

On Marton's Inner Bound for the General Broadcast Channel

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Abstract—We establish several new results on Marton's coding scheme and its corresponding inner bound on the capacity region of the general broadcast channel. We show that unlike the Gaussian case, Marton's coding scheme without superposition coding is not optimal in general even for a degraded broadcast channel with no common message. We then establish properties of Marton's inner bound that help restrict the search space for computing the sum-rate. Next, we show that the inner bound is optimal along certain directions. Finally, we propose a coding scheme that may lead to a larger inner bound.

I. INTRODUCTION

In this paper, we consider the general two-receiver broadcast channel with an input alphabet \mathcal{X} , output alphabets \mathcal{Y} and \mathcal{Z} , and conditional probability distribution function $q(y, z|x)$. The capacity region of this channel is defined as the set of rate triples (R_0, R_1, R_2) such that the sender X can reliably communicate a common message at rate R_0 to both receivers and two private messages at rates R_1 and R_2 to receivers Y and Z respectively, see [1] or [2]. The capacity region of this channel is known for several special cases but unknown in general. The best known general inner bound to the capacity region is due to Marton [3][6].

In this paper, we study Marton's inner bound. Marton's inner bound for a general two-receiver discrete-memoryless broadcast channel is as follows:

Marton's Inner bound [3][2][4][6]: The union of non-negative rate triples (R_0, R_1, R_2) satisfying the inequalities

$$R_0 + R_1 \leq I(UW; Y), \quad (1)$$

$$R_0 + R_2 \leq I(VW; Z), \quad (2)$$

$$R_0 + R_1 + R_2 \leq I(UW; Y) + I(V; Z|W) - I(U; V|W), \quad (3)$$

$$R_0 + R_1 + R_2 \leq I(U; Y|W) + I(VW; Z) - I(U; V|W), \quad (4)$$

$$2R_0 + R_1 + R_2 \leq I(UW; Y) + I(VW; Z) - I(U; V|W), \quad (5)$$

for some random variables $(U, V, W, X, Y, Z) \sim p(u, v, w, x)q(y, z|x)$ constitutes an inner bound to the capacity region. Further to compute this region it suffices to consider $|U|, |V| \leq |X|, |W| \leq |X| + 4$ and assume that X is a deterministic function of (U, V, W) [8].

In this paper we prove various results related to this inner bound.

Insufficiency of Marton's coding scheme without a superposition variable: Random variable W corresponds to the “superposition-coding” aspect of the bound, and the random variables U and V correspond to the “Marton-coding” aspect of the bound. Necessity of the “superposition-coding” aspect of the inner bound had previously been observed for a non-degraded broadcast channel [10]. For degraded channels, it is known that W is unnecessary for achieving the capacity region of Gaussian broadcast channels (through dirty paper coding) [14]. We show that, unlike in the Gaussian broadcast channel case, “Marton's coding scheme” alone is not sufficient to achieve the capacity region of the general degraded broadcast channel.

Computing the sum rate in Marton's inner bound: Given a broadcast channel $q(y, z|x)$ the maximum sum-rate achievable via Marton's strategy is given by

$$\max_{p(u, v, w, x)} \min\{I(W; Y), I(W; Z)\} + I(U; Y|W) + I(V; Z|W) - I(U; V|W). \quad (6)$$

Further it suffices to consider $|U|, |V| \leq |X|, |W| \leq |X| + 1$ and assume that X is a deterministic function of (U, V, W) [8]. Note that $\min\{I(W; Y), I(W; Z)\}$ depends only on $p(w, x)$. The last three terms $I(U; Y|W) + I(V; Z|W) - I(U; V|W)$ can be written as

$$\sum_w p(w) (I(U; Y|W = w) + I(V; Z|W = w) - I(U; V|W = w)).$$

Let us write the above optimization as follows:

$$\max_{p(w, x)} \left[\min\{I(W; Y), I(W; Z)\} + \sum_w p(w) \max_{p(u, v|w, x)} [I(U; Y|W = w) + I(V; Z|W = w) - I(U; V|W = w)] \right].$$

One can think of the maximization in the following way

$$\max_{p(w, x)} \min\{I(W; Y), I(W; Z)\} + \sum_w p(w) T(p(x|w))$$

where $T(p(x))$ is the maximum of $I(U; Y) + I(V; Z) - I(U; V)$ over all $p(u, v|x)$ where $X \sim p(x)$ and X is a deterministic function of (U, V) .

It is shown in [13] that the latter maximization problem concerning $T(p(x))$ has a remarkable solution for all binary input broadcast channels: it suffices to take $U = X$ and $V = \text{constant}$, or $V = X$ and $U = \text{constant}$. In other words, for all binary input broadcast channels we have

$$I(U; Y) + I(V; Z) - I(U; V) \leq \max\{I(X; Y), I(X; Z)\}. \quad (7)$$

To prove this, authors of [13] consider different mappings from $\mathcal{U} \times \mathcal{V} \mapsto \mathcal{X}$. Because of the cardinality bound of two on \mathcal{U} and \mathcal{V} , the authors argue that the XOR mapping (i.e. $X = U \oplus V \bmod 2$) and the AND mapping (i.e. $X = U \wedge V$) cannot occur in any maximizer of $I(U; Y) + I(V; Z) - I(U; V)$.

We believe that finding the correct extension of equation (7) to larger alphabets can be useful in (a) computing Marton's inner bound efficiently for a given channel, and (b) comparing the Marton inner bound with its multi-letter characterizations to see if Marton's inner bound is optimal or not (see [11] for a discussion of this line of attack).

One of the main results of this part is to generalize to larger alphabets the statement that the XOR mapping cannot occur. We show that one cannot find distinct u_0, u_1 in \mathcal{U} , distinct v_0, v_1 in \mathcal{V} and distinct x_0, x_1 in \mathcal{X} such that $p(x_0|u_0, v_0) = p(x_0|u_1, v_1) = p(x_1|u_1, v_0) = p(x_1|u_0, v_1) = 1$.

Optimality of Marton's inner bound along certain directions: We compute the maximum of $\lambda_0 R_0 + \lambda_1 R_1 + \lambda_2 R_2$ over all (R_0, R_1, R_2) in the capacity region where λ_0, λ_1 and λ_2 are real numbers such that $\lambda_0 \geq \lambda_1 + \lambda_2$. We observe that Marton's inner bound is tight along these directions.

An achievable region: Since capacity is defined in the limit of large block length, it is natural to expect that optimal coding schemes have an invariant structure with respect to shifts in time. This suggests that capacity should be expressed via a formula that has a *fixed-point character*, namely it should involve joint distributions that are invariant under a time shift. Following this general idea, we propose a new inner bound for the capacity region. We don't know if the proposed inner bound is strictly better than Marton's inner bound.

The rest of the paper is organized as follows. Section II contains the main results of the paper, and section III contains the proofs of these results, with some of the details relegated to the appendices.

II. MAIN

Let $\mathcal{C}(q(y, z|x))$ denote the capacity region of the broadcast channel $q(y, z|x)$, and $\mathcal{C}_M(q(y, z|x))$ denote Marton's inner bound for the channel $q(y, z|x)$, defined in the introduction by equations (1)-(5). The notation X^i is used to denote the vector (X_1, X_2, \dots, X_i) , and X_i^n to denote $(X_i, X_{i+1}, \dots, X_n)$.

A. Insufficiency of Marton's coding scheme without a superposition variable

In Marton's inner bound the auxiliary random variable W corresponds to the "superposition-coding" aspect of the bound,

and the random variables U and V correspond to the "Marton-coding" aspect of the bound. When $R_0 = 0$ (private messages only) and $W = \emptyset$, Marton's inner bound reduces to the set of non-negative rate pairs (R_1, R_2) satisfying

$$R_1 \leq I(U; Y|Q), \quad (8)$$

$$R_2 \leq I(V; Z|Q), \quad (9)$$

$$R_1 + R_2 \leq I(U; Y|Q) + I(V; Z|Q) - I(U; V|Q), \quad (10)$$

for some random variables $(Q, U, V, X, Y, Z) \sim p(q)p(u, v, x|q)q(y, z|x)$.

It is known that this inner bound is tight for Gaussian broadcast channels (through dirty paper coding), implying that W is unnecessary for achieving the capacity region of this class of degraded broadcast channels [14]. We show through an example that this is not the case in general.

Claim 1: There are degraded broadcast channels for which Marton's private message inner bound without W is strictly contained in the capacity region of the channel (which is known to equal the Marton region with superposition variable in the case of degraded channels).

B. Computing the sum-rate for Marton's Inner Bound

1) Extensions of the binary inequality: In this subsection we are concerned with the following maximization problem that is tightly related to the calculation of the sum rate for Marton's inner bound: given $p(x)$, maximize $I(U; Y) + I(V; Z) - I(U; V)$ over all $p(u, v|x)$ where X is a function of (U, V) .

To state our main result we need the following two definitions:

Definition 1: The input symbols x_0 and x_1 are said to be *indistinguishable* by the channel if $q(y|x_0) = q(y|x_1)$ for all y , and $q(z|x_0) = q(z|x_1)$ for all z . A channel $q(y, z|x)$ is said to be *irreducible* if no two of its inputs symbols are indistinguishable by the channel.

Definition 2: Let $\mathcal{U} = \{u_1, u_2, \dots, u_{|\mathcal{U}|}\}$, $\mathcal{V} = \{v_1, \dots, v_{|\mathcal{V}|}\}$ be finite sets, and ξ be a deterministic mapping from $\mathcal{U} \times \mathcal{V}$ to \mathcal{X} . One can represent the mapping by a table having $|\mathcal{U}|$ rows and $|\mathcal{V}|$ columns; the rows are indexed by $u_1, u_2, \dots, u_{|\mathcal{U}|}$ and the columns are indexed by $v_1, v_2, \dots, v_{|\mathcal{V}|}$. In the cell (i, j) , we write $\xi(u_i, v_j)$, for the symbol x that (u_i, v_j) is being mapped to. The profile of the i^{th} row is defined to be a vector of size $|\mathcal{X}|$ counting the number of occurrences of the elements of \mathcal{X} in the i^{th} row. In other words if $\mathcal{X} = \{x_1, x_2, \dots, x_{|\mathcal{X}|}\}$, the k^{th} element of the profile of the i^{th} row is the number of times that x_k shows up in the i^{th} row of the table. The profile of the j^{th} column is defined similarly. Define the profile of the table to be a vector of size $(|\mathcal{U}| + |\mathcal{V}|)|\mathcal{X}|$ formed by concatenating the profile vectors of the rows and the columns of the table. The profile vector of the mapping ξ is denoted by \vec{v}_ξ .

We now state the main result of this subsection.

Theorem 1: Take an arbitrary irreducible broadcast channel $q(y, z|x)$ where $q(y|x) > 0, q(z|x) > 0$ for all x, y, z . Fix some $p(x)$. Take any $p(u, v|x)$ maximizing $I(U; Y) + I(V; Z) - I(U; V)$ where X is a function of (U, V) . Without

loss of generality assume that $p(u) > 0$ for all $u \in \mathcal{U}$, and $p(v) > 0$ for all $v \in \mathcal{V}$. Let $x = \xi(u, v)$ denote the deterministic mapping from $\mathcal{U} \times \mathcal{V}$ to \mathcal{X} . Then all of the following conditions must hold:

- $p(u, v) > 0$, $p(u, y) > 0$, and $p(v, z) > 0$ for all u, v, y and z .
- The profile vector of the mapping ξ , \vec{v}_ξ , cannot be written as

$$\sum_{t=1}^M \alpha_t \vec{v}_{\xi_t},$$

where ξ_t (for $t = 1, 2, 3, \dots, M$) are deterministic mappings from $\mathcal{U} \times \mathcal{V}$ to \mathcal{X} not equal to ξ , and α_t are non-negative numbers adding up to one, i.e. $\sum_{t=1}^M \alpha_t = 1$.

- Let the functions

$$\begin{aligned} f_u : \mathcal{X} &\rightarrow \mathbb{R} \text{ for every } u \in \mathcal{U}, \\ g_v : \mathcal{X} &\rightarrow \mathbb{R} \text{ for every } v \in \mathcal{V}, \\ \text{and } h : \mathcal{X} &\rightarrow \mathbb{R}, \end{aligned}$$

be defined by

$$\begin{aligned} f_u(x) &= \sum_y q(y|x) \log p(u, y), \\ g_v(x) &= \sum_z q(z|x) \log p(v, z), \end{aligned}$$

$$\begin{aligned} h(x) &= \min_{u' \in \mathcal{U}, v' \in \mathcal{V}} \left(\log(p(u', v')) \right. \\ &\quad \left. - f_{u'}(x) - g_{v'}(x) \right). \end{aligned}$$

These definitions make sense because of the first bullet of this theorem. Then, for any u and v , the following two equations hold:

$$\log(p(u, v)) = \max_x [f_u(x) + g_v(x) + h(x)],$$

and

$$\begin{aligned} p(x_0|u, v) &= 1 \text{ for some } x_0 \in \mathcal{X} \Rightarrow \\ x_0 &\in \arg\max_x [f_u(x) + g_v(x) + h(x)]. \end{aligned}$$

Discussion 1: These constraints imply restrictions on the maximizers. The second bullet implies that one cannot find distinct u_0, u_1 in \mathcal{U} , distinct v_0, v_1 in \mathcal{V} and distinct x_0, x_1 in \mathcal{X} such that $p(x_0|u_0, v_0) = p(x_0|u_1, v_1) = p(x_1|u_1, v_0) = p(x_1|u_0, v_1) = 1$.¹

Next, assume that all we know about the mapping pattern is that $x_0 = \xi(u_0, v_0) = \xi(u_1, v_1)$ for some x_0 . Then the third

¹Let the mapping ξ_1 be equal to ξ except that (u_0, v_0) and (u_1, v_1) are mapped to x_1 (instead of x_0), and (u_1, v_0) and (u_0, v_1) are mapped to x_0 (instead of x_1). Figure 1 illustrates this. The mapping ξ_1 has the same profile vector as ξ . Thus we can write the original profile as a convex combination of other profiles (the condition in the displayed equation of the second bullet is violated for the choice of $M = 1$, ξ_1 and $\alpha_1 = 1$). Thus the second bullet implies that it cannot happen. Similarly the mapping shown in Figure 2 cannot occur because there is another mapping with the same profile.

	x_0		x_1	
	x_1		x_0	

	x_1		x_0	
	x_0		x_1	

Fig. 1. If we have a mapping with the XOR structure, we can get another mapping with the same profile by switching x_0 and x_1 of four cells of the mappings.

	x_0		x_1		x_2
	x_1		x_2		x_0

	x_1		x_2		x_0
	x_0		x_1		x_2

Fig. 2. Another mapping that cannot occur because one can find another mapping with the same profile.

bullet implies that $p(u_0, v_0)p(u_1, v_1) \leq p(u_1, v_0)p(u_0, v_1)$. This holds since

$$\begin{aligned} \log p(u_0, v_0) + \log p(u_1, v_1) &= \\ f_{u_0}(x_0) + g_{v_0}(x_0) + h(x_0) + \\ f_{u_1}(x_0) + g_{v_1}(x_0) + h(x_0) &= \\ f_{u_0}(x_0) + g_{v_1}(x_0) + h(x_0) + \\ f_{u_1}(x_0) + g_{v_0}(x_0) + h(x_0) &\leq \\ \max_x f_{u_0}(x) + g_{v_1}(x) + h(x) + \\ \max_x f_{u_1}(x) + g_{v_0}(x) + h(x) &= \\ \log p(u_0, v_1) + \log p(u_1, v_0). \end{aligned}$$

2) *Sum rate evaluation:* In this subsection we turn to evaluation of the whole sum-rate expression of Marton's inner bound (including the W terms). We need the following definition:

For any $\lambda \in [0, 1]$, let

$$\begin{aligned} T_\lambda &= \max_{p(u, v, w, x)} (\lambda I(W; Y) + (1 - \lambda) I(W; Z) + \\ &\quad I(U; Y|W) + I(V; Z|W) - I(U; V|W)). \end{aligned}$$

Computing the sum-rate for Marton's inner bound is closely related to the above maximization problem for $\lambda \in [0, 1]$:

Claim 2: The maximum of the sum-rate for Marton's inner bound is equal to $\min_{\lambda \in [0, 1]} T_\lambda$.

Since the original submission of the conference version of this paper, some interesting properties of T_λ such as its convexity in λ , and its connection to the outer bound and its factorization have been investigated in [11] and [12]. An alternative proof of Claim 2 using a theorem by Terkelsen is also reported in [11].

The main theorem of this section restricts the search space for computing T_λ . In this section, we only deal with broadcast channels $q(y, z|x)$ with strictly positive transition matrices, i.e. when $q(y|x) > 0, q(z|x) > 0$ for all x, y, z . In order to evaluate T_λ when $q(y|x)$ or $q(z|x)$ become zero for some y or z , one can use the continuity of T_λ in $q(y, z|x)$ and take the limit of T_λ for a sequence of channels with positive entries converging to the desired channel. The reason for dealing with this class of broadcast channels should become clear by the following lemma which is a corollary to the first bullet of Theorem 1.

Lemma 1: Take an arbitrary broadcast channel $q(y, z|x)$ with strictly positive transition matrices (i.e. $q(y|x) > 0, q(z|x) > 0$ for all x, y, z). Let $p(u, v, w, x)$ be an arbitrary joint distribution maximizing T_λ for some $\lambda \in [0, 1]$ where $H(X|U, V, W) = 0$. If $p(u, w)$ and $p(v, w)$ are positive for some triple (u, v, w) , then it must be the case that $p(u, v, w) > 0, p(u, w, y) > 0$ and $p(v, w, z) > 0$ for all y and z .

Theorem 2: Take an arbitrary irreducible broadcast channel $q(y, z|x)$ with strictly positive transition matrices. In computing T_λ for some $\lambda \in [0, 1]$, it suffices to take the maximum over auxiliary random variables $p(u, v, w, x)q(y, z|x)$ simultaneously satisfying the following constraints:

- $|\mathcal{U}| \leq \min(|\mathcal{X}|, |\mathcal{Y}|), |\mathcal{V}| \leq \min(|\mathcal{X}|, |\mathcal{Z}|), |\mathcal{W}| \leq |\mathcal{X}|$.
- $H(X|UVW) = 0$. Given w where $p(w) > 0$, we use $x = \xi^{(w)}(u, v)$ to denote the deterministic mapping from $\mathcal{U}_w \times \mathcal{V}_w$ to \mathcal{X} . Here \mathcal{U}_w is the set of $u \in \mathcal{U}$ such that $p(u|w) > 0$ and \mathcal{V}_w is the set of $v \in \mathcal{V}$ such that $p(v|w) > 0$.
- For arbitrary w such that $p(w) > 0$, the profile vector of the mapping $\xi^{(w)}, \overrightarrow{v_{\xi^{(w)}}}$, cannot be written as

$$\sum_{t=1}^M \alpha_t \overrightarrow{v_{\xi_t}},$$

where ξ_t (for $t = 1, 2, 3, \dots, M$) are deterministic mappings from $\mathcal{U}_w \times \mathcal{V}_w$ to \mathcal{X} not equal to $\xi^{(w)}$, and α_t are non-negative numbers adding up to one, i.e. $\sum_{t=1}^M \alpha_t = 1$.

- For arbitrary w such that $p(w) > 0$, let the functions

$$\begin{aligned} f_{u,w} &: \mathcal{X} \rightarrow \mathbb{R} \text{ for every } u \in \mathcal{U}_w, \\ g_{v,w} &: \mathcal{X} \rightarrow \mathbb{R} \text{ for every } v \in \mathcal{V}_w, \\ \text{and } h_w &: \mathcal{X} \rightarrow \mathbb{R}, \end{aligned}$$

be defined by

$$\begin{aligned} f_{u,w}(x) &= \sum_y q(y|x) \log p(uy|w), \\ g_{v,w}(x) &= \sum_z q(z|x) \log p(vz|w), \\ h_w(x) &= \min_{u' \in \mathcal{U}_w, v' \in \mathcal{V}_w} \left(\log(p(u'v'|w)) \right. \\ &\quad \left. - f_{u',w}(x) - g_{v',w}(x) \right). \end{aligned}$$

These definitions make sense because of Lemma 1. Then, for any $u \in \mathcal{U}_w$ and $v \in \mathcal{V}_w$, the following two equations hold:

$$\log(p(uv|w)) = \max_x [f_{u,w}(x) + g_{v,w}(x) + h_w(x)],$$

and

$$\begin{aligned} p(x_0|u, v, w) &= 1 \text{ for some } x_0 \in \mathcal{X} \Rightarrow \\ x_0 &\in \operatorname{argmax}_x [f_{u,w}(x) + g_{v,w}(x) + h_w(x)]. \end{aligned}$$

- Given any w , random variables U_w, V_w, X_w, Y_w, Z_w distributed according to $p(u, v, x, y, z|w)$ satisfy the following:

$$\begin{aligned} I(\overline{U}; Y_w) &\geq I(\overline{U}; V_w Z_w) \text{ for any } \overline{U} \rightarrow U_w \rightarrow V_w X_w Y_w Z_w, \\ I(\overline{V}; Z_w) &\geq I(\overline{V}; U_w Y_w) \text{ for any } \overline{V} \rightarrow V_w \rightarrow U_w X_w Y_w Z_w. \end{aligned}$$

Discussion 2: The first constraint imposes cardinality bounds on $|\mathcal{U}|$ and $|\mathcal{V}|$ that are better than those reported in [8]. However, we only claim the improved cardinality bounds for T_λ and not the whole capacity region. The second constraint is not new, and can be found in [8]. The other constraints are useful in restricting the search space due to the constraints imposed on $p(u, v, w, x)$. For instance, the third and fourth bullet restrict the set of possible mappings, as discussed in the previous subsection.

C. Optimality along certain directions

In order to state the main result of this section we need the following definition:

Definition 3: [7] Let $\mathcal{C}_{d_1}(q(y, z|x))$ and $\mathcal{C}_{d_2}(q(y, z|x))$ denote the degraded message set capacity regions, i.e. when $R_1 = 0$ and $R_2 = 0$, respectively. The capacity region $\mathcal{C}_{d_1}(q(y, z|x))$ is the set of non-negative rate pairs (R_0, R_2) satisfying

$$\begin{aligned} R_0 &\leq I(W; Y), \\ R_2 &\leq I(X; Z|W), \\ R_0 + R_2 &\leq I(X; Z), \end{aligned}$$

for some random variables $(W, X, Y, Z) \sim p(w, x)q(y, z|x)$. The capacity region $\mathcal{C}_{d_2}(q(y, z|x))$ is defined similarly.

We now state the main result of this section:

Theorem 3: For a broadcast channel $q(y, z|x)$ and real numbers λ_0, λ_1 and λ_2 such that $\lambda_0 \geq \lambda_1 + \lambda_2$,

$$\begin{aligned} \max_{(R_0, R_1, R_2) \in \mathcal{C}(q(y, z|x))} (\lambda_0 R_0 + \lambda_1 R_1 + \lambda_2 R_2) &= \\ \max \{ \max_{(R_0, R_2) \in \mathcal{C}_{d_1}(q(y, z|x))} (\lambda_0 R_0 + \lambda_2 R_2), \\ \max_{(R_0, R_1) \in \mathcal{C}_{d_2}(q(y, z|x))} (\lambda_0 R_0 + \lambda_1 R_1) \}, \end{aligned}$$

where $\mathcal{C}_{d_1}(q(y, z|x))$ and $\mathcal{C}_{d_2}(q(y, z|x))$ are the degraded message set capacity regions for the given channel.

Corollary 1: The above observation essentially says that if $\lambda_0 \geq \lambda_1 + \lambda_2$, then a maximum of $\lambda_0 R_0 + \lambda_1 R_1 + \lambda_2 R_2$ over triples (R_0, R_1, R_2) in the capacity region occurs when either $R_1 = 0$ or $R_2 = 0$.

Remark 1: Since $\mathcal{C}_{d_1}(q(y, z|x)) \cup \mathcal{C}_{d_2}(q(y, z|x)) \subset \mathcal{C}_M(q(y, z|x)) \subset \mathcal{C}(q(y, z|x))$, the above lemma implies that Marton's inner bound is tight along the direction of such

$(\lambda_0, \lambda_1, \lambda_2)$, i.e.

$$\begin{aligned} & \max_{(R_0, R_1, R_2) \in \mathcal{C}(q(y, z|x))} (\lambda_0 R_0 + \lambda_1 R_1 + \lambda_2 R_2) = \\ & \max_{(R_0, R_1, R_2) \in \mathcal{C}_M(q(y, z|x))} (\lambda_0 R_0 + \lambda_1 R_1 + \lambda_2 R_2), \\ & \text{whenever } \lambda_0 \geq \lambda_1 + \lambda_2. \end{aligned}$$

D. An achievable region

Since capacity is defined in the limit of large block length, it is natural to expect that optimal coding schemes have an invariant structure with respect to shifts in time. This suggests that capacity should be expressed via a formula that has a fixed-point character, namely it should involve joint distributions that are invariant under a time shift. The following theorem is a proposed inner bound along these lines.

Theorem 4: For a broadcast channel $q(y, z|x)$, consider two i.i.d. copies (U_1, V_1, W_1) and (U_2, V_2, W_2) and a conditional pmf $r(x|u_1, v_1, w_1, u_2, v_2, w_2)$. Assume that $U_1, V_1, W_1, U_2, V_2, W_2, X_1, X_2, Y_1, Y_2, Z_1, Z_2$ are distributed according to

$$\begin{aligned} & p(u_1, v_1, w_1, u_2, v_2, w_2, x_1, y_1, z_1, x_2, y_2, z_2) = \\ & r(u_1, v_1, w_1) r(u_2, v_2, w_2) \cdot \\ & r(x_2|u_1, v_1, w_1, u_2, v_2, w_2) q(y_2, z_2|x_2) \cdot \\ & \tilde{r}(x_1|u_1, v_1, w_1) q(y_1, z_1|x_1), \end{aligned}$$

where $\tilde{r}(x|u, v, w)$ is defined as

$$\sum_{u' \in \mathcal{U}, v' \in \mathcal{V}, w' \in \mathcal{W}} r(x|u', v', w', u, v, w) r(u', v', w').$$

Then a rate triple (R_0, R_1, R_2) is achievable if

$$\begin{aligned} & R_0, R_1, R_2 \geq 0, \\ & R_0 + R_1 < I(U_2 W_2; Y_1 Y_2 U_1 W_1), \\ & R_0 + R_2 < I(V_2 W_2; Z_1 Z_2 V_1 W_1), \\ & R_0 + R_1 + R_2 < I(V_2; Z_1 Z_2 V_1 W_1 | W_2) \\ & \quad + I(U_2 W_2; Y_1 Y_2 U_1 W_1) - I(U_2; V_2 | W_2), \\ & R_0 + R_1 + R_2 < I(U_2; Y_1 Y_2 U_1 W_1 | W_2) \\ & \quad + I(V_2 W_2; Z_1 Z_2 V_1 W_1) - I(U_2; V_2 | W_2), \\ & 2R_0 + R_1 + R_2 < I(U_2 W_2; Y_1 Y_2 U_1 W_1) \\ & \quad + I(V_2 W_2; Z_1 Z_2 V_1 W_1) - I(U_2; V_2 | W_2), \end{aligned}$$

for some $U_1, V_1, W_1, U_2, V_2, W_2, X_1, X_2$ that satisfy the above conditions.

Remark 2: The above inner bound reduces to Marton's inner bound if the conditional distribution $r(x|u_1, v_1, w_1, u_2, v_2, w_2) = r(x|u_2, v_2, w_2)$, i.e. $U_1 V_1 W_1 \rightarrow U_2 V_2 W_2 \rightarrow X$ form a Markov chain.

III. PROOFS

Proof of Claim 1: Consider the degraded broadcast channel $p(y, z|x) = p(y|x)p(z|y)$, where the channel from X to Y is a BSC(0.3) and the channel from Y to Z is as follows: $p_{Z|Y}(0|0) = 0.6$, $p_{Z|Y}(1|0) = 0.4$, $p_{Z|Y}(0|1) = 0$, $p_{Z|Y}(1|1) = 1$. We show that the private message capacity

region for this channel is strictly larger than Marton's inner bound without W .

We first intuitively sketch outline of the proof: take a non-negative real α and consider the maximum of $R_1 + \alpha R_2$ over the pairs (R_1, R_2) in the capacity region. Since the broadcast channel is degraded, the maximum is equal to $\max_{V \rightarrow X \rightarrow YZ} I(X; Y|V) + \alpha I(V; Z)$. Since $X \rightarrow Y \rightarrow Z$, when the weight of the degraded receiver is less than or equal to 1, an optimum V will be equal to a constant (corresponding to $R_2 = 0$). As we gradually increase α beyond one, the optimum V gradually moves from a constant random variable to X (corresponding to $R_1 = 0$). Now, let us consider the maximum of $R_1 + \alpha R_2$ over the pairs (R_1, R_2) in Marton's inner bound without the auxiliary random variable W . The latter maximum is equal to $I(U; Y) + \alpha I(V; Z) - I(U; V)$. When $\alpha \leq 1$, it is optimum to take $U = X$, $V = \text{constant}$ and dedicate all the rate to the stronger receiver. Simulation results however indicate that as we increase α beyond one in the problem of maximizing $I(U; Y) + \alpha I(V; Z) - I(U; V)$, $U = X$, $V = \text{constant}$ continues to be optimal up to a threshold. Beyond this threshold, suddenly $U = \text{constant}$, $V = X$ becomes the optimizing choice, and stays as the optimizing choice afterwards. In other words, unlike the gradual transition of the maximizing V for the actual region, there is a sharp transition in the maximizing V for Marton's inner bound without W .

In the following, we provide a more detailed proof: the maximum of $R_1 + 2.4R_2$ over pairs (R_1, R_2) in the capacity region, is equal to $\max_{V \rightarrow X \rightarrow YZ} I(X; Y|V) + 2.4I(V; Z)$. Take the joint pmf of $p(v, x)$ to be as follows: $P(V = 0, X = 0) = 0$, $P(V = 0, X = 1) = 0.41$, $P(V = 1, X = 0) = 0.48$, $P(V = 1, X = 1) = 0.11$. For this choice of $p(v, x)$, $I(X; Y|V) + 2.4I(V; Z) = 0.1229\dots$. Therefore the maximum of $R_1 + 2.4R_2 \geq 0.1229\dots$. The maximum of $R_1 + 2.4R_2$ over Marton's inner bound without W is equal to $\sup_{UV \rightarrow X \rightarrow YZ} I(U; V) + 2.4I(V; Z) - I(U; V)$. Using the perturbation method of [8], one can bound the cardinality of U and V from above by $|\mathcal{X}|$, and further assume that X is a deterministic function of (U, V) . This makes the domain compact, implying that the above supremum is indeed a maximum.

Since X is a binary random variable, we need to search over binary random variables U, V . Numerical simulations show that the maximum is equal to $0.1215\dots < 0.1229\dots$ and occurs when $X = V$ and $U = \text{constant}$. Therefore Marton's inner bound without W is not tight for this broadcast channel. ■

Proof of Claim 2: In order to prove the observation, one needs to argue that the following exchange of max and min is legitimate:

$$\begin{aligned} & \max_{p(u, v, w, x)} \min_{\lambda \in [0, 1]} \lambda I(W; Y) + (1 - \lambda) I(W; Z) + \\ & \quad I(U; Y|W) + I(V; Z|W) - I(U; V|W) = \\ & \min_{\lambda \in [0, 1]} \max_{p(u, v, w, x)} \lambda I(W; Y) + (1 - \lambda) I(W; Z) + \\ & \quad I(U; Y|W) + I(V; Z|W) - I(U; V|W). \end{aligned}$$

Let $R_{\text{Marton-Sum}}$ denote the sum-rate for Marton's inner bound. We would like to show that $R_{\text{Marton-Sum}}$ is equal to $\min_{0 \leq \lambda \leq 1} T_\lambda$.

Let \mathcal{D} be the union over all $p(u, v, w, x)$ of real pairs (d_1, d_2) satisfying

$$\begin{aligned} d_1 &\leq I(W; Y) + I(U; Y|W) + I(V; Z|W) - I(U; V|W), \\ d_2 &\leq I(W; Z) + I(U; Y|W) + I(V; Z|W) - I(U; V|W). \end{aligned}$$

We claim that this region is convex. Take two points (d_1, d_2) and (d'_1, d'_2) in the region. Corresponding to these are joint distributions $p(u_1, v_1, w_1, x_1)q(y_1, z_1|x_1)$ and $p(u_2, v_2, w_2, x_2)q(y_2, z_2|x_2)$. Take a uniform binary random variable Q independent of all the previously defined random variables. Set $U = U_Q$, $V = V_Q$, $W = (Q, W_Q)$, $X = X_Q$, $Y = Y_Q$, $Z = Z_Q$. We will then have

$$\begin{aligned} I(W; Y) + I(U; Y|W) + I(V; Z|W) - I(U; V|W) &= \\ I(W_Q, Q; Y_Q) + I(U_Q; Y_Q|W_Q, Q) + \\ I(V_Q; Z_Q|W_Q, Q) - I(U_Q; V_Q|W_Q, Q) &\geq \\ I(W_Q; Y_Q|Q) + I(U_Q; Y_Q|W_Q, Q) + \\ I(V_Q; Z_Q|W_Q, Q) - I(U_Q; V_Q|W_Q, Q) &= \\ \frac{1}{2}(I(W_1; Y_1) + I(U_1; Y_1|W_1) + \\ I(V_1; Z_1|W_1) - I(U_1; V_1|W_1)) + \\ \frac{1}{2}(I(W_2; Y_2) + I(U_2; Y_2|W_2) + \\ I(V_2; Z_2|W_2) - I(U_2; V_2|W_2)) &\geq \\ \frac{1}{2}(d_1 + d'_1). \end{aligned}$$

Similarly,

$$I(W; Z) + I(U; Y|W) + I(V; Z|W) - I(U; V|W) \geq \frac{1}{2}(d_2 + d'_2).$$

Thus, the point $(\frac{1}{2}(d_1 + d'_1), \frac{1}{2}(d_2 + d'_2))$ is in the region. Thus, \mathcal{D} is convex.

Next, note that the point $(R_{\text{Marton-Sum}}, R_{\text{Marton-Sum}})$ is in \mathcal{D} . We claim that it is a boundary point of \mathcal{D} . If it is an interior point, there must exist $\epsilon > 0$ such that $(R_{\text{Marton-Sum}} + \epsilon, R_{\text{Marton-Sum}} + \epsilon)$ is in \mathcal{D} . This implies the existence of some $p(u, v, w, x)$ where

$$\begin{aligned} R_{\text{Marton-Sum}} + \epsilon &\leq \\ I(W; Y) + I(U; Y|W) + I(V; Z|W) - I(U; V|W), \\ R_{\text{Marton-Sum}} + \epsilon &\leq \\ I(W; Z) + I(U; Y|W) + I(V; Z|W) - I(U; V|W). \end{aligned}$$

This implies that

$$\begin{aligned} R_{\text{Marton-Sum}} + \epsilon &\leq \\ \min(I(W; Y), I(W; Z)) + I(U; Y|W) + \\ I(V; Z|W) - I(U; V|W) \end{aligned}$$

for some $p(u, v, w, x)$, which is a contradiction.

Using the supporting hyperplane theorem and the fact that \mathcal{D} is convex and closed, one can conclude that there exists a supporting hyperplane to \mathcal{D} at the boundary point

$(R_{\text{Marton-Sum}}, R_{\text{Marton-Sum}})$. We claim that this supporting hyperplane must have the equation $\lambda^* d_1 + (1 - \lambda^*) d_2 = T(\lambda^*)$ for some $\lambda^* \in [0, 1]$. The proof is as follows: any supporting hyperplane has the formula $\lambda^* d_1 + (1 - \lambda^*) d_2 = k$ for some real λ^* and real k . We claim that λ^* must be in $[0, 1]$ and $k = T(\lambda^*)$. Assume that for instance $\lambda^* < 0$. We know that \mathcal{D} must be entirely contained in one of the two closed half-spaces determined by the hyperplane. Note that the points $(0, 0)$, $(-\infty, 0)$ and $(0, -\infty)$ are in \mathcal{D} (take $p(u, v, w, x)$ satisfying $I(U; V|W) = 0$ in the definition of \mathcal{D}). The value of $\lambda^* d_1 + (1 - \lambda^*) d_2$ at these points is equal to 0, $+\infty$ and $-\infty$ respectively. Thus, \mathcal{D} cannot possibly be entirely contained in one of the two closed half-spaces determined by the hyperplane. Similarly the case $1 - \lambda^* < 0$ can be refuted. Therefore λ^* must be in $[0, 1]$. Since the points $(-\infty, 0)$ and $(0, -\infty)$ are in \mathcal{D} , the half-space determined by the hyperplane that contains \mathcal{D} is the one determined by the equation $\lambda^* d_1 + (1 - \lambda^*) d_2 \leq k$ for some k . Since the half-space has at least one point of \mathcal{D} , the value of k must be equal to $\max_{(d_1, d_2) \in \mathcal{D}} \lambda^* d_1 + (1 - \lambda^*) d_2$. The latter is equal to $T(\lambda^*)$. Thus, the supporting hyperplane at the boundary point $(R_{\text{Marton-Sum}}, R_{\text{Marton-Sum}})$ has the equation $\lambda^* d_1 + (1 - \lambda^*) d_2 = T(\lambda^*)$ for some $\lambda^* \in [0, 1]$.

Since $(R_{\text{Marton-Sum}}, R_{\text{Marton-Sum}})$ lies on this hyperplane, $\lambda^* R_{\text{Marton-Sum}} + (1 - \lambda^*) R_{\text{Marton-Sum}} = T(\lambda^*)$ implies that $R_{\text{Marton-Sum}} = T(\lambda^*)$ for some $\lambda^* \in [0, 1]$. Therefore

$$\min_{0 \leq \lambda \leq 1} T_\lambda \leq R_{\text{Marton-Sum}}.$$

On the other hand, for every λ , $T_\lambda \geq R_{\text{Marton-Sum}}$. Therefore

$$\min_{0 \leq \lambda \leq 1} T_\lambda \geq R_{\text{Marton-Sum}}.$$

Proof of Theorem 1: We begin by proving the **first bullet**.

$p(u, y)$ is positive for all u, y because there must exist some x such that $p(u, x) > 0$. Since the transition matrices have positive entries and $p(u, y) \geq p(u, x)q(y|x)$, $p(u, y)$ will be positive for all y . A similar argument proves that $p(v, z) > 0$ for all v, z . Next assume that $p(u, v) = 0$ for some (u, v) . Take some u', v' such that $p(u', v') > 0$. Let us reduce $p(u', v')$ by ϵ and increase $p(u, v)$ by ϵ . Furthermore, have (u, v) mapped to the same x that (u', v') was mapped to; this ensures that the marginal distribution of X is preserved. One can write

$$\begin{aligned} I(U; Y) + I(V; Z) - I(U; V) &= \\ H(Y) + H(Z) + H(UV) - H(UY) - H(VZ). \end{aligned}$$

The only change in this expression comes from the change in $H(UV) - H(UY) - H(VZ)$. The derivative of $H(UV)$ with respect to ϵ , at $\epsilon = 0$, will be infinity. But the derivative of $H(UY)$ and $H(VZ)$ will be finite since $p(u, y)$, $p(u', y)$, $p(v, z)$ and $p(v', z)$ are positive for all y and z . So, the first derivative of $H(UV) - H(UY) - H(VZ)$ with respect to ϵ , at $\epsilon = 0$, will be positive. This is a contradiction since $p(u, v|x)$ was assumed to maximize $I(U; Y) + I(V; Z) - I(U; V)$.

We now prove the **second bullet**. Assume that $\mathcal{U} = \{u_1, u_2, \dots, u_{|\mathcal{U}|}\}$ and $\mathcal{V} = \{v_1, v_2, \dots, v_{|\mathcal{V}|}\}$. Let $\pi_{i,j} = p(u_i, v_j)$ for $i = 1, \dots, |\mathcal{U}|$, $j = 1, \dots, |\mathcal{V}|$. From the first bullet we know that $\pi_{i,j} > 0$ for all i and j . Let $\bar{\epsilon} = \min_{i,j} \pi_{i,j}$. Take some $\epsilon \in (0, \bar{\epsilon})$. Let $x = \xi_0(u, v)$ denote the deterministic mapping from $\mathcal{U} \times \mathcal{V}$ to \mathcal{X} .

We prove the statement by contradiction. Assume that

$$\vec{v}_{\xi_0} = \sum_{t=1}^M \alpha_t \vec{v}_{\xi_t},$$

for some mappings ξ_t ($t = 1, 2, \dots, M$) distinct from ξ_0 and non-negative numbers α_t adding up to one.

Let random variables $T_{i,j}$ (for $i = 1, \dots, |\mathcal{U}|$, $j = 1, 2, 3, \dots, |\mathcal{V}|$) be $M+1$ -ary random variables mutually independent of each other, and of U, V, X, Y, Z satisfying:

- $p(T_{i,j} = 0) = 1 - \frac{\epsilon}{\pi_{i,j}}$,
- $p(T_{i,j} = 1) = \frac{\epsilon}{\pi_{i,j}} \alpha_1$,
- $p(T_{i,j} = 2) = \frac{\epsilon}{\pi_{i,j}} \alpha_2$,
- $p(T_{i,j} = 3) = \frac{\epsilon}{\pi_{i,j}} \alpha_3$,
- ...
- $p(T_{i,j} = M) = \frac{\epsilon}{\pi_{i,j}} \alpha_M$.

Let \tilde{X} be defined as follows:

- On the event that $(U, V) = (u_i, v_j)$, let \tilde{X} be equal to $\xi_{T_{i,j}}(u_i, v_j)$. In other words, if $T_{i,j} = 0$, \tilde{X} is equal to $\xi_0(u_i, v_j)$; if $T_{i,j} = 1$, \tilde{X} is equal to $\xi_1(u_i, v_j)$, etc.

We claim that $p(\tilde{X} = x|U = u_i) = p(X = x|U = u_i)$ for all $i = 1, 2, 3, \dots, |\mathcal{U}|$ and x ; and similarly $p(\tilde{X} = x|V = v_j) = p(X = x|V = v_j)$ for all $j = 1, 2, 3, \dots, |\mathcal{V}|$ and x . This is proved in Appendix B. Note that the above property implies that \tilde{X} and X have the same marginal distributions.

Let \tilde{Y} and \tilde{Z} be defined such that $UV(T_{i,j})_{i:1,2,\dots,j=1,2,\dots} \rightarrow \tilde{X} \rightarrow \tilde{Y}\tilde{Z}$, and the conditional law of \tilde{y} and \tilde{z} given \tilde{x} is the same as $q(y, z|x)$. Here $(T_{i,j})_{i:1,2,\dots,j=1,2,\dots}$ denotes the collection of all $T_{i,j}$ for all i and j .

Without loss of generality, let us assume $\alpha_1 \neq 0$. Since the mapping $\xi_0(\cdot, \cdot)$ is not equal to $\xi_1(\cdot, \cdot)$, there must exist (i, j) such that $\xi_0(u_i, v_j) \neq \xi_1(u_i, v_j)$. Let us label the input symbol $\xi_0(u_i, v_j)$ by x_0 , and the input symbol $\xi_1(u_i, v_j)$ by x_1 . We know that the channel is irreducible. Let us then assume that there is some y such that $q(y|x_0) \neq q(y|x_1)$; the proof for the case when there is some z such that $q(z|x_0) \neq q(z|x_1)$ is similar. Let $\tilde{U} = (U, T_{i,j})$ and $\tilde{V} = V$.

Since $p(\tilde{X} = x|U = u) = p(X = x|U = u)$ for all u and x , and $p(\tilde{X} = x|V = v) = p(X = x|V = v)$ for all v and x , we have

- $I(U; \tilde{Y}) = I(U; Y)$,
- $I(V; \tilde{Z}) = I(V; Z)$.

Therefore $I(\tilde{V}; \tilde{Z}) = I(V; Z)$ and $I(\tilde{U}; \tilde{Y}) = I(U; Y) + I(T_{i,j}; \tilde{Y}|U)$. Furthermore since $T_{i,j}$ is independent of U, V , we have $I(\tilde{U}; \tilde{V}) = I(U; V)$. Therefore

$$\begin{aligned} I(\tilde{U}; \tilde{Y}) + I(\tilde{V}; \tilde{Z}) - I(\tilde{U}; \tilde{V}) &= (I(U; Y) + I(V; Z) - I(U; V)) \\ &= I(T_{i,j}; \tilde{Y}|U). \end{aligned}$$

Since $p(u, v, x)$ was maximizing $I(U; Y) + I(V; Z) - I(U; V)$ under the fixed marginal distribution on x , we must have $I(T_{i,j}; \tilde{Y}|U) = 0$. Therefore $I(T_{i,j}; \tilde{Y}|U = u_i) = 0$ holds as well.

In Appendix C, we prove that the following are true

$$p(\tilde{X} = x_0|U = u_i, T_{i,j} = 0) \neq p(\tilde{X} = x_0|U = u_i, T_{i,j} = 1),$$

$$p(\tilde{X} = x_1|U = u_i, T_{i,j} = 0) \neq p(\tilde{X} = x_1|U = u_i, T_{i,j} = 1).$$

But for any $x \notin \{x_0, x_1\}$,

$$p(\tilde{X} = x|U = u_i, T_{i,j} = 0) = p(\tilde{X} = x|U = u_i, T_{i,j} = 1).$$

Remember that we assumed that there is some y such that $q(y|x_0) \neq q(y|x_1)$. In Appendix D, we show that

$$p(\tilde{Y} = y|U = u_i, T_{i,j} = 0) \neq p(\tilde{Y} = y|U = u_i, T_{i,j} = 1).$$

This implies that \tilde{Y} and $T_{i,j}$ are not conditionally independent given $U = u_i$. Therefore $I(T_{i,j}; \tilde{Y}|U = u_i) \neq 0$ which is a contradiction.

We now prove the **third bullet**. The proof begins by noting that the definition of $h(x)$ implies that for any (u, v, x) ,

$$h(x) \leq \log(p(u, v)) - f_u(x) - g_v(x).$$

Therefore, for any (u, v, x) ,

$$\log(p(u, v)) \geq f_u(x) + g_v(x) + h(x).$$

Thus,

$$\log(p(u, v)) \geq \max_x (f_u(x) + g_v(x) + h(x)). \quad (11)$$

Note that the first partial derivative of $H(UV) - H(UY) - H(VZ)$ with respect to $p(u, v, x)$ is proportional to

$$\begin{aligned} & -\log p(u, v) - 1 + \sum_y q(y|x) \log p(u, y) + 1 + \\ & \sum_z q(z|x) \log p(v, z) + 1 = \\ & -\log p(u, v) + f_u(x) + g_v(x) + 1. \end{aligned}$$

Assume that the triple (u, v, x) is such that $p(u, v, x) > 0$. Take some arbitrary u' and v' . Reducing $p(u, v, x)$ by a small ϵ and increasing $p(u', v', x)$ by ϵ does not affect the marginal distribution of X and hence should not increase the expression $H(UV) - H(UY) - H(VZ)$. Therefore the first derivative of $H(UV) - H(UY) - H(VZ)$ with respect to $p(u, v, x)$ must be greater than or equal to the first derivative of $H(UV) - H(UY) - H(VZ)$ with respect to $p(u', v', x)$. Thus,

$$\begin{aligned} & -\log p(u, v) + f_u(x) + g_v(x) + 1 \geq \\ & -\log p(u', v') + f_{u'}(x) + g_{v'}(x) + 1. \end{aligned}$$

In other words, for any arbitrary u' and v' , we have

$$\begin{aligned} & \log p(u, v) - f_u(x) - g_v(x) \leq \\ & \log p(u', v') - f_{u'}(x) - g_{v'}(x). \end{aligned}$$

Therefore

$$\begin{aligned} & \log p(u, v) - f_u(x) - g_v(x) \leq \\ & \min_{u', v'} \log p(u', v') - f_{u'}(x) - g_{v'}(x) = h(x). \end{aligned}$$

Thus, $\log p(u, v) \leq f_u(x) + g_v(x) + h(x)$ whenever $p(u, v, x) > 0$. This together with equation (11) imply that

$$\log(p(u, v)) = \max_x f_u(x) + g_v(x) + h(x),$$

and

$$p(x_0|u, v) = 1 \text{ for some } x_0 \in \mathcal{X} \Rightarrow x_0 \in \operatorname{argmax}_x f_u(x) + g_v(x) + h(x).$$

Proof of Lemma 1: This is a consequence of bullet one of Theorem 1. ■

Proof of Theorem 2: From the set of pmfs $p(u, v, w, x)$ that maximize the expression $\lambda I(W; Y) + (1 - \lambda)I(W; Z) + I(U; Y|W) + I(V; Z|W) - I(U; V|W)$, let $p_0(u, v, w, x)$ be the one that achieves the largest value of $I(W; Y) + I(W; Z)$. In Appendix A, we prove that one can find $p(\hat{u}, \hat{v}, \hat{w}, \hat{x})$ such that

- $\lambda I(W; Y) + (1 - \lambda)I(W; Z) + I(U; Y|W) + I(V; Z|W) - I(U; V|W)$ is equal to $\lambda I(\hat{W}; \hat{Y}) + (1 - \lambda)I(\hat{W}; \hat{Z}) + I(\hat{U}; \hat{Y}|\hat{W}) + I(\hat{V}; \hat{Z}|\hat{W}) - I(\hat{U}; \hat{V}|\hat{W})$,
- $I(W; Y) + I(W; Z)$ is equal to $I(\hat{W}; \hat{Y}) + I(\hat{W}; \hat{Z})$,
- $|\hat{\mathcal{U}}| \leq \min(|\mathcal{X}|, |\mathcal{Y}|)$,
- $|\hat{\mathcal{V}}| \leq \min(|\mathcal{X}|, |\mathcal{Z}|)$,
- $|\hat{\mathcal{W}}| \leq |\mathcal{X}|$,
- $H(\hat{X}|\hat{U}\hat{V}\hat{W}) = 0$.

Thus the constraints in the first and second bullets are satisfied by $p(\hat{u}, \hat{v}, \hat{w}, \hat{x})$. The second and third bullets of Theorem 1 imply that $p(\hat{u}, \hat{v}, \hat{w}, \hat{x})$ will automatically satisfy the third and fourth bullet of Theorem 2. In Appendix E, we show that the fifth bullet of Theorem 2 holds for any joint distribution that maximizes the expression $\lambda I(W; Y) + (1 - \lambda)I(W; Z) + I(U; Y|W) + I(V; Z|W) - I(U; V|W)$, and at the same time has the largest possible value of $I(W; Y) + I(W; Z)$. Thus it must also hold for $p(\hat{u}, \hat{v}, \hat{w}, \hat{x})$. ■

Proof of Theorem 3: It suffices to show that

$$\begin{aligned} & \max_{(R_0, R_1, R_2) \in \mathcal{C}(q(y, z|x))} (\lambda_0 R_0 + \lambda_1 R_1 + \lambda_2 R_2) \leq \\ & \max_{(R_0, R_2) \in \mathcal{C}_{d_1}(q(y, z|x))} \left\{ \max_{(R_0, R_1) \in \mathcal{C}_{d_2}(q(y, z|x))} (\lambda_0 R_0 + \lambda_1 R_1) \right\}. \end{aligned}$$

The key step is to show that if (R_0, R_1, R_2) is in the capacity region of a broadcast channel, then $(R_0 + \min\{R_1, R_2\}, R_1 - \min\{R_1, R_2\}, R_2 - \min\{R_1, R_2\})$ is also in the capacity region. Since $\lambda_0 \geq \lambda_1 + \lambda_2$, we then have that $\lambda_0(R_0 + \min\{R_1, R_2\}) + \lambda_1(R_1 - \min\{R_1, R_2\}) + \lambda_2(R_2 - \min\{R_1, R_2\}) \geq \lambda_0 R_0 + \lambda_1 R_1 + \lambda_2 R_2$, so at the maximum we must have $\min(R_1, R_2) = 0$. One can prove this property using the result of Willems [9], which shows that the *maximal* probability of error capacity region is equal to the *average* probability of error capacity region. Willems's proof of his result, however, is rather involved. Instead, we provide a simple direct proof. Consider an arbitrary code $(M_0, M_1, M_2, X^n, \epsilon)$.

We show that

$$\begin{aligned} & \frac{\lambda_0}{n} H(M_0) + \frac{\lambda_1}{n} H(M_1) + \frac{\lambda_2}{n} H(M_2) - O(\epsilon) \leq \\ & \max_{(R_0, R_2) \in \mathcal{C}_{d_1}(q(y, z|x))} \lambda_0 R_0 + \lambda_2 R_2, \\ & \max_{(R_0, R_1) \in \mathcal{C}_{d_2}(q(y, z|x))} \lambda_0 R_0 + \lambda_1 R_1, \end{aligned}$$

where $O(\epsilon)$ denotes a constant (depending only on $|\mathcal{X}|, |\mathcal{Y}|, |\mathcal{Z}|$) times ϵ .

Assume without loss of generality that $H(M_2) \leq H(M_1)$, i.e. $R_2 \leq R_1$. Let $\hat{W} = M_0 M_2$, $\hat{X} = X^n$, $\hat{Y} = Y^n$, $\hat{Z} = Z^n$. Note that $q(\hat{y}, \hat{z}|\hat{x})$ is the n -fold version of $q(y, z|x)$. Let us look at $\mathcal{C}_{d_1}(q(\hat{y}, \hat{z}|\hat{x}))$, evaluated at the joint pmf $p(\hat{w}, \hat{x})$:

$$\begin{aligned} \hat{R}_0 & \leq I(\hat{W}; \hat{Z}), \\ \hat{R}_1 & \leq I(\hat{X}; \hat{Y}|\hat{W}), \\ \hat{R}_0 + \hat{R}_1 & \leq I(\hat{X}; \hat{Y}). \end{aligned}$$

Note that, by Fano's inequality,

$$\begin{aligned} I(\hat{W}; \hat{Z}) &= I(M_0 M_2; Z^n) = H(M_0) + H(M_2) - O(n\epsilon), \\ I(\hat{X}; \hat{Y}|\hat{W}) &= I(X^n; Y^n|M_0 M_2) = H(M_1) - O(n\epsilon), \\ I(\hat{X}; \hat{Y}) &= I(X^n; Y^n) \geq H(M_0) + H(M_1) - O(n\epsilon). \end{aligned}$$

Therefore $\hat{R}_0 = H(M_0) + H(M_2) - O(n\epsilon) = n(R_0 + R_2) - O(n\epsilon)$ and $\hat{R}_1 = H(M_1) - H(M_2) = n(R_1 - R_2) - O(n\epsilon)$ is in $\mathcal{C}_{d_1}(q(\hat{y}, \hat{z}|\hat{x}))$. Since $q(\hat{y}, \hat{z}|\hat{x})$ is the n -fold version of $q(y, z|x)$ and $\mathcal{C}_{d_1}(q(\hat{y}, \hat{z}|\hat{x}))$ is the degraded message set capacity region for $q(\hat{y}, \hat{z}|\hat{x})$, we must have: $\mathcal{C}_{d_1}(q(\hat{y}, \hat{z}|\hat{x})) = n \cdot \mathcal{C}_{d_1}(q(y, z|x))$, where the multiplication here is pointwise. Thus, $(\frac{\hat{R}_0}{n}, \frac{\hat{R}_1}{n}) \in \mathcal{C}_{d_1}(q(y, z|x))$. We can complete the proof by letting $\epsilon \rightarrow 0$, and conclude that $(R_0 + R_2, R_1 - R_2, 0) \in \mathcal{C}_{d_1}(q(y, z|x))$, and thus also in the capacity region. ■

Proof of Theorem 4: Consider a natural number n , and define the super symbols $\tilde{X} = X_1 X_2 \dots X_n$, $\tilde{Y} = Y_1 Y_2 \dots Y_n$, $\tilde{Z} = Z_1 Z_2 \dots Z_n$ representing n -inputs and n -outputs of the product broadcast channel

$$q^n(y_1 y_2 \dots y_n, z_1 z_2 \dots z_n | x_1 x_2 \dots x_n) = \prod_{i=1}^n q(y_i, z_i | x_i).$$

Since the capacity region of the product channel $q^n(\tilde{y}, \tilde{z}|\tilde{x})$ is n times the capacity region of $q(y, z|x)$, we have $\frac{1}{n} \mathcal{C}_M(q^n(y_1 y_2 \dots y_n, z_1 z_2 \dots z_n | x_1 x_2 \dots x_n)) \subset \mathcal{C}(q(y, z|x))$. Given an *arbitrary* joint pmf $p(u^n, v^n, w^n, x^n)$, one can then show that the following region is an inner bound to

$\mathcal{C}(q(y, z|x))$:

$$R_0, R_1, R_2 \geq 0, \quad (12)$$

$$R_0 + R_1 \leq \frac{1}{n} I(U^n W^n; Y^n), \quad (12)$$

$$R_0 + R_2 \leq \frac{1}{n} I(V^n W^n; Z^n), \quad (13)$$

$$R_0 + R_1 + R_2 \leq \frac{1}{n} [I(U^n W^n; Y^n) + I(V^n; Z^n|W^n) - I(U^n; V^n|W^n)], \quad (14)$$

$$R_0 + R_1 + R_2 \leq \frac{1}{n} [I(U^n; Y^n|W^n) + I(V^n W^n; Z^n) - I(U^n; V^n|W^n)], \quad (15)$$

$$2R_0 + R_1 + R_2 \leq \frac{1}{n} [I(U^n W^n; Y^n) + I(V^n W^n; Z^n) - I(U^n; V^n|W^n)], \quad (16)$$

where $U^n, V^n, W^n, X^n, Y^n, Z^n$ are distributed according to $p(u^n, v^n, w^n, x^n)q(y^n, z^n|x^n)$. Clearly if we assume that (U^n, V^n, W^n, X^n) is n i.i.d. copies of $p(u, v, w, x)$ we get back the one-letter version of Marton's inner bound.

Assume that

$$p(u^n, v^n, w^n) = \prod_{i=1}^n r(u_i, v_i, w_i).$$

Note that U_i, V_i, W_i are i.i.d. copies of (U, V, W) distributed according to $r(u, v, w)$. We further use the given conditional law $r(x|u_1, v_1, w_1, u_2, v_2, w_2)$ to define the joint distribution of X^n given U^n, V^n, W^n as

$$p(x_2^n|u^n, v^n, w^n) = \prod_{i=2}^n r(x_i|u_{i-1}, v_{i-1}, w_{i-1}, u_i, v_i, w_i),$$

$X_1 = \text{constant}$.

We then have

$$\begin{aligned} I(U^n W^n; Y^n) &= H(U^n W^n) - H(U^n W^n|Y^n) = \\ &= \sum_{i=1}^n H(U_i W_i) - H(U_i W_i|U^{i-1} W^{i-1} Y^n) = \\ &= \sum_{i=1}^n I(U_i W_i; U^{i-1} W^{i-1} Y^n) \geq \sum_{i=2}^n I(U_i W_i; U_{i-1} W_{i-1} Y_i Y_{i-1}) \\ &= (n-1)I(U_2 W_2; Y_1 Y_2 U_1 W_1). \end{aligned}$$

Similarly $I(V^n W^n; Z^n) \geq (n-1)I(V_2 W_2; V_1 W_1 Z_1 Z_2)$.

Next, note that

$$\begin{aligned} I(V^n; Z^n|W^n) &= H(V^n|W^n) - H(V^n|W^n Z^n) = \\ &= \sum_{i=1}^n H(V_i|W_i) - H(V_i|V^{i-1} W^n Z^n) = \\ &= \sum_{i=1}^n I(V_i; V^{i-1} W^n Z^n|W_i) \geq \\ &= \sum_{i=2}^n I(V_i; V_{i-1} W_{i-1} Z_i Z_{i-1}|W_i) = \\ &= (n-1)I(V_2; V_1 W_1 Z_2 Z_1|W_2). \end{aligned}$$

Similarly, $I(U^n; Y^n|W^n) \geq (n-1)I(U_2; Y_1 Y_2 U_1 W_1|W_2)$. Lastly, note that $I(U^n; V^n|W^n) = n \cdot I(U; V|W)$. We obtain the desired result by substituting these values into equations (12)-(16), and letting $n \rightarrow \infty$. ■

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APPENDIX A

Suppose $p_0(u, v, w, x)$ is a joint distribution that maximizes $\lambda I(W; Y) + (1 - \lambda)I(W; Z) + I(U; Y|W) + I(V; Z|W) - I(U; V|W)$, and among all such joint distributions has the largest value of $I(W; Y) + I(W; Z)$. In this appendix, we prove that one can find $p(\hat{u}, \hat{v}, \hat{w}, \hat{x})$ such that

- $\lambda I(W; Y) + (1 - \lambda)I(W; Z) + I(U; Y|W) + I(V; Z|W) - I(U; V|W)$ is equal to $\lambda I(\hat{W}; \hat{Y}) + (1 - \lambda)I(\hat{W}; \hat{Z}) + I(\hat{U}; \hat{Y}|\hat{W}) + I(\hat{V}; \hat{Z}|\hat{W}) - I(\hat{U}; \hat{V}|\hat{W})$,
- $I(W; Y) + I(W; Z)$ is equal to $I(\hat{W}; \hat{Y}) + I(\hat{W}; \hat{Z})$,
- $|\hat{\mathcal{U}}| \leq \min(|\mathcal{X}|, |\mathcal{Y}|)$,

- $|\widehat{\mathcal{V}}| \leq \min(|\mathcal{X}|, |\mathcal{Z}|)$,
- $|\widehat{\mathcal{W}}| \leq |\mathcal{X}|$,
- $H(\widehat{X}|\widehat{U}\widehat{V}\widehat{W}) = 0$.

We begin by reducing the cardinality of W . Assume that $|\mathcal{W}| > |\mathcal{X}|$ and $p(w) \neq 0$ for all w . There must therefore exist a function $L : \mathcal{W} \rightarrow \mathbb{R}$ where

$$\mathbb{E}[L(W)|X] = 0,$$

$$\exists w : p(w) \neq 0, \quad L(w) \neq 0.$$

Let us perturb $p_0(u, v, w, x)$ along L as follows:

$$p_\epsilon(u, v, w, x, y, z) = p_0(u, v, w, x, y, z) \cdot [1 + \epsilon L(w)],$$

where ϵ is a real number in some interval $[-\bar{\epsilon}_1, \bar{\epsilon}_2]$ for some positive reals $\bar{\epsilon}_1$ and $\bar{\epsilon}_2$.

Consider the expression $\lambda I(W; Y) + (1 - \lambda)I(W; Z) + I(U; Y|W) + I(V; Z|W) - I(U; V|W)$ at $p_\epsilon(u, v, w, x, y, z)$. It can be verified that the expression is a linear function of ϵ under this perturbation. Since a maximum of this expression occurs at $\epsilon = 0$, which is a point strictly inside the interval $[-\bar{\epsilon}_1, \bar{\epsilon}_2]$, it must be the case that this expression is a constant function of ϵ . Next consider the expression $I(W; Y) + I(W; Z)$ at $p_\epsilon(u, v, w, x, y, z)$. It can be verified that the expression is a linear function of ϵ under this perturbation. Note that $p_0(u, v, w, x)$ is a joint distribution that has the largest value of $I(W; Y) + I(W; Z)$ among all joint distributions that maximize $\lambda I(W; Y) + (1 - \lambda)I(W; Z) + I(U; Y|W) + I(V; Z|W) - I(U; V|W)$. Thus a maximum of $I(W; Y) + I(W; Z)$ occurs at $\epsilon = 0$, which is a point strictly inside the interval $[-\bar{\epsilon}_1, \bar{\epsilon}_2]$. But this can only happen when $I(W; Y) + I(W; Z)$ is a constant function of ϵ . Now, taking $\epsilon = -\bar{\epsilon}_1$ or $\epsilon = \bar{\epsilon}_2$ gives us a joint distribution with the same values of $\lambda I(W; Y) + (1 - \lambda)I(W; Z) + I(U; Y|W) + I(V; Z|W) - I(U; V|W)$ and $I(W; Y) + I(W; Z)$, but with a smaller support on \mathcal{W} . Using this argument, one can reduce the cardinality of W to $|\mathcal{X}|$.

Next, we show how one can reduce the cardinality of U to find $p(\widehat{u}, \widehat{v}, \widehat{w}, \widehat{x})$ such that

- $\lambda I(W; Y) + (1 - \lambda)I(W; Z) + I(U; Y|W) + I(V; Z|W) - I(U; V|W)$ is equal to $\lambda I(\widehat{W}; \widehat{Y}) + (1 - \lambda)I(\widehat{W}; \widehat{Z}) + I(\widehat{U}; \widehat{Y}|\widehat{W}) + I(\widehat{V}; \widehat{Z}|\widehat{W}) - I(\widehat{U}; \widehat{V}|\widehat{W})$,
- $I(W; Y) + I(W; Z)$ is equal to $I(\widehat{W}; \widehat{Y}) + I(\widehat{W}; \widehat{Z})$,
- $|\widehat{\mathcal{U}}| \leq \min(|\mathcal{X}|, |\mathcal{Y}|)$,
- $|\widehat{\mathcal{W}}| \leq |\mathcal{X}|$.

We can repeat a similar procedure to impose the constraint $|\widehat{\mathcal{V}}| \leq \min(|\mathcal{X}|, |\mathcal{Z}|)$. Imposing the extra constraint $H(\widehat{X}|\widehat{U}\widehat{V}\widehat{W}) = 0$ will be discussed at the end.

If $|\mathcal{X}| \leq |\mathcal{Y}|$, establishing the cardinality bound of $|\mathcal{X}|$ on U suffices. This cardinality bound is proved in Theorem 1 of [8]. This cardinality bound can be shown using perturbations of the type $L : \mathcal{U} \times \mathcal{W} \rightarrow \mathbb{R}$ where

$$\mathbb{E}[L(U, W)|WX] = 0.$$

Note that these perturbations preserve the marginal distribution of $p(w, x)$, and thus also $I(W; Y) + I(W; Z)$. The interesting

case is therefore when $|\mathcal{X}| > |\mathcal{Y}|$. Assume that $|\mathcal{U}| > |\mathcal{Y}|$. If for every $w \in \mathcal{W}$, $p(u|w) \neq 0$ for at most $|\mathcal{Y}|$ elements u , we are done, since we can relabel the elements in the range of U to ensure that only an alphabet of size at most $|\mathcal{Y}|$ is used, without affecting any of the mutual information terms in the expression of interest. There must therefore exist a function $L : \mathcal{U} \times \mathcal{W} \rightarrow \mathbb{R}$ where

$$\mathbb{E}[L(U, W)|WY] = 0,$$

$$\exists(u, w) : p_0(u, w) \neq 0, \quad L(u, w) \neq 0.$$

Let us perturb $p_0(u, v, w, x)$ along the random variable $L : \mathcal{U} \times \mathcal{W} \rightarrow \mathbb{R}$. Random variables $\widetilde{U}, \widetilde{V}, \widetilde{W}, \widetilde{X}, \widetilde{Y}, \widetilde{Z}$ are distributed according to $p_\epsilon(\widetilde{u}, \widetilde{v}, \widetilde{w}, \widetilde{x}, \widetilde{y}, \widetilde{z})$ defined as follows

$$p_\epsilon(\widetilde{u}, \widetilde{v}, \widetilde{w}, \widetilde{x}, \widetilde{y}, \widetilde{z}) = p_0(\widetilde{u}, \widetilde{v}, \widetilde{w}, \widetilde{x}, \widetilde{y}, \widetilde{z}) \cdot [1 + \epsilon L(\widetilde{u}, \widetilde{w})],$$

where ϵ is a real number in some interval $[-\bar{\epsilon}_1, \bar{\epsilon}_2]$.

The first derivative of $\lambda I(W; Y) + (1 - \lambda)I(W; Z) + I(U; Y|W) + I(V; Z|W) - I(U; V|W)$ with respect to ϵ , at $\epsilon = 0$ should be zero. Since

$$\begin{aligned} & \lambda I(W; Y) + (1 - \lambda)I(W; Z) + I(U; Y|W) \\ & + I(V; Z|W) - I(U; V|W) = \\ & \lambda(H(W) + H(Y) - H(WY)) \\ & + (1 - \lambda)(H(W) + H(Z) - H(WZ)) + \\ & H(YW) + H(ZW) - H(UYW) \\ & - H(VZW) + H(UVW) - H(W), \end{aligned}$$

we will have:

$$\begin{aligned} & \lambda(H_L(W) + H_L(Y) - H_L(WY)) + \\ & (1 - \lambda)(H_L(W) + H_L(Z) - H_L(WZ)) \\ & + H_L(YW) + H_L(ZW) - H_L(UYW) \\ & - H_L(VZW) + H_L(UVW) - H_L(W) = 0, \end{aligned}$$

where $H_L(W)$ denotes $\sum_w E[L|W = w]p(w) \log \frac{1}{p(w)}$ and similarly for the other terms. Using Lemma 2 of [8], we have:

$$\begin{aligned} & \lambda I(\widetilde{W}; \widetilde{Y}) + (1 - \lambda)I(\widetilde{W}; \widetilde{Z}) + I(\widetilde{U}; \widetilde{Y}|\widetilde{W}) \\ & + I(\widetilde{V}; \widetilde{Z}|\widetilde{W}) - I(\widetilde{U}; \widetilde{V}|\widetilde{W}) = \\ & \lambda I(W; Y) + (1 - \lambda)I(W; Z) + I(U; Y|W) \\ & + I(V; Z|W) - I(U; V|W) + \\ & \lambda(-\mathbb{E}[r(\epsilon \cdot \mathbb{E}[L|W])] - \mathbb{E}[r(\epsilon \cdot \mathbb{E}[L|Y])]) \\ & + \mathbb{E}[r(\epsilon \cdot \mathbb{E}[L|WY])] + \\ & (1 - \lambda)(-\mathbb{E}[r(\epsilon \cdot \mathbb{E}[L|W])] - \mathbb{E}[r(\epsilon \cdot \mathbb{E}[L|Z])]) \\ & + \mathbb{E}[r(\epsilon \cdot \mathbb{E}[L|WZ])] + \\ & -\mathbb{E}[r(\epsilon \cdot \mathbb{E}[L|YW])] - \mathbb{E}[r(\epsilon \cdot \mathbb{E}[L|ZW])] \\ & + \mathbb{E}[r(\epsilon \cdot \mathbb{E}[L|UYW])] + \mathbb{E}[r(\epsilon \cdot \mathbb{E}[L|VWZ])] \\ & - \mathbb{E}[r(\epsilon \cdot \mathbb{E}[L|UVW])] + \mathbb{E}[r(\epsilon \cdot \mathbb{E}[L|W])], \end{aligned}$$

where $r(x) = (1+x)\log(1+x)$. Since $\mathbb{E}[L(U, W)|WY] = 0$, and L is a function of UW , we have:

$$\begin{aligned} & \lambda I(\tilde{W}; \tilde{Y}) + (1-\lambda)I(\tilde{W}; \tilde{Z}) + I(\tilde{U}; \tilde{Y}|\tilde{W}) \\ & + I(\tilde{V}; \tilde{Z}|\tilde{W}) - I(\tilde{U}; \tilde{V}|\tilde{W}) = \\ & \lambda I(W; Y) + (1-\lambda)I(W; Z) + I(U; Y|W) \\ & + I(V; Z|W) - I(U; V|W) + \\ & (1-\lambda)(-\mathbb{E}[r(\epsilon \cdot \mathbb{E}[L|Z])] + \mathbb{E}[r(\epsilon \cdot \mathbb{E}[L|WZ])]) \\ & - \mathbb{E}[r(\epsilon \cdot \mathbb{E}[L|ZW])] + \mathbb{E}[r(\epsilon \cdot \mathbb{E}[L|VWZ])]. \end{aligned}$$

Since $r(x) = (1+x)\log(1+x)$ is a convex function, we have

$$\begin{aligned} & -\mathbb{E}[r(\epsilon \cdot \mathbb{E}[L|Z])] + \mathbb{E}[r(\epsilon \cdot \mathbb{E}[L|WZ])] \geq 0, \\ & -\mathbb{E}[r(\epsilon \cdot \mathbb{E}[L|ZW])] + \mathbb{E}[r(\epsilon \cdot \mathbb{E}[L|VWZ])] \geq 0. \end{aligned}$$

Therefore for any $\epsilon \in [-\bar{\epsilon}_1, \bar{\epsilon}_2]$, we have

$$\begin{aligned} & \lambda I(\tilde{W}; \tilde{Y}) + (1-\lambda)I(\tilde{W}; \tilde{Z}) + I(\tilde{U}; \tilde{Y}|\tilde{W}) \\ & + I(\tilde{V}; \tilde{Z}|\tilde{W}) - I(\tilde{U}; \tilde{V}|\tilde{W}) \geq \\ & \lambda I(W; Y) + (1-\lambda)I(W; Z) + I(U; Y|W) \\ & + I(V; Z|W) - I(U; V|W). \end{aligned}$$

This implies that $\lambda I(\tilde{W}; \tilde{Y}) + (1-\lambda)I(\tilde{W}; \tilde{Z}) + I(\tilde{U}; \tilde{Y}|\tilde{W}) + I(\tilde{V}; \tilde{Z}|\tilde{W}) - I(\tilde{U}; \tilde{V}|\tilde{W})$ is a constant function of ϵ . The maximum of $I(\tilde{W}; \tilde{Y}) + I(\tilde{W}; \tilde{Z})$ as a function of ϵ occurs at $\epsilon = 0$. Therefore

$$I_L(W; Y) + I_L(W; Z) = 0,$$

where $I_L(W; Y)$ denotes $\sum_{u,w,y} p(u, w, y) L(u, w) \log \frac{p(w,y)}{p(w)p(y)}$, etc. (see Lemma 2 of [8]).

Using Lemma 2 of [8], one can observe that $[I(\tilde{W}; \tilde{Y}) + I(\tilde{W}; \tilde{Z})] - [I(W; Y) + I(W; Z)]$ equals

$$-\mathbb{E}[r(\epsilon \cdot \mathbb{E}[L|Z])] + \mathbb{E}[r(\epsilon \cdot \mathbb{E}[L|WZ])] \geq 0.$$

But this can only happen when $I(\tilde{W}; \tilde{Y}) + I(\tilde{W}; \tilde{Z})$ is a constant function of ϵ . Now, taking $\epsilon = -\bar{\epsilon}_1$ or $\epsilon = \bar{\epsilon}_2$ gives us auxiliary random variable (\tilde{U}, \tilde{W}) with smaller support than that of (U, W) . We can continue this process as long as there exists $w \in \mathcal{W}$, such that $p(u|w) \neq 0$ for more than $|\mathcal{Y}|$ elements u .

It remains to show that one can impose the extra constraint $H(\tilde{X}|\tilde{U}\tilde{V}\tilde{W}) = 0$. Fix $p(u, v, w)$. Consider the expressions $\lambda I(W; Y) + (1-\lambda)I(W; Z) + I(U; Y|W) + I(V; Z|W) - I(U; V|W)$ and $I(W; Y) + I(W; Z)$ as functions of the conditional distribution of $r(x|u, v, w)$. We know that for instance that the former expression is maximized at $p(x|u, v, w)$. Further, the extreme points of the corresponding region for $r(x|u, v, w)$ satisfy $r(x|u, v, w) \in \{0, 1\}$. Both of the expressions are convex functions of $r(x|u, v, w)$. This is because $I(W; Y)$ is convex in the conditional distribution $p(y|w)$; similarly $I(U; Y|W = w)$ is convex for any fixed value of w . The term $I(U; V|W)$ that appears with a negative sign is constant since the joint distribution of $p(u, v, w)$ is fixed.

We can express $p(x|u, v, w)$ as a linear combination of the extreme points of the region formed by all conditional

distributions $r(x|u, v, w)$. Since the maximum of $\lambda I(W; Y) + (1-\lambda)I(W; Z) + I(U; Y|W) + I(V; Z|W) - I(U; V|W)$ occurs at some $p(x|u, v, w)$ and the expression is convex in $r(x|u, v, w)$, the maximum must also occur at all the extreme points showing up in the linear combination. One can use the convexity of $I(W; Y) + I(W; Z)$ in $r(x|u, v, w)$ to show that the value of $I(W; Y) + I(W; Z)$ at all these extreme points must be also equal to that at $p(x|u, v, w)$.

APPENDIX B

In this Appendix we complete the proof of Theorem 1 by proving that $p(\tilde{X} = x|U = u_i) = p(X = x|U = u_i)$ for all $i = 1, 2, 3, \dots, |\mathcal{U}|$ and x ; and similarly $p(\tilde{X} = x|V = v_j) = p(X = x|V = v_j)$ for all $j = 1, 2, 3, \dots, |\mathcal{V}|$ and x .

Note that

$$\begin{aligned} & p(\tilde{X} = x|U = u_i) = \\ & \sum_j p(V = v_j|U = u_i) p(\tilde{X} = x|U = u_i, V = v_j) = \\ & \sum_j p(V = v_j|U = u_i) \sum_{k=0}^M p(T_{i,j} = k) \mathbf{1}[\xi_k(u_i, v_j) = x] = \\ & \sum_j p(V = v_j|U = u_i) (1 - \frac{\epsilon}{\pi_{i,j}}) \mathbf{1}[\xi_0(u_i, v_j) = x] + \\ & \sum_j p(V = v_j|U = u_i) \sum_{k=1}^M \frac{\epsilon}{\pi_{i,j}} \alpha_k \mathbf{1}[\xi_k(u_i, v_j) = x] = \\ & \sum_j p(V = v_j|U = u_i) (\frac{\pi_{i,j} - \epsilon}{\pi_{i,j}}) \mathbf{1}[\xi_0(u_i, v_j) = x] + \\ & \sum_{k=1}^M \sum_j p(V = v_j|U = u_i) \frac{\epsilon}{\pi_{i,j}} \alpha_k \mathbf{1}[\xi_k(u_i, v_j) = x]. \end{aligned}$$

Note that $p(V = v_j|U = u_i) = \frac{p(V=v_j, U=u_i)}{p(U=u_i)} = \frac{\pi_{i,j}}{p(U=u_i)}$. Therefore

$$\begin{aligned} & p(\tilde{X} = x|U = u_i) = \\ & \sum_j \frac{\pi_{i,j} - \epsilon}{p(U=u_i)} \mathbf{1}[\xi_0(u_i, v_j) = x] + \\ & \sum_{k=1}^M \sum_j \frac{\epsilon}{p(U=u_i)} \alpha_k \mathbf{1}[\xi_k(u_i, v_j) = x] = \\ & \sum_j \frac{\pi_{i,j}}{p(U=u_i)} \mathbf{1}[\xi_0(u_i, v_j) = x] \\ & - \frac{\epsilon}{p(U=u_i)} \sum_j \mathbf{1}[\xi_0(u_i, v_j) = x] + \\ & \frac{\epsilon}{p(U=u_i)} \sum_{k=1}^M \alpha_k \sum_j \mathbf{1}[\xi_k(u_i, v_j) = x]. \end{aligned}$$

But since

$$\vec{v}_{\xi_0} = \sum_{t=1}^M \alpha_t \vec{v}_{\xi_t},$$

the profiles of the i^{th} rows must also satisfy the same property:

$$\sum_j \mathbf{1}[\xi_0(u_i, v_j) = x] = \sum_{k=1}^M \alpha_k \sum_j \mathbf{1}[\xi_k(u_i, v_j) = x].$$

Therefore,

$$\begin{aligned} & p(\tilde{X} = x|U = u_i) = \\ & \sum_j \frac{\pi_{i,j}}{p(U=u_i)} \mathbf{1}[\xi_0(u_i, v_j) = x] + 0 - 0 = \\ & \sum_j \frac{\pi_{i,j}}{p(U=u_i)} \mathbf{1}[\xi_0(u_i, v_j) = x] = p(X = x|U = u_i). \end{aligned}$$

The equation $p(\tilde{X} = x|V = v_j) = p(X = x|V = v_j)$ for all $j = 1, 2, 3, \dots, |\mathcal{V}|$ and x can be proved similarly.

APPENDIX C

Note that

$$\begin{aligned} p(\tilde{X} = x_0|U = u_i, T_{i,j} = 0) = \\ p(\tilde{X} = x_0|U = u_i, T_{i,j} = 0, V = v_j)p(V = v_j|U = u_i, T_{i,j} = 0) \\ + \\ p(\tilde{X} = x_0|U = u_i, T_{i,j} = 0, V \neq v_j)p(V \neq v_j|U = u_i, T_{i,j} = 0). \end{aligned}$$

Since under the event $(U, V) = (u_i, v_j)$ and $T_{i,j} = 0$, \tilde{X} is equal to x_0 , the term $p(\tilde{X} = x_0|U = u_i, T_{i,j} = 0, V = v_j)$ will be equal to one. Since (U, V) is independent of $T_{i,j}$, we have

$$\begin{aligned} p(V = v_j|U = u_i, T_{i,j} = 0) &= p(V = v_j|U = u_i), \\ p(V \neq v_j|U = u_i, T_{i,j} = 0) &= p(V \neq v_j|U = u_i). \end{aligned}$$

Lastly $p(\tilde{X} = x_0|U = u_i, T_{i,j} = 0, V \neq v_j)$ is equal to $p(\tilde{X} = x_0|U = u_i, V \neq v_j)$ since under the event that $(U = u_i, V \neq v_j)$, \tilde{X} will be independent of $T_{i,j}$ (note that $T_{i,j}$ random variables were mutually independent of each other). Therefore,

$$\begin{aligned} p(\tilde{X} = x_0|U = u_i, T_{i,j} = 0) &= \\ p(V = v_j|U = u_i) + \\ p(\tilde{X} = x_0|U = u_i, V \neq v_j)p(V \neq v_j|U = u_i). \end{aligned} \quad (17)$$

Next, note that

$$\begin{aligned} p(\tilde{X} = x_0|U = u_i, T_{i,j} = 1) = \\ p(\tilde{X} = x_0|U = u_i, T_{i,j} = 1, V = v_j)p(V = v_j|U = u_i, T_{i,j} = 1) + \\ p(\tilde{X} = x_0|U = u_i, T_{i,j} = 1, V \neq v_j)p(V \neq v_j|U = u_i, T_{i,j} = 1). \end{aligned}$$

Since under the event $(U, V) = (u_i, v_j)$ and $T_{i,j} = 1$, \tilde{X} is equal to x_1 , the term $p(\tilde{X} = x_0|U = u_i, T_{i,j} = 1, V = v_j)$ will be equal to zero. Following an argument like above, one can show that

$$\begin{aligned} p(\tilde{X} = x_0|U = u_i, T_{i,j} = 1) = \\ 0 + p(\tilde{X} = x_0|U = u_i, V \neq v_j)p(V \neq v_j|U = u_i). \end{aligned} \quad (18)$$

Comparing equations (17) and (18), and noting that $p(V = v_j|U = u_i) > 0$, we conclude that

$$p(\tilde{X} = x_0|U = u_i, T_{i,j} = 0) \neq p(\tilde{X} = x_0|U = u_i, T_{i,j} = 1).$$

The proof for

$$p(\tilde{X} = x_1|U = u_i, T_{i,j} = 0) \neq p(\tilde{X} = x_1|U = u_i, T_{i,j} = 1)$$

is similar.

It remains to show that for any $x \notin \{x_0, x_1\}$,

$$p(\tilde{X} = x|U = u_i, T_{i,j} = 0) = p(\tilde{X} = x|U = u_i, T_{i,j} = 1).$$

Note that

$$\begin{aligned} p(\tilde{X} = x|U = u_i, T_{i,j} = 1) = \\ p(\tilde{X} = x|U = u_i, T_{i,j} = 1, V = v_j)p(V = v_j|U = u_i, T_{i,j} = 1) + \\ p(\tilde{X} = x|U = u_i, T_{i,j} = 1, V \neq v_j)p(V \neq v_j|U = u_i, T_{i,j} = 1) = \\ 0 + p(\tilde{X} = x|U = u_i, V \neq v_j)p(V \neq v_j|U = u_i) = \\ p(\tilde{X} = x|U = u_i, T_{i,j} = 0). \end{aligned}$$

APPENDIX D

We prove the statement by contradiction. Assume that

$$p(\tilde{Y} = y|U = u_i, T_{i,j} = 0) = p(\tilde{Y} = y|U = u_i, T_{i,j} = 1).$$

We have

$$\begin{aligned} p(\tilde{Y} = y|U = u_i, T_{i,j} = 0) = \\ p(\tilde{Y} = y|U = u_i, T_{i,j} = 0, \tilde{X} = x_0)p(\tilde{X} = x_0|U = u_i, T_{i,j} = 0) + \\ p(\tilde{Y} = y|U = u_i, T_{i,j} = 0, \tilde{X} = x_1)p(\tilde{X} = x_1|U = u_i, T_{i,j} = 0) + \\ \sum_{x \in \mathcal{X}, x \notin \{x_0, x_1\}} (p(\tilde{Y} = y|U = u_i, T_{i,j} = 0, \tilde{X} = x) \times \\ p(\tilde{X} = x|U = u_i, T_{i,j} = 0)) = \\ p(\tilde{Y} = y|\tilde{X} = x_0)p(\tilde{X} = x_0|U = u_i, T_{i,j} = 0) + \\ p(\tilde{Y} = y|\tilde{X} = x_1)p(\tilde{X} = x_1|U = u_i, T_{i,j} = 0) + \\ \sum_{x \in \mathcal{X}, x \notin \{x_0, x_1\}} (p(\tilde{Y} = y|\tilde{X} = x)p(\tilde{X} = x|U = u_i, T_{i,j} = 0)). \end{aligned}$$

Similarly,

$$\begin{aligned} p(\tilde{Y} = y|U = u_i, T_{i,j} = 1) = \\ p(\tilde{Y} = y|\tilde{X} = x_0)p(\tilde{X} = x_0|U = u_i, T_{i,j} = 1) + \\ p(\tilde{Y} = y|\tilde{X} = x_1)p(\tilde{X} = x_1|U = u_i, T_{i,j} = 1) + \\ \sum_{x \in \mathcal{X}, x \notin \{x_0, x_1\}} (p(\tilde{Y} = y|\tilde{X} = x)p(\tilde{X} = x|U = u_i, T_{i,j} = 1)). \end{aligned}$$

It was shown in Appendix C that

$$\begin{aligned} p(\tilde{X} = x_0|U = u_i, T_{i,j} = 0) &\neq p(\tilde{X} = x_0|U = u_i, T_{i,j} = 1), \\ p(\tilde{X} = x_1|U = u_i, T_{i,j} = 0) &\neq p(\tilde{X} = x_1|U = u_i, T_{i,j} = 1). \end{aligned}$$

But for any $x \notin \{x_0, x_1\}$,

$$\begin{aligned} p(\tilde{X} = x|U = u_i, T_{i,j} = 0) &= \\ p(\tilde{X} = x|U = u_i, T_{i,j} = 1). \end{aligned} \quad (19)$$

Thus, we must have

$$\begin{aligned} p(\tilde{Y} = y|\tilde{X} = x_0)p(\tilde{X} = x_0|U = u_i, T_{i,j} = 0) + \\ p(\tilde{Y} = y|\tilde{X} = x_1)p(\tilde{X} = x_1|U = u_i, T_{i,j} = 0) = \\ p(\tilde{Y} = y|\tilde{X} = x_0)p(\tilde{X} = x_0|U = u_i, T_{i,j} = 1) + \\ p(\tilde{Y} = y|\tilde{X} = x_1)p(\tilde{X} = x_1|U = u_i, T_{i,j} = 1). \end{aligned}$$

This implies that

$$\frac{p(\tilde{X}=x_0|U=u_i, T_{i,j}=1)-p(\tilde{X}=x_0|U=u_i, T_{i,j}=0)}{p(\tilde{X}=x_1|U=u_i, T_{i,j}=0)-p(\tilde{X}=x_1|U=u_i, T_{i,j}=1)} = \frac{p(\tilde{Y}=y|\tilde{X}=x_1)}{p(\tilde{Y}=y|\tilde{X}=x_0)}.$$

Note that the nominator and denominator are positive by what was proved in Appendix C.

On the other hand, we also have by equation (19):

$$\begin{aligned} p(\tilde{X} = x_0|U = u_i, T_{i,j} = 0) + \\ p(\tilde{X} = x_1|U = u_i, T_{i,j} = 0) = \\ p(\tilde{X} = x_0|U = u_i, T_{i,j} = 1) + \\ p(\tilde{X} = x_1|U = u_i, T_{i,j} = 1). \end{aligned}$$

This implies that

$$\frac{p(\tilde{X}=x_0|U=u_i, T_{i,j}=1) - p(\tilde{X}=x_0|U=u_i, T_{i,j}=0)}{p(\tilde{X}=x_1|U=u_i, T_{i,j}=0) - p(\tilde{X}=x_1|U=u_i, T_{i,j}=1)} = 1.$$

Hence,

$$\frac{p(\tilde{Y}=y|\tilde{X}=x_1)}{p(\tilde{Y}=y|\tilde{X}=x_0)} = 1.$$

But we know that $p(\tilde{Y}=y|\tilde{X}=x_0) \neq p(\tilde{Y}=y|\tilde{X}=x_1)$ since the input values x_0 and x_1 are distinguishable by the Y receiver. This is a contradiction.

APPENDIX E

The proof follows from the following two statements:

Statement 1: Assume that $p^*(u, v, w, x)$ is an arbitrary joint distribution maximizing $\lambda I(W; Y) + (1 - \lambda)I(W; Z) + I(U; Y|W) + I(V; Z|W) - I(U; V|W)$, and having the largest value of $I(W; Y) + I(W; Z)$ among all maximizing joint distributions. For every w , $p^*(x|w)$ must belong to the set $\mathcal{T}(q(y, z|x))$ defined as follows. Let $\mathcal{T}(q(y, z|x))$ be the set of pmfs on \mathcal{X} , $t(x)$, such that

$$\begin{aligned} & \max_{p(u, v, w|x)t(x)q(y, z|x)} \{ \lambda I(W; Y) + (1 - \lambda)I(W; Z) \\ & + I(U; Y|W) + I(V; Z|W) - I(U; V|W) \} \\ & = \max_{p(u, v|x)t(x)q(y, z|x)} (I(U; Y) + I(V; Z) - I(U; V)), \end{aligned}$$

and $I(W; Y) = I(W; Z) = 0$ for any² pmf $p(u, v, w|x)t(x)$ that maximizes the expression $\lambda I(W; Y) + (1 - \lambda)I(W; Z) + I(U; Y|W) + I(V; Z|W) - I(U; V|W)$.³

Statement 2: Let $q(y, z|x)$ be a general broadcast channel, and $t(x) \in \mathcal{T}(q(y, z|x))$. Consider the maximization problem: $\max_{p(u, v|x)t(x)q(y, z|x)} (I(U; Y) + I(V; Z) - I(U; V))$. Assume that a maximum occurs at $p^*(u, v|x)$. Then the following holds for random variables $(U, V, X, Y, Z) \sim p^*(u, v|x)t(x)q(y, z|x)$:

- $I(\bar{U}; Y) \geq I(\bar{U}; VZ)$ for every $\bar{U} \rightarrow U \rightarrow VXYZ$.
- $I(\bar{V}; Z) \geq I(\bar{V}; UY)$ for every $\bar{V} \rightarrow V \rightarrow UXYZ$.

A. Proof of Statement 1:

Assume that the marginal pmf of X given $W = w$ does not belong to \mathcal{T} for some w . By the definition then, at least one of the following must hold:

Case 1: Corresponding to $p_{X|W=w}^*(x)$ is the conditional distribution $p(\hat{u}, \hat{v}, \hat{w}|\hat{x})$ such that

$$\begin{aligned} & I(U; Y|W = w) + I(V; Z|W = w) - I(U; V|W = w) < \\ & \lambda I(\hat{W}; \hat{Y}) + (1 - \lambda)I(\hat{W}; \hat{Z}) + I(\hat{U}; \hat{Y}|\hat{W}) \\ & + I(\hat{V}; \hat{Z}|\hat{W}) - I(\hat{U}; \hat{V}|\hat{W}) \end{aligned} \quad (20)$$

where $p(\hat{u}, \hat{v}, \hat{w}, \hat{x}, \hat{y}, \hat{z}) = p(\hat{u}, \hat{v}, \hat{w}|\hat{x})p_{X|W=w}^*(\hat{x})q(\hat{y}, \hat{z}|\hat{x})$.

Case 2: Corresponding to $p_{X|W=w}^*(x)$ is the conditional distribution $p(\hat{u}, \hat{v}, \hat{w}|\hat{x})$ such that

$$\begin{aligned} & I(U; Y|W = w) + I(V; Z|W = w) - I(U; V|W = w) = \\ & \lambda I(\hat{W}; \hat{Y}) + (1 - \lambda)I(\hat{W}; \hat{Z}) + I(\hat{U}; \hat{Y}|\hat{W}) \\ & + I(\hat{V}; \hat{Z}|\hat{W}) - I(\hat{U}; \hat{V}|\hat{W}) \end{aligned}$$

but $I(\hat{W}; \hat{Y}) + I(\hat{W}; \hat{Z}) > 0$, where $p(\hat{u}, \hat{v}, \hat{w}, \hat{x}, \hat{y}, \hat{z}) = p(\hat{u}, \hat{v}, \hat{w}|\hat{x})p_{X|W=w}^*(\hat{x})q(\hat{y}, \hat{z}|\hat{x})$.

Define $\tilde{U}, \tilde{V}, \tilde{W}$ jointly distributed with U, V, W, X, Y, Z as follows: whenever $W \neq w$, the random variables $\tilde{U} = U, \tilde{V} = V, \tilde{W} = W$. For $W = w$, the Markov chain $\tilde{U}\tilde{V}\tilde{W} \rightarrow X \rightarrow UVWYZ$ holds, and $p(\tilde{u}, \tilde{v}, \tilde{w}|x) = p(\hat{u}, \hat{v}, \hat{w}|\hat{x})$. Next, assume that $U' = \tilde{U}, V' = \tilde{V}, W' = W\tilde{W}$.

If case 1 holds, we prove that $\lambda I(W'; Y) + (1 - \lambda)I(W'; Z) + I(U'; Y|W') + I(V'; Z|W') - I(U'; V'|W') > \lambda I(W; Y) + (1 - \lambda)I(W; Z) + I(U; Y|W) + I(V; Z|W) - I(U; V|W)$, which results in a contradiction. If case 2 holds, we prove that $\lambda I(W'; Y) + (1 - \lambda)I(W'; Z) + I(U'; Y|W') + I(V'; Z|W') - I(U'; V'|W') = \lambda I(W; Y) + (1 - \lambda)I(W; Z) + I(U; Y|W) + I(V; Z|W) - I(U; V|W)$ but that $I(W'; Y) + I(W'; Z) > I(W; Y) + I(W; Z)$, which results in a contradiction.

Assume that case 1 holds. Since $W' = W\tilde{W}$, $I(W'; Y) = I(W; Y) + I(\tilde{W}; Y|W)$ and $I(W'; Z) = I(W; Z) + I(\tilde{W}; Z|W)$, we need to show that

$$\begin{aligned} & \lambda I(\tilde{W}; Y|W) + (1 - \lambda)I(\tilde{W}; Z|W) + I(\tilde{U}; Y|W\tilde{W}) + \\ & I(\tilde{V}; Z|W\tilde{W}) - I(\tilde{U}; \tilde{V}|W\tilde{W}) > \\ & I(U; Y|W) + I(V; Z|W) - I(U; V|W) \end{aligned}$$

Remember that whenever $W \neq w$, random variables $\tilde{U}, \tilde{V}, \tilde{W}$ were defined to be equal to U, V, W . Therefore we need to show that

$$\begin{aligned} & p(W = w) [\lambda I(\tilde{W}; Y|W = w) + (1 - \lambda)I(\tilde{W}; Z|W = w) + \\ & I(\tilde{U}; Y|W = w, \tilde{W}) + I(\tilde{V}; Z|W = w, \tilde{W}) \\ & - I(\tilde{U}; \tilde{V}|W = w, \tilde{W})] > \\ & p(W = w) [I(U; Y|W = w) + I(V; Z|W = w) \\ & - I(U; V|W = w)]. \end{aligned}$$

On the event $W = w$, random variables $\tilde{U}, \tilde{V}, \tilde{W}$ were defined so that $p(\tilde{u}, \tilde{v}, \tilde{w}|x)$ is equal to $p(\hat{u}, \hat{v}, \hat{w}|\hat{x})$. Furthermore the marginal distribution of $p(\hat{x})$ is $p^*(x|W = w)$. Therefore $I(\hat{W}; Y|W = w) = I(\hat{W}; \hat{Y})$, $I(\hat{W}; Z|W = w) = I(\hat{W}; \hat{Z})$, $I(\hat{U}; Y|W = w, \tilde{W}) = I(\hat{U}; \hat{Y}|\hat{W})$, etc. Thus it remains to show that

$$\begin{aligned} & \lambda I(\hat{W}; \hat{Y}) + (1 - \lambda)I(\hat{W}; \hat{Z}) + I(\hat{U}; \hat{Y}|\hat{W}) \\ & + I(\hat{V}; \hat{Z}|\hat{W}) - I(\hat{U}; \hat{V}|\hat{W}) > \end{aligned}$$

$$I(U; Y|W = w) + I(V; Z|W = w) - I(U; V|W = w).$$

This holds because of equation (20). This concludes the proof for case 1.

²Note that such a pmf may not unique.

³We have used maximum and not supremum in the above conditions since cardinality bounds on the auxiliary random variables exist [8].

Now, assume that case 2 holds. Following, the above proof for case 1, one can get

$$\begin{aligned} & \lambda I(W'; Y) + (1 - \lambda)I(W'; Z) + I(U'; Y|W') \\ & + I(V'; Z|W') - I(U'; V'|W') \geq \\ & \lambda I(W; Y) + (1 - \lambda)I(W; Z) + I(U; Y|W) \\ & + I(V; Z|W) - I(U; V|W). \end{aligned}$$

Note that $I(W'; Y) + I(W'; Z) = I(W; Y) + I(\widetilde{W}; Y|W) + I(W; Z) + I(\widetilde{W}; Z|W)$. Thus, we need to show that $I(\widetilde{W}; Y|W) + I(\widetilde{W}; Z|W) > 0$. Note that

$$\begin{aligned} & I(\widetilde{W}; Y|W) + I(\widetilde{W}; Z|W) = \\ & p(W = w)(I(\widetilde{W}; Y|W = w) + I(\widetilde{W}; Z|W = w)) \\ & = p(W = w)(I(\widehat{W}; \widehat{Y}) + I(\widehat{W}; \widehat{Z})) > 0. \end{aligned}$$

B. Proof of Statement 2:

Take an arbitrary \overline{U} satisfying $\overline{U} \rightarrow U \rightarrow VXYZ$. Let $\widehat{W} = \overline{U}$, $\widehat{U} = U$, $\widehat{V} = V$. Since $t(x) \in \mathcal{T}(q(y, z|x))$, and $p^*(u, v|x)$ maximizes $I(U; Y) + I(V; Z) - I(U; V)$, we can write:

$$\begin{aligned} & I(U; Y) + I(V; Z) - I(U; V) \geq \\ & \lambda I(\widehat{W}; Y) + (1 - \lambda)I(\widehat{W}; Z) + I(\widehat{U}; Y|\widehat{W}) + I(\widehat{V}; Z|\widehat{W}) \\ & - I(\widehat{U}; \widehat{V}|\widehat{W}), \end{aligned} \quad (21)$$

and furthermore if equality holds, we must have $I(\widehat{W}; Y) = I(\widehat{W}; Z) = 0$. We prove that this implies that $I(\overline{U}; Y) \geq I(\overline{U}; VZ)$.

We can write:

$$\begin{aligned} & I(U; Y) + I(V; Z) - I(U; V) \geq \\ & \lambda I(\widehat{W}; Y) + (1 - \lambda)I(\widehat{W}; Z) + I(\widehat{U}; Y|\widehat{W}) + I(\widehat{V}; Z|\widehat{W}) \\ & - I(\widehat{U}; \widehat{V}|\widehat{W}) = \\ & \lambda I(\overline{U}; Y) + (1 - \lambda)I(\overline{U}; Z) + I(U; Y|\overline{U}) \\ & + I(V; Z|\overline{U}) - I(U; V|\overline{U}). \end{aligned}$$

Since $\overline{U} \rightarrow U \rightarrow VXYZ$, we have $I(U; Y) = I(\overline{U}U; Y)$ and $I(U; V) = I(\overline{U}U; V)$. This implies that

$$\begin{aligned} & I(\overline{U}; Y) + I(V; Z) - I(\overline{U}; V) \geq \\ & \lambda I(\overline{U}; Y) + (1 - \lambda)I(\overline{U}; Z) + I(V; Z|\overline{U}) \end{aligned}$$

or,

$$I(\overline{U}; Y) + I(V; Z) \geq \lambda I(\overline{U}; Y) + (1 - \lambda)I(\overline{U}; Z) + I(V; Z|\overline{U})$$

or,

$$(1 - \lambda)I(\overline{U}; Y) \geq (1 - \lambda)I(\overline{U}; Z) + I(V; \overline{U}|Z).$$

In other words

$$(1 - \lambda)I(\overline{U}; Y) \geq (1 - \lambda)I(\overline{U}; VZ) + \lambda I(V; \overline{U}|Z). \quad (22)$$

Let us consider the following two cases:

- $\lambda < 1$: In this case, equation (22) implies that $I(\overline{U}; Y) \geq I(\overline{U}; VZ) + \frac{\lambda}{1-\lambda}I(V; \overline{U}|Z)$. This inequality implies the desired inequality $I(\overline{U}; Y) \geq I(\overline{U}; VZ)$.

- $\lambda = 1$: In this case, equation (22) implies that $I(V; \overline{U}|Z) = 0$. Furthermore equation (21) will hold with equality. Since $t(x) \in \mathcal{T}$, we must have $I(\overline{U}; Y) = I(\overline{U}; Z) = 0$.

The fact that $I(V; \overline{U}|Z) = I(\overline{U}; Y) = I(\overline{U}; Z) = 0$ implies that $I(\overline{U}; Y) = I(\overline{U}; ZV) = 0$. Therefore the inequality $I(\overline{U}; Y) \geq I(\overline{U}; ZV)$ also holds in this case.

In each case, we are done. The proof for the inequality $I(\overline{V}; Z) \geq I(\overline{V}; YU)$ is similar.