# Consensus Answers for Queries over Probabilistic Databases

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#### Abstract

We address the problem of finding a "best" deterministic query answer to a query over a probabilistic database. For this purpose, we propose the notion of a consensus world (or a consensus answer) which is a deterministic world (answer) that minimizes the expected distance to the possible worlds (answers). This problem can be seen as a generalization of the well-studied inconsistent information aggregation problems (e.g. rank aggregation) to probabilistic databases. We consider this problem for various types of queries including SPJ queries, *Top*-k queries, group-by aggregate queries, and clustering. For different distance metrics, we obtain polynomial time optimal or approximation algorithms for computing the consensus answers (or prove NP-hardness). Most of our results are for a general probabilistic database model, called *and/xor tree model*, which significantly generalizes previous probabilistic database models like x-tuples and block-independent disjoint models, and is of independent interest.

## **1** Introduction

There is an increasing interest in uncertain and probabilistics databases arising in application domains such as information retrieval [11, 35], recommendation systems [32, 33], mobile object data management [8], information extraction [20], data integration [3] and sensor networks [13]. Supporting complex queries and decision-making on probabilistic databases is significantly more difficult than in deterministic databases, and the key challenges include defining proper and intuitive semantics for queries over them, and developing efficient query processing algorithms.

The common semantics in probabilistic databases are the "possible worlds" semantics, where a probabilistic database is considered to correspond to a probability distribution over a set of deterministic databases called "possible worlds". Therefore, posing queries over such a probabilistic database generates a probability distribution over a set of deterministic results which we call "possible answers". However, a full list of possible answers together with their probabilities is not desirable in most cases since the size of the list could be exponentially large, and the probability associated with each single answer is extremely small. One approach to addressing this issue is to "combine" the possible answers somehow to obtain a more compact representation of the result. For select-project-join queries, for instance, one proposed approach is to union all the possible answers, and compute the probability of each result tuple by adding the probabilities of all possible answers it belongs to [11]. This approach, however, can not be easily extended to other types of queries like ranking or aggregate queries.

Furthermore, from the user or application perspective, despite the probabilistic nature of the data, a single, deterministic query result would be desirable in most cases, on which further analysis or decision-making could be based. For SPJ queries, this is often achieved by "thresholding", i.e., returning only the result tuples with a sufficiently high probability of being true. For aggregate queries, often expected values are returned instead [24]. For ranking queries, on the other hand, a range of different approaches have been proposed to find the true ranking of the tuples. These include *UTop*-k, *URank*-k [37], probabilistic threshold *Top*-k function [22], Global *Top*-k [43], *expected rank* [9], and so on. Although these definitions seem to reason about the ranking over probabilistic databases in some "natural" ways, there is a lack of a unified and systematic analysis framework to justify their semantics and to discriminate the usefulness of one from another.

In this paper, we consider the problem of combining the results for all possible worlds in a systematic way by putting it in the context of *inconsistent information aggregation* which has been studied extensively in numerous

contexts over the last half century. In our context, the set of different query answers returned from possible worlds can be thought as inconsistent information which we need to aggregate to obtain a single representative answer. To the best of our knowledge, this connection between query processing in probabilistic databases and inconsistent information aggregation, though natural, has never been realized before in any formal and mathematical way. Concretely, we propose the notion of *the consensus answer*. Roughly speaking, the consensus answer is a answer that is *closest* to the answers of the possible worlds in expectation. To measure the closeness of two answers  $\tau_1$  and  $\tau_2$ , we have to define suitable distance function  $d(\tau_1, \tau_2)$  over the answer space. For example, if an answer is a vector, we can simply use the  $L_2$  norm; whereas in other cases, for instance, *Top*-k queries, the definition of d is more involved. If the most consensus answer can be taken from any point in the answer space, we refer it as the *mean answer*. A *median answer* is defined similarly except that the median answer must be the answer of some possible world with non-zero probability.

From a mathematical perspective, if the distance function is properly defined to reflect the closeness of the answers, the most consensus answer is perhaps the best deterministic representative of the set of all possible answers since it can be thought as the centroid of the set of points corresponding to the possible answers. Our key results can be summarized as follows:

- (Probabilistic And/Xor Tree) We propose a new model for modeling correlations, called the *probabilistic and/xor tree* model, that can capture two types of correlations, mutual exclusion and coexistence. This model generalizes the previous models such as x-tuples, and block-independent disjoint tuples model. More important, this model admits an elegant generating functions based framework for many types of probability computations.
- (Set Distance Metrics) We show that the mean and the median world can be found in polynomial time for the *symmetric difference* metric for and/xor tree model. For the Jaccard distance metric, we present a polynomial time algorithm to compute the mean and median world for tuple independent database.
- (*Top*-k ranking Queries) The problem of aggregating inconsistent rankings has been well-studied under the name of *rank aggregation* [14]. We develop polynomial time algorithms for computing mean and median *Top*-k answers under the symmetric difference metric, and the mean answers under *intersection metric* and *generalized Spearman's footrule distance* [16], for the and/xor tree model.
- (Groupby Aggregates) For group by count queries, we present a 4-approximation to the problem of finding a median answer (finding mean answers is trivial).
- (Consensus Clustering) We also consider the consensus clustering problem for the and/xor tree model and get a constant approximation by extending a previous result [2].

**Outline:** We begin with a discussion of the related work (Section 2). We then define the probabilistic and/xor tree model (Section 3), and present a generating functions-based method to do probability computations on them (Section 3.3). The bulk of our key results are presented in Sections 4 and 5 where we address the problem of finding consensus worlds for different set distance metrics and for *Top*-k ranking queries respectively. We then briefly discuss finding consensus worlds for group-by *count* aggregate queries and clustering queries in Section 6.

### 2 Related Work

There has been much work on managing probabilistic, uncertain, incomplete, and/or fuzzy data in database systems and this area has received renewed attention in the last few years (see e.g. [23, 5, 28, 19, 17, 7, 8, 11, 40, 18]). This work has spanned a range of issues from theoretical development of data models and data languages, to practical implementation issues such as indexing techniques. In terms of representation power, most of this work has either assumed independence between the tuples [17, 11], or has restricted the correlations that can be modeled [5, 28, 3, 34]. Several approaches for modeling complex correlations in probabilistic databases have also been proposed [35, 4, 36, 39].

For efficient query evaluation over probabilistic databases, one of the key results is the dichotomy of conjunctive query evaluation on tuple-independent probabilistic databases by Dalvi and Suciu [11, 12]. Briefly the result states that the complexity of evaluating a conjunctive query over tuple-independent probabilistic databases is either PTIME

or #P-complete. For the former case, Dalvi and Suciu [11] also present an algorithm to find what are called *safe query plans*, that permit correct *extensional* evaluation of the query. Unfortunately the problem of finding consensus answers appears to be much harder; this is because even if a query has a safe plan, the result tuples may still be arbitrarily correlated.

In recent years, there has also been much work on efficiently answering different types of queries over probabilistic databases. Soliman et al. [37] first considered the problem of ranking over probabilistic databases, and proposed two ranking functions to combine the tuple scores and probabilities. Yi et al. [41] presented improved algorithms for the same ranking functions. Zhang and Chomicki [43] presented a desiderata for ranking functions and propose Global *Top*-k queries. Ming Hua et al. [21, 22] recently presented a different ranking function called *Probabilistic threshold* Top-k *queries*. Finally, Cormode et al. [9] also present a semantics of ranking functions and a new ranking function called *expected rank*. In a recent work, we proposed a parameterized ranking function, and presented general algorithms for evaluating them [29] Other types of queries have also been recently considered over probabilistic databases (e.g. clustering [10], nearest neighbors [6] etc.).

The problem of aggregating inconsistent information from different sources arises in numerous disciplines and has been studied in different contexts over decades. Specifically, the RANK-AGGREGATION problem aims at combining k different complete ranked lists  $\tau_1, \ldots, \tau_k$  on the same set of objects into a single ranking, which is the best description of the combined preferences in the given lists. This problem was considered as early as 18th century when Condorcet and Borda proposed a voting system for elections [31, 25]. In the late 50's, Kemeny proposed the first mathematical criterion for choosing the best ranking [26]. Namely, the Kemeny optimal aggregation  $\tau$  is the ranking that minimizes  $\sum_{i=1}^{k} d(\tau, \tau_i)$ , where  $d(\tau_i, \tau_j)$  is the number of pairs of elements that are ranked in different order in  $\tau_i$  and  $\tau_j$  (also called Kendall's tau distance). While computing the Kemeny optimal is shown to be NP-hard [15], 2-approximation can be easily achieved by picking the best one from k given ranking lists. The other well-known 2-approximation is from the fact the Spearman footrule distance, defined to be  $d_F(\tau_i, \tau_j) = \sum_t |\tau_i(t) - \tau_j(t)|$ , is within twice the Kendall's tau distance and the footrule aggregation can be done optimally in polynomial time [14]. Ailon et al. [2] improve the approximation ratio to 4/3. We refer the readers to [27] for a survey on this problem. For aggregating *Top*-k answers, Ailon [1] recently obtained an 3/2-approximation based on rounding an LP solution.

The CONSENSUS-CLUSTERING problem asks for the best clustering of a set of elements which minimizes the number of pairwise disagreements with the given k clusterings. It is known to be NP-hard [42] and a 2-approximation can also be obtained by picking the best one from the given k clusterings. The best known approximation ratio is 4/3due to Ailon et al. [2]. Recently Cormode et al. [10] proposed approximation algorithms for k-center and k-median clustering problems under attribute-level uncertainty in probabilistic databases.

### **3** Preliminaries

We begin with reviewing the possible worlds semantics, and introduce the probabilistic and/xor tree model.

### **3.1** Possible World Semantics

We consider probabilistic databases with both tuple-level uncertainty (the existence of a tuple is uncertain) and attribute-level uncertainty (a tuple attribute value is uncertain). Specifically, we denote a probabilistic relation by  $R^P(K; A)$ , where K is the key attribute, and A is the value attribute<sup>1</sup>. For a particular tuple in  $R^P$ , its key attribute is certain and is sometimes called the possible worlds key.  $R^P$  is assumed to correspond to a probability space (PW, Pr) where the set of outcomes is a set of deterministic relations, which we call *possible worlds*,  $PW = \{pw_1, pw_2, ..., pw_N\}$ . Note that two tuples can not have the same value for the key attribute in a single possible world. Because of the typically exponential size of PW, an explicit possible worlds representation is not feasible, and hence the semantics are usually captured implicitly by probabilistic models with polynomial size specification.

Let T denote the set of tuples in all possible worlds. For ease of notation, we will use  $t \in pw$  in place of "t appears in the possible world pw", Pr(t) to denote Pr(t is present) and  $Pr(\neg t)$  to denote Pr(t is not present).

<sup>&</sup>lt;sup>1</sup>For clarity, we will assume singleton key and value attributes.



Figure 1: (i) The and/xor tree representation of a set of block-independent disjoint tuples; the generating function obtained by assigning the same variable x to all leaves gives us the distribution over the sizes of the possible worlds. (ii) Example of a highly correlated probabilistic database with 3 possible worlds and (iii) the and/xor tree that captures the correlation; the coefficient of y (0.3) is  $Pr(r(t_3, 6) = 1)$  (i.e., prob. that that alternative of  $t_3$  is ranked at position 1).

Further, for a tuple  $t^P \in R^P$ , we call the certain tuples corresponding to it (with the same key value) in the union of the possible worlds, its *alternatives*.

**Block-Independent Disjoint (BID) Scheme:** BID is one of the more popular models for probabilistic databases, and assumes that different probabilistic tuples (with different key values) are independent of each other [11, 40, 12, 38]. Formally, a BID scheme has the relational schema of the from R(K; A; Pr) where K is the possible worlds key, A is the value attribute, and Pr captures the probability of the corresponding tuple alternative.

### 3.2 Probabilistic And/Xor Tree

We generalize the block-independent disjoint tuples model, which can capture *mutual exclusion* between tuples, by adding support for *mutual co-existence*, and allowing these to be specified in a hierarchical manner. Two events satisfy the mutual co-existence correlation if in any possible world, either both happen or neither occurs. We model such correlations using a *probabilistic and/xor tree* (or and/xor tree for short), which also generalizes the notions of *x-tuples* [34, 41], *p*-or-sets [12] and tuple independent databases. We first considered this model for tuple-level uncertainty in an earlier paper [29], and generalize it here to handle attribute-level uncertainty.

We use  $\bigotimes$  (or) to denote mutual exclusion and  $\bigotimes$  (and) for coexistence. Figure 1 shows two examples of probabilistic and/xor trees. Briefly, the leaves of the tree correspond to the tuple alternatives (we abuse the notation somewhat and use  $t_i$  to denote both the tuple, and its key value). The first tree captures a relation with four independent tuples,  $t_1, t_2, t_3, t_4$ , each with two alternatives, whereas the second tree shows how we can capture arbitrary possible worlds using an and/xor tree (Figure 1(ii) shows the possible worlds corresponding to that tree).

Now, let us formally define a probabilistic and/xor tree. In tree  $\mathcal{T}$ , we denote the set of children of node v by  $Ch_{\mathcal{T}}(v)$  and the least common ancestor of two leaves  $l_1$  and  $l_2$  by  $LCA_{\mathcal{T}}(l_1, l_2)$ . We omit the subscript if the context is clear.

**Definition 1** A probabilistic and/xor tree  $\mathcal{T}$  represents the mutual exclusion and co-existence correlations in a probabilistic relation  $\mathbb{R}^{P}(K; A)$ , where K is the possible worlds key, and A is the value attribute. In  $\mathcal{T}$ , each leaf is a key-attribute pair (a tuple alternative), and each inner node has a mark,  $\bigcirc$  or  $\bigcirc$ . For each  $\bigcirc$  node u and each of its children  $v \in Ch(u)$ , there is a nonnegative value  $\Pr(u, v)$  associated with the edge (u, v). Moreover, we require

• (Probability Constraint)  $\sum_{v:v \in Ch(u)} \Pr(u, v) \leq 1$ .

• (Key Constraint) For any two different leaves  $l_1, l_2$  holding the same key,  $LCA(l_1, l_2)$  is a  $\bigcirc$  node<sup>2</sup>.

Let  $\mathcal{T}_v$  be the subtree rooted at v and  $Ch(v) = \{v_1, \ldots, v_\ell\}$ . The subtree  $\mathcal{T}_v$  inductively defines a random subset  $S_v$  of its leaves by the following independent process:

<sup>&</sup>lt;sup>2</sup> The key constraint is imposed to avoid two leaves with the same key coexisting in a possible world.

- If v is a leaf,  $S_v = \{v\}$ .
- If  $\mathcal{T}_v$  roots at a  $\bigotimes$  node, then
  - $S_{v} = \begin{cases} S_{v_{i}} & \text{with prob} \operatorname{Pr}(v, v_{i}) \\ \emptyset & \text{with prob} \ 1 \sum_{i=1}^{\ell} \operatorname{Pr}(v, v_{i}) \end{cases}$
- If  $\mathcal{T}_v$  roots at a  $\bigotimes$  node, then  $S_v = \bigcup_{i=1}^{\ell} S_{v_i}$

Probabilistic and/xor trees can capture more complicated correlations than the prior models such as the BID model or x-tuples. We remark that Markov or Bayesian network models are able to capture more general correlations [35], however, the structure of the model is more complex and probability computations on them (inference) is typically exponential in the treewidth of the model. The treewidth of an and/xor tree (viewing it as a Markov network) is not bounded, and hence the techniques developed for those models can not be used to obtain a polynomial time algorithms for and/xor trees.

### 3.3 Computing Probabilities on And/Xor Trees

Aside from the representational power of the and/xor tree model, perhaps its best feature is that many types of probability computations can be done efficiently and elegantly on them using *generating functions*. In our prior work [29], we used a similar technique for computing ranking functions for tuple-level uncertainty model. Here we generalize the idea to a broader range of probability computations.

We denote the and/xor tree by  $\mathcal{T}$ . Suppose  $\mathcal{X} = \{x_1, x_2, \ldots\}$  is a set of variables. Define a mapping s which associates each leaf  $l \in \mathcal{T}$  with a variable  $s(l) \in \mathcal{X}$ . Let  $\mathcal{T}_v$  denote the subtree rooted at v and let  $v_1, \ldots, v_l$  be v's children. For each node  $v \in \mathcal{T}$ , we define a generating function  $\mathcal{F}_v$  recursively:

- If v is a leaf,  $\mathcal{F}_v^i(\mathcal{X}) = s(v)$ .
- If v is a  $\bigotimes$  node,

$$\mathcal{F}_{v}(\mathcal{X}) = (1 - \sum_{h=1}^{l} p(v, v_{h})) + \sum_{h=1}^{l} \mathcal{F}_{v_{h}}(\mathcal{X}) \cdot p(v, v_{h})$$

• If v is a  $\bigotimes$  node,  $\mathcal{F}_v^i(\mathcal{X}) = \prod_{h=1}^l \mathcal{F}_{v_h}(\mathcal{X}).$ 

The generating function  $\mathcal{F}(\mathcal{X})$  for tree  $\mathcal{T}$  is the one defined above for the root. It is easy to see, if we have a constant number of variables, the polynomial can be expanded in the form of  $\sum_{i_1,i_2,\ldots} c_{i_1,i_2,\ldots} x_1^{i_1} x_2^{i_2} \ldots$  in polynomial time.

Now recall that each possible world pw contains a subset of the leaves of  $\mathcal{T}$  (as dictated by the  $\otimes$  and  $\otimes$  nodes). The following theorem characterizes the relationship between the coefficients of  $\mathcal{F}$  and the probabilities we are interested in.

**Theorem 1** The coefficient of the term  $\prod_j x_j^{i_j}$  in  $\mathcal{F}(\mathcal{X})$  is the total probability of the possible worlds for which, for all *j*, there are exactly  $i_j$  leaves associated with variable  $x_j$ .

The proof is by induction on the tree structure and is omitted.

**Example 1** If we associate all leaves with the same variable x, the coefficient of  $x^i$  is equal to  $\Pr(|pw| = i)$ .

The above can be used to obtain a distribution on the possible world sizes (Figure 1(i)).

**Example 2** If we associate a subset S of the leaves with variable x, and other leaves with constant 1, the coefficient of  $x^i$  is equal to  $\Pr(|pw \cap S| = i)$ .

**Example 3** Next we show how to compute Pr(r(t) = i) (i.e., the probability t is ranked at position i), where r(t) denote the rank of the tuple in a possible world by some score metric. Assume t only has one alternative, (t, a), and hence only one possible value of score, s. Then, in the and/xor tree T, we associate all leaves with key other than t and score value larger than s with variable x, and the leaf (t, a) with variable y, and the rest of leaves with constant 1. Then, the coefficient of  $x^{j-1}y$  in the generating function is exactly Pr(r(t) = i). If the tuple has multiple alternatives, we can compute Pr(r(t) = i) for it by summing up the probabilities for the alternatives.

See Figure 1(iii) for an example.

#### **3.4 Problem Definition**

We denote the domain of answers for a query by  $\Omega$  and the distance function between two answers by d(). Formally, we define the most consensus answer  $\tau$  to be a feasible answer to the query such that the expected distance between  $\tau$  and the answer  $\tau_{pw}$  of the (random) world pw is minimized, i.e,  $\tau = \arg \min_{\tau' \in \Omega} \{ \mathsf{E}[\mathsf{d}(\tau', \tau_{pw})] \}$ .

We call the most consensus answer in  $\Omega$  the mean answer when  $\Omega$  is the set of all feasible answers. If  $\Omega$  is restricted to be the set of possible answers (answers of some possible worlds with non-zero probability), we call the most consensus answer in  $\Omega$  the median answer. Taking the example of the *Top*-k queries, the median answer must be the *Top*-k answer of some possible world while the mean answer can be any sorted list of size k.

### 4 Set Distance Measures

We first consider the problem of finding the consensus world for a given probabilistic database, under two set distance measures: symmetric difference, and Jaccard distance.

#### 4.1 Symmetric Difference

The symmetric difference distance between two sets  $S_1$ ,  $S_2$  is defined to be  $d_{\Delta}(S_1, S_2) = |S_1 \Delta S_2| = |(S_1 \setminus S_2) \cup (S_2 \setminus S_1)|$ . Note that two different alternatives of a tuple are treated as different tuples here.

**Theorem 2** The mean world under the symmetric difference distance is the set of all tuples with probability > 0.5.

**Proof:** Suppose S is a fixed set of tuples and  $\bar{S} = T - S$ . Let  $\delta(p) = \begin{cases} 1, & \text{if } p = true \\ 0, & \text{if } p = false \end{cases}$  be the indicator function. We write  $E_{pw \in PW}[\mathsf{d}_{\Delta}(S, pw)]$  as follows:

$$\begin{split} \mathsf{E}[\mathsf{d}_{\Delta}(S, pw)] &= \mathsf{E}[\sum_{t \in S} \delta(t \notin pw) + \sum_{t \in \bar{S}} \delta(t \in pw)] \\ &= \sum_{t \in S} \mathsf{E}[\delta(t \notin pw)] + \sum_{t \in \bar{S}} \mathsf{E}[\delta(t \in pw)] = \sum_{t \in S} \mathsf{Pr}(\neg t) + \sum_{t \in \bar{S}} \mathsf{Pr}(t) \end{split}$$

Thus, each tuple t contributes  $Pr(\neg t)$  to the expected distance if  $t \in S$  and Pr(t) otherwise, and hence the minimum is achieved by the set of tuples with probability 0.5.

Finding the consensus median world is somewhat trickier, with the main concern being that the world that contains all tuples with probability > 0.5 may not be a possible world.

**Corollary 1** If the correlation can be modeled by a probabilistic and/xor tree, the median world is the set contains all tuples with probability greater than 0.5.

The proof is by induction on the height of the tree, and is omitted for space constraints. This however does not hold for arbitrary correlations, and it is easy to see that finding a median world is NP-Hard even if result tuple probability computation is easy. We show a reduction to MAX-2-SAT for a simple 2-relation query. Let the MAX-2-SAT instance consists of n literals,  $x_1, \ldots, x_n$ , and k clauses. Consider a query  $R \bowtie S$ , where S(x, b) = $\{(x_1, 0), (x_1, 1), (x_2, 0), (x_2, 1), \ldots\}$  contains two mutually exlusive tuples each for n literals; all tuples are equiprobable with probability 0.5. R(C, x, b) is a certain table, and contains two tuples for each clause: for the clause  $c_1 = x_1 \lor x_2$ , it contains tuples  $(c_1, x_1, 1)$  and  $(c_1, x_2, 0)$ . The result of  $\pi_C(R \bowtie S)$  contains one tuple for each clause, associated with a probability of 0.75. So the median answer is the possible answer containing maximum number of tuples, which corresponds to finding the assignment to  $x_i$ 's that maximizes the number of satisfied clauses.

### 4.2 Jaccard Distance

The Jaccard distance between two sets  $S_1, S_2$  is defined to be  $d_J(S_1, S_2) = \frac{|S_1 \Delta S_2|}{|S_1 \cup S_2|}$ . Jaccard distance always lies in [0, 1] and is a real metric, i.e, satisfies triangle inequality. Next we present polynomial time algorithms for finding the mean and median worlds for tuple independent databases, and median world for the BID model.

**Lemma 1** Given an and/xor tree,  $\mathcal{T}$  and a possible world for it, W (corresponding to a set of leaves of  $\mathcal{T}$ ), we can compute  $\mathsf{E}[\mathsf{d}(W, pw)]$  in polynomial time.

**Proof:** A generating function  $\mathcal{F}_{\mathcal{T}}$  is constructed with the variables associated with leaves as follows: for  $t \in W$   $(t \notin W)$ , the associated variable is x(y). For example, in a tuple independent database, the generating function is:

$$\mathcal{F}(x,y) = \prod_{t \in W} \left( \mathsf{Pr}(\neg t) + \mathsf{Pr}(t)x \right) \prod_{t \notin W} \left( \mathsf{Pr}(\neg t) + \mathsf{Pr}(t)y \right)$$

From Theorem 1, the coefficient  $c_{i,j}$  of term  $x^i y^j$  in generating function  $\mathcal{F}$  is equal to the total probability of the worlds such that the Jaccard distance between those worlds and W is exactly  $\frac{|W|-i+j}{|W|+j}$ . Thus, the distance is  $\sum_{i,j} c_{i,j} \frac{|W|-i+j}{|W|+j}$ .

**Lemma 2** For tuple independent databases, if the mean world contains tuple  $t_1$  but not tuple  $t_2$ , then  $Pr(t_1) \ge Pr(t_2)$ .

**Proof:** Say  $W_1$  is the mean world and the lemma is not true, i.e,  $\exists t_1 \in W_1, t_2 \notin W_1$  s.t.  $\Pr(t_1) < \Pr(t_2)$ . Let  $W = W_1 - \{t_1\}, W_2 = W + \{t_2\}$  and  $W' = T - W - \{t_1\} - \{t_2\}$ . We will prove  $W_2$  has a smaller expected Jaccard distance, thus rendering contradiction. Suppose  $|W_1| = |W_2| = k$ . We let matrix  $\mathbf{M} = [m_{i,j}]_{i,j}$  where  $m_{i,j} = \frac{k-i+j}{k+j}$ . We construct generating functions as we did in Lemma 1. Suppose  $\mathcal{F}_1$  and  $\mathcal{F}_2$  are the generating functions for  $W_1$  and  $W_2$ , respectively. We write  $||\mathbf{A}|| = \sum_{i,j} a_{i,j}$  for any matrix  $\mathbf{A}$  and let  $\mathbf{A} \otimes \mathbf{B}$  the Hadamard product of  $\mathbf{A}$  and  $\mathbf{B}$  (take product entrywise). We denote:

$$\begin{split} \mathcal{F}'(x,y) &= \prod_{t \in W} \left( \mathsf{Pr}(\neg t) + \mathsf{Pr}(t)x \right) \prod_{t \in W'} \left( \mathsf{Pr}(\neg t) + \mathsf{Pr}(t)y \right) \\ \text{We can easily see:} \\ \mathcal{F}_1(x,y) &= \mathcal{F}'(x,y) \left( \mathsf{Pr}(\neg t_1) + \mathsf{Pr}(t_1)x \right) \left( \mathsf{Pr}(\neg t_2) + \mathsf{Pr}(t_2)y \right) \end{split}$$

$$\mathcal{F}_{2}(x,y) = \mathcal{F}'(x,y) \left(\mathsf{Pr}(\neg t_{1}) + \mathsf{Pr}(t_{1})y\right) \left(\mathsf{Pr}(\neg t_{2}) + \mathsf{Pr}(t_{2})x\right)$$

Then, taking the difference, we get  $\overline{\mathcal{F}} = \mathcal{F}_1(x, y) - \mathcal{F}_2(x, y)$  is equal to:

$$\mathcal{F}'(x,y)\left(\mathsf{Pr}(\neg t_1)\mathsf{Pr}(t_2) - \mathsf{Pr}(t_1)\mathsf{Pr}(\neg t_2)\right)(y-x) \tag{1}$$

Let  $C_{\mathcal{F}} = [c_{i,j}]$  be the coefficient matrix of  $\mathcal{F}$  where  $c_{i,j}$  is the coefficient of term  $x^i y^j$ . Using the proof of Lemma 1:

$$\mathsf{E}[\mathsf{d}(W_1, pw)] - \mathsf{E}[\mathsf{d}(W_2, pw)] = ||\mathbf{C}_{\mathcal{F}_1} \otimes \mathbf{M}|| - ||\mathbf{C}_{\mathcal{F}_2} \otimes \mathbf{M}||$$
$$= ||\mathbf{C}_{\bar{\mathcal{F}}} \otimes \mathbf{M}||$$

Let  $c'_{i,j}$  and  $\bar{c}_{i,j}$  be the coefficient of  $x^i y_j$  in  $\mathcal{F}'$  and  $\bar{\mathcal{F}}$ , respectively. It is not hard to see  $\bar{c}_{i,j} = (c'_{i,j-1} - c'_{i-1,j})p$  from (1) where  $p = (\Pr(\neg t_1)\Pr(t_2) - \Pr(t_1)\Pr(\neg t_2)) > 0$ . Then we have

$$\begin{aligned} ||\mathbf{C}_{\bar{\mathcal{F}}} \otimes \mathbf{M}|| &= p \sum_{i,j} \left( (c'_{i,j-1} - c'_{i-1,j})m_{i,j} \right) \\ &= p \sum_{i,j} c'_{i,j} \left( m_{i,j+1} - m_{i+1,j} \right) \\ &= p \sum_{i,j} c'_{i,j} \left( \frac{k-i+j+1}{k+j+1} - \frac{k-i-1+j}{k+j} \right) \end{aligned}$$

Due to the fact that  $\frac{k-i+j+1}{k+j+1} - \frac{k-i-1+j}{k+j} > 0$  for any  $i, j \ge 0$ , the proof is complete.

The above two lemmas can be used to efficiently find the mean world for tuple-independent databases, by sorting the tuples in the decreasing order by probabilities, and computing the expected distance for every prefix of the sorted order.

A similar algorithm can be used to find the median world for the BID model (by only considering the highest probability alternative for each tuple). Finding mean worlds or median worlds under more general correlation models remains an open problem.

### 5 Top-k Queries

In this section, we consider *Top*-k queries in probabilistic databases. Each tuple  $t_i$  has a score  $s(t_i)$ . In the tuple-level uncertainty model,  $s(t_i)$  is fixed for each  $t_i$ , while in the attribute-level uncertainty model, it is an random variable. In the and/xor tree model, we assume that the attribute field is the score (uncertain attributes that don't contribute to the score can be ignored). We further assume no two tuples can take the same score for avoiding ties. We use r(t) to denote the random variable indicating the rank of t and  $r_{pw}(t)$  to denote the rank of t in possible world pw. If t does not appear in the possible world pw, then  $r_{pw}(t) = \infty$ . So,  $\Pr(r(t) > i)$  includes the probability that t's rank is larger than i and that t doesn't exist. We say  $t_1$  ranks higher than  $t_2$  in possible world pw if  $r_{pw}(t_1) < r_{pw}(t_2)$ .

Finally, we use the symbol  $\tau$  to denote rankings, and  $\tau^i$  to denote the restriction of the *Top*-k list  $\tau$  to the first *i* items. We use  $\tau(i)$  to denote the *i*<sup>th</sup> item in the list  $\tau$  for positive integer *i*, and  $\tau(t)$  to denote the position of  $t \in T$  in  $\tau$ .

### 5.1 Distance between Two Top-k Answers

Fagin et al. [16] provide a comprehensive analysis of the problem of comparing two *Top*-k lists. They present extensions of the Kendall's tau and Spearman footrule metrics (defined on full rankings) to *Top*-k lists and propose several other natural metrics, such as the intersection metric and Goodman and Kruskal's gamma function. In our paper, we consider three of the metrics discussed in that paper: the symmetric difference metric, the intersection metric and one particular extension to Spearman's footrule distance. We briefly recall some definitions here. For more details and the relation between different definitions, please refer to [16].

Given two *Top*-k lists,  $\tau_1$  and  $\tau_2$ , the normalized symmetric difference metric is defined as:

$$d_{\Delta}(\tau_1, \tau_2) = \frac{1}{2k} |\tau_1 \Delta \tau_2| = \frac{1}{2k} |(\tau_1 \setminus \tau_2) \cup (\tau_2 \setminus \tau_1)|.$$

While  $d_{\Delta}$  focuses only on the membership, the intersection metric  $d_I$  also takes the order of tuples into consideration. It is defined to be:

$$\mathsf{d}_I(\tau_1, \tau_2) = \frac{1}{k} \sum_{i=1}^k \mathsf{d}_\Delta(\tau_1^i, \tau_2^i)$$

Both  $d_{\Delta}$  and  $d_I()$  values are always between 0 and 1.

The original Spearman's Footrule metric is defined as the  $L_1$  distance between two permutations  $\sigma_1$  and  $\sigma_2$ . Formally,  $F(\sigma_1, \sigma_2) = \sum_{t \in T} |\sigma_1(t) - \sigma_2(t)|$ . Let  $\ell$  be a integer greater than k. The *footrule distance with location* parameter  $\ell$ , denoted  $F^{(\ell)}$  generalizes the original footrule metric. It is obtained by placing all missing elements in each list at position  $\ell$  and then computing the usual footrule distance between them. A natural choice of  $\ell$  is k + 1 and we denote  $F^{(\ell+1)}$  by  $d_F$ . It is also proven that  $d_F$  is a real metric and a member of a big and important equivalence class <sup>3</sup> [16].

It is shown in [16] that:

$$d_F(\tau_1, \tau_2) = (\mathbf{k} + 1)|\tau_1 \Delta \tau_2| + \sum_{t \in \tau_1 \cap \tau_2} |\tau_1(t) - \tau_2(t)| - \sum_{t \in \tau_1 \setminus \tau_2} \tau_1(t) - \sum_{t \in \tau_2 \setminus \tau_1} \tau_2(t).$$

Next we consider the problem of evaluating consensus answers for these distance metrics.

<sup>&</sup>lt;sup>3</sup> All distance functions in one equivalence class are bounded by each other within a constant factor. This class includes several extensions of Spearman's footrule and Kendall's tau metrics.

### **5.2** Symmetric Difference and PT - k function

In this section, we show how to find mean and median *Top*-k answers under symmetric difference metric in the and/xor tree model. The probabilistic threshold *Top*-k (PT - k) query [22] has been proposed for evaluating ranking queries over probabilistic databases, and essentially returns all tuples t for which  $Pr(r(t) \le k)$  is greater than a given threshold. If we set the threshold carefully so that the PT - k query returns k tuples, we can show that the answer returned is the mean answer under symmetric difference metric.

**Theorem 3** If  $\tau = \{\tau(1), \tau(2), \dots, \tau(k)\}$  is the set of k tuples with the largest  $\Pr(r(t) \leq k)$ , then  $\tau$  is the mean Top-k answer under metric  $d_{\Delta}$ , i.e., the answer minimizes  $\mathsf{E}[\mathsf{d}_{\Delta}(\tau, \tau_{pw})]$ .

**Proof:** Suppose  $\tau$  is fixed. We write  $\mathsf{E}[\mathsf{d}_{\Delta}(\tau, \tau_{pw})]$  as follows:  $\mathsf{E}[\mathsf{d}_{\Delta}(\tau, \tau_{pw})] = \mathsf{E}[\sum_{t \in T} \delta(t \in \tau \land t \notin \tau_{pw}) + \delta(t \in \tau_{pw} \land t \notin \tau)]$   $= \sum_{t \in T \setminus \tau} \mathsf{E}[\delta(t \in \tau_{pw})] + \sum_{t \in \tau} \mathsf{E}[\delta(t \notin \tau_{pw})]$   $= \sum_{t \in T \setminus \tau} \mathsf{Pr}(r(t) \le \mathsf{k}) + \sum_{t \in \tau} \mathsf{Pr}(r(t) > \mathsf{k})$   $= \mathsf{k} + \sum_{t \in T} \mathsf{Pr}(r(t) \le \mathsf{k}) - 2\sum_{t \in \tau} \mathsf{Pr}(r(t) \le \mathsf{k})$ 

The first two terms are invariant with respect to  $\tau$ . Therefore, it is clear that the set of k tuples with the largest  $\Pr(r(t) \leq k)$  minimizes the expectation.

To find a median answer, we essentially need to find the *Top*-k answer  $\tau$  of some possible world such that  $\sum_{t \in \tau} \Pr(r(t) \le k)$  is maximum. Next we show how to do this given an and/xor tree in polynomial time.

We write  $P(t) = \Pr(r(t) \le k)$  for ease of notation. We use dynamic programming over the tree structure. For each possible attribute value  $a \in A$ , let  $\mathcal{T}^a$  be the tree which contains all leaves with attribute value at least a. We recursively compute the set of tuples  $pw_{v,i}^a$ , which maximizes the value  $\sum_{t \in pw_{v,i}^a} P(t)$  among all possible worlds generated by the subtree  $\mathcal{T}_v^a$  rooted at v and of size i, for each node v in  $\mathcal{T}^a$  and  $1 \le i \le k$ . We compute this for all different a values, and the optimal solution can be chosen to be  $\min_a(pw_{r,k}^a)$ .

Suppose  $v_1, v_2, \ldots, v_l$  are v's children. The recursion formula is:

• If v is a  $\bigotimes$  node,  $pw^a_{v,i} = \arg \max_{pw \in PW(\mathcal{T}^a_{v_i})} \sum_{t \in pw} P(t)$ .

• If v is a  $\bigcirc$  node,  $pw_{v,i}^a = \bigcup_j pw_j$  such that  $\sum_j |pw_j| = i, pw_j \in PW(\mathcal{T}_{v_j}^a)$  and  $\sum_{t \in \bigcup_j pw_j} P(t)$  is maximized.

In the latter case, the maximum value can be computed by dynamic programming again as follows. Let  $pw_{[v_1,...v_h],i}^a = \bigcup_{j=1}^h pw_j$  such that  $\sum_{j=1}^h |pw_j| = i$ ,  $pw_j \in PW(\mathcal{T}_{v_j}^a)$  and  $\sum_{t \in \bigcup_{j=1}^h pw_j} P(t)$  is maximized. It can be computed recursive by seeing  $pw_{[v_1,...v_h],i}^a = pw_{[v_1,...v_{h-1}],p}^a \cup pw_{v_h,q}^a$  for p, q such that p + q = i and  $\sum_{t \in pw_{[v_1,...v_{h-1}],p}^a \cup pw_{v_h,q}^a} P(t)$  is maximized. Then, it is easy to see  $pw^a(v,i)$  is simply  $pw^a([v_1,...,v_l],i)$ .

**Theorem 4** The median Top-k answer under symmetric difference metric can be found in polynomial time for a probabilistic and/xor tree.

### 5.3 Intersection Metric

Note that the intersection metric  $d_I$  is a linear combination of the normalized asymmetric difference metric  $d_{\Delta}$ . Using a similar approach used in the proof of Theorem 3, we can show that:

$$\begin{split} \mathsf{E}[\mathsf{d}_{I}(\tau,\tau_{pw})] &= \frac{1}{\mathbf{k}} \sum_{i=1}^{\mathbf{k}} \mathsf{E}[\mathsf{d}_{\Delta}(\tau^{i},\tau_{pw}^{i})] \\ &= \frac{1}{\mathbf{k}} \sum_{i=1}^{\mathbf{k}} \frac{1}{i} \left( \mathbf{k} + \sum_{t \in T} \mathsf{Pr}(r(t) \le \mathbf{k}) - 2 \sum_{t \in \tau^{i}} \mathsf{Pr}(r(t) \le i) \right) \end{split}$$

Thus we need to find  $\tau$  which maximizes the last term,  $A(\tau) = \sum_{i=1}^{k} \left(\frac{1}{i} \sum_{t \in \tau^{i}} \Pr(r(t) \leq i)\right)$ . We first rewrite the objective as follows, using the indicator ( $\delta$ ) function:

$$\begin{split} A(\tau) &= \sum_{i=1}^{k} \left( \frac{1}{i} \sum_{t \in T} \Pr(r(t) \le i)) \delta(t \in \tau^{i}) \right) \\ &= \sum_{t \in T} \left( \sum_{i=1}^{k} \frac{1}{i} \Pr(r(t) \le i) \sum_{j=1}^{i} \delta(t = \tau(j)) \right) \\ &= \sum_{t \in T} \sum_{j=1}^{k} \left( \delta(t = \tau(j)) \sum_{i=j}^{k} \frac{1}{i} \Pr(r(t) \le i) \right) \end{split}$$

The last equality holds since  $\sum_{i=1}^{k} \sum_{j=1}^{i} a_{ij} = \sum_{j=1}^{k} \sum_{i=j}^{k} a_{ij}$ .

The optimization task can thus be written as an *assignment problem*, with each tuple t acting as an agent and each of the *Top*-k positions j as a task. Assigning task j to agent t gains a profit of  $\sum_{i=j}^{k} \frac{1}{i} \Pr(r(t) \le i)$  and the goal is to find an assignment such that each task is assigned to at most one agent, and the profit is maximized. The best known algorithm for computing the optimal assignment runs in  $O(nk\sqrt{n})$  time, via computing a maximum weight matching on bipartite graph [30].

Approximating the Intersection Metric: We define the following ranking function, where  $H_k$  denotes the  $k^{th}$  Harmonic number:

$$\Upsilon_H(t) = \sum_{i=1}^{k} (H_k - H_{i-1}) \Pr(r(t) = i) = \sum_{i=1}^{k} \frac{\Pr(r(t) \le i)}{i}.$$

This is a special case of the parameterized ranking function proposed in [29] and can be computed in  $O(nk \log^2 n)$  time for all tuples in the and/xor tree. We claim that the *Top*-k answer  $\tau_H$  returned by  $\Upsilon_H$  function, i.e., the k tuples with the highest  $\Upsilon_H$  values, is a good approximation of the mean answer with respect to the intersection metric by arguing that  $\tau_H = \{t_1, t_2, \ldots, t_k\}$  is actually an approximated maximizer of  $A(\tau)$ . Indeed, we prove the fact that  $A(\tau_H) \geq \frac{1}{H_k}A(\tau^*)$  where  $\tau^*$  is the optimal mean *Top*-k answer.

Let  $B(\tau) = \sum_{t \in \tau} \Upsilon_H(t)$  for any *Top*-k answer  $\tau$ . It is easy to see  $A(\tau^*) \leq B(\tau^*) \leq B(\tau_H)$  since  $\tau_H$  maximizes the B() function. Then, we can get:

$$A(\tau_{H}) = \sum_{j=1}^{k} \sum_{i=j}^{k} \frac{1}{i} \Pr(r(t_{j}) \le i)$$
  

$$\geq \sum_{j=1}^{k} \left(\frac{H_{k} - H_{j-1}}{H_{k}}\right) \sum_{i=1}^{k} \frac{1}{i} \Pr(r(t_{j}) \le i)$$
  

$$= \sum_{j=1}^{k} \left(\frac{H_{k} - H_{j-1}}{H_{k}}\right) \Upsilon_{H}(t_{j})$$
  

$$\geq \frac{1}{k} \sum_{i=1}^{k} \left(\frac{H_{k} - H_{i-1}}{H_{k}}\right) \sum_{i=1}^{k} \Upsilon_{H}(t_{i})$$
  

$$= \frac{1}{H_{k}} B(\tau_{H}) \ge \frac{1}{H_{k}} A(\tau^{*}).$$

The second inequality holds because for non-decreasing sequences  $a_i(1 \le i \le n)$  and  $c_i(1 \le i \le n)$ ,  $\sum_{i=1}^n a_i c_i \ge \frac{1}{n} (\sum_{i=1}^n a_i) (\sum_{i=1}^n c_i)$ 

$$\begin{split} \mathsf{E}[F^{*}(\tau,\tau_{pw})] &= \mathsf{E}\left[(\mathsf{k}+1)|\tau\Delta\tau_{pw}| + \sum_{t\in\tau\cap\tau_{pw}}|\tau(t) - \tau_{pw}(t)| - \sum_{t\in\tau\setminus\tau_{pw}}\tau(t) - \sum_{t\in\tau_{pw}\setminus\tau}\tau_{pw}(t)\right] \\ &= (\mathsf{k}+1)\mathsf{E}[|\tau\Delta\tau_{pw}|] + \sum_{t\in T}\mathsf{E}\left[\delta(t\in\tau\cap\tau_{pw})|\tau(t) - \tau_{pw}(t)|\right] - \sum_{t\in T}\mathsf{E}\left[\delta(t\in\tau\setminus\tau_{pw})\tau(t)\right] - \mathsf{E}\left[\sum_{t\in\tau_{pw}\setminus\tau}\tau_{pw}(t)\right] \\ &= (\mathsf{k}+1)\mathsf{E}[|\tau\Delta\tau_{pw}|] + \sum_{t\in T}\sum_{i=1}^{\mathsf{k}}\sum_{j=1}^{\mathsf{k}}\mathsf{E}\left[\delta(t\in\tau\cap\tau_{pw})\delta(t=\tau_{pw}(i))\delta(t=\tau(j))|i-j|\right] \\ &- \sum_{t\in T}\sum_{i=1}^{\mathsf{k}}\mathsf{E}\left[\delta(t\in\tau\setminus\tau_{pw})\delta(t=\tau(i))i\right] - \sum_{t\in T\setminus\tau}\Upsilon_{2}(t) \\ &= (\mathsf{k}+1)\mathsf{E}[|\tau\Delta\tau_{pw}|] + \sum_{t\in T}\sum_{i=1}^{\mathsf{k}}\left(\delta(t=\tau(i))\sum_{j=1}^{\mathsf{k}}\mathsf{Pr}(r(t)=j)|i-j|\right) - \sum_{t\in T}\sum_{i=1}^{\mathsf{k}}\left(\delta(t=\tau(i))i\mathsf{Pr}(r(t)>\mathsf{k})\right) - \sum_{t\in T\setminus\tau}\Upsilon_{2}(t) \\ &= (\mathsf{k}+1)(\mathsf{k}+\sum_{t\in T}\Upsilon_{1}(t) - 2\sum_{t\in\tau}\Upsilon_{1}(t)) + \sum_{t\in T}\sum_{i=1}^{\mathsf{k}}\delta(t=\tau(i))\Upsilon_{3}(t,i) - \sum_{t\in T\setminus\tau}\Upsilon_{2}(t) \\ &= (\mathsf{k}+1)\mathsf{k} + \sum_{t\in T}\left((\mathsf{k}+1)\Upsilon_{1}(t) - \Upsilon_{2}(t)\right) + \sum_{t\in T}\sum_{i=1}^{\mathsf{k}}\delta(t=\tau(i))(\Upsilon_{3}(t,i) + \Upsilon_{2}(t) - 2(\mathsf{k}+1)\Upsilon_{1}(t)) \end{split}$$

Figure 2: Derivation for Spearman's Footrule Distance

### 5.4 Spearman's Footrule

For a *Top*-k answer  $\tau = \{\tau(1), \tau(2), \dots, \tau(k)\}$ , we define:

- $\Upsilon_1(t) = \sum_{i=1}^k \Pr(r(t=i))$
- $\Upsilon_2(t) = \sum_{i=1}^k \Pr(r(t=i)) \cdot i$
- $\Upsilon_3(t,i) = \sum_{j=1}^k \Pr(r(t)=j)) |i-j| + i \Pr(r(t)>k).$

It is easy to see  $\Upsilon_1(t), \Upsilon_2(t), \Upsilon_3(t)$  can be computed in polynomial time for a probabilistic and/xor tree using our generating functions method.

A careful and non-trivial rewriting of  $E_{pw \in PW}[F^*(\tau, \tau_{pw})]$  shows that it also has the form (Figure 2):

$$E_{pw \in PW}[F^*(\tau, \tau_{pw})] = C + \sum_{t \in T} \sum_{i=1}^{k} \delta(t = \tau(i)) f(t, i)$$

where C is a constant independent of  $\tau$ , and f(t, i) is a function of t and i, which is polynomially computable. Figure 2 shows the exact derivation.

Thus, we only need to minimize the second term, which can be modeled as the assignment problem and can be solved in polynomial time.

### 5.5 Kendall's Tau Distance

Then *Kendall's tau* distance (also called Kemeny distance)  $d_K$  between two *Top*-k lists  $\tau_1$  and  $\tau_2$  is defined to be the number of unordered pairs  $(t_i, t_j)$  such that the order of *i* and *j* disagree in any full rankings extended from  $\tau_1$  and  $\tau_2$ , respectively. It is shown that  $d_F$  and  $d_K$  and a few other generalizations of Spearman's footrule and Kendall's tau metrics form a big equivalence class, i.e., they are within a constant factor of each other [16]. Therefore, the optimal solution for  $d_F$  implies constant approximations for all metrics in this class (the constant for  $d_K$  is 2).

However, we can also easily obtain a 3/2-approximation for d<sub>K</sub> by extending the 3/2-approximation for partial rank aggregation problem due to Ailon [1]. The only information used in their algorithm is the proportion of lists

where  $t_i$  is ranked higher than  $t_j$  for all i, j. In our case, this corresponds to  $Pr(r(t_i) < r(t_j))$ . This can be easily computed in polynomial time using the generating functions method.

We also note that the problem of optimally computing the mean answer is NP-hard for probabilistic and/xor trees. This follows from the fact that probabilistic and/xor trees can simulate arbitrary possible worlds, and previous work has shown that aggregating even 4 rankings under this distance metric is NP-Hard [14].

### **6** Other Types of Queries

We briefly extend the notion of consensus answers to two other types of queries and present some initial results.

#### 6.1 Aggregate Queries

Consider a query of the type:

select groupname, count(\*) from R group by groupname

Suppose there are *m* potential groups (indexed by groupname) and *n* independent tuples with attribute uncertainty. The probabilistic database can be specified by the matrix  $P = [p_{i,j}]_{n \times m}$  where  $p_{i,j}$  is the probability that tuple *i* takes groupname *j* and  $\sum_{j=1}^{m} p_{i,j} = 1$  for any  $1 \le i \le n$ . A query result (on a deterministic relation) is a *m*-dimensional vector **r** where the *i*<sup>th</sup> entry is the number of tuples having groupname *i*. The natural distance metric to use is the squared vector distance.

Computing the mean answer is easy in this case, because of linearity of expectation: we simply take the mean for each aggregate separately, i.e.,  $\bar{\mathbf{r}} = \mathbf{1}P$  where  $\mathbf{1} = (1, 1, ..., 1)$ . We note the mean answer minimizes the expected squared vector distance to any possible answer.

The median world requires that the returned answer be a possible answer. It is not clear how to solve this problem optimally in polynomial time. To enumerate all worlds is obviously not computationally feasible. Rounding entries of  $\bar{\mathbf{r}}$  to the nearest integers may not result in a possible answer.

Next we present a polynomial time algorithm to find a closest possible answer to the mean world  $\bar{\mathbf{r}}$ . This yields a 4-approximation for finding the median answer. We can model the problem as follows: Consider the bipartite graph B(U, V, E) where each node in U is a tuple, each node in V is a groupname, and an edge  $(u, v), u \in U, v \in V$  indicates that tuple u takes groupname v with non-zero probability. We call a subgraph G' such that  $deg_{G'}(u) = 1$  for all  $u \in U$  and  $deg_{G'}(v) = \mathbf{r}[v]$ , an **r**-matching of B for some m-dimensional integral vector **r**. Given this, our objective is to find an **r**-matching of B such that  $||\mathbf{r} - \bar{\mathbf{r}}||_2$  is minimized. Before presenting the main algorithm, we need the following lemma.

**Lemma 3** The possible world  $\mathbf{r}^*$  that is closest to  $\bar{\mathbf{r}}$  is of the following form:  $\mathbf{r}^*[i]$  is either  $\lfloor \bar{\mathbf{r}}[i] \rfloor$  or  $\lceil \bar{\mathbf{r}}[i] \rceil$  for each  $1 \leq i \leq m$ .

**Proof:** Let  $M^*$  be the corresponding  $\mathbf{r}^*$ -matching. Suppose the lemma is not true, and there exists i such that  $|\mathbf{r}^*[i] - \bar{\mathbf{r}}[i]| > 1$ . W.l.o.g, we assume  $\mathbf{r}^*[i] > \bar{\mathbf{r}}[i]$ . The other case can be proved the same way. Consider the connected component  $K = \{U', V', E(U', V')\}$  containing i. We claim that there exists  $j \in V'$  such that  $\mathbf{r}^*[j] < \bar{\mathbf{r}}[j]$  and there is an alternating path P with respect to  $M^*$  connecting i and j. Therefore,  $M' = M^* \oplus P$  is also a valid matching. Suppose M' is a  $\mathbf{r}'$ -matching. But:

 $||\mathbf{r}'|$ 

$$\begin{aligned} \mathbf{r}' - \bar{\mathbf{r}} ||_{2}^{2} &= \sum_{v=1}^{m} (\mathbf{r}'[v] - \bar{\mathbf{r}}[v])^{2} \\ &= \sum_{v=1}^{m} (\mathbf{r}^{*}[v] - \bar{\mathbf{r}}[v])^{2} - (\mathbf{r}^{*}[i] - \bar{\mathbf{r}}[i])^{2} - (\mathbf{r}^{*}[j] - \bar{\mathbf{r}}[j])^{2} + (\mathbf{r}'[j] - \bar{\mathbf{r}}[j])^{2} \\ &= ||\mathbf{r}^{*} - \bar{\mathbf{r}}||_{2}^{2} - (\mathbf{r}^{*}[i] - \bar{\mathbf{r}}[i])^{2} - (\mathbf{r}^{*}[j] - \bar{\mathbf{r}}[j])^{2} \\ &+ (\mathbf{r}^{*}[i] - 1 - \bar{\mathbf{r}}[i])^{2} + (\mathbf{r}^{*}[j] + 1 - \bar{\mathbf{r}}[j])^{2} \\ &= ||\mathbf{r}^{*} - \bar{\mathbf{r}}||_{2}^{2} + 2 - 2\mathbf{r}^{*}[i] + 2\bar{\mathbf{r}}[i] + 2\mathbf{r}^{*}[j] - 2\bar{\mathbf{r}}[j] \\ &< ||\mathbf{r}^{*} - \bar{\mathbf{r}}||_{2}^{2}. \end{aligned}$$

This contradicts the assumption  $r^*$  is the vector closest to  $\bar{r}$ .

Now, we prove the claim. We grow a alternating path (w.r.t.  $M^*$ ) tree rooted at i in a BFS manner: at odd depth, we extend all edges in  $M^*$  and at even depth, we extend all edge not in  $M^*$ . Let  $O \subseteq V$  be the set of nodes at odd depth (i is at depth 1) and  $E \subseteq U$  the set of nodes at even depth. It is easy to see  $N_B(E) = O$ ,  $E \subseteq N_B(O)$  and  $\sum_{v \in O} \mathbf{r}^*[v] = |E|$ . Suppose  $\mathbf{r}^*[v] \ge \bar{\mathbf{r}}[v]$  for all v and  $\mathbf{r}^*[i] \ge \bar{\mathbf{r}}[i]$ . However, the contradiction follows since:

$$|E| = \sum_{v \in O} \mathbf{r}^*[v] > \sum_{v \in O} \bar{\mathbf{r}}[v] = \sum_{v \in O} \sum_{u \in N_B(O)} \mathbf{P}[u, v]$$
$$= \sum_{v \in O} \sum_{u \in E} \mathbf{P}[u, v] = |E|.$$

With Lemma 3 at hand, we can construct the following min-cost network flow instance to compute the vector  $\mathbf{r}^*$  closest to  $\mathbf{\bar{r}}$ . Add to B a source s and a sink t. Add edges (s, u) with capacity upper bound 1 for all  $u \in U$ . For each  $v \in V$  and  $\mathbf{\bar{r}}[v]$  is not integer, add two edges  $e_1(v, t)$  and  $e_2(v, t)$ .  $e_1(v, t)$  has both lower and upper bound of capacity  $\lfloor \mathbf{\bar{r}}[v] \rfloor$  and  $e_2(v, t)$  has capacity upper bound 1 and cost  $(\lceil \mathbf{\bar{r}}[v] \rceil - \mathbf{\bar{r}}[v])^2 - (\lfloor \mathbf{\bar{r}}[v] \rfloor - \mathbf{\bar{r}}[v])^2$ . If  $\mathbf{\bar{r}}[v]$  is a integer, we only add  $e_1(v, t)$ . We find a min-cost integral flow of value n on this network. For any v such that  $e_2(v, t)$  is saturated, we set  $\mathbf{r}^*[v]$  to be  $\lceil \mathbf{\bar{r}} \rceil$  and  $\lfloor \mathbf{\bar{r}} \rfloor$  otherwise. Such a flow with minimum cost suggests the optimality of the vector  $\mathbf{r}^*$  due to Lemma 3.

**Theorem 5** There is a polynomial time algorithm for finding the vector  $\mathbf{r}^*$  to  $\bar{\mathbf{r}}$  such that  $\mathbf{r}^*$  corresponds to some possible answer with non-zero probability.

Finally, we can prove that:

**Corollary 2** There is a polynomial time deterministic 4-approximation for finding the median aggregate answer.

**Proof:** Suppose  $\mathbf{r}^*$  is the answer closest to the mean answer  $\bar{\mathbf{r}}$  and  $\mathbf{r}^m$  is the median answer. Let  $\mathbf{r}$  be the vector corresponding to the random answer. Then:

$$\begin{split} \mathsf{E}[\mathsf{d}(\mathbf{r}^*,\mathbf{r})] &\leq & \mathsf{E}[2(\mathsf{d}(\mathbf{r}^*,\bar{\mathbf{r}}) + \mathsf{d}(\bar{\mathbf{r}},\mathbf{r}))] = 2\left(\mathsf{d}(\mathbf{r}^*,\bar{\mathbf{r}}) + \mathsf{E}[\mathsf{d}(\bar{\mathbf{r}},\mathbf{r})]\right) \\ &\leq & 4\mathsf{E}[\mathsf{d}(\bar{\mathbf{r}},\mathbf{r})] \leq 4\mathsf{E}[\mathsf{d}(\mathbf{r}^m,\mathbf{r})]. \end{split}$$

#### 6.2 Clustering

The CONSENSUS-CLUSTERING problem is defined as follows: given k clusterings  $C_1, \ldots, C_k$  of V, find a clustering C that minimizes  $\sum_{i=1}^k d(C, C_i)$ . In the setting of probabilistic databases, the given clusterings are the clusterings in the possible worlds, weighted by the existence probability. The main problem with extending the notion of consensus answers to clustering is that the input clusterings are not well-defined (unlike ranking where the score function defines the ranking in any world). We consider a somewhat simplified version of the problem, where we assume that two tuples  $t_i$  and  $t_j$  are clustered together in a possible world, if and only if they take the same value for the value attribute A (which is uncertain). Thus, a possible world pw uniquely determines a clustering  $C_{pw}$ . We define the distance between two clustering  $C_1$  and  $C_2$  to be the number of unordered pairs of tuples that are clustered together in  $C_1$ , but separated in the other (the CONSENSUS-CLUSTERING metric). To deal with nonexistent keys in a possible world, we artifically create a cluster containing all of those.

Our task is to find a mean clustering C such that  $E[d(C, C_{pw})]$ . Approximation with factor of 4/3 is known for CONSENSUS-CLUSTERING [2], and can be adapted to our problem in a straightforward manner. In fact, that approximation algorithm simply needs  $w_{t_i,t_j}$  for all  $t_i, t_j$ , where  $w_{t_i,t_j}$  is the fraction of input clusters that cluster  $t_i$  and  $t_j$  together, and can be computed as:  $w_{t_i,t_j} = \sum_{a \in A} Pr(i.A = a \land j.A = a)$ .

To compute these quantities given an and/xor tree, we associate a variable x with all leaves with value (i, a) and (j, a), and constant 1 with the other leaves. From Theorem 1,  $\Pr(i.A = a \land j.A = a)$  is simply the coefficient of  $x^2$  in the corresponding generating function.

### 7 Conclusion

We addressed the problem of finding a single representative answer to a query over probabilistic databases by generalizing the notion of inconsistent information integration. We believe this approach provides a systematic and formal way to reason about the semantics of probabilistic query answers, especially for *Top*-k queries. Our initial work has opened up many interesting avenues for future work. These include design of efficient exact and approximate algorithms for finding consensus answers for other types of queries, exploring connections to safe plans, and understanding the semantics of the other previously proposed ranking functions using this framework.

### References

- [1] Nir Ailon. Aggregation of partial rankings, p-ratings and top-m lists. In SODA, pages 415-424, 2007.
- [2] Nir Ailon, Moses Charikar, and Alantha Newman. Aggregating inconsistent information: Ranking and clustering. In J.ACM, volume 55(5), 2008.
- [3] Periklis Andritsos, Ariel Fuxman, and Renee J. Miller. Clean answers over dirty databases. In ICDE, 2006.
- [4] Lyublena Antova, Christoph Koch, and Dan Olteanu. From complete to incomplete information and back. In SIGMOD, 2007.
- [5] B., H. Garcia-Molina, and D. Porter. The management of probabilistic data. IEEE TKDE, 1992.
- [6] George Beskales, Mohamed A. Soliman, and Ihab F. Ilyas. Efficient search for the top-k probable nearest neighbors in uncertain databases. In VLDB, 2008.
- [7] B. Buckles and F. E. Petry. A fuzzy model for relational databases. Intl. Journal of Fuzzy Sets and Syst., 1982.
- [8] Reynold Cheng, Dmitri Kalashnikov, and Sunil Prabhakar. Evaluating probabilistic queries over imprecise data. In SIGMOD, 2003.
- [9] Graham Cormode, Feifei Li, and Ke Yi. Semantics of ranking queries for probabilistic data and expected ranks. In ICDE, 2009.
- [10] Graham Cormode and Andrew McGregor. Approximation algorithms for clustering uncertain data. In PODS, 2008.
- [11] Nilesh Dalvi and Dan Suciu. Efficient query evaluation on probabilistic databases. In VLDB, 2004.
- [12] Nilesh Dalvi and Dan Suciu. Management of probabilistic data: Foundations and challenges. In PODS, 2007.
- [13] Amol Deshpande, Carlos Guestrin, Sam Madden, Joseph M. Hellerstein, and Wei Hong. Model-driven data acquisition in sensor networks. In VLDB, 2004.
- [14] C. Dwork, R. Kumar, M. Naor, and D. Sivakumar. Rank aggregation methods for the web. In *Proceedings of the Tenth International Conference on the World Wide Web (WWW)*, pages 613–622, 2001.
- [15] C. Dwork, R. Kumar, M. Naor, and D. Sivakumar. Rank aggregation revistied. In Manuscript, 2001.
- [16] Ronald Fagin, Ravi Kumar, and D. Sivakumar. Comparing top k lists. SIAM J. Discrete Mathematics, 17(1):134-160, 2003.
- [17] N. Fuhr and T. Rolleke. A probabilistic relational algebra for the integration of information retrieval and database systems. ACM Trans. on Info. Syst., 1997.
- [18] Minos Garofalakis and Dan Suciu, editors. IEEE Data Engineering Bulletin Special Issue on Probabilistic Data Management. March 2006.
- [19] Gosta Grahne. Horn tables an efficient tool for handling incomplete information in databases. In PODS, 1989.
- [20] Rahul Gupta and Sunita Sarawagi. Creating probabilistic databases from information extraction models. In VLDB, Seoul, Korea, 2006.
- [21] M. Hua, J. Pei, W. Zhang, and X. Lin. Efficiently answering probabilistic threshold top-k queries on uncertain data. In ICDE, 2008.
- [22] M. Hua, J. Pei, W. Zhang, and X. Lin. Ranking queries on uncertain data: A probabilistic threshold approach. In SIGMOD, 2008.
- [23] T. Imielinski and W. Lipski, Jr. Incomplete information in relational databases. Journal of the ACM, 1984.
- [24] T. S. Jayram, Andrew McGregor, S. Muthukrishnan, and Erik Vee. Estimating statistical aggregates on probabilistic data streams. In PODS, pages 243–252, 2007.
- [25] J.C.Borda. Mémoire sur les élections au scrutin. Histoire de l'Académie Royale des Sciences, 1781.
- [26] J.G.Kemeny. Mathematics without numbers. Daedalus, 88:571-591, 1959.
- [27] J.Hodge and R.E.Klima. The mathematics of voting and elections: a hands-on approach. AMS, 2000.
- [28] L. Lakshmanan, N. Leone, R. Ross, and V. S. Subrahmanian. Probview: a flexible probabilistic database system. ACM Trans. on DB Syst., 1997.
- [29] Jian Li, Barna Saha, and Amol Deshpande. Ranking and clustering in probabilistic databases. http://www.cs.umd.edu/~lijian/paper/clusterrank\_tr.pdf, 2008. Unpublished manuscript.
- [30] Silvio Micali and Vijay V. Vazirani. An o(sqrt(|v|)|e|) algorithm for finding maximum matching in general graphs. In FOCS '80: Proceedings of the 21th Annual Symposium on Foundations of Computer Science, pages 17–27, 1980.

- [31] M.J.Condorcet. Éssai sur l'application de l'analyse à la probabilité des décisions rendues à la pluralité des voix. 1785.
- [32] Christopher Re, Nilesh Dalvi, and Dan Suciu. Efficient top-k query evaluation on probabilistic data. In ICDE, 2007.
- [33] Christopher Re and Dan Suciu. Materialized views in probabilistic databases for information exchange and query optimization. In *VLDB*, Vienna, Austria, 2007.
- [34] A. Sarma, O. Benjelloun, A. Halevy, and J. Widom. Working models for uncertain data. In ICDE, 2006.
- [35] Prithviraj Sen and Amol Deshpande. Representing and querying correlated tuples in probabilistic databases. In ICDE, 2007.
- [36] Prithviraj Sen, Amol Deshpande, and Lise Getoor. Exploiting shared correlations in probabilistic databases. In VLDB, 2008.
- [37] M. Soliman, I. Ilyas, and K. C. Chang. Top-k query processing in uncertain databases. In ICDE, 2007.
- [38] Christopher Réand Dan Suciu. Efficient evaluation of having queries on a probabilistic database. In DBPL, 2007.
- [39] Daisy Zhe Wang, Eirinaios Michelakis, Minos Garofalakis, and Joseph M. Hellerstein. BayesStore: Managing large, uncertain data repositories with probabilistic graphical models. In VLDB, Auckland, New Zealand, 2008.
- [40] J. Widom. Trio: A system for integrated management of data, accuracy, and lineage. In CIDR, 2005.
- [41] Ke Yi, Feifei Li, Divesh Srivastava, and George Kollios. Efficient processing of top-k queries in uncertain databases. In ICDE, 2008.
- [42] Y.Wakabayashi. The complexity of computing medians of relations. In Resenhas, volume 3(3), pages 323–349, 1998.
- [43] Xi Zhang and Jan Chomicki. On the semantics and evaluation of top-k queries in probabilistic databases. In DBRank, 2008.