

Stochastic convergence of regularized solutions and their finite element approximations to inverse source problems

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Abstract

In this work, we investigate the regularized solutions and their finite element solutions to the inverse source problems governed by partial differential equations, and establish the stochastic convergence and optimal finite element convergence rates of these solutions, under pointwise measurement data with random noise. Unlike most existing regularization theories, the regularization error estimates are derived without any source conditions, while the error estimates of finite element solutions show their explicit dependence on the noise level, regularization parameter, mesh size, and time step size, which can guide practical choices among these key parameters in real applications. The error estimates also suggest an iterative algorithm for determining an optimal regularization parameter. Numerical experiments are presented to demonstrate the effectiveness of the analytical results.

Key words. Inverse source problems, regularization, finite element approximation, stochastic error estimates.

AMS subject classifications. 35R30, 65J20, 65M60, 65N21, 65N30

1 Introduction

This work presents a quantitative understanding of stochastic convergence of the regularized solutions and their finite element approximations to the inverse source problems governed by partial differential equations, under the measurement data with random noise. The inverse source problems may arise from very different applications and modeling, e.g., diffusion or groundwater flow processes [1, 4, 6, 21, 5, 29, 30], heat conduction or convection-diffusion processes [3, 20, 21, 33, 40], or acoustic problems [7, 36]. Pollutant source inversion can find many applications, e.g., indoor and outdoor air pollution, detecting and monitoring underground water pollution. Physical, chemical and biological measures have been developed for the identification of sources and source strengths [4, 48, 49]. Due to the important applications of ill-posed inverse source problems, stable numerical solutions have been widely studied, both deterministically and statistically [34, 39, 38]. A popular approach for inverse

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source problems is the least-squares optimization with appropriate regularizations [3, 21, 47], which will be also the formulation we take in this work.

Our first main result is the establishment of the optimal stochastic error estimates of the regularized solutions in terms of the noise level, without any source conditions. This presents a brand new idea in error estimates of approximate solutions to ill-posed inverse problems achieved by regularization, and it is very different from the existing regularization theories and their approximation error estimates nearly all of which were established under some source conditions. Regularization and convergence of regularized solutions have been widely studied under various source conditions. The classical source condition requires the existence of a small source function [15]. One source condition was proposed in [16] for an inverse conductivity problem to relax the restrictive requirement on the smallness of the source function in the classical convergence theory [15]. A variational source condition was proposed in [25], and were further extended in [9, 19, 22]. It is still a hot topic how to verify the classical or variational source conditions for most inverse problems under reasonable physical assumptions on the forward solutions and identifying parameters. It appears that the analytical techniques in all existing verifications of source conditions are quite different for each concrete inverse problem [11, 12, 26, 27, 32]. The current work makes a very promising first attempt to achieve the error estimates of regularized solutions, without any source conditions, hence gets rid of the technical difficulties in convergence analysis.

The second main contribution of this work is to derive the stochastic convergence and error estimates of finite element approximations to the inverse source problems. The error estimates of finite element solutions to inverse problems have been known to be quite challenging and still open to most practically important inverse problems. There have been various efforts on error estimates of finite element solutions for inverse problems, especially for inverse elliptic and parabolic equations. But most existing studies have been carried out only for some not so frequently used mathematical formulations of inverse problems; see [45] for a detailed review and related references therein. We are not aware of any error estimates of finite element solutions to the frequently used least-squares formulations with Tikhonov regularizations, especially when the observation data are treated as random variables. We had a recent study in [28] for a modified regularization formulation for an inverse stationary source problem, where error estimates were achieved under some negative norms, which, however, may be rather inconvenient to realize in applications. One of our main focuses in this work is to make an attempt to fill the gap, to provide error estimates of finite element solutions to the least-squares formulations with Tikhonov regularizations, and more importantly, the observation data will be treated fully as random variables in the entire analysis. As we shall demonstrate, the new error estimates are not only optimal but also presents explicit dependence on the critical parameters like noise level, regularization parameter, mesh size and time stepsize. Results of this type are highly desirable in real applications as they can provide explicit guidance in choosing these key parameters, and are also the major challenge and difficulty in error estimates of finite element solutions to regularized inverse problems.

We would like to mention a very important by-product from our convergence analysis, namely, it suggests a deterministic iterative algorithm for finding an effective regularization parameter. The choice of an effective regularization parameter is essential to the success of all output least-squares minimization approaches with Tikhonov regularizations, but it has remained to be a big challenge how to find an effective regularization parameter for most inverse problems.

Another feature of this work is that the entire analysis is carried out for a very practical scenario, i.e., the scattered data. We shall assume the measurement data is collected pointwise, with noise, otherwise no any additional regularity assumption is made. This is unlike analyses and results in most existing regularization theories.

We studied in a recent work [13] the stochastic convergence of a nonconforming finite element

method for the thin plate spline smoother for observational data. The spline model for scattered data has attracted considerable attention in the literature. The convergence rate in expectation of the error between the solution of the spline model and the true solution was established in [42]. Under the condition that measurement noise are sub-Gaussian random variables, the stochastic convergence of the empirical error was obtained by the peeling argument in [43] ($d = 1$) and [13] ($d = 2, 3$). We shall borrow some analytical tools from [13, 42] to study the stochastic convergence in expectation when the measurement noise is random variables having bounded variance in subsection 2.1. The peeling argument is used in subsection 2.2 to show that the empirical error has an exponential decaying tail when the measurement noise is sub-Gaussian random variables. The discretization and its error estimates are considered in section 2.3 both in the expectation and in the Orlicz norm for sub-Gaussian measurement noise. The general results developed in section 2 are applied to study an inverse nonstationary source problem in section 3. And numerical examples are presented in section 4 to demonstrate the effectiveness of our analytical results.

2 Inverse source problem

Let Ω be a bounded domain in \mathbf{R}^d ($d = 1, 2, 3$), and X and Y be two real Hilbert spaces such that Y is continuously embedded in $C(\bar{\Omega})$ and compactly embedded in $L^2(\Omega)$. The inner product and the norm of a Hilbert space H are denoted as $(\cdot, \cdot)_H$ and $\|\cdot\|_H$, respectively; but (\cdot, \cdot) is used if $H = L^2(\Omega)$. Throughout the paper, we shall use C , with or without subscript, to denote a generic constant independent of the mesh size h , the time step size τ , and it may take a different value at each occurrence.

Let S be a linear bounded operator from X to Y and $f^* \in X$ be an unknown source. We are interested in the inverse source problem of the general form:

(SIP) Given the measurement data of Sf^* , recover the source f^* .

There are many examples of inverse source problems of this type. Our studies will focus on a very important physical scenario, assuming that the pointwise measurement data is collected on a set of distributed sensors located at $\{x_i\}_{i=1}^n$ ($x_i \neq x_j$ for $i \neq j$) inside the physical domain Ω [3, 20, 5, 29, 35, 36, 37]. We assume that the measurements come with noise and takes the form

$$m_i = (Sf^*)(x_i) + e_i, \quad i = 1, 2, \dots, n, \quad (1)$$

where $e = (e_1, e_2, \dots, e_n)^T$ is the data noise vector, with $\{e_i\}_{i=1}^n$ being independent and identically distributed random variables on a probability space $(\mathfrak{X}, \mathcal{F}, \mathbb{P})$. We shall denote $m = (m_1, m_2, \dots, m_n)^T$ to be the vector of scattering data. Throughout this work, we write $\mathbb{E}[A]$ for the expectation of a random variable A .

We look for an approximate solution f_n of the unknown source function f^* through the least-squares regularized minimization:

$$\min_{f \in X} \frac{1}{n} \sum_{i=1}^n |(Sf)(x_i) - m_i|^2 + \lambda_n \|f\|_X^2, \quad (2)$$

where $\lambda_n > 0$ is called a regularization parameter.

We shall consider that the set of discrete points $\{x_i\}_{i=1}^n$ are scattered but quasi-uniformly distributed in Ω , i.e., there exists a constant $B > 0$ such that $d_{\max}/d_{\min} \leq B$, where d_{\max} and d_{\min} are defined by

$$d_{\max} = \sup_{x \in \Omega} \inf_{1 \leq i \leq n} |x - x_i| \quad \text{and} \quad d_{\min} = \inf_{1 \leq i \neq j \leq n} |x_i - x_j|. \quad (3)$$

For any $u, v \in C(\bar{\Omega})$ and $y \in \mathbb{R}^n$, we define

$$(y, v)_n = \frac{1}{n} \sum_{i=1}^n y_i v(x_i), \quad (u, v)_n = \frac{1}{n} \sum_{i=1}^n u(x_i) v(x_i),$$

and the semi-norm $\|u\|_n = (\sum_{i=1}^n u^2(x_i)/n)^{1/2}$ for any $u \in C(\bar{\Omega})$.

Throughout the work, we consider two kinds of random noises $\{e_i\}_{i=1}^n$:

(R1) $\{e_i\}_{i=1}^n$ are independent random variables satisfying $\mathbb{E}[e_i] = 0$ and $\mathbb{E}[e_i^2] \leq \sigma^2$;

(R2) $\{e_i\}_{i=1}^n$ are independent sub-Gaussian random variables with parameter σ ,

and provide two different techniques to analyse the stochastic convergence and a practical approach to choose the parameter λ_n in each case. We study the convergence under the expectation \mathbb{E} in the case **(R1)**, and establish a stronger convergence in the case **(R2)**, where the errors have exponential decay tails.

2.1 Stochastic convergence for noisy data of variables with bounded variance

We consider the measurement data of type **(R1)** in this section, and study the stochastic convergence of the error under the expectation \mathbb{E} .

Assumption 2.1. *We assume that*

(1) *There exists a constant $\beta > 1$ such that for all $u \in Y$,*

$$\|u\|_{L^2(\Omega)}^2 \leq C(\|u\|_n^2 + n^{-\beta} \|u\|_Y^2), \quad \|u\|_n^2 \leq C(\|u\|_{L^2(\Omega)}^2 + n^{-\beta} \|u\|_Y^2). \quad (4)$$

(2) *The first n eigenvalues, $0 < \eta_1 \leq \eta_2 \leq \dots \leq \eta_n$, of the eigenvalue problem*

$$(\psi, v)_X = \eta(S\psi, Sv) \quad \forall v \in X,$$

satisfy that $\eta_k \geq Ck^\alpha$ ($k = 1, 2, \dots, n$) for some constant C depending only on the operator $S : X \rightarrow Y$. The constant α satisfies $1 < \alpha \leq \beta$.

The following observation is inspired by [42], where it was shown that the solution of a thin plate spline smoother model is attained in a finite dimensional subset.

Lemma 2.1. *For a given $m \in \mathbb{R}^n$, let f be the solution to the optimization problem*

$$\min_{f \in X, (Sf)(x_i) = m_i} \|f\|_X^2, \quad (5)$$

then $f \in V_n$, where V_n is an n -dimensional subset of X .

Proof. Let V be a subset of X such that

$$V = \{v \in X : (Sv)(x_i) = 0, i = 1, 2, \dots, n\}.$$

Define the projection operator $P_V : X \rightarrow V$,

$$(P_V[f], v)_X = (f, v)_X \quad \forall v \in V.$$

Choose $\phi_i \in X$ such that $(S\phi_i)(x_j) = \delta_{ij}$, where δ_{ij} is the Kronecker delta function. Let $\psi_i = -P_V[\phi_i] + \phi_i$ and $V_n = \text{span}\{\psi_1, \dots, \psi_n\}$. It's easy to check that $(S\psi_i)(x_j) = \delta_{ij}$ also holds. For any $f \in X$, define the interpolation operator I :

$$If = \sum_{i=1}^n (Sf)(x_i) \psi_i.$$

We can easily see that $If \in V_n$ and $f - If \in V$, hence we derive

$$\begin{aligned} (f - If, If)_X &= (f - If, \sum_{i=1}^n (Sf)(x_i) (\phi_i - P_V[\phi_i]))_X \\ &= \sum_{i=1}^n (Sf)(x_i) (f - If, \phi_i - P_V[\phi_i])_X = 0, \end{aligned}$$

where we have used the fact that $(v, \phi_i - P_V[\phi_i])_X = 0$ for all $v \in V$.

We see directly from the above equality that $(If, If)_X \leq (f, f)_X$, hence we have

$$\min_{f \in V_n, (Sf)(x_i) = m_i} \|f\|_X^2 = \min_{f \in X, Sf(x_i) = m_i} \|f\|_X^2.$$

This completes the proof. \square

Lemma 2.2. Assume Assumption 2.1 is fulfilled. Let V_n be defined as in Lemma 2.1, then the eigenvalue problem

$$(\psi, v)_X = \rho (S\psi, Sv)_n \quad \forall v \in V_n, \quad (6)$$

has n eigenvalues $\rho_1 \leq \rho_2 \leq \dots \leq \rho_n$, and all the eigenfunctions form an orthogonal basis of V_n with respect to the norm $\|S \cdot\|_n$. Moreover, there exists a constant $C > 0$ independent of k such that $\rho_k \geq Ck^\alpha$ for $k = 1, 2, \dots, n$.

Proof. Consider $V_n = \text{span}\{\psi_i\}_{i=1}^n$ as defined in the proof of Lemma 2.1, and $(S\psi_i)(x_j) = \delta_{ij}$. We can write $\psi = \sum_{i=1}^n (S\psi)(x_i) \psi_i$ for any $\psi \in V_n$. This implies $\|S \cdot\|_n$ is a norm of V_n . Therefore, the generalized eigenvalue problem (6) has n finite eigenvalues $\rho_1 \leq \rho_2 \leq \dots \leq \rho_n$ and all eigenfunctions form an orthogonal basis of V_n with respect to the norm $\|S \cdot\|_n$.

We are now ready to give a lower bound of the eigenvalues ρ_k . Using the min-max principle of the Rayleigh quotient for the eigenvalues and (4), we can derive

$$\begin{aligned} \rho_k &= \min_{\dim(X)=k, X \subset V_n} \max_{u \in X} \frac{(u, u)_X}{(Su, Su)_n} \\ &\geq C \min_{\dim(X)=k, X \subset V_n} \max_{u \in X} \frac{(u, u)_X}{(Su, Su) + n^{-\beta}(u, u)_X} \\ &\geq C \min_{\dim(X)=k, X \subset L^2(\Omega)} \max_{u \in X} \frac{(u, u)_X}{(Su, Su) + n^{-\beta}(u, u)_X} \\ &= C \frac{1}{\eta_k^{-1} + n^{-\beta}} \geq C \frac{1}{k^{-\alpha} + n^{-\beta}}, \end{aligned}$$

where we have used the fact that $\eta_k \geq Ck^\alpha$ by Assumption 1. Now $k^\alpha n^{-\beta} \leq n^{\alpha-\beta} \leq 1$ for all $k \leq n$ and $\alpha \leq \beta$. We conclude that $\rho_k \geq Ck^\alpha$. This completes the proof. \square

Theorem 2.3. Assume Assumption 2.1 is fulfilled. Let $f_n \in X$ be the unique solution of (2). Then there exist constants $\lambda_0 > 0$ and $C > 0$ such that for any $\lambda_n \leq \lambda_0$,

$$\mathbb{E}[\|Sf_n - Sf^*\|_n^2] \leq C\lambda_n\|f^*\|_X^2 + C\sigma^2/(n\lambda_n^{1/\alpha}), \quad (7)$$

$$\mathbb{E}[\|f_n - f^*\|_X^2] \leq C\|f^*\|_X^2 + C\sigma^2/(n\lambda_n^{1+1/\alpha}). \quad (8)$$

More over if we assume the eigenfunctions $\{\phi_k\}_{k=1}^\infty$ of S form an orthonormal basis of X , and define the space Z as

$$Z = \{v \in X : v = \sum_{k=1}^\infty v_k \phi_k, \text{ with } v_k = (v, \phi_k)_{L^2(\Omega)} \text{ and } \sum_{k=1}^\infty \eta_k^{1/2} v_k^2 < \infty\}.$$

Then we have the following weaker convergence result for $n^{-\beta} \leq \lambda_n$:

$$\mathbb{E}[\|f_n - f^*\|_{Z'}^2] \leq C\lambda_n^{1/2}\|f^*\|_X^2 + C\sigma^2/(n\lambda_n^{1/2+1/\alpha}), \quad (9)$$

where Z' is the dual space of Z .

Proof. By deriving the necessary condition of the quadratic minimization (2), we can readily see that the unique minimizer $f_n \in X$ satisfies the variational equation

$$\lambda_n(f_n, v)_X + (Sf_n, Sv)_n = (m, Sv)_n \quad \forall v \in X. \quad (10)$$

For any $v \in X$, we introduce the energy norm $\|v\|_{\lambda_n}^2 := \lambda(v, v)_X + \|Sv\|_n^2$. By taking $v = f_n - f^*$ in (10) along with (1), we obtain

$$\|f_n - f^*\|_{\lambda_n} \leq \lambda_n^{1/2}\|f^*\|_X + \sup_{v \in L^2(\Omega)} \frac{(e, Sv)_n}{\|v\|_{\lambda_n}}. \quad (11)$$

It remains to estimate the supremum term in (11). Using Lemma 2.1, we can rewrite this supremum term equivalently as

$$\begin{aligned} \sup_{v \in X} \frac{(e, Sv)_n^2}{\|v\|_{\lambda_n}^2} &= \sup_{v \in X} \frac{(e, Sv)_n^2}{\lambda_n(v, v)_X + \|Sv\|_n^2} \\ &\leq \sup_{v \in X} \frac{(e, Sv)_n^2}{\lambda_n \min_{u \in X, Su(x_i)=Sv(x_i)} (u, u)_X + \|Sv\|_n^2} \\ &= \sup_{v \in X} \frac{(e, Sv)_n^2}{\lambda_n \min_{u \in V_n, Su(x_i)=Sv(x_i)} (u, u)_X + \|Sv\|_n^2} \\ &= \sup_{v \in V_n} \frac{(e, Sv)_n^2}{\lambda_n(v, v)_X + \|Sv\|_n^2}. \end{aligned}$$

Let $\rho_1 \leq \rho_2 \leq \dots \leq \rho_n$ be the eigenvalues of the problem

$$(\psi, v)_X = \rho(S\psi, Sv)_n \quad \forall v \in V_n, \quad (12)$$

with the corresponding eigenfunctions $\{\psi_k\}_{k=1}^n$, which is an orthonormal basis of V_n under the inner product $(S\cdot, S\cdot)_n$. Thus $(S\psi_k, S\psi_l)_n = \delta_{kl}$ and consequently, $(\psi_k, \psi_l)_X = \rho_k \delta_{kl}$, $k, l = 1, 2, \dots, n$.

Now for any $v \in V_n$, we have the expansion $v(x) = \sum_{k=1}^n v_k \psi_k(x)$, where $v_k = (Sv, S\psi_k)_n$ for $k = 1, 2, \dots, n$. Thus $\|v\|_{\lambda_n}^2 = \sum_{k=1}^n (\lambda_n \rho_k + 1) v_k^2$. By the Cauchy-Schwarz inequality we can readily get

$$\begin{aligned} (e, Sv)_n^2 &= \frac{1}{n} \sum_{i=1}^n e_i \sum_{k=1}^n v_k \psi_k(x_i) = \frac{1}{n} \sum_{k=1}^n v_k \sum_{i=1}^n e_i \psi_k(x_i) \\ &\leq \frac{1}{n^2} \sum_{k=1}^n (1 + \lambda_n \rho_k) v_k^2 \cdot \sum_{k=1}^n (1 + \lambda_n \rho_k)^{-1} \left(\sum_{i=1}^n e_i (S\psi_k)(x_i) \right)^2. \end{aligned}$$

This, along with the fact that $\|S\psi_k\|_n = 1$, implies

$$\begin{aligned} \mathbb{E} \left[\sup_{v \in V_n} \frac{(e, Sv)_n^2}{\|v\|_{\lambda_n}^2} \right] &\leq \frac{1}{n^2} \sum_{k=1}^n (1 + \lambda_n \rho_k)^{-1} \mathbb{E} \left(\sum_{i=1}^n e_i (S\psi_k)(x_i) \right)^2 \\ &\leq \sigma^2 n^{-1} \sum_{k=1}^n (1 + \lambda_n \rho_k)^{-1}. \end{aligned}$$

In the last inequality, we use the fact that the random variables $\{e_i\}_{i=1}^n$ are independent and identically distributed, i.e. $\mathbb{E}[e_i e_j] = \delta_{ij}$.

Now by Assumption 2.1 we readily derive

$$\mathbb{E} \left[\sup_{v \in X} \frac{(e, Sv)_n^2}{\|v\|_{\lambda_n}^2} \right] \leq C \sigma^2 n^{-1} \sum_{k=1}^n (1 + \lambda_n \rho_k)^{-1} \leq C \frac{\sigma^2}{n \lambda_n^{1/\alpha}}.$$

This completes the proof by using (11).

Furthermore, if the eigenfunctions $\{\phi_k\}_{k=1}^\infty$ of S form an orthonormal basis of X , i.e. $(\phi_k, \phi_l) = \delta_{kl}$, then $(S\phi_k, S\phi_l) = \eta_k^{-1} \delta_{kl}$. For any $v \in X$, we have the expansion $v = \sum_{k=1}^\infty v_k \phi_k$ with $v_k = (v, \phi_k)$. Obviously, $\|Sv\|_{L^2(\Omega)}^2 = \sum_{k=1}^\infty \eta_k^{-1} v_k^2$ and $\|v\|_X^2 = \sum_{k=1}^\infty v_k^2$. By definition of dual space and $\|g\|_Z^2 = \sum_{k=1}^\infty \eta_k^{1/2} g_k^2$,

$$\begin{aligned} \|v\|_{Z'} &= \sup_{0 \neq g \in Z} \frac{|(v, g)_X|}{\|g\|_X} = \sup_{0 \neq g \in Z} \frac{|\sum_{k=1}^\infty g_k v_k|}{\|g\|_X} \\ &\leq \frac{(\sum_{k=1}^\infty \eta_k^{-1/2} g_k^2)^{1/2} (\sum_{k=1}^\infty \eta_k^{1/2} g_k^2)^{1/2}}{\|g\|_X} \\ &= \left(\sum_{k=1}^\infty \eta_k^{-1/2} v_k^2 \right)^{1/2} \\ &\leq \left(\sum_{k=1}^\infty \eta_k^{-1} v_k^2 \right)^{1/4} \left(\sum_{k=1}^\infty v_k^2 \right)^{1/4} = \|Sv\|_{L^2(\Omega)}^{1/2} \|v\|_X^{1/2}. \end{aligned}$$

Take v to be $f^* - f_n$ in the above inequality, we derive that,

$$\|f^* - f_n\|_{Z'}^2 \leq \|Sf^* - Sf_n\|_{L^2(\Omega)} \|f^* - f_n\|_X.$$

From Assumption 2.1 (1), $\|Sf^* - Sf_n\|_{L^2(\Omega)}^2 \leq \|Sf^* - Sf_n\|_n^2 + n^{-\beta} \|f^* - f_n\|_X^2$, along with (7) and (8), we finally have,

$$\mathbb{E}[\|f_n - f^*\|_{Z'}^2] \leq C(1 + \lambda_n^{-1} n^{-\beta})^{1/2} \left(\lambda_n^{1/2} \|f^*\|_X^2 + \sigma^2 / (n \lambda_n^{1/2+1/\alpha}) \right).$$

With $n^{-\beta} \leq \lambda_n$, we prove the weaker convergence (9). \square

2.2 Stochastic convergence for noisy data being sub-gaussian random variables

We consider in this section the case **(R2)** for the data (1), that is,

$$\mathbb{E}\left[\exp(\lambda(e_i - \mathbb{E}[e_i]))\right] \leq \exp\left(\frac{1}{2}\sigma^2\lambda^2\right) \quad \forall \lambda \in \mathbb{R}, \quad (13)$$

and study the stochastic convergence of the error $\|Sf^* - Sf_n\|_n$.

We first give a brief introduction of sub-Gaussian random variables and the theory of empirical processes that will be used in our subsequent analysis; see [13, 44, 43] for more details. The probability distribution function of a sub-Gaussian random variable Z has an exponentially decaying tail, that is,

$$\mathbb{P}(|Z - \mathbb{E}[Z]| \geq z) \leq 2 \exp\left(-\frac{z^2}{2\sigma^2}\right) \quad \forall z > 0. \quad (14)$$

We shall also use the Orlicz norm. For a monotonically increasing convex function ψ satisfying $\psi(0) = 0$, the Orlicz norm $\|Z\|_\psi$ of a random variable Z is defined as

$$\|Z\|_\psi = \inf \left\{ C > 0 : \mathbb{E}\left[\psi\left(\frac{|Z|}{C}\right)\right] \leq 1 \right\}. \quad (15)$$

For most of our analyses, we will use the Orlicz norm $\|Z\|_{\psi_2}$, with $\psi_2(t) = e^{t^2} - 1$ for $t > 0$. Through some calculations, we have the estimate (see, e.g., [13, (4.5)])

$$\mathbb{P}(|Z| \geq z) \leq 2 \exp\left(-\frac{z^2}{\|Z\|_{\psi_2}^2}\right) \quad \forall z > 0. \quad (16)$$

Consider a semi-metric space \mathbb{T} with a semi-metric d and the random process $\{Z_t : t \in \mathbb{T}\}$ indexed by \mathbb{T} . The random process $\{Z_t : t \in \mathbb{T}\}$ is called sub-Gaussian if

$$\mathbb{P}(|Z_s - Z_t| > z) \leq 2 \exp\left(-\frac{z^2}{2d(s, t)^2}\right) \quad \forall s, t \in \mathbb{T}, \quad z > 0. \quad (17)$$

For a semi-metric space (\mathbb{T}, d) and $\varepsilon > 0$, the covering number $N(\varepsilon, \mathbb{T}, d)$ is the minimum number of ε -balls that cover \mathbb{T} ; and $\log N(\varepsilon, \mathbb{T}, d)$ is called the covering entropy that is a crucial quantity to characterize the complexity of space \mathbb{T} . We assume

Assumption 2.2. *For a unit ball SY in Y and any $\varepsilon > 0$, there exists a constant $\gamma < 2$ such that the covering entropy is controlled by*

$$\log N(\varepsilon, SY, \|\cdot\|_{L^\infty(\Omega)}) \leq C\varepsilon^{-\gamma}.$$

Important estimates of the covering entropy for Sobolev spaces can be found in [8]. We shall often need the following maximal inequality [44, Section 2.2.1].

Lemma 2.4. *If $\{Z_t : t \in \mathbb{T}\}$ is a separable sub-Gaussian random process, then it holds for some constant $K > 0$ that*

$$\left\| \sup_{s, t \in T} |Z_s - Z_t| \right\|_{\psi_2} \leq K \int_0^{\text{diam } \mathbb{T}} \sqrt{\log N\left(\frac{\varepsilon}{2}, T, d\right)} d\varepsilon.$$

The useful results in the following two lemmas can be found in [13].

Lemma 2.5. $\{E_n(f) := (e, Sf)_n : f \in X\}$ is a sub-Gaussian random process with respect to the semi-distance $d(f, v) = \sigma n^{-1/2} \|Sf - Sv\|_n$ for any $f, v \in X$.

Lemma 2.6. Let $C_1 > 0$ and $K_1 > 0$ be two constants, and Z be any random variable satisfying

$$\mathbb{P}(|Z| > \alpha(1+z)) \leq C_1 \exp\left(-\frac{z^2}{K_1^2}\right) \quad \forall \alpha > 0, \quad z \geq 1,$$

then there exists a constant $C(C_1, K_1) > 0$ depending on C_1 and K_1 such that

$$\|Z\|_{\psi_2} \leq C(C_1, K_1) \alpha.$$

Theorem 2.7. Assume Assumption 2.2 is fulfilled. Let $\rho_0 = \|f^*\|_X + \sigma n^{-1/2}$, and $f_n \in X$ be the solution of the minimization (2). If we take $\lambda_n^{1/2+\gamma/4} = O(\sigma n^{-1/2} \rho_0^{-1})$, then there exists a constant $C > 0$ such that

$$\mathbb{P}(\|Sf_n - Sf^*\|_n \geq \lambda_n^{1/2} \rho_0 z) \leq 2e^{-Cz^2} \quad \text{and} \quad \mathbb{P}(\|f_n\|_X \geq \rho_0 z) \leq 2e^{-Cz^2}.$$

More over, with the same assumptions and notations in Theorem 2.3, we have,

$$\mathbb{P}(\|f_n - f^*\|_{Z'} \geq \lambda_n^{1/4} \rho_0 z) \leq 2e^{-Cz^2}. \quad (18)$$

Proof. By using the estimate (16), it suffices to prove

$$\|Sf_n - Sf^*\|_n \|_{\psi_2} \leq C \lambda_n^{1/2} \rho_0 \quad \text{and} \quad \|f_n\|_X \|_{\psi_2} \leq C \rho_0. \quad (19)$$

Because of similarity, we will prove only the first estimate in (19) by the peeling argument. It follows from (2) that

$$\|Sf_n - Sf^*\|_n^2 + \lambda_n \|f_n\|_X^2 \leq 2(e, Sf_n - Sf^*)_n + \lambda_n \|f^*\|_X^2. \quad (20)$$

Let $\delta > 0$, $\rho > 0$ be two constants to be determined later, and we set for $i, j \geq 1$,

$$A_0 = [0, \delta), \quad A_i = [2^{i-1}\delta, 2^i\delta), \quad B_0 = [0, \rho), \quad B_j = [2^{j-1}\rho, 2^j\rho). \quad (21)$$

For $i, j \geq 0$, we further define

$$F_{ij} = \{v \in X : \|Sv\|_n \in A_i, \|v\|_X \in B_j\},$$

then we can readily see

$$\mathbb{P}(\|Sf_n - Sf^*\|_n > \delta) \leq \sum_{i=1}^{\infty} \sum_{j=0}^{\infty} \mathbb{P}(f_n - f^* \in F_{ij}). \quad (22)$$

Now we estimate $\mathbb{P}(f_n - f^* \in F_{ij})$ for each pair $\{i, j\}$. By Lemma 2.5, we know $\{(e, Sv)_n : v \in X\}$ is a sub-Gaussian random process with respect to the semi-distance $d(f, v)$. With this semi-distance, it is easy to see that $\text{diam}(F_{ij}) \leq 2\sigma n^{-1/2} \cdot 2^i \delta$, then we can deduce by using Lemma 2.4 that

$$\begin{aligned} \left\| \sup_{f-f^* \in F_{ij}} |(e, Sf - Sf^*)_n| \right\|_{\psi_2} &\leq K \int_0^{\sigma n^{-1/2} \cdot 2^{i+1} \delta} \sqrt{\log N\left(\frac{\varepsilon}{2}, F_{ij}, d\right)} d\varepsilon \\ &= K \int_0^{\sigma n^{-1/2} \cdot 2^{i+1} \delta} \sqrt{\log N\left(\frac{\varepsilon}{2\sigma n^{-1/2}}, F_{ij}, \|S \cdot\|_n\right)} d\varepsilon. \end{aligned}$$

By Assumption 2.2, we have the estimate for the covering entropy

$$\begin{aligned} \log N\left(\frac{\varepsilon}{2\sigma n^{-1/2}}, F_{ij}, \|S \cdot\|_n\right) &\leq \log N\left(\frac{\varepsilon}{2\sigma n^{-1/2}}, F_{ij}, \|S \cdot\|_{L^\infty(\Omega)}\right) \\ &= \log N\left(\frac{\varepsilon}{2\sigma n^{-1/2}}, S(F_{ij}), \|\cdot\|_{L^\infty(\Omega)}\right) \leq C\left(\frac{2\sigma n^{-1/2} \cdot 2^j \rho}{\varepsilon}\right)^\gamma, \end{aligned}$$

where we have used the fact that $S(F_{ij})$ is included in the ball in Y of radius $C(2^j \rho)$ since $S : X \rightarrow Y$ is a bounded operator. Using this, we can further derive

$$\begin{aligned} \left\| \sup_{f-f^* \in F_{ij}} |(e, Sf - Sf^*)_n|_{\psi_2} \right\| &\leq K \int_0^{\sigma n^{-1/2} \cdot 2^{i+1} \delta} \left(\frac{2\sigma n^{-1/2} \cdot 2^j \rho}{\varepsilon}\right)^{\gamma/2} d\varepsilon \\ &= C\sigma n^{-1/2} (2^j \rho)^{\gamma/2} (2^i \delta)^{1-\gamma/2}. \end{aligned} \quad (23)$$

Then by using the estimates (20) and (16), we have for $i, j \geq 1$,

$$\begin{aligned} \mathbb{P}(f_n - f^* \in F_{ij}) &\leq \mathbb{P}\left(2^{2(i-1)}\delta^2 + \lambda_n 2^{2(j-1)}\rho^2 \leq 2 \sup_{f-f^* \in F_{ij}} |(e, f - f^*)_n| + \lambda_n \rho_0^2\right) \\ &= \mathbb{P}\left(2 \sup_{f-f^* \in F_{ij}} |(e, Sf - Sf^*)_n| \geq 2^{2(i-1)}\delta^2 + \lambda_n 2^{2(j-1)}\rho^2 - \lambda_n \rho_0^2\right) \\ &\leq 2 \exp\left[-\frac{1}{C\sigma^2 n^{-1}} \left(\frac{2^{2(i-1)}\delta^2 + \lambda_n 2^{2(j-1)}\rho^2 - \lambda_n \rho_0^2}{(2^i \delta)^{1-\gamma/2} (2^j \rho)^{\gamma/2}}\right)^2\right]. \end{aligned}$$

Now for $z \geq 1$, we take $\delta^2 = \lambda_n \rho_0^2 (1+z)^2$, $\rho = \rho_0$, then with the choice that $\lambda_n^{\frac{1}{2} + \frac{\gamma}{4}} = O(\sigma n^{-1/2} \rho_0^{-1})$ and direct computing, we readily obtain for $i, j \geq 1$ that

$$\mathbb{P}(f_n - f^* \in F_{ij}) \leq 2 \exp\left[-C\left(\frac{2^{2(i-1)}z(1+z) + 2^{2(j-1)}}{(2^i(1+z))^{1-\gamma/2} (2^j)^{\gamma/2}}\right)^2\right]. \quad (24)$$

To simplify the above estimate, we use Young's inequality that $ab \leq a^p/p + b^q/q$ for any $a, b > 0$ and $p, q > 1$ such that $p^{-1} + q^{-1} = 1$ to obtain

$$(2^i(1+z))^{1-\gamma/2} (2^j)^{\gamma/2} \leq C((1+z)2^i + 2^j).$$

Therefore we get from (24) for $i, j \geq 1$ that

$$\mathbb{P}(f_n - f^* \in F_{ij}) \leq 2 \exp\left[-C(2^{2i}z^2 + 2^{2j})\right].$$

Similarly, one can show for $i \geq 1, j = 0$ that

$$\mathbb{P}(f_n - f^* \in F_{i0}) \leq 2 \exp\left[-C(2^{2i}z^2)\right].$$

Collecting the above estimates for all $i, j \geq 0$ and using the facts that

$$\sum_{j=1}^{\infty} \exp(-C(2^{2j})) \leq \exp(-C) < 1 \quad \text{and} \quad \sum_{i=1}^{\infty} \exp(-C(2^{2i}z^2)) \leq \exp(-Cz^2),$$

we come to the conclusion that

$$\sum_{i=1}^{\infty} \sum_{j=0}^{\infty} \mathbb{P}(f_n - f^* \in F_{ij}) \leq 2 \sum_{i=1}^{\infty} \sum_{j=1}^{\infty} \exp(-C(2^{2i}z^2 + 2^{2j})) + 2 \sum_{i=1}^{\infty} \exp(-C(2^{2i}z^2)).$$

The above estimate can be further bounded by $4\exp(-Cz^2)$. Using this, we get from (22) that

$$\mathbb{P}(\|Sf_n - Sf^*\|_n > \lambda_n^{1/2}\rho_0(1+z)) \leq 4\exp(-Cz^2) \quad \forall z \geq 1. \quad (25)$$

This, along with Lemma 2.6, implies that $\|Sf_n - Sf^*\|_n \leq C\lambda_n^{1/2}\rho_0$, which is the first estimate in (19). The second estimate is similar to the first one by taking $i \geq 0$ and $j \geq 1$ in the summation above (25). Using the very same technique in Theorem 2.3, one could directly get (18). \square

2.3 Convergence of the discrete solutions

In this section we consider the approximation to the optimal control problem (2), i.e.,

$$\min_{f \in X} \|Sf - m\|_n^2 + \lambda_n \|f\|_X^2.$$

We can directly verify that the solution $f_n \in X$ satisfies the weak formulation

$$\lambda_n(f_n, v)_X + (Sf_n, Sv)_n = (m, Sv)_n \quad \forall v \in X. \quad (26)$$

Let $V_h \subset X$ and $Y_h \subset Y$ be two discrete function spaces (e.g., finite element spaces) with dimensions N_h and M_h respectively, and $S_h : X \rightarrow Y_h \subset Y$ be the discrete approximation of the operator $S : X \rightarrow Y$. We make the following standard assumptions on the discretization space V_h and the approximation operator S_h .

Assumption 2.3. *For the discrete operator $S_h : X \rightarrow Y_h \subset Y$,*

(1) *there exists an error estimate $e(h)$ such that the discrete operator S_h satisfies*

$$\|Sf - S_h f\|_n^2 \leq C e(h) \|f\|_X^2 \quad \forall f \in X.$$

(2) *For any $f \in X$, there exists $v_h \in V_h$ such that*

$$\lambda_n \|f - v_h\|_X^2 + \|S_h f - S_h v_h\|_n^2 \leq C(\lambda_n + e(h)) \|f\|_X^2.$$

We can now look for the discrete solution to the problem (2):

$$\min_{f_h \in V_h} \|S_h f_h - m\|_n^2 + \lambda_n \|f_h\|_X^2.$$

Obviously, f_h satisfies the weak formulation:

$$\lambda_n(f_h, v_h)_X + (S_h f_h, S_h v_h)_n = (m, S_h v_h)_n \quad \forall v_h \in V_h. \quad (27)$$

2.3.1 Convergence for noisy data from random variables with bounded variance

We study in this section the expectational convergence of the discrete solution to (27) in the case **(R1)** for the data (1), with the main results stated below.

Theorem 2.8. *Assume Assumption 2.1 and 2.3 are fulfilled. Let $f_h \in V_h$ be the solution of (27). Then there exist constants $\lambda_0 > 0$ and $C > 0$ such that for any $\lambda_n \leq \lambda_0$,*

$$\mathbb{E}[\|Sf^* - S_h f_h\|_n^2] \leq C(\lambda_n + e(h)) \|f^*\|_X^2 + C \left[1 + \frac{e(h)}{\lambda_n} + \frac{N_h e(h)}{\lambda_n^{1-1/\alpha}} \right] \frac{\sigma^2}{n \lambda_n^{1/\alpha}}, \quad (28)$$

$$\mathbb{E}[\|f^* - f_h\|_X^2] \leq C \frac{\lambda_n + e(h)}{\lambda_n} \|f^*\|_X^2 + C \left[1 + \frac{e(h)}{\lambda_n} + \frac{N_h e(h)}{\lambda_n^{1-1/\alpha}} \right] \frac{\sigma^2}{n \lambda_n^{1+1/\alpha}}. \quad (29)$$

More over, with the same assumption and notations in Theorem 2.3, we have,

$$\begin{aligned}\mathbb{E}[\|f^* - f_h\|_{Z'}^2] &\leq C(\lambda_n^{1/2} + e^{1/2}(h)) \frac{\lambda_n + e(h)}{\lambda_n} \|f^*\|_X^2 \\ &\quad + C(\lambda_n^{1/2} + e^{1/2}(h)) \left[1 + \frac{e(h)}{\lambda_n} + \frac{N_h e(h)}{\lambda_n^{1-1/\alpha}}\right] \frac{\sigma^2}{n\lambda_n^{1/\alpha}}.\end{aligned}\quad (30)$$

In particular, if $e(h) \leq C\lambda_n$ and $N_h e(h) \leq C\lambda_n^{1-1/\alpha}$, we have

$$\mathbb{E}[\|Sf^* - S_h f_h\|_n^2] \leq C\lambda_n \|f^*\|_X^2 + C\sigma^2/(n\lambda_n^{1/\alpha}), \quad (31)$$

$$\mathbb{E}[\|f^* - f_h\|_X^2] \leq C\|f^*\|_X^2 + C\sigma^2/(n\lambda_n^{1+1/\alpha}), \quad (32)$$

$$\mathbb{E}[\|f^* - f_h\|_{Z'}^2] \leq C\lambda_n^{1/2} \|f^*\|_X^2 + C\sigma^2/(n\lambda_n^{1/2+1/\alpha}). \quad (33)$$

Proof. For any $f, v \in X$, we denote $a_h(f, v) = \lambda_n(f, v)_X + (S_h f, S_h v)_n$ and $\|f\|_{a_h}^2 = a_h(f, f)$. For any $w_h \in V_h$, by taking $v = w_h$ in (26) and $v_h = w_h$ in (27), we readily obtain

$$\begin{aligned}a_h(f_h - v_h, w_h) &= a_h(f_n - v_h, w_h) + ((S - S_h)f_n, S_h w_h)_n + (Sf^* - Sf_n, (S_h - S)w_h)_n \\ &\quad + (e, (S_h - S)w_h)_n \equiv a_h(f_n - v_h, w_h) + F(w_h) \quad \forall v_h, w_h \in V_h.\end{aligned}$$

By the triangle inequality, we can further derive

$$\|f_n - f_h\|_{a_h} \leq C \inf_{v_h \in V_h} \|f_n - v_h\|_{a_h} + C \sup_{w_h \in V_h} \frac{|F(w_h)|}{\|w_h\|_{a_h}}. \quad (34)$$

But from Assumption 2.3 (1), we have

$$\sup_{w_h \in V_h} \frac{|((S - S_h)f_n, S_h w_h)_n|}{\|w_h\|_{a_h}} \leq \|Sf_n - S_h f_n\|_n \leq Ce(h)^{1/2} \|f_n\|_X, \quad (35)$$

$$\sup_{w_h \in V_h} \frac{|(Sf^* - Sf_n, (S_h - S)w_h)_n|}{\|w_h\|_{a_h}} \leq C\|Sf^* - Sf_n\|_n \frac{e(h)^{1/2}}{\lambda_n^{1/2}}. \quad (36)$$

Now we estimate $\mathbb{E}(\sup_{w_h \in V_h} |(e, Sw_h - S_h w_h)_n|^2 / \|w_h\|_{a_h}^2)$. Let $\{\psi_k\}_{k=1}^{N_h}$ be the orthogonal basis of V_h (with $N_h = \dim(V_h)$) such that $(\psi_i, \psi_j) = \delta_{ij}$. Then for any $w_h \in V_h$, we have $w_h = \sum_{j=1}^{N_h} (w_h, \psi_j) \psi_j$, and $\|w_h\|_{L^2(\Omega)}^2 = \sum_{j=1}^{N_h} (w_h, \psi_j)^2$. Applying the Cauchy-Schwarz inequality,

$$\begin{aligned}(e, (S - S_h)w_h)_n^2 &\leq \frac{1}{n^2} \sum_{j=1}^{N_h} (w_h, \psi_j)^2 \sum_{j=1}^{N_h} \left(\sum_{i=1}^n e_i (S - S_h) \psi_j(x_i) \right)^2 \\ &= \frac{1}{n^2} \|w_h\|_{L^2(\Omega)}^2 \sum_{j=1}^{N_h} \left(\sum_{i=1}^n e_i (S - S_h) \psi_j(x_i) \right)^2,\end{aligned}$$

we derive

$$\begin{aligned}\mathbb{E} \left(\sup_{w_h \in V_h} \frac{|(e, Sw_h - S_h w_h)_n|^2}{\|w_h\|_{a_h}^2} \right) &\leq \frac{1}{\lambda_n n^2} \sum_{j=1}^{N_h} \mathbb{E} \left(\sum_{i=1}^n e_i (S - S_h) \psi_j(x_i) \right)^2 \\ &= \frac{1}{\lambda_n n} \sum_{j=1}^{N_h} \sigma^2 \|(S - S_h) \psi_j\|_n^2 \leq C \frac{\sigma^2}{\lambda_n n} N_h e(h).\end{aligned}\quad (37)$$

This completes the desired estimates by substituting (35), (36), (37) into (34) and using Assumption 2.3 (2) and Theorem 2.3.

With same notations in Theorem 2.3 and apply the estimate therein, we have that

$$\begin{aligned}\|f^* - f_h\|_{Z'}^2 &\leq \|Sf^* - Sf_h\|_{L^2(\Omega)} \|f^* - f_h\|_X \\ &\leq C (\|Sf^* - Sf_h\|_n + n^{-\beta} \|f^* - f_h\|_X) \|f^* - f_h\|_X \\ &\leq C (\|Sf^* - S_h f_h\|_n + \|S_h f_h - Sf_h\|_n + n^{-\beta} \|f^* - f_h\|_X) \|f^* - f_h\|_X.\end{aligned}$$

Apply the estimates (28), (29) and assumption 2.3 (1) to the above inequality, we finally have (30). \square

2.3.2 Convergence for noisy data being sub-gaussian random variables

We consider in this subsection the convergence of the discrete solution in the case **(R2)** for the data (1). We start by recalling the following lemma in [43, Corollary 2.6] about the estimation of the covering entropy of finite dimensional subsets.

Lemma 2.9. *Let G be a finite dimensional subspace of X of dimension $N_G > 0$ and $G_R = \{f \in G : \|f\|_X \leq R\}$. Then it holds that*

$$N(\varepsilon, G_R, \|\cdot\|_X) \leq (1 + 4R/\varepsilon)^{N_G} \quad \forall \varepsilon > 0.$$

Lemma 2.10. *Assume Assumption 2.3 is fulfilled. Let $G_h := \{w_h \in V_h : \|w_h\|_{a_h} \leq 1\}$. Assume that $e(h) \leq C\lambda_n$ and $N_h e(h) \leq C\lambda_n^{1-\gamma/2}$. Then it holds that*

$$\left\| \sup_{w_h \in G_h} |(e, Sw_h - S_h w_h)_n| \right\|_{\psi_2} \leq C\sigma n^{-1/2} \lambda_n^{-\gamma/4}.$$

Proof. By Lemma 2.5 we know that $\{\hat{E}_n(v_h) := (e, Sw_h - S_h w_h)_n \mid w_h \in G_h\}$ is a sub-Gaussian random process with respect to the semi-distance $\hat{d}(v_h, w_h) = \sigma n^{-1/2} \|(Sw_h - S_h w_h) - (Sw_h - S_h w_h)\|_n$. By Assumption 2.3 and the condition that $e(h) \leq C\lambda_n$, we derive for any $w_h \in G_h$ that $\|Sw_h - S_h w_h\|_n \leq Ce^{1/2}(h) \|w_h\|_X \leq Ce^{1/2}(h) \lambda_n^{-1/2} \leq C$. This implies that the diameter of G_h is bounded by $C\sigma n^{-1/2}$. Now we deduce by the maximal inequality in Lemma 2.4,

$$\left\| \sup_{w_h \in G_h} |(e, Sw_h - S_h w_h)_n| \right\|_{\psi_2} \leq K \int_0^{C\sigma n^{-1/2}} \sqrt{\log N\left(\frac{\varepsilon}{2}, G_h, \hat{d}\right)} d\varepsilon. \quad (38)$$

By Assumption 2.3, we know

$$\hat{d}(v_h, w_h) \leq C\sigma n^{-1/2} e^{1/2}(h) \|v_h - w_h\|_X \quad \forall v_h, w_h \in V_h.$$

Thus we can see

$$\log N\left(\frac{\varepsilon}{2}, G_h, \hat{d}\right) = \log N\left(\frac{\varepsilon}{C\sigma n^{-1/2} e^{1/2}(h)}, G_h, \|\cdot\|_X\right). \quad (39)$$

Now we estimate the covering entropy of G_h . First, we have $\|w_h\|_X \leq \lambda_n^{-1/2}$ for any $w_h \in G_h$. Noting the dimension N_h of V_h , we obtain by Lemma 2.9 and (39) that

$$\log N\left(\frac{\varepsilon}{2}, G_h, \hat{d}\right) \leq CN_h (1 + \sigma n^{-1/2} e^{1/2}(h) \lambda_n^{-1/2} / \varepsilon).$$

Inserting this estimate in (38),

$$\begin{aligned} \left\| \sup_{v_h \in G_h} |(e, \hat{v}_h - \Pi_h v_h)_n| \right\|_{\psi_2} &\leq C \int_0^{C\sigma n^{-1/2}} \sqrt{CN_h(1 + \sigma n^{-1/2} e^{1/2}(h) \lambda_n^{-1/2}/\varepsilon)} d\varepsilon \\ &\leq C \sqrt{N_h \sigma n^{-1/2} e^{1/2}(h) \lambda_n^{-1/2}}. \end{aligned}$$

This completes the proof using the condition that $N_h e(h) \leq C \lambda_n^{1-\gamma/2}$. \square

The following theorem presents the main results of this section.

Theorem 2.11. *Assume Assumption 2.2 and 2.3 are fulfilled. Let $f_h \in V_h$ be the solution of (27). Denote by $\rho_0 = \|f^*\|_X + \sigma n^{-1/2}$. If we take $e(h) \leq C \lambda_n$, $N_h e(h) \leq C \lambda_n^{1-\gamma/2}$ and $\lambda_n^{1/2+\gamma/4} = O(\sigma n^{-1/2} \rho_0^{-1})$, then there exists a constant $C > 0$ such that for any $z > 0$,*

$$\mathbb{P}(\|S_h f_h - S f^*\|_n \geq \lambda_n^{1/2} \rho_0 z) \leq 2e^{-Cz^2} \quad \text{and} \quad \mathbb{P}(\|f_h\|_X \geq \rho_0 z) \leq 2e^{-Cz^2}.$$

More over, with the same assumption and notations in Theorem 2.3, we have,

$$\mathbb{P}(\|f_h - f^*\|_{Z'} \geq \lambda_n^{1/4} \rho_0 z) \leq 2e^{-Cz^2}. \quad (40)$$

Proof. We first derive from (34) that

$$\| \|f_n - f_h\|_{a_h} \|_{\psi_2} \leq C \left\| \inf_{v_h \in V_h} \|f_n - v_h\|_{a_h} \right\|_{\psi_2} + C \left\| \sup_{w_h \in V_h} \frac{|F(w_h)|}{\|w_h\|_{a_h}} \right\|_{\psi_2}.$$

But we know $\sup_{w_h \in V_h} |F(w_h)|/\|w_h\|_{a_h} = \sup_{w_h \in G_h} |F(w_h)|$ from the proof of Theorem 2.8, hence it suffices to estimate $\left\| \sup_{w_h \in G_h} |(e, S w_h - S_h w_h)_n| \right\|_{\psi_2}$. Then the desired estimates follow readily from (19), Lemma 2.10, the assumption that $\sigma n^{-1/2} = O(\lambda_n^{1/2+\gamma/4} \rho_0)$ and (16).

The proof of (40) is similar to (30), this completes the proof. \square

3 An inverse nonstationary source problem

In this section, we apply the theory developed in the previous section 2 to study the regularized solutions to an inverse nonstationary source problem associated with the heat conduction system

$$\begin{cases} u_t + Lu = F(x, t) & \text{in } \Omega \times (0, T), \\ u(x, t) = 0 & \text{on } \partial\Omega \times (0, T), \quad u(x, 0) = 0 & \text{in } \Omega, \end{cases} \quad (41)$$

where L is a second order elliptic operator of the form $Lu = -\nabla \cdot (a(x) \nabla u) + c(x)u$, and $\Omega \subset \mathbb{R}^d$ ($d = 1, 2, 3$) is a bounded domain with C^2 boundary or a convex polyhedral domain. We assume $a \in C^1(\bar{\Omega})$, $c \in C(\bar{\Omega})$ with $c(x) \geq 0$ in Ω , and that the source is of the separable form $F(x, t) = f(x)g(t)$ for $(x, t) \in \Omega \times (0, T)$, where the temporal component $g \in H^1(0, T)$ is known and satisfies that $g(t) \geq 0 \quad \forall t \in (0, T)$, while $f(x)$ is unknown to be recovered.

For the subsequent analysis, we first recall some standard results for parabolic equations (cf., e.g., [17, §7.1]). For $F \in H^1(0, T; L^2(\Omega))$, we know the solution u to (41) satisfies $\partial_t u \in C([0, T]; L^2(\Omega)) \cap L^2(0, T; H_0^1(\Omega))$ and the a priori estimate

$$\|\partial_t u\|_{C([0, T]; L^2(\Omega))} \leq C \|F\|_{H^1(0, T; L^2(\Omega))} \leq C \|f\|_{L^2(\Omega)}.$$

It follows then from the equation (41) and the regularity theory of elliptic equations that $u \in C([0, T]; H^2(\Omega))$ and there exists a constant C such that

$$\|u\|_{C([0, T]; H^2(\Omega))} \leq C\|f\|_{L^2(\Omega)}. \quad (42)$$

Let $X = L^2(\Omega)$, $Y = H^2(\Omega)$, and the forward operator $S : X \rightarrow Y$ be defined by $Sf = u(\cdot, T)$. By (42) we know that $S : X \rightarrow Y$ is a bounded operator

$$\|Sf\|_{H^2(\Omega)} \leq C\|f\|_{L^2(\Omega)} \quad \forall f \in L^2(\Omega).$$

We are mainly interested in the following inverse nonstationary source problem:

(TIP) Given the measurement data of $u(\cdot, t)$ at the terminal $t = T$, recover the spatial source distribution $f^*(x)$ in the entire domain Ω .

We focus on an important physical scenario, i.e., measurement data is collected pointwise on a set of distributed sensors located at $\{x_i\}_{i=1}^n$ inside the domain Ω [3, 20, 5, 29, 35, 36, 37]. Again, we assume the data is of the noisy form (1), where $\{x_i\}_{i=1}^n$ is quasi-uniformly distributed in the sense of (3).

We then look for an approximate solution of the true source f^* through the following least-squares regularized minimization:

$$\min_{f \in X} \|Sf - m\|_n^2 + \lambda_n \|f\|_X^2. \quad (43)$$

3.1 Stochastic convergence for the inverse heat source problem

In this subsection we apply the results in section 2 to study the stochastic convergence of the solution of the problem (43) to the exact source f^* . We first recall an important property about the eigenvalue distribution for the elliptic operator L [2, 18].

Lemma 3.1. *Suppose Ω is a bounded domain in \mathbb{R}^d and $a, c \in C^0(\bar{\Omega})$, $c \geq 0$, then the eigenvalue problem*

$$L\psi = \mu\psi \text{ with } \psi|_{\partial\Omega} = 0 \quad (44)$$

has a countable set of positive eigenvalues $\mu_1 \leq \mu_2 \leq \dots$, with its corresponding eigenfunctions $\{\phi_k\}_{k=1}^\infty$ forming an orthogonal basis of $L^2(\Omega)$. Moreover, there exist constants $C_1, C_2 > 0$ such that $C_1 k^{2/d} \leq \mu_k \leq C_2 k^{2/d}$ for all $k = 1, 2, \dots$.

With Lemma 3.1, we can derive the important spectral property of operator S .

Theorem 3.2. *Let $g \in H^1(0, T)$ and $g > 0$. Then the eigenvalue problem*

$$(\psi, v) = \rho(S\psi, Sv) \quad \forall v \in X \quad (45)$$

has a countable set of positive eigenvalues $0 < \rho_1 \leq \rho_2 \leq \dots$. Moreover, there exists a constant $C > 0$ such that $\rho_k \geq Ck^{4/d}$ for all $k = 1, 2, \dots$.

Proof. We first consider the eigenvalue problem

$$\psi = \eta S\psi. \quad (46)$$

Let $\{\phi_k\}_{k=1}^\infty$ be eigenfunctions of the problem (44) which forms an orthogonal basis of $L^2(\Omega)$. We write $f = \sum_{k=1}^\infty f_k \phi_k$ for a set of coefficients f_k . Let $u = \sum_{k=1}^\infty u_k(t) \phi_k$ be the solution of the problem (41). Plugging these two expressions of f and u into the first equation of (41), we get by noting the fact

that $L\phi_k = \mu_k\phi_k$ and comparing the coefficients of ϕ_k on both sides of the equation that $u_k(0) = 0$ and

$$u'_k(t) + \mu_k u_k = f_k g(t) \quad \text{in } (0, T).$$

We can write the solution as $u_k(T) = \alpha_k f_k$, with $\alpha_k = e^{-\mu_k T} \int_0^T e^{\mu_k s} g(s) ds$. Since $g > 0$ in $(0, T)$, we know $\alpha_1 \geq \alpha_2 \geq \dots > 0$. Moreover, we can easily see that $|\alpha_k| \leq C\mu_k^{-1}$. Noting that $Sf = u(\cdot, T) = \sum_{k=1}^{\infty} u_k(T)\phi_k$, we can formally write

$$S\left(\sum_{k=1}^{\infty} f_k \phi_k\right) = \sum_{k=1}^{\infty} \alpha_k f_k \phi_k.$$

Since $\{\phi_k\}_{k=1}^{\infty}$ is an orthogonal basis of $L^2(\Omega)$, we can readily see that the eigenvalue problem (46) has a countable set of positive eigenvalues $\{\eta_k = \alpha_k^{-1}\}_{k=1}^{\infty}$, with $\{\phi_k\}_{k=1}^{\infty}$ being their corresponding eigenfunctions. By Lemma 3.1, we have $\eta_k = \alpha_k^{-1} \geq C\mu_k \geq C_1 k^{2/d}$. Therefore, the eigenvalue problem (45) has a countable set of eigenvalues $\{\rho_k\}_{k=1}^{\infty}$ that satisfies $\rho_k = \eta_k^2 \geq Ck^{4/d}$. This completes the proof. \square

Next, we will certify that the abstract function space Z in Theorem 2.3 is actually a subspace of $H^1(\Omega)$ for the inverse problem discussed in this section. So that the weaker convergence of the inverse problem corresponding to this section is H^{-1} convergence under a certain assumption in the following Lemma.

Lemma 3.3. *With the same notations in Lemma 3.1 and Theorem 3.2. Then abstract function space*

$$Z = \{v \in L^2(\Omega) : v = \sum_{k=1}^{\infty} v_k \phi_k, v_k = (v, \phi_k)_{L^2(\Omega)} \text{ and } \sum_{k=1}^{\infty} \rho_k^{1/2} v_k^2 < \infty\}$$

is actually a subspace of the Sobolev space $H^1(\Omega)$, so the dual space $H^{-1}(\Omega) \subset Z'$. Moreover, if the eigenvalues ρ_k of (45) satisfy that $\rho_k = O(k^{4/d})$ for all $k = 1, 2, \dots$, then $Z = H^1(\Omega)$ and $Z' = H^{-1}(\Omega)$.

Proof. Since the eigenfunctions $\{\phi_k\}_{k=1}^{\infty}$ forms an orthogonal basis of $L^2(\Omega)$, then for any $v \in L^2(\Omega)$ can be expanded as

$$v = \sum_{k=1}^{\infty} v_k \phi_k \text{ with } v_k = (v, \phi_k)_{L^2(\Omega)}.$$

From the definition of $\{\phi_k\}_{k=1}^{\infty}$ in (44), integrating by part, we have

$$a(\phi_k, q) = \mu_k(\phi_k, q), \quad \forall q \in H^1(\Omega),$$

where $a(p, q) = (ap, q) + (cp, q)$. From the ellipticity of the operator L and take $q = \phi_j$, we could derive that

$$\|\phi_k\|_{H^1(\Omega)} = O(\|\phi_k\|_{L^2(\Omega)}), \quad (\phi_k, \phi_j)_{H^1(\Omega)} = \delta_{kj}.$$

With the expansion $v = \sum_{k=1}^{\infty} v_k \phi_k$,

$$\|v\|_{H^1(\Omega)}^2 = O\left(\sum_{k=1}^{\infty} \mu_k v_k^2\right) \leq C \sum_{k=1}^{\infty} \rho_k^{1/2} v_k^2,$$

this will give $\|v\|_{H^1(\Omega)} \leq C\|v\|_Z$.

Moreover, if the eigenvalues ρ_k satisfy that $\rho_k = O(k^{4/d})$, i.e. $\rho_k = O(\mu_k^2)$, we could derive from the above estimate that $\|v\|_{H^1(\Omega)} = O(\|v\|_Z)$. That is to say $Z = H^1(\Omega)$ and $Z' = H^{-1}(\Omega)$. \square

Remark: In the general case, we could only conclude the eigenvalues satisfy $\rho_k \geq Ck^{4/d}$ from Theorem 3.2. But from the Lemma above, if we expect the space $Z = H^1(\Omega)$, we need an upper bound, i.e. $\rho_k \leq C_1 k^{4/d}$. This is actually not a strict condition, for example, we could just assume the right hand side $g(x) \geq g_{\min} > 0$ in Theorem 3.2. With the same notations in proof of Theorem 3.2, one could get

$$\begin{aligned} |\alpha_i| &= |e^{-\mu_i T} \int_0^T e^{\mu_i s} g(s) ds| \geq g_{\min} |e^{-\mu_i T} \int_0^T e^{\mu_i s} ds| \\ &= g_{\min} \frac{1 - e^{-\mu_i T}}{\mu_i} \geq g_{\min} \frac{1}{2\mu_i}. \end{aligned}$$

Here one could take T_0 such that $e^{-\mu_1 T_0} = \frac{1}{2}$, then for $T \geq T_0$, $u_i(T) = 1 - e^{-\mu_i T} \geq \frac{1}{2}$. This will readily give $\rho_k = O(k^{4/d})$. Hence $Z = H^1(\Omega)$ and $Z' = H^{-1}(\Omega)$ as conclusion of Theorem 3.2. In the following section, we will always assume $\rho_k = O(k^{4/d})$, i.e. $Z = H^1(\Omega)$ and $Z' = H^{-1}(\Omega)$.

Verification of Assumptions 2.1 and 2.2. We first know Assumption 2.1(1) holds with $\beta = 4/d$ from [42, Theorems 3.3-3.4]. This, along with Theorem 3.2, verifies Assumption 2.1(2) with $\alpha = \beta = 4/d$. Assumption 2.2 (with $\gamma = d/2$) is a consequence of the following important estimate about the covering entropy [8].

Lemma 3.4. *Let Q be the unit cube in \mathbf{R}^d and $SW^{s,p}(Q)$ be the unit sphere of space $W^{s,p}(Q)$ for $s > 0$ and $p \geq 1$. Then it holds for sufficient small $\varepsilon > 0$ that*

$$\log N(\varepsilon, SW^{s,p}(Q), \|\cdot\|_{L^q(Q)}) \leq C\varepsilon^{-d/s},$$

where $1 \leq q \leq \infty$ for $sp > d$, and $1 \leq q \leq q^*$ with $q^* = p(1 - sp/d)^{-1}$ for $sp \leq d$.

Under Assumptions 2.1 and 2.2, the following two main results are direct consequences of Theorems 2.3 and 2.7, respectively, for the noisy data of type **(R1)** (random variables with bounded variance) and the noisy data of type **(R2)** (sub-Gaussian random variables).

Theorem 3.5. *For the minimizer $f_n \in L^2(\Omega)$ to the problem (43), there exist constants $\lambda_0 > 0$ and $C > 0$ such that the following estimates hold for any $\lambda_n \leq \lambda_0$:*

$$\begin{aligned} \mathbb{E}[\|Sf_n - Sf^*\|_n^2] &\leq C\lambda_n \|f^*\|_{L^2(\Omega)}^2 + C\sigma^2/(n\lambda_n^{d/4}), \\ \mathbb{E}[\|f_n\|_{L^2(\Omega)}^2] &\leq C\|f^*\|_{L^2(\Omega)}^2 + C\sigma^2/(n\lambda_n^{1+d/4}). \end{aligned}$$

Moreover, if the eigenvalues ρ_k of (45) satisfy that $\rho_k = O(k^{4/d})$, then

$$\mathbb{E}[\|f_n - f^*\|_{H^{-1}(\Omega)}^2] \leq C\lambda_n^{1/2} \|f^*\|_{L^2(\Omega)}^2 + C\sigma^2/(n\lambda_n^{1/2+d/4}).$$

Theorem 3.6. *Let $f_n \in L^2(\Omega)$ be the solution of (43) and $\rho_0 = \|f^*\|_{L^2(\Omega)} + \sigma n^{-1/2}$. If we take λ_n such that $\lambda_n^{1/2+d/8} = O(\sigma n^{-1/2} \rho_0^{-1})$, then the following estimates hold for some constant $C > 0$:*

$$\mathbb{P}(\|Sf_n - Sf^*\|_n \geq \lambda_n^{1/2} \rho_0 z) \leq 2e^{-Cz^2}, \quad \mathbb{P}(\|f_n\|_{L^2(\Omega)} \geq \rho_0 z) \leq 2e^{-Cz^2}.$$

Moreover, if the eigenvalues ρ_k of (45) satisfy that $\rho_k = O(k^{4/d})$, then

$$\mathbb{P}(\|f_n - f^*\|_{L^2(\Omega)} \geq \lambda_n^{1/4} \rho_0 z) \leq 2e^{-Cz^2}.$$

3.2 Finite element method for the inverse heat source problem

In this section we consider a finite element approximation to the optimal control problem (43) associated with the inverse heat source problem (TIP). For convenience, we assume Ω is a polygonal or polyhedral domain in \mathbf{R}^d ($d = 2, 3$). Let \mathcal{M}_h be a family of shape-regular and quasi-uniform finite element meshes over the domain Ω , and $V_h \subset H_0^1(\Omega)$ be the conforming linear finite element space over the mesh \mathcal{M}_h . We divide the time interval $(0, T)$ into a uniform grid with time step size $\tau = T/N$ and write $t^i = i\tau$ for $i = 0, 1, \dots, N$.

We will use the backward Euler scheme in time and the linear finite element method in space to approximate the heat conduction problem (41): Find $u_h^i \in V_h$, $i = 1, 2, \dots, N$, such that

$$\left(\frac{u_h^i - u_h^{i-1}}{\tau}, v_h \right) + a(u_h^i, v_h) = (fg^i, v_h) \quad \forall v_h \in V_h, \quad (47)$$

where $a(v, w) = (a\nabla v, \nabla w) + (cv, w)$ for any $v, w \in H_0^1(\Omega)$. We approximate the forward solution Sf by $S_{\tau,h}f = u_h^N$. The inverse problem (43) can be approximated by the following least-squares problem

$$\min_{f \in V_h} \|S_{\tau,h}f - m\|_n^2 + \lambda_n \|f\|_{L^2(\Omega)}^2. \quad (48)$$

We shall make use of the results in section 3.1 to study the stochastic convergence of the solution $f_{\tau,h}$ of the problem (48) to the true solution $f^* \in L^2(\Omega)$.

Verification of Assumption 2.3. Let $P_h : L^2(\Omega) \rightarrow V_h$ be the orthogonal projection operator in the L^2 inner product. For any $f \in X = L^2(\Omega)$, we know from (47) that $S_{\tau,h}f = S_{\tau,h}(P_h f)$. Therefore, Assumption 2.3 (2) is trivially satisfied. It remains to check Assumption 2.3 (1), which amounts to derive the error estimate of the fully discrete method (47). The classical theory for the implicit Euler scheme in time and finite element method in space for solving parabolic equations requires the regularity $\partial_{tt}u \in L^1(0, T; L^2(\Omega))$ of the solution of the problem (41) (see e.g., [41, Chapter 1]). This regularity requires the compatibility condition $F(x, 0) = f(x)g(0) = 0$ on $\partial\Omega$, which may not be convenient to meet in practice. Instead, we will derive an error estimate in the remaining part of this section, without this compatibility condition, by adapting some arguments in [41, Chapter 3] for the error estimates of finite element solutions to parabolic equations with rough initial data.

We start with the weak $W^{2,1}(0, T; L^2(\Omega))$ regularity for the solution to (41).

Lemma 3.7. *Let $F(x, t) = f(x)g(t)$ for $(x, t) \in \Omega \times (0, T)$, with $g \in H^2(0, T)$. Then there exists a generic constant C such that the solution u to (41) satisfies*

$$\begin{aligned} \|\partial_t u\|_{C([0, T]; L^2(\Omega))} &\leq C\|F(\cdot, 0)\|_{L^2(\Omega)} + C \int_0^T \|\partial_t F\|_{L^2(\Omega)} dt, \\ \|t\partial_{tt}u\|_{C([0, T]; L^2(\Omega))} &\leq C\|F(\cdot, 0)\|_{L^2(\Omega)} + C \int_0^T (\|\partial_t F\|_{L^2(\Omega)} + t\|\partial_{tt}F\|_{L^2(\Omega)}) dt. \end{aligned}$$

Proof. The proof follows from the standard energy argument, so only an outline is given here. We differentiate the first equation in (41) in time to see that $v(x, t) = \partial_t u$ satisfies the conditions that $v = 0$ on $\partial\Omega \times (0, T)$ and $v(x, 0) = F(x, 0)$ in Ω , and

$$\partial_t v + Lv = \partial_t F(x, t) \quad \text{in } \Omega \times (0, T). \quad (49)$$

Then the first estimate in the lemma follows by multiplying both sides of equation (49) by v and integrating by parts.

Next we multiply both sides of (49) by $t\partial_t v$, then integrate by parts and apply the first estimate in the lemma to get

$$\int_0^t t \|\partial_t v\|_{L^2(\Omega)}^2 dt \leq C \|F(\cdot, 0)\|_{L^2(\Omega)}^2 + C \left(\int_0^T \|\partial_t F\|_{L^2(\Omega)} dt \right)^2 + C \int_0^T t \|\partial_t F\|_{L^2(\Omega)}^2 dt. \quad (50)$$

Finally, we differentiate the equation (49) in time to get

$$\partial_{tt} v + L(\partial_t v) = \partial_{tt} F(x, t) \quad \text{in } \Omega \times (0, T).$$

By multiplying both sides of the equation by $t^2 \partial_t v$, integrating by parts again and applying (50), we obtain

$$\begin{aligned} t \|\partial_t v\|_{L^2(\Omega)} &\leq C \|F(\cdot, 0)\|_{L^2(\Omega)} + C \int_0^T (\|\partial_t F\|_{L^2(\Omega)} + t \|\partial_{tt} F\|_{L^2(\Omega)}) dt \\ &\quad + C \left(\int_0^T t \|\partial_t F\|_{L^2(\Omega)}^2 dt \right)^{1/2}, \end{aligned}$$

which implies the second estimate of the lemma by noticing that

$$\begin{aligned} \int_0^T t \|\partial_t F\|_{L^2(\Omega)}^2 dt &\leq \sup_{t \in (0, T)} \|t \partial_t F\|_{L^2(\Omega)} \cdot \int_0^T \|\partial_t F\|_{L^2(\Omega)} dt \\ &= \sup_{t \in (0, T)} \left\| \int_0^t \partial_s (s \partial_s F(s)) ds \right\|_{L^2(\Omega)} \cdot \int_0^T \|\partial_t F\|_{L^2(\Omega)} dt \\ &\leq \int_0^T (\|\partial_t F\|_{L^2(\Omega)} + t \|\partial_{tt} F\|_{L^2(\Omega)}) dt \cdot \int_0^T \|\partial_t F\|_{L^2(\Omega)} dt. \end{aligned}$$

This completes the proof. \square

Lemma 3.8. *Let $u_h \in H^1(0, T; V_h)$ be the following semi-discrete finite element solution of the problem (41):*

$$(\partial_t u_h, v_h) + a(u_h, v_h) = (F, v_h) \quad \forall v_h \in V_h \quad \text{a.e. in } (0, T). \quad (51)$$

Then there exists a constant C independent of the mesh size h such that

$$\|u - u_h\|_{C([0, T]; L^2(\Omega))} \leq Ch^2 \max_{t \in [0, T]} (\|\partial_t u\|_{L^2(\Omega)} + \|t \partial_{tt} u\|_{L^2(\Omega)} + \|F\|_{L^2(\Omega)} + \|t \partial_t F\|_{L^2(\Omega)}),$$

where $h = \max_{K \in \mathcal{M}} h_K$ and h_K is the diameter of the element $K \in \mathcal{M}$.

Proof. We follow the argument in [41, Chapter 3]. Define $G : L^2(\Omega) \rightarrow H_0^1(\Omega)$ and $G_h : L^2(\Omega) \rightarrow V_h$ such that for any $w \in L^2(\Omega)$, $Gw \in H_0^1(\Omega)$, $G_h w \in V_h$ satisfy

$$a(Gw, v) = (w, v) \quad \forall v \in H_0^1(\Omega); \quad a(G_h w, v_h) = (w, v_h) \quad \forall v_h \in V_h.$$

The equations (41) and (51) can be reformulated as

$$\partial_t(Gu) + u = GF, \quad \partial_t(G_h u_h) + u_h = G_h F.$$

Writing $e = u - u_h$, then we know e satisfies

$$G_h(\partial_t e) + e = \rho \quad \text{a.e. in } (0, T), \quad (G_h e)(\cdot, 0) = 0 \quad \text{in } \Omega,$$

where $\rho = (G_h - G)(\partial_t u) + (G - G_h)F$. By the argument in the proof of Lemma 3.7 we can obtain (see [41, Lemma 3.4]) that

$$\max_{t \in [0, T]} \|e\|_{L^2(\Omega)} \leq C \max_{t \in [0, T]} (\|\rho(t)\|_{L^2(\Omega)} + \|t\partial_t \rho(t)\|_{L^2(\Omega)}).$$

This completes the proof by noting that $\|Gw - G_h w\|_{L^2(\Omega)} \leq Ch^2 \|w\|_{L^2(\Omega)} \quad \forall w \in L^2(\Omega)$, which follows by the Aubin-Nitsche argument since the domain Ω is convex. \square

The following lemma for the error estimate of the fully discrete finite element method was not covered by the general results in [41, Chapter 8] since we do not have the condition that $F(x, 0) = 0$ on $\partial\Omega$ here, which was critical in [41].

Lemma 3.9. *Let $u_h \in H^1(0, T; V_h)$ be the solution of the problem (51) and $u_h^i \in V_h, i = 1, 2, \dots, N$, be the solution of the problem (47). Then there exists a constant C independent of h, τ such that*

$$\max_{1 \leq i \leq N} \|u_h(\cdot, t_i) - u_h^i\|_{L^2(\Omega)} \leq C\tau(1 + \ln N)(\|F\|_{C([0, T]; L^2(\Omega))} + \|\partial_t F\|_{C([0, T]; L^2(\Omega))}).$$

Proof. Let $\{\lambda_j\}_{j=1}^M$ be the eigenvalues of the eigenvalue problem

$$a(\phi_h, v_h) = \lambda(\phi_h, v_h) \quad \forall v_h \in V_h,$$

and $\{\phi_j\}_{j=1}^M$ be the corresponding eigenfunctions which form an orthonormal basis of V_h in the $L^2(\Omega)$ norm. By the Poincaré inequality, we know that $\lambda_j \geq C, j = 1, 2, \dots, M$, for some constant C independent of the mesh size h .

We write $u_h(x, t) = \sum_{j=1}^M u_j(t)\phi_j(x)$ and $F(x, t) = \sum_{j=1}^M F_j(t)\phi_j(x)$, where $u_j(t) = (u_h(\cdot, t), \phi_j)$ and $F_j(t) = (F(\cdot, t), \phi_j)$. Then it follows from (51) that

$$u_j'(t) + \lambda_j u_j = F_j(t) \quad \text{a.e. in } (0, T),$$

whose solution can be written as

$$u_j(t^i) = \int_0^{t^i} e^{\lambda_j(s-t^i)} F_j(s) ds = \int_0^{t^i} e^{-\lambda_j t} F_j(t^i - t) dt. \quad (52)$$

Similarly, we write $u_h^i = \sum_{j=1}^M U_j^i \phi_j$, where $U_j^i = (u_h^i, \phi_j), i = 1, 2, \dots, N, j = 1, 2, \dots, M$. From (47) we know that

$$\frac{1}{\tau}(U_j^i - U_j^{i-1}) + \lambda_j U_j^i = F_j^i := F_j(t^i), \quad i = 1, 2, \dots, N, j = 1, 2, \dots, M.$$

This implies that $U_j^i = r(\lambda_j)U_j^{i-1} + \tau r(\lambda_j \tau)F_j^i$, where $r(t) = (1 + t)^{-1} \quad \forall t \geq 0$, hence

$$U_j^i = \sum_{k=1}^i \tau r(\lambda_j \tau)^k F_j^{i-k+1}. \quad (53)$$

For any $j = 1, \dots, M$, we distinguish two cases. If $\lambda_j \tau \geq 1$, we know from (52) that

$$|u_j(t^i)| \leq \|F_j\|_{C[0, T]} \int_0^{t^i} e^{-\lambda_j t} dt = \lambda_j^{-1}(1 - e^{-\lambda_j t^i}) \|F_j\|_{C[0, T]} \leq \tau \|F_j\|_{C[0, T]}.$$

On the other hand, we obtain from (53) that

$$|U_j^i| \leq \left(\sum_{k=1}^i 2^{-k} \right) \tau \|F_j\|_{C[0,T]} \leq 2\tau \|F_j\|_{C[0,T]}.$$

Therefore, we derive for $\lambda_j \tau \geq 1$ that

$$|u_j^i(t^i) - U_j^i| \leq C\tau \|F_j\|_{C[0,T]}. \quad (54)$$

Now we consider the case when $\lambda_j \tau \leq 1$. By (52) we have

$$\begin{aligned} u_j(t^i) &= \sum_{k=1}^i \int_{t^{k-1}}^{t^k} e^{-\lambda_j t} (F_j(t^i - t) - F(t^i - t^{k-1})) dt + \sum_{k=1}^i \int_{t^{k-1}}^{t^k} e^{-\lambda_j t} F_j^{i-k+1} dt \\ &= \sum_{k=1}^i \int_{t^{k-1}}^{t^k} e^{-\lambda_j t} (F_j(t^i - t) - F(t^i - t^{k-1})) dt + \sum_{k=1}^i \tau \frac{e^{\lambda_j \tau} - 1}{\lambda_j \tau} e^{-k\lambda_j \tau} F_j^{i-k+1}, \end{aligned}$$

which, together with (53), yields

$$\begin{aligned} u_j(t^i) - U_j^i &= \sum_{k=1}^i \tau \left(\frac{e^{\lambda_j \tau} - 1}{\lambda_j \tau} e^{-k\lambda_j \tau} - r(\lambda_j \tau)^k \right) F_j^{i-k+1} \\ &\quad + \sum_{k=1}^i \int_{t^{k-1}}^{t^k} e^{-\lambda_j t} (F_j(t^i - t) - F(t^i - t^{k-1})) dt := \text{I} + \text{II}. \end{aligned} \quad (55)$$

Recalling the following elementary estimate in [41, (7.22)],

$$|e^{-kt} - r(t)^k| \leq Ck^{-1} \quad \forall t \geq 0, \forall k = 1, 2, \dots,$$

and using the fact that $(t^{-1}(e^t - 1) - 1)/(1 - e^{-t})$ is bounded for $0 \leq t \leq 1$, we obtain

$$\begin{aligned} |\text{I}| &\leq \sum_{k=1}^i \tau \left| \left(\frac{e^{\lambda_j \tau} - 1}{\lambda_j \tau} - 1 \right) e^{-k\lambda_j \tau} + (e^{-k\lambda_j \tau} - r(\lambda_j \tau)^k) \right| |F_j^{i-k+1}| \\ &\leq C\tau \left[\left(\frac{e^{\lambda_j \tau} - 1}{\lambda_j \tau} - 1 \right) \frac{1}{1 - e^{-\lambda_j \tau}} + \sum_{k=1}^i k^{-1} \right] \|F_j\|_{C[0,T]} \\ &\leq C(1 + \ln i) \tau \|F_j\|_{C[0,T]}. \end{aligned}$$

The term II can be bounded by the standard argument as follows:

$$\text{II} \leq C\tau \|\partial_t F_j\|_{C[0,T]} \int_0^{t^i} e^{-\lambda_j t} dt \leq C\lambda_j^{-1} \tau \|\partial_t F_j\|_{C[0,T]} \leq C\tau \|\partial_t F_j\|_{C[0,T]},$$

where we have used the fact that $\lambda_j \geq C$ for some constant C independent of h .

Combining (54), (55) and the above two estimates we obtain

$$|u_j(t^i) - U_j^i| \leq C\tau(1 + \ln N)(\|F_j\|_{C[0,T]} + \|\partial_t F_j\|_{C[0,T]}).$$

This completes the proof. \square

By Lemmata 3.7-3.9, we know that under the condition $g \in H^2(0, T)$,

$$\|S_{\tau,h}f - Sf\|_{L^2(\Omega)} \leq C(h^2 + \tau|\ln \tau|)\|f\|_{L^2(\Omega)}, \quad (56)$$

for some constant C which depends possibly on T , $\|g\|_{H^2(0,T)}$ but is independent of h and τ .

Assumption 2.3 (1) is now a consequence of the following lemma.

Lemma 3.10. *If $g \in H^2(0, T)$, $S_{\tau,h}f = u_h^N$ with u_h^N being the solution of the problem (47), then for any $f \in L^2(\Omega)$, there exists a constant C independent of h and τ such that*

$$\|Sf - S_{\tau,h}f\|_n \leq C(h^2 + \tau|\ln \tau|)\|f\|_{L^2(\Omega)}.$$

Proof. Let $\Pi_h : C(\bar{\Omega}) \rightarrow V_h$ be the canonical finite element interpolant, then we know from the standard interpolation theory of finite element methods [14] that

$$\begin{aligned} \|Sf - \Pi_h(Sf)\|_{L^\infty(K)} &\leq Ch^{2-d/2}\|Sf\|_{H^2(K)} \quad \forall K \in \mathcal{M}_h, \\ \|Sf - \Pi_h(Sf)\|_{L^2(K)} &\leq Ch^2\|Sf\|_{H^2(K)} \quad \forall K \in \mathcal{M}_h. \end{aligned}$$

Let $\mathbb{T}_K = \{x_i : x_i \in K, 1 \leq i \leq n\}$. By the assumption that $\{x_i\}_{i=1}^n$ is quasi-uniformly distributed and the mesh \mathcal{M}_h is quasi-uniform, we know that the cardinal $\#\mathbb{T}_K \leq Cnh^d$. Thus we have

$$\|Sf - \Pi_h(Sf)\|_n^2 \leq \frac{1}{n} \sum_{K \in \mathcal{M}_h} \#\mathbb{T}_K \|Sf - \Pi_h(Sf)\|_{L^\infty(K)}^2 \leq Ch^4\|Sf\|_{H^2(\Omega)}^2.$$

On the other hand, we can derive by making use of inverse estimates that

$$\begin{aligned} \|S_{\tau,h}f - \Pi_h(Sf)\|_n^2 &\leq \frac{1}{n} \sum_{K \in \mathcal{M}_h} \#\mathbb{T}_K \|S_{\tau,h}f - \Pi_h(Sf)\|_{L^\infty(K)}^2 \\ &\leq \frac{1}{n} \sum_{K \in \mathcal{M}_h} \#\mathbb{T}_K |K|^{-1} \|S_{\tau,h}f - \Pi_h(Sf)\|_{L^2(K)}^2 \\ &\leq C \|S_{\tau,h}f - \Pi_h(Sf)\|_{L^2(\Omega)}^2 \\ &\leq C \|S_{\tau,h}f - Sf\|_{L^2(\Omega)}^2 + C \|\Pi_h(Sf) - Sf\|_{L^2(\Omega)}^2 \\ &\leq C \|S_{\tau,h}f - Sf\|_{L^2(\Omega)}^2 + Ch^4\|Sf\|_{H^2(\Omega)}^2. \end{aligned}$$

Therefore,

$$\|Sf - S_{\tau,h}f\|_n \leq C \|S_{\tau,h}f - Sf\|_{L^2(\Omega)} + Ch^2\|f\|_{L^2(\Omega)}.$$

This completes the proof by (56). \square

After the verification of Assumption 2.3, the following stochastic convergence of the finite element method to the inverse heat source problem follows readily from Theorem 2.8.

Theorem 3.11. *Let $g \in H^2(0, T)$ and the measurement data (1) be of the type **(R1)**. Then there exist constants $\lambda_0 > 0$ and $C > 0$ such that for any $\lambda_n \leq \lambda_0$ and $\tau|\ln \tau| = O(h^2)$, the following estimates hold for the solutions $f_n \in L^2(\Omega)$ to (43) and $f_h \in V_h$ to (48):*

$$\begin{aligned} \mathbb{E}[\|Sf^* - S_{\tau,h}f_h\|_n^2] &\leq C(\lambda_n + h^4)\|f^*\|_{L^2(\Omega)}^2 + C \left(1 + \frac{h^4}{\lambda_n}\right) \frac{\sigma^2}{n\lambda_n^{d/4}}, \\ \mathbb{E}[\|f^* - f_h\|_{L^2(\Omega)}^2] &\leq C \left(1 + \frac{h^4}{\lambda_n}\right) \|f^*\|_{L^2(\Omega)}^2 + C \left(1 + \frac{h^4}{\lambda_n}\right) \frac{\sigma^2}{n\lambda_n^{1+d/4}}. \end{aligned}$$

$$\mathbb{E}[\|f^* - f_h\|_{H^{-1}(\Omega)}^2] \leq C(\lambda_n^{1/2} + h^2) \left(1 + \frac{h^4}{\lambda_n}\right) \|f^*\|_{L^2(\Omega)}^2 + C(\lambda_n^{1/2} + h^2) \left(1 + \frac{h^4}{\lambda_n}\right) \frac{\sigma^2}{n\lambda_n^{1+d/4}}.$$

Proof. Since the mesh is assumed to be quasi-uniform, the dimension N_h of the linear finite element space V_h is bounded by $N_h \leq Ch^{-d}$. By Theorem 3.2, we know that $\alpha = 4/d$. Take $\tau|\ln \tau| = O(h^2)$, then we know from Theorem 2.8 that

$$\mathbb{E}[\|Sf^* - S_{\tau,h}f_h\|_n^2] \leq C(\lambda_n + h^4) \|f^*\|_{L^2(\Omega)}^2 + C \left[1 + \frac{h^4}{\lambda_n} + \left(\frac{h^4}{\lambda_n}\right)^{1-\frac{d}{4}}\right] \frac{\sigma^2}{n\lambda_n^{1+d/4}}.$$

We can easily check that $(h^4/\lambda_n)^{1-\frac{d}{4}} \leq 1$ for $h^4/\lambda_n \leq 1$, and $(h^4/\lambda_n)^{1-\frac{d}{4}} \leq h^4/\lambda_n$ for $h^4/\lambda_n \geq 1$. Therefore, we have $(h^4/\lambda_n)^{1-\frac{d}{4}} \leq 1 + h^4/\lambda_n$. This leads to the conclusions of Theorem 3.11. \square

We end this section with the following convergence of the finite element method to the inverse heat source problem **(TIP)**, directly following from Theorem 2.11 by noticing that $N_h \leq Ch^{-d} \leq C\lambda_n^{-\gamma/2}$ with $\gamma = d/2$.

Theorem 3.12. *Let $g \in H^2(0, T)$, the measurement data (1) is of the type **(R2)** and $\rho_0 = \|f^*\|_{L^2(\Omega)} + \sigma n^{-1/2}$. If we take $h = O(\lambda_n^{1/4})$, $\tau|\ln \tau| = O(\lambda_n^{1/2})$, and $\lambda_n^{1/2+d/8} = O(\sigma n^{-1/2} \rho_0^{-1})$, then there exists a constant $C > 0$ such that for any $z > 0$,*

$$\mathbb{P}(\|S_{\tau,h}f_h - Sf^*\|_n \geq \lambda_n^{1/2} \rho_0 z) \leq 2e^{-Cz^2}, \quad \mathbb{P}(\|f_h\|_{L^2(\Omega)} \geq \rho_0 z) \leq 2e^{-Cz^2}.$$

$$\mathbb{P}(\|f_h - f^*\|_{H^{-1}(\Omega)} \geq \lambda_n^{1/4} \rho_0 z) \leq 2e^{-Cz^2}.$$

4 Numerical examples

In this section, we present several numerical examples to confirm the theoretical results in previous sections. We take the domain $\Omega = (0, 1) \times (0, 1)$ and a set of uniformly distributed measurement locations $\{x_i\}_{i=1}^n$ in Ω . In all examples below, we take the coefficients $a(x) = 1, c(x) = 0$, which fulfill the uniform ellipticity condition, and $g(t) \equiv 1, T = 1$. The finite element mesh \mathcal{M}_h of Ω is constructed by first dividing Ω into $h^{-1} \times h^{-1}$ uniform rectangles and then connecting the lower left and upper right vertices of each rectangle. We set the noise e_1, \dots, e_n in the dataset (1) to be the normal random variables with variance σ .

Motivated by Theorem 3.5, we propose a self-consistent algorithm to determine the regularization parameter λ_n in (48) based on the rule

$$\lambda_n^{1/2+d/8} = \sigma n^{-1/2} \|f^*\|_{L^2(\Omega)}^{-1}. \quad (57)$$

This choice requires the knowledge of the true source function f^* and the noise level σ . We propose now a self-consistent algorithm to determine the parameter λ_n , without knowing the true source function f^* and the noise level σ . To do so, we estimate $\|f^*\|_{L^2(\Omega)}$ by $\|f_h\|_{L^2(\Omega)}$ and σ by $\|S_{\tau,h}f_h - m\|_n$ since $\|Sf^* - m\|_n = \|e\|_n$. This is expected to yield a good estimate of the variance by the law of large numbers.

Algorithm 4.1 (Computing an estimate of the regularization parameter λ_n). 1° *Given an initial guess of $\lambda_{n,0}$; for $j = 0, 1, \dots$, do the following*

2° *Solve (48) for f_h with λ_n replaced by $\lambda_{n,j}$ over the mesh \mathcal{M}_h ;*

3° *Update $\lambda_{n,j+1}$: $\lambda_{n,j+1}^{1/2+d/8} = n^{-1/2} \|S_{\tau,h}f_h - m\|_n \|f_h\|_{L^2(\Omega)}^{-1}$.*

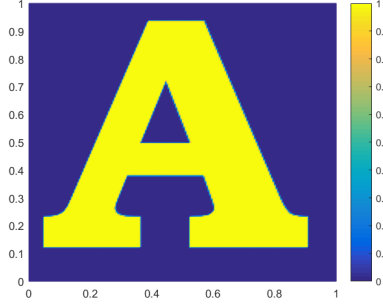


Figure 1: The surface plot of the exact solution f^* .

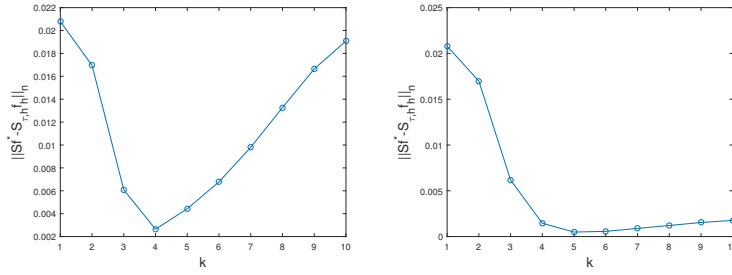


Figure 2: The empirical errors $\|Sf^* - S_{\tau,h}f_h\|_n$ with $\lambda_n = 10^{-k}$ ($k = 1, \dots, 10$) for $\sigma = 0.1$ (left) and $\sigma = 0.01$ (right).

A natural choice of the initial guess is $\lambda_{n,0} = n^{-4/(d+4)}$ since f^* and σ are unknown, which is used in our numerical examples.

Example 4.1. This example is used to verify the nearly optimality of the choice of the smoothing parameter λ_n suggested by (57). We choose $n = 10^4$, $\sigma = 0.1$ or $\sigma = 0.01$, and the mesh size $h = 0.05$ and the time step size $\tau = 0.01$, which are sufficiently small so that the finite element errors are negligible. We take the true source f^* to be the function whose surface is given as in Fig. 1.

Example 4.1 demonstrates the nearly optimality of the choice of the smoothing parameter λ_n suggested by (57). In fact, we have $\|f^*\|_{L^2(\Omega)} \approx 0.54$, then (57) suggests $\lambda_n \approx 2.3 \times 10^{-4}$ (for $\sigma = 0.1$) and $\lambda_n \approx 1.1 \times 10^{-5}$ (for $\sigma = 0.01$). These two approximate λ_n 's are indeed very close to the optimal $\lambda_n = 1 \times 10^{-4}$ (for $\sigma = 0.1$) and $\lambda_n = 1 \times 10^{-5}$ (for $\sigma = 0.01$), which we have estimated by computing the errors $\|Sf^* - S_{\tau,h}f_h\|_n$ with 10 different choices of regularization parameter: $\lambda_{n,k} = 10^{-k}$ ($k = 1, 2, \dots, 10$), see Fig. 2.

Example 4.2. This example is presented to verify if the probability density function of the empirical error $\|Sf^* - S_{\tau,h}f_h\|_n$ has an exponentially decaying tail. We set the variance $\sigma = 0.001$, $n = 25 \times 10^4$, and choose the mesh size h and time step size τ to be small enough so that the finite element errors are negligible. We take 10,000 samples and compute the empirical error $\|Sf^* - S_{\tau,h}f_h\|_n$ for each sampling.

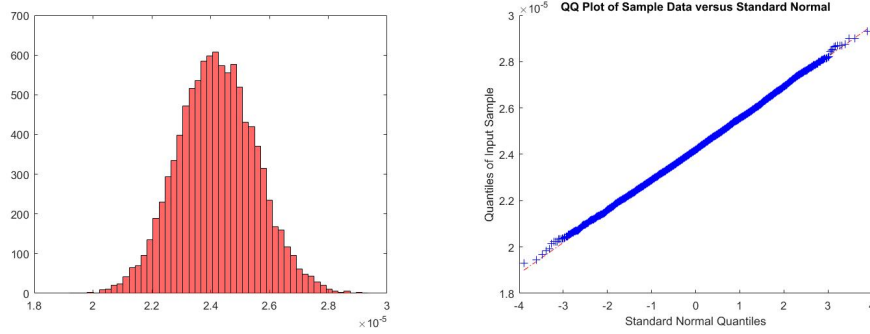


Figure 3: The histogram (left) and quantile-quantile (right) plots of the empirical error $\|S_{\tau,h}f_h - Sf^*\|_n$ with 10,000 samples.

In Example 4.2, we can compute that $\|Sf^*\|_{L^\infty(\Omega)} \approx 0.04$, so the relative noise level $\sigma/\|Sf^*\|_{L^\infty(\Omega)}$ is about 2.5% for this example. Figure 3(a) shows the histogram plot of the empirical errors, while Figure 3(b) plots the quantile-quantile (Q-Q) plot to compare the sample distribution of the empirical error with the standard normal distribution. The Q-Q plot is a standard graphic tool in statistics to check the data distribution [46]. If the sample distribution is indeed normal, the Q-Q plot should give a scattered plot, where the points show a linear relationship between the sample and the theoretical quantiles. We can observe from Figure 3 (right) that almost all the points are concentrated around the dotted line, which implies that the overall distribution of the error is very close to a normal distribution. Moreover, the points around the two ends are also not far from the line, which indicates that the tail distribution of the error is also close to a Gaussian tail, as indicated in Theorem 3.12. The probability density function is computed by the Matlab function 'qqplot'.

Example 4.3. This example is to confirm Theorems 3.11 and 3.12, namely, to verify if the empirical error $\|Sf^* - S_{\tau,h}f_h\|_n$ depends linearly on $\lambda_n^{1/2}$ when the regularization parameter λ_n is taken by the optimal choice (57). The mesh size $h = \lambda_n^{1/4}$ and the time step size $\tau|\ln \tau| = \lambda_n^{1/2}$ are chosen according to Theorems 3.11 and 3.12. We take the true source f^* to be the function given in Figure 1, and n to change from 25×10^2 to 25×10^4 .

We can see from Figure 4 clearly the linear dependence of the empirical error $\|Sf^* - S_{\tau,h}f_h\|_n$ on $\lambda_n^{1/2}$ for $\sigma = 0.01$ and 0.04 . We can compute that $\|Sf^*\|_{L^\infty(\Omega)} \approx 0.04$, so the relative noise levels $\sigma/\|Sf^*\|_{L^\infty(\Omega)}$ are about 25% and 100% for $\sigma = 0.01$ and 0.04 , respectively.

Through the previous 3 examples, we have verified the optimality of the choice rule (57) for λ_n , the stochastic convergence (Theorem 3.12), and the convergence order of the finite element method. But we do not know the exact solution and the variance of the noise in most applications, so we use the next example to show the efficiency of Algorithm 4.1 to determine an optimal regularization parameter λ_n iteratively, without the knowledge of f^* and σ .

Example 4.4. We choose $n = 25 \times 10^4$ and set the noise e_1, \dots, e_n in the dataset (1) to be independent normal random variables with variance $\sigma = 0.001$. Algorithm 4.1 is terminated when the absolute difference between two consecutive iterates $\lambda_{n,k}$ and $\lambda_{n,k+1}$ is less than 10^{-10} .

We can compute that $\|Sf^*\|_{L^\infty(\Omega)} \approx 0.04$, so the relative noise level $\sigma/\|Sf^*\|_{L^\infty(\Omega)}$ is about 2.5% in this example. Figure 5 shows clearly the convergence of the sequence $\{\lambda_{n,k}\}$ generated by Algorithm

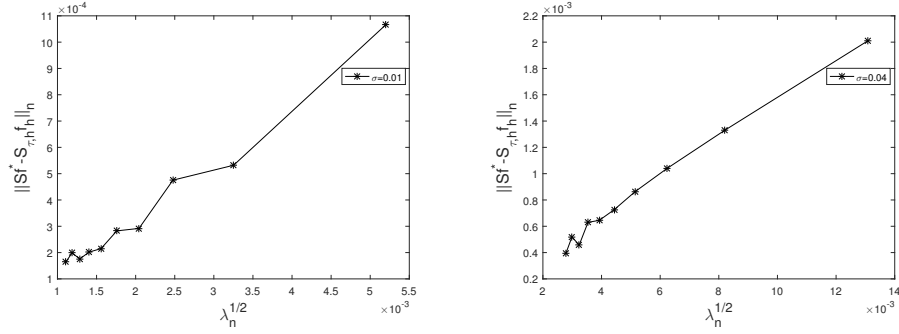


Figure 4: *Linear dependence of the empirical error $\|Sf^* - S_{\tau,h}f_h\|_n$ on $\lambda_n^{1/2}$ with $\sigma = 0.01$ (left) and $\sigma = 0.04$ (right).*

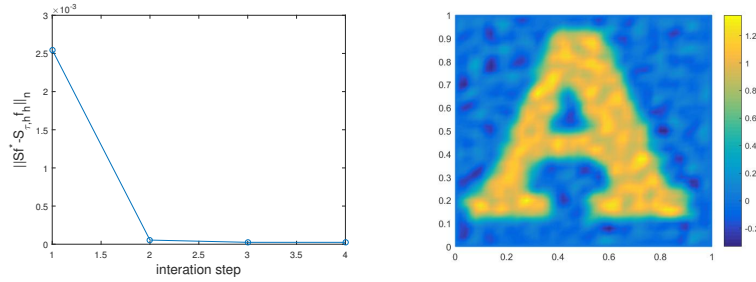


Figure 5: *The relative empirical error $\|Sf^* - S_{\tau,h}f_h\|_n$ at each iteration (left); The computed solution f_h at the end of iterations (right).*

4.1. The numerical computation gives $\lambda_{n,4} = 5.53 \times 10^{-8}$ that agrees very well with the optimal choice 5.33×10^{-8} given by (57). Furthermore, $\|m - S_{\tau,h}f_h\|_n = 9.99 \times 10^{-4}$ provides also a good estimate of the variance σ .

References

- [1] B. Abdelaziz, A. El Badia and A. El Hajj, *Reconstruction of extended sources with small supports in the elliptic equation $\Delta u + \mu u = F$ from a single Cauchy data*, Comptes Rendus Mathematique, 351 (2013), pp. 797-801.
- [2] S. Agmon, *Lectures on Elliptic Boundary Problems*, Van Norstrand, Princeton, NJ, 1965.
- [3] V. Akcelik, G. Biros, A. Draganescu, O. Ghattas, J. Hill, and B. Waanders, *Dynamic data-driven inversion for terascale simulations: Real-time identification of airborne contaminants*, Proceedings of Supercomputing, Seattle, Washington, 2005.
- [4] J. Atmadja and A. Bagtzoglou, *State of the art report on mathematical methods for groundwater pollution source identification*, Environmental Forensics, 2 (2001), pp. 205-214.

- [5] A. El Badia, A. El Hajj, M. Jazar, and H. Moustafa, *Lipchitz stability estimates for an inverse source problem in an elliptic equation from interior measurements*, *Applicable Analysis*, 95 (2016), pp. 1873-1890.
- [6] A. El Badia, T. Ha Duong, and F. Moutazaim, *Numerical solution for the identification of source terms from boundary measurements*, *Inverse Problems in Engineering*, 8 (2000), pp. 345-364
- [7] A. El Badia and T. Nara, *An inverse source problem for Helmholtz's equation from the Cauchy data with a single wave number*, *Inverse Problems*, 27 (2011), 105001.
- [8] M.S. Birman and M.Z. Solomyak, *Piecewise polynomial approximations of functions of the classes W_{α}^k* , *Mat. Sb.*, 73 (1967), pp. 331-355.
- [9] R.I. Bot and B. Hofmann, *An extension of the variational inequality approach for nonlinear ill-posed problems*, *Journal of Integral Equations and Applications*, 22 (2010), pp. 369-392.
- [10] S. Boucheron, G. Lugosi, and P. Massart, *Concentration inequalities using the entropy method*, *Ann. Probab.*, 31 (2003), pp. 1583-1614.
- [11] D.H. Chen, D. Jiang, and J. Zou, *Convergence rates of Tikhonov regularizations for elliptic and parabolic inverse radiativity problems*, *Inverse Problems*, to appear.
- [12] D.H. Chen and I. Yousept, *Variational source condition for ill-posed backward nonlinear Maxwell's equations*, *Inverse Problems*, 35 (2019), 025001.
- [13] Z. Chen, R. Tuo, and W. Zhang, *Stochastic convergence of a nonconforming finite element method for the thin plate spline smoother for observational data*, *SIAM J. Numer. Anal.*, 56 (2018), pp. 635-659.
- [14] P.G. Ciarlet, *The Finite Element Method for Elliptic Problems*, North-Holland, Amsterdam, 1978.
- [15] H.W. Engl, K. Kunisch, and A. Neubauer, *Convergence rates for Tikhonov regularization of nonlinear ill-posed problems*, *Inverse problems*, 5 (1989), pp. 523-540.
- [16] H.W. Engl and J. Zou, *A new approach to convergence rate analysis of Tikhonov regularization for parameter identification in heat conduction*, *Inverse Problems*, 16 (2000), pp. 1907-1923.
- [17] L.C. Evans, *Partial Differential Equations*, American Mathematical Society, Providence, Rhode Island, 1998.
- [18] J. Fleckinger and M. Lapidus, *Eigenvalues of elliptic boundary value problems with an indefinite weight function*, *Transactions of the American Mathematical Society*, 295 (1986), pp. 305-324.
- [19] J. Flemming, *Theory and examples of variational regularization with non-metric fitting functionals*, *J. Inverse Ill-Posed Probl.*, 18 (2010), pp. 677-699.
- [20] G. Garcia, A. Osses, and M. Tapia, *A heat source reconstruction formula from single internal measurements using a family of null controls*, *J. Inverse Ill-Posed Probl.*, 21 (2013), pp. 755-779.
- [21] S. Gorelick, B. Evans, and I. Remson, *Identifying sources of groundwater pollution: an optimization approach*, *Water Resources Research*, 19 (1983), pp. 779-790.
- [22] M. Grasmair, *Generalized Bregman distances and convergence rates for non-convex regularization methods*, *Inverse Problems*, 26 (2010), 115014.

- [23] A. Hamdi, *The recovery of a time-dependent point source in a linear transport equation: application to surface water pollution*, Inverse Problems, 24 (2009), pp. 1-18.
- [24] A. Hamdi, *Identification of a time-varying point source in a system of two coupled linear diffusion-advection- reaction equations: application to surface water pollution*, Inverse Problems, 25 (2009), pp. 1-21.
- [25] B. Hofmann, B. Kaltenbacher, C. Pöschl, and O. Scherzer, *A convergence rates result for Tikhonov regularization in Banach spaces with non-smooth operators*, Inverse Problems, 23 (2007), pp. 987–1010.
- [26] T. Hohage and F. Weilding, *Verification of a variational source condition for acoustic inverse medium scattering problems*, Inverse Problems, 31 (2015), 075006.
- [27] T. Hohage and F. Weilding, *Variational source condition and stability estimates for inverse electromagnetic medium scattering problems*, Inverse Problems Imaging, 11 (2017), pp. 203-220.
- [28] Q. Hu, S. Shu, and J. Zou, *A new variational approach for inverse source problems*, Numer. Math. Theor. Meth. Appl., 12 (2019), pp. 331-347.
- [29] V. Isakov, *Inverse source problems for partial differential equations*, Springer-Verlag, New York, 1998.
- [30] V. Isakov, S. Leung, and J. Qian, *A Three-dimensional inverse gravimetry problem for ice with snow caps*, Inverse Problems and Imaging, 7 (2013), pp. 523-544.
- [31] B. Jin and Z. Zhou, *Error analysis of finite element approximations of diffusion coefficient identification for elliptic and parabolic problems*, arXiv:2010.02447.
- [32] J. Krebs, A.K. Louis and H. Wendland, *Sobolev error estimates and a priori parameter selection for semi-discrete Tikhonov regularization*, J. Inverse Ill-Posed Probl. 17 (2009), pp. 845-869.
- [33] C.-S. Liu, *An integral equation method to recover non-additive and non-separable heat source without initial temperature*, Intern. J. Heat Mass Transfer, 97 (2016), pp. 943-953.
- [34] X. Liu and Z. Zhai, *Inverse modeling methods for indoor airborne pollutant tracking literature review and fundamentals*, Indoor Air, 17 (2007), pp. 419-438.
- [35] X. Liu, *Identification of Indoor Airborne Contaminant Sources with Probability-Based Inverse Modeling Methods*, Ph.D. Thesis, Department of Civil, Environmental and Architectural Engineering, University of Colorado, 2008.
- [36] P. Nelson and S.H. Yoon, *Estimation of acoustic source strength by inverse methods: Part I, conditioning of the inverse problem*, J. Sound Vibration, 233 (2000), pp. 639-664.
- [37] G. Nunnari, A. Nucifora, and C. Randieri, *The application of neural techniques to the modelling of time-series of atmospheric pollution data*, Ecological Modelling, 111 (1998), pp. 187-205.
- [38] T. Skaggs and Z. Kabala, *Recovering the history of a groundwater contaminant plume: method of quasi-reversibility*, Water Resources Research, 31 (1995), pp. 2669-2673.
- [39] M. Snodgrass and P. Kitanidis, *A geostatistical approach to contaminant source identification*, Water Resources Research, 33 (1997), pp. 537-546.

- [40] M. Tadi, *Inverse heat conduction based on boundary measurements*, Inverse Problems, 13 (1997), pp. 1585-1605.
- [41] V. Thomee, *Galerkin Finite Element Methods for Parabolic Problems*, Springer Verlag, Berlin, 2006.
- [42] F.I. Utreras, *Convergence rates for multivariate smoothing spline functions*, J. Approx.Theory, 52 (1988), pp. 1-27.
- [43] S.A. van de Geer, *Empirical process in M-estimation*, Cambridge University Press, Cambridge, 2000.
- [44] A.W. van der Vaart and J.A. Wellner, *Weak Convergence and Empirical Processes: with Applications to Statistics*, Springer, New York, 1996.
- [45] L. Wang and J. Zou, *Error estimates of finite element methods for parameter identifications in elliptic and parabolic systems*, Disc. Cont. Dyn. Sys. B, 14 (2010), pp. 1641-1670.
- [46] M.B. Wilk and R. Gnanadesikan, *Probability plotting methods for the analysis of data*, Biometrika, 55 (1968), pp. 1-17.
- [47] J. Wong and P. Yuan, *A FE-based algorithm for the inverse natural convection problem*, Intern. J. Numer. Methods Fluids, 68 (2012), pp. 48-82.
- [48] X. Zhang, C.X. Zhu, G. Feng, H. Zhu, and P. Guo, *Potential use of bacteroidales specific 16S rRNA in tracking the rural pond-drinking water pollution*, J. Agro-Envir. Sci., 30 (2011), pp. 1880-1887.
- [49] B. Zhu, Y. Chen, and J. Peng, *Lead isotope geochemistry of the urban environment in the Pearl River Delta*, Appl. Geochem., 16 (2011), pp. 409-417.