, then, $\ell_{\leq k} \geq$ $\epsilon q \ln q$.

Proof. (The series $\left\{\ell_{\leq k}\right\}$ tends to $q \ln q$.)
Finally, we need a lemma where we glue all the steps mentioned in the introduction. At each iteration $k$, we maintain an ensemble $\mathcal{R}^{k}$ satisfying the requisite properties.

We call a distribution $\mathbf{D}$ on $\mathcal{P}([m])$ a $\left(g_{1}, \ldots, g_{k}\right)$-phased coupon collector distribution if $\mathbf{D}=D_{1}^{\left\{g_{1}\right\}} \vee D_{2}^{\left\{g_{2}\right\}} \ldots \vee D_{k}^{\left\{g_{k}\right\}}$ where each $D_{i}$ is a probability mass function on $[m]$. The following lemma tracks how the parameters change with each iteration.
Lemma II. 5 (Iteration Lemma). Let $k \leq q^{\epsilon}$ for some $\epsilon<1 / 5$. Let $\gamma=\gamma^{\prime}=q^{-2 \epsilon} / 2$ and $\delta^{\prime}=q^{-5 \epsilon} / 4$. Assume $n \leq \exp \left(\delta^{\prime} q\right) \log _{2} m /\left(48 \cdot q^{\epsilon} \log _{2} q\right)$. Then, there exists a partition $\mathcal{H}_{1}(k) \sqcup \mathcal{H}_{2}(k)$ of the columns of $C$, a universe of rows $U_{k} \subseteq[m]$, an ensemble $\mathcal{R}^{k}=\left(R_{1}, R_{2}, \ldots R_{L_{k}}\right)$, integers $\left(g_{1}, \ldots, g_{k}\right)$ and a $\left(g_{1}, \ldots, g_{k}\right)$-phased coupon collector distribution $\mathbf{D}_{k}$ such that:
a $g_{1}=q-2$, and $g_{i+1}=q-\mu_{q}^{c c}\left(g_{i}\right)-(i+1) \gamma q-2$
$\mathrm{b} \forall i \in\left[L_{k}\right],\left|R_{i}\right|=g_{\leq k} \geq \ell_{\leq k}-2 k-k^{2} \gamma q / 2$
c $\forall h \in \mathcal{H}_{2}(k), \forall i \in\left[L_{k}\right],\left|h\left(R_{i} \cup U_{k}\right)\right| \leq q-1$
$\mathrm{d} \forall h \in \mathcal{H}_{1}(k), \mathcal{R}^{k}$ is a $\left((k+1) \gamma, \gamma^{2}-k \delta^{\prime}\right)$-sampler for $\mathbf{D}_{k}$ wrt h
$\mathrm{e} \forall h \in \mathcal{H}_{1}(k), \forall i \in\left[L_{k}\right]| | h\left(R_{i}\right)\left|-\mathbb{E}\left[h\left(\mathbf{D}_{k}\right)\right]\right| \leq(k+$ 1) $\gamma q$
f $\log _{2}\left|U_{k}\right| \geq \log _{2} m-k \log _{2} q \cdot 24 \exp \left(-\delta^{\prime} q\right) n$.

Proof. We will use induction on $k$. For $k=1$ we have $g_{1}=$ $q-2$. We use Lemma $\boxed{I I} 2$ with $\mathbf{R}$ being the constant $\emptyset$, and $\mathcal{R}=\{\emptyset\}$. Clearly, $\mathcal{R}$ is a $\left(\gamma, \gamma^{2}\right)$-sampler for $\mathbf{R}$. Let $D$ be the uniform distribution over $[\mathrm{m}]$ and let $\mathcal{R}^{\prime}=\left(R_{1}^{\prime}, R_{2}^{\prime}, \ldots, R_{s}^{\prime}\right)$ be obtained by taking $s=\exp \left(\left(\gamma^{2}-\delta^{\prime}\right) q\right)$ independent samples according to $\mathbf{R}^{\prime} \sim D^{\{q-2\}}$. So, $\mathbf{D}_{1}=D^{\{q-2\}}$. For a fixed column $h$ we have the following: with probability $1-12 \exp \left(-\delta^{\prime} q\right)$ over the random choices, the composed ensemble

$$
\widetilde{\mathcal{R}}=\left(R_{1}^{\prime}, R_{2}^{\prime}, \ldots, R_{s}^{\prime}\right)
$$

is good wrt $h$, i.e., $\widetilde{\mathcal{R}}$ is a $\left(2 \gamma, \gamma^{2}-\delta^{\prime}\right)$-sampler for $\mathbf{R}^{\prime}$ wrt the column $h$, and furthermore $\forall i \in[s]$,

$$
\left|\left|h\left(R_{i}^{\prime}\right)\right|-\mathbb{E}\left[\left|h\left(\mathbf{R}^{\prime}\right)\right|\right]\right| \leq 2 \gamma q
$$

Hence, on expectation only $12 \exp \left(-\delta^{\prime} q\right) n$ columns are bad. Therefore, with probability at least $1 / 2$ at most $24 \exp \left(-\delta^{\prime} q\right) n$ columns are bad. Also, the probability of an $R_{i}^{\prime} \in \mathcal{R}^{\prime}$ having size less than $q-2$ (because some two of our $q-2$ choices of rows picked the same row) is at most $q^{2} / m$. Thus, by the union bound the probability of (b) not holding is at most $\underset{\widetilde{R}}{ } \cdot q^{2} / m$ which is less than $1 / 2$. Therefore, there is choice of $\widetilde{\mathcal{R}}$, which we call $\mathcal{R}^{1}$, such that at most $24 \exp \left(-\delta^{\prime} q\right) n$ columns are bad and (b) holds. The set of bad columns is $\mathcal{H}_{2}(1)$ and the set of good columns is $\mathcal{H}_{1}(1)$. Then, clearly (d) and (e) are true.
Let $\mathcal{H}_{2}(1)=\left\{h_{1}, \ldots, h_{b}\right\}$ where $b \leq 24 \exp \left(-\delta^{\prime} q\right) n$ and WLOG assume that 1 is the most frequent symbol in $h_{1}$. Retain only those rows in $U$ that correspond to the symbol 1 in $h_{1}$. Call this pruned universe $U^{\prime}$ : we have ensured that so long as we add rows to $R_{i} \in \mathcal{R}^{1}$ only from $U^{\prime}$, the image size in $h_{1}$ is at most $h_{1}(R)+1 \leq q-1$. Thus, by taking a multiplicative hit of at most $1 / q$ we have rendered $h_{1}$ ineffective. Iterating this over $\mathcal{H}_{2}(1)$ we take a multiplicative hit of $\left(\frac{1}{q}\right)^{b}$. Hence, we obtain a universe $U^{\prime}$, which will be $U_{1}$, such that $\log _{2}\left|U^{\prime}\right|=\log _{2}\left|U_{1}\right| \geq \log _{2} m-24 \exp \left(-\delta^{\prime} q\right) n \log _{2} q$. This establishes (c) and (f). This establishes the claims for $k=1$; the induction step in general is similar.
Now, as our IH let us assume that for $(k-1)$ we have the partition $\mathcal{H}_{1}(k-1) \sqcup \mathcal{H}_{2}(k-1), U_{k-1} \subseteq[m], \mathcal{R}^{k-1}$, integers $\left(g_{1}, \ldots, g_{k-1}\right)$ and $\mathbf{D}_{k-1}$ such that (a) through (f) are satisfied. Then, we repeat the above argument. We have $g_{k}=q-\mu_{q}^{c c}\left(g_{k-1}\right)-k \gamma q-2$. We use Lemma 【I. 2 for $h \in \mathcal{H}_{1}(k-1)$ with $\mathbf{R}$ being $\mathbf{D}_{k-1}$, and $\mathcal{R}=\mathcal{R}^{k-1}$ which is a $\left(k \gamma, \gamma^{2}-(k-1) \delta^{\prime}\right)$-sampler for $\mathbf{D}_{k-1}$ wrt $h$. Let $\left(R_{1}, \ldots, R_{s}\right)$ be obtained by $s=\exp \left(\gamma^{2}-k \delta^{\prime}\right)$ independent samples from $\mathcal{R}^{k-1}$. Let $D$ be the uniform distribution over $U_{k-1}$ and let $\mathcal{R}^{\prime}=\left(R_{1}^{\prime}, R_{2}^{\prime}, \ldots, R_{s}^{\prime}\right)$ be obtained by taking $s$ independent samples according to $\mathbf{R}^{\prime} \sim D^{\left\{g_{k}\right\}}$. We let $\mathbf{D}_{k}=\mathbf{D}_{k-1} \vee D^{\left\{g_{k}\right\}}$. For a fixed column $h$ we have the following: wp $1-12 \exp \left(-\delta^{\prime} q\right)$ over the random choices, the composed ensemble

$$
\widetilde{\mathcal{R}}=\left(R_{1} \cup R_{1}^{\prime}, R_{2} \cup R_{2}^{\prime}, \ldots, R_{s} \cup R_{s}^{\prime}\right)
$$

is good wrt $h$, i.e., $\widetilde{\mathcal{R}}$ is a $\left((k+1) \gamma, \gamma^{2}-k \delta^{\prime}\right)$-sampler for $\mathbf{D}_{k}$ wrt $h$, and furthermore $\forall i \in[s]$,

$$
\left|\left|h\left(R_{i} \cup R_{i}^{\prime}\right)\right|-\mathbb{E}\left[\left|h\left(\mathbf{D}_{k}\right)\right|\right]\right| \leq(k+1) \gamma q .
$$

Hence, on expectation only $12 \exp \left(-\delta^{\prime} q\right) n$ columns of $\mathcal{H}_{1}(k-1)$ are bad. Therefore, with probability at least $1 / 2$ at most $24 \exp \left(-\delta^{\prime} q\right) n$ columns of $\mathcal{H}_{1}(k-1)$ are bad. Also, the probability of an $R_{i} \cup R_{i}^{\prime} \in \widetilde{\mathcal{R}}$ having size less than $g_{\leq k}$ (because some two of our $q-\mu_{q}^{c c}\left(g_{k-1}\right)-k \gamma q-2$ choices of rows for $R_{i}^{\prime}$ picked the same row of collided with some row in
$\left.R_{i}\right)$ is at most $(q \ln q)^{2} /\left|U_{k-1}\right|$. Thus, by the union bound the probability of (b) not holding is at most $s \cdot(q \ln q)^{2} /\left|U_{k-1}\right|$ which is less than $1 / 2$. Therefore, there is choice of $\widetilde{\mathcal{R}}$, which we call $\mathcal{R}^{k}$, such that at most $24 \exp \left(-\delta^{\prime} q\right) n$ columns of $\mathcal{H}_{1}(k-1)$ are bad and (b) holds. Combining these bad columns with $\mathcal{H}_{2}(k-1)$ we obtain $\mathcal{H}_{2}(k)$ and the columns not in $\mathcal{H}_{2}(k)$ form the set $\mathcal{H}_{1}(k)=\mathcal{H}_{2}(k)$. Then, clearly (d) and (e) are true.

Let $\mathcal{H}_{2}(k) \backslash \mathcal{H}_{2}(k-1)=\left\{h_{1}, \ldots, h_{b}\right\}$ where $b \leq$ $24 \exp \left(-\delta^{\prime} q\right) n$ and WLOG assume that 1 is the most frequent symbol in $h_{1}$. Retain only those rows in $U_{k-1}$ that correspond to the symbol 1 in $h_{1}$. Call this pruned universe $U^{\prime}$ : this pruning ensures that so long as we add rows to $R_{i} \in R^{k}$ only from $U^{\prime}$, the image size in $h_{1}$ is at most $q-1$. Thus, by taking a multiplicative hit of at most $1 / q$ we have rendered $h_{1}$ ineffective. Iterating this over $\mathcal{H}_{2}(k)$ we take a multiplicative hit of $\left(\frac{1}{q}\right)^{b}$. Hence, we obtain a universe $U^{\prime}$, which will be $U_{k}$, such that $\log _{2}\left|U^{\prime}\right|=\log _{2}\left|U_{k}\right| \geq \log _{2}\left|U_{k-1}\right|-$ $24 \exp \left(-\delta^{\prime} q\right) n \log _{2} q \geq \log _{2} m-k \log _{2} q \cdot 24 \exp \left(-\delta^{\prime} q\right) n$. Together with property (c) of $U_{k-1}$ this establishes (c) and (f). This completes the induction step.

Proof of Theorem 1.3 (main result of the paper). Fix an $\epsilon^{\prime}<$ $1 / 6$ and let $C$ be an $\epsilon^{\prime} q \ln q$-list-decoding-code for the $q /(q-1)$ channel. Choose $\lambda \ll \epsilon^{\prime}$ and let $\epsilon=\epsilon^{\prime}+\lambda$. Let $q$ be sufficiently large so that $k=q^{\epsilon} \geq e q^{\epsilon^{\prime}+\lambda / 2}$. We will appeal to Lemma II.5 (with $k$ and $\epsilon$ ) and assume that $n \leq \exp \left(\delta^{\prime} q\right) \log _{2} m /(48$. $q^{\epsilon} \log _{2} q$ ). Then, by choosing a set of rows $R$ in the ensemble $\mathcal{R}^{k}$ and using (b) and Proposition II.4 we obtain that $|R| \geq \epsilon^{\prime} q \ln q$. However, using (c) we have that for all columns $h \in \mathcal{H}_{2}(k),|h(R)| \leq q-1$. Also, using (e) and Lemma II.3 we obtain that for all $h \in \mathcal{H}_{1}(k),|h(R)|<q$. This is a contradiction and hence $n>\exp \left(\delta^{\prime} q\right) \log _{2} m /\left(48 \cdot q^{\epsilon} \log _{2} q\right)$ or for sufficiently large $q$ we have $n>\Omega\left(\exp \left(q^{1-6 \epsilon^{\prime}} / 8\right) \log _{2} m\right)$.
We note that it is possible by a more careful analysis to improve the bound of $\Omega\left(\exp \left(q^{1-6 \epsilon^{\prime}} / 8\right) \log _{2} m\right)$ to $\Omega\left(\exp \left(q^{1-4 \epsilon^{\prime}} / 8\right) \log _{2} m\right)$ in which case we may apply the bound till a list size of $q \ln q / 4$. This bound is obtained by modifying Lemma $\Pi .5$ to accommodate $\gamma^{\prime}$ and $\delta^{\prime}$ which vary across the induction steps and being more scrupulous about the argument in the preceding paragraph.

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## Appendix

Proof of Lemma II. 3 Consider a constant $\lambda \ll \epsilon$. For $i=$ $1,2, \ldots, k$, let $B_{i}$ be the set of $q^{1-2 \epsilon-\lambda}$ elements of $[q]$ taking the topmost values in $\pi_{i}$. Let $B=B_{1} \cup B_{2} \cup \cdots \cup B_{k}$; note that $|B| \leq k q^{1-2 \epsilon-\lambda}=o\left(q^{1-\epsilon}\right)$. Then, $\mathbb{E}\left[\left|\mathbf{A}_{1} \cup \mathbf{A}_{2} \cup \ldots \mathbf{A}_{k}\right|\right]$ is at most

$$
|B|+\sum_{x \notin B}\left(1-\prod_{i=1}^{k}\left(1-\pi_{i}(x)\right)^{a_{i}}\right)
$$

Now, for $x \notin B$, we have $\pi_{i}(x) \leq 1 / q^{1-2 \epsilon-\lambda}$, and

$$
\begin{array}{r}
1-\pi_{i}(x) \geq \exp \left(-\pi_{i}(x) /\left(1-\pi_{i}(x)\right)\right) \\
\geq \exp \left(-\pi_{i}(x)\left(1+2 / q^{1-2 \epsilon-\lambda}\right)\right)
\end{array}
$$

Then, by the AM-GM inequality we have the upper-bound

$$
|B|+q-q \exp \left(-\left(1+2 / q^{1-2 \epsilon-\lambda}\right)(1 / q) \sum_{i, x} a_{i} \pi_{i}(x)\right)
$$

Our claim follows from this because $\exp \left(-(1+2 / \sqrt{q})(1 / q) \sum_{i, x} a_{i} \pi_{i}(x)\right) \geq \exp (-a / q)-$ $o\left(1 / q^{1-2 \epsilon}\right) \geq \exp (-a / q)-o\left(q^{1-\epsilon}\right) / q$.

Proof of Proposition [I.4 Suppose $\ell_{\leq i} \in[j q,(j+1) q]$ for some $j \geq 0$, then, $\ell_{i+1} \geq q / e^{j+1}$. Therefore, the number of $i$ 's for which $\ell_{\leq i} \in[j q,(j+1) q] \leq e^{j+1}$. Suppose $\ell_{\leq k}<\epsilon q \ln q$, then as a contradiction we have

$$
k<e+e^{2}+\cdots+e^{\epsilon \ln q} \leq e q^{\epsilon} .
$$

