

Comments and Corrections

Corrections to “Satisficing in Multiarmed Bandit Problems”

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Abstract—An unfortunate mistake in the proof of Theorem 8 of the above article is corrected.

We correct an error in the published proof of [2, Th. 8]. The error arises from an incorrect application of concentration inequalities. The correction follows the same structure as that published in [3, Appendix G], which corrects the proofs of performance bounds for UCL algorithms in [4] and thus in [2, Ths. 7 and 8]. For simplicity of presentation, we first state the correction and then provide the associated proof.

The heuristic value Q_i^t in [2, (27)] is

$$Q_i^t = \mu_i^t + \sigma_i^t \Phi^{-1}(1 - \alpha_t). \quad (C1)$$

To correct of [2, Th. 8], set $\alpha_t = 1/(Kt^a)$ with $a > 4/(3(1 - \epsilon^2/16))$, $\epsilon \in (0, 4)$, and $K = \sqrt{2\pi e}$. The last part of the statement of [2, Th. 8] should be replaced by

“Then, the following statements hold for the satisfaction-in-mean-reward UCL algorithm with uncorrelated uninformative prior and $K = \sqrt{2\pi e}$:

- 1) The expected number of times a nonsatisfying arm i is chosen until time T satisfies.

$$\mathbb{E}[n_i^T] \leq \left(\frac{8a}{(\Delta_i^M)^2} \right) \log T + o(\log T).$$

- 2) The cumulative expected satisfaction-in-mean-reward regret until time T satisfies

$$J_{SM} \leq \sum_{i=1}^N \Delta_i^M \left(\frac{8a}{(\Delta_i^M)^2} \right) \log T + o(\log T).”$$

For the δ -sufficing and $(\mathcal{M}, \delta)$ -satisficing UCL algorithms of [2], similar corrections also hold with Q_i^t defined by (C1) and a modification to α_t . For these algorithms, the modification to α_t and its consequences can be succinctly stated by referring to the following Lemma, which is a straightforward application of Theorem 2 above.

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TABLE I

SUMMARY OF THE CORRECTIONS FOR THE SATISFICING UCL ALGORITHMS. DEFINE Q_i^t BY (C1) WITH $\epsilon \in (0, 4)$, $K = \sqrt{2\pi e}$ AND SET α_t AS FOLLOWS. THE CORRECTED PERFORMANCE BOUNDS WITH f IS DEFINED BY (C3)

Algorithm	α_t	Bound
Deterministic UCL	$\alpha_t = 1/Kt^a$, $a > \frac{4}{3(1-\epsilon^2/16)}$	$\mathbb{E}[n_i^T] \leq \frac{8a\sigma_i^2}{\Delta_i^M} \log T + o(\log T)$
Satisfaction-in-mean-reward UCL	$\alpha_t = 1/Kt^a$, $a > \frac{4}{3(1-\epsilon^2/16)}$	$\mathbb{E}[n_i^T] \leq \frac{8a}{(\Delta_i^M)^2} \log T + o(\log T)$
δ -Sufficing UCL	α_t from Equation (C2), $\delta \mapsto \delta/2$	$n_i^T \leq f(\delta/2, \Delta_i)$
$(\mathcal{M}, \delta)$ -Satisficing UCL	α_t from Equation (C2), $\delta \mapsto \delta/3$	$n_i^T \leq f(\delta/3, \Delta_i^M)$

Lemma 1: Let $\epsilon \in (0, 4)$ and define

$$\alpha_t = 1 - \Phi \left(\sqrt{\frac{2}{1 - \epsilon^2/16} \log \frac{\log((1 + \epsilon)t)}{\delta \log(1 + \epsilon)}} \right). \quad (C2)$$

Then, at time t

$$\Pr[[2, (40)] \text{ holds}] = \Pr \left[\frac{\mu_i^t - m_i}{\sigma_i^t} \geq \Phi^{-1}(1 - \alpha_t) \right] \leq \delta.$$

The corrections to the four algorithms published in [2] and the corresponding corrected expressions for the performance bounds are summarized in Table I. For δ -sufficing and $(\mathcal{M}, \delta)$ -satisficing UCL, the bounds take the form

$$f(\delta, \Delta) := \frac{8\sigma_s^2}{\Delta^2(1 - \epsilon^2/16)} \log \frac{\log((1 + \epsilon)T)}{\delta \log(1 + \epsilon)} + 1. \quad (C3)$$

Note that with the correction, which accounts for the dependence of n_i^t on rewards accrued, the upper bound functional form (C3) is no longer independent of T . However, the dependence on T is of the form $\log \log T$, which is a very slowly increasing function of T . Therefore, in any realistic application the upper bound will effectively be constant and the qualitative result of [2] does not change.

REVISED PROOF

We employ the following concentration inequality from Garivier and Moulines [1] to fix the proof. Let $(X_t)_{t \geq 1}$ be a sequence of independent sub-Gaussian random variables with $\mathbb{E}[X_t] = \mu_t$, i. e., $\mathbb{E}[\exp(\lambda(X_t - \mu_t))] \leq \exp(\lambda^2 \sigma^2 / 2)$ for some variance parameter $\sigma > 0$. Consider a previsible sequence $(\epsilon_t)_{t \geq 1}$ of Bernoulli variables, i. e., for all $t > 0$, ϵ_t is deterministically known given $\{X_\tau\}_{0 < \tau < t}$. Let

$$s^t = \sum_{s=1}^t X_s \epsilon_s, m^t = \sum_{s=1}^t \mu_s \epsilon_s, n^t = \sum_{s=1}^t \epsilon_s.$$

Theorem 2 (See [1, Th. 22] and [3, Th. 11]): Let $(X_t)_{t \geq 1}$ be a sequence of sub-Gaussian¹ independent random variables with common variance parameter σ and let $(\epsilon_t)_{t \geq 1}$ be a previsible sequence of Bernoulli variables. Then, for all integers t and all $\delta, \epsilon > 0$

$$\begin{aligned} \Pr \left[\frac{s^t - m^t}{\sqrt{n^t}} > \delta \right] \\ \leq \left[\frac{\log t}{\log(1 + \epsilon)} \right] \exp \left(-\frac{\delta^2}{2\sigma^2} \left(1 - \frac{\epsilon^2}{16} \right) \right). \end{aligned} \quad (C4)$$

We will also use the following lower bound for $\Phi^{-1}(1 - \alpha)$, the quantile function of the normal distribution.

Proposition 3: For any $t \in \mathbb{N}$ and $a > 1$, the following holds:

$$\Phi^{-1} \left(1 - \frac{1}{\sqrt{2\pi e t^a}} \right) \geq \sqrt{\nu \log t^a} \quad (C5)$$

for any $0 < \nu \leq 1.59$.

Proof: We begin with the inequality $\Phi^{-1}(1 - \alpha) > \sqrt{-\log(2\pi\alpha^2(1 - \log(2\pi\alpha^2)))}$ established in [4]. It suffices to show that

$$-\log \left(\frac{1}{e t^{2a}} \left(1 - \log \left(\frac{1}{e t^{2a}} \right) \right) \right) - \nu \log t^a \geq 0$$

for $\nu \in (0, 1.59]$. The left-hand side of the above inequality is

$$g(t) := 1 - \log 2 + a(2 - \nu) \log t - \log(1 + a \log t).$$

It can be verified that g admits a unique minimum at $t = e^{(\nu-1)/(a(2-\nu))}$ and the minimum value is $\nu - \log 2 + \log(2 - \nu)$, which is positive for $0 < \nu \leq 1.59$. ■

In the following, we choose $\nu = 3/2$.

Correction to the proof of [2, Th. 8]: The structure of the published proof carries through. Let i be a nonsatisfying arm, i.e., $m_i < \mathcal{M}$, and recall that i^* denotes the arm with maximum mean reward. Let η be a positive integer and let $\epsilon \in (0, 4)$ and $a > 4/(3(1 - \epsilon^2/16))$.

We first analyze the probability that [2, eq. (31)] holds by applying Theorem 2. Let $\{X_\tau\}_{0 < \tau < t}$ be the sequence of rewards associated with arm i , and let $(\epsilon_t)_{t \geq 1}$ equal 1 if the algorithm chooses arm i at time t . Note that, for an uncorrelated uninformative prior, $\mu_i^t = \bar{m}_i^t = s^t/n^t$, $\sigma_i^t = 1/\sqrt{n_i^t}$, $m_i = m^t/n^t$, and $n_i^t = n^t$. [2, eq. (31)] is thus equivalent to

$$\frac{s^t}{n^t} - \frac{m^t}{n^t} \geq \frac{1}{\sqrt{n^t}} \Phi^{-1}(1 - \alpha_t) \Rightarrow \frac{s^t - m^t}{\sqrt{n^t}} \geq \Phi^{-1}(1 - \alpha_t).$$

Letting $\delta = \Phi^{-1}(1 - \alpha_t)$ and applying (C4) yields

$$\begin{aligned} \Pr [2, \text{eq. (31)}] \text{ holds} &= \Pr \left[\frac{s^t - m^t}{\sqrt{n^t}} \geq \delta \right] \\ &\leq \left[\frac{\log t}{\log(1 + \epsilon)} \right] \exp \left(-\frac{3 \log t^a}{4} \left(1 - \frac{\epsilon^2}{16} \right) \right) \\ &= \left[\frac{\log t}{\log(1 + \epsilon)} \right] t^{-\frac{3a(1-\epsilon^2)/16}{4}} \end{aligned}$$

where the second inequality follows from (C5). The same bound holds for [2, eq. (32)].

¹The result in [1, Th. 22] is stated for bounded rewards, but it extends immediately to sub-Gaussian rewards by noting that the upper bound on the moment generating function for a bounded random variable obtained using a Hoeffding inequality has the same functional form as the sub-Gaussian random variable.

It can be verified that for the corrected Q_i^t in (C1), the constant “8” in [2, eqs. (35), (38), and (39)] will be replaced by $8a$. Following the proof in [2] with the above corrections

$$\mathbb{E} [n_i^T] \leq \left[\frac{8a}{(\Delta_i^M)^2} \log T \right] + \sum_{t=1}^T 3 \left[\frac{\log t}{\log(1 + \epsilon)} \right] t^{-\frac{3a(1-\epsilon^2)/16}{4}}.$$

The sum can be bounded by the integral

$$\int_1^T \left(\frac{\log t}{\log(1 + \epsilon)} + 1 \right) t^{-\frac{3a(1-\epsilon^2)/16}{4}} dt + 1. \quad (C6)$$

It can be verified that the integral (C6) is of class $o(\log T)$ as long as the exponent $3a(1 - \epsilon^2/16)/4 > 1$. Putting everything together, we have

$$\mathbb{E} [n_i^T] \leq \frac{8a\sigma_s^2}{\Delta_i^2} \log T + o(\log T).$$

The second statement follows from the definition of the cumulative expected regret. ■

The corrections to the proofs of [2, Th. 10] (δ -sufficing UCL) and [2, Th. 11] ($(\mathcal{M}, \delta)$ -satisficing UCL) follow the same structure.

Correction to proof of [2, Th. 10]: For the corrected α_t defined in (C2) with $\delta \mapsto \frac{\delta}{2}$, [2, eq. (42)] is equivalent to

$$\Delta_i = m_{i^*} - m_i < 2C_i^t = \frac{2\sigma_s}{\sqrt{n_i^t}} \Phi^{-1}(1 - \alpha_t).$$

Squaring, rearranging, and applying (C2), we see that this never holds if

$$n_i^t > \frac{8\sigma_s^2}{\Delta_i^2(1 - \epsilon^2/16)} \log \frac{2 \log((1 + \epsilon)t)}{\delta \log(1 + \epsilon)} = \eta.$$

Then, Lemma 1 implies that [2, eqs. (40), (41)] each hold with probability at most $\delta/2$. Therefore, for $n_i^t > \eta + 1 = f(\delta/2, \Delta_i)$, a nonsatisfying arm is selected with probability at most δ . ■

Correction to proof of [2, Th. 11]: For the corrected α_t defined in (C2) with $\delta \mapsto \frac{\delta}{2}$, an argument analogous to that for [2, eq. (42)] above shows that [2, eq. (44)] never holds for $n_i^t > \eta + 1 = f(\delta/3, \Delta_i^M) - 1$.

Applying Lemma 1 implies that [2, eq. (43)] holds with probability at most $\delta/3$. Similarly to the corrected proof for [2, Th. 10] above, for $n_i^t > \eta + 1 = f(\delta/3, \Delta_i^M)$, $Q_i^t \geq Q_{i^*}^t$ with probability at most $\frac{2\delta}{3}$. Thus, a nonsatisfying arm is selected with probability at most δ . ■

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