

Corrections to “Generalization Bounds via Information Density and Conditional Information Density”

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Abstract—An error in the proof of the data-dependent tail bounds on the generalization error presented in Hellström and Durisi (2020) is identified, and a correction is proposed. Furthermore, we note that the absolute continuity requirements in Hellström and Durisi (2020) need to be strengthened to avoid measurability issues.

I. DATA-DEPENDENT BOUNDS IN [1, EQS. (26), (34), (95), AND (98)]

IN THE proof of [1, eq. (26)], we incorrectly claimed that [1, eq. (32)] implies [1, eq. (26)]. The issue is that [1, eq. (32)] holds for a *fixed* λ , whereas, for [1, eq. (26)] to hold, [1, eq. (32)] needs to hold uniformly over all $\lambda \in \mathbb{R}$.

This issue can be fixed as follows. Since $\text{gen}(w, \mathbf{Z})$ is $\sigma/\sqrt{n}$ -sub-Gaussian with zero mean under $P_{\mathbf{Z}}$ for all w , we can apply [2, Th. 2.6.(IV)] (with $\lambda = 1 - 1/n$ therein) to conclude that

$$\mathbb{E}_{P_{\mathbf{Z}}} \left[\exp \left(\frac{n-1}{2\sigma^2} (\text{gen}(w, \mathbf{Z}))^2 \right) \right] \leq \sqrt{n}. \quad (1)$$

Taking the expectation with respect to P_W , changing measure to $P_{W\mathbf{Z}}$, and rearranging terms, we obtain

$$\mathbb{E}_{P_{W\mathbf{Z}}} \left[\exp \left(\frac{n-1}{2\sigma^2} (\text{gen}(W, \mathbf{Z}))^2 - \log \sqrt{n} - \iota(W, \mathbf{Z}) \right) \right] \leq 1. \quad (2)$$

Proceeding as in [1, Corollary 2], with an additional use of Jensen’s inequality, we find that with probability at least $1 - \delta$ under $P_{\mathbf{Z}}$,

$$\begin{aligned} & \left| \mathbb{E}_{P_{W|\mathbf{Z}}} [\text{gen}(W, \mathbf{Z})] \right| \\ & \leq \sqrt{\frac{2\sigma^2}{n-1} \left(D(P_{W|\mathbf{Z}} \| P_W) + \log \frac{\sqrt{n}}{\delta} \right)}. \end{aligned} \quad (3)$$

Similarly, proceeding as in the proof of [1, eq. (34)], we find that with probability at least $1 - \delta$ under $P_{W\tilde{\mathbf{Z}}S}$,

$$|\text{gen}(W, \mathbf{Z})| \leq \sqrt{\frac{2\sigma^2}{n-1} \left(\iota(W, \mathbf{Z}) + \log \frac{\sqrt{n}}{\delta} \right)}. \quad (4)$$

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The issue reported in this note also affects the data-dependent tail bounds for the random-subset setting reported in [1, eqs. (95) and (98)]. To fix it, we use that for any fixed $(w, \tilde{\mathbf{z}})$, the random variable $\widehat{\text{gen}}(w, \tilde{\mathbf{z}}, S)$ is $1/\sqrt{n}$ -sub-Gaussian with zero mean under P_S . Applying [2, Th. 2.6.(IV)] with $\lambda = 1 - 1/n$ we obtain

$$\mathbb{E}_{P_S} \left[\exp \left(\frac{n-1}{2} (\widehat{\text{gen}}(w, \tilde{\mathbf{z}}, S))^2 \right) \right] \leq \sqrt{n}. \quad (5)$$

Taking the expectation with respect to $P_{W\tilde{\mathbf{Z}}}$, changing measure to $P_{W\tilde{\mathbf{Z}}S}$, and rearranging terms, we conclude that

$$\mathbb{E}_{P_{W\tilde{\mathbf{Z}}S}} \left[\exp \left(\frac{n-1}{2} (\widehat{\text{gen}}(W, \tilde{\mathbf{Z}}, S))^2 - \log \sqrt{n} - \iota(W, S|\tilde{\mathbf{Z}}) \right) \right] \leq 1. \quad (6)$$

Proceeding as in [1, Corollary 6], we conclude that with probability at least $1 - \delta$ under $P_{\tilde{\mathbf{Z}}S}$,

$$\begin{aligned} & \mathbb{E}_{P_{W|\tilde{\mathbf{Z}}S}} [\widehat{\text{gen}}(W, \tilde{\mathbf{Z}}, S)] \\ & \leq \sqrt{\frac{2}{n-1} \left(D(P_{W|\tilde{\mathbf{Z}}S} \| P_{W|\tilde{\mathbf{Z}}}) + \log \frac{\sqrt{n}}{\delta} \right)}. \end{aligned} \quad (7)$$

Furthermore, with probability at least $1 - \delta$ under $P_{W\tilde{\mathbf{Z}}S}$,

$$|\widehat{\text{gen}}(W, \tilde{\mathbf{Z}}, S)| \leq \sqrt{\frac{2}{n-1} \left(\iota(W, S|\tilde{\mathbf{Z}}) + \log \frac{\sqrt{n}}{\delta} \right)}. \quad (8)$$

To summarize, the data-dependent tail bounds reported in [1, eqs. (26), (34), (95), and (98)] should be replaced with (3), (4), (7), and (8) respectively.

Note that the data-independent tail bounds that we provide in [1, eqs. (27), (35), (41), (42), (96), (99), (101), and (102)] still hold verbatim, although their proofs need to be modified. Specifically, for a fixed λ , one needs to first replace the information measure appearing in the bounds with its data-independent relaxation. The desired bounds then follow by setting λ equal to a suitably chosen, data-independent constant. Consider for example the data-independent bound in [1, eq. (27)]. To obtain it, we first use [1, eq. (33)] in [1, eq. (32)], which results in

$$\begin{aligned} & P_{\mathbf{Z}} \left[\frac{\lambda^2 \sigma^2}{2n} - \lambda \mathbb{E}_{P_{W|\mathbf{Z}}} [\text{gen}(W, \mathbf{Z})] + \frac{\mathbb{E}_{P_{\mathbf{Z}}} [D(P_{W|\mathbf{Z}} \| P_W)^t]}{\delta^{1/t}} \right. \\ & \quad \left. + \log \frac{1}{\delta} \geq 0 \right] \geq 1 - 2\delta. \end{aligned} \quad (9)$$

The desired result follows by setting $\lambda = \pm\sqrt{\frac{a}{b}}$, where $a = \mathbb{E}_{P_Z}^{1/t} [D(P_W|Z || P_W)^t] / \delta^{1/t} + \log \frac{1}{\delta}$ and $b = \sigma^2/(2n)$, and then replacing δ with $\delta/2$.

II. ABSOLUTE CONTINUITY ASSUMPTION

In the statement of [1, Th. 1], we assumed that $P_{WZ} \ll P_W P_Z$. To avoid measurability issues, we should also assume that $P_W P_Z \ll P_{WZ}$. Similarly, in [1, Th. 4], we should also assume that $P_{W|\tilde{Z}} P_{\tilde{Z}} P_S \ll P_{W\tilde{Z}S}$.

REFERENCES

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