Minimal Frames and Transparent Frames for Risk, Time, and Uncertainty¹ (author names blinded for peer review)

Behavior differs between transparent and nontransparent presentations of decisions, but 'transparent presentation' has not been precisely defined. We formally define 'transparent frames' for risk and time, establish their uniqueness, provide algorithms for constructing them, and compare them to 'standard' presentation formats. A logic emerges for predicting systematic shifts in choice under risk and over time, and how violations of rational choice theory will depend on frames. An experiment verifies most of those predictions in choice under risk. We extend results to choice under uncertainty and also predict frame dependence of ambiguity aversion, a result supported by recent experimental evidence.

1. Introduction

Among challenges to the descriptive validity of the expected utility (or EU) hypothesis, none is more widely studied than Allais' (1953) common consequence effect, an example of which is shown below.

G1:	\$500,000	with probability 1	G2:	\$2,500,000 \$500,000 0	with probability 0.10 with probability 0.89 with probability 0.01
G3:	\$500,000 0	with probability 0.11 with probability 0.89	G4:	\$2,500,000	with probability 0.10 with probability 0.90

When choosing between gambles G1 and G2 and between G3 and G4, many people,² including prominent decision theorists like Leonard Savage, exhibit the choice pattern (G1, G4) in violation of EU.

Savage (1954, p. 102) suggests that people choose G1 over G2 because "they do not find the chance of winning a very large fortune in place of receiving a large fortune outright adequate compensation for even a small risk of being left at the status quo" and select G4 over G3 because "the chance of winning is nearly the same in both gambles, so the larger prize appears preferable." Savage then reframes the choice problems as a state-based matrix where states across choice alternatives are correlated (shown below).

	Т	icket Number	`s
	1	2 - 11	12 - 100
G1:	\$500,000	\$500,000	\$500,000
G2:	0	\$2,500,000	\$500,000
G3:	\$500,000	\$500,000	0
G4:	0	\$2,500,000	

In these reframed options, the prize depends on a ticket drawn from a bag containing 100 numbered tickets. Given this reframing, Savage chose in accordance with the independence axiom and in a

¹ We thank anonymous referees and an associate editor for many comments that improved this article.

² Common consequence effects are not always observed in some experiments that pay subjects according to their choices (using outcomes feasible for actual payment; see e.g. Conlisk 1989, Fan 2002, and Burke et al. 1996).

consistently risk averse manner, expressing a preference for G1 over G2 and for G3 over G4. Savage felt that his revision of the latter choice, brought about by this new frame, had "corrected an error" (p. 103).

Three decades later, Tversky and Kahneman (1986) presented further evidence that framing mediates decision errors. All 88 of their experimental subjects chose B over A in the choice below, in which probabilities are expressed as percentages of marbles of different colors in boxes governing each option:

Option B stochastically dominates A since it offers a 1% chance of a larger gain (\$45 versus \$30) and a 1% chance of a smaller loss (-\$10 versus -\$15). However, they observed that a majority (58%) of subjects chose C over D with the alternatives presented differently, as below:

One obtains Option C from A by combining the 1% blue and 2% yellow chances of losing \$15, and D from B by combining the 6% red and 1% green chances of winning \$45.³ This produces a 'minimal' presentation in that the matrix has fewer columns; yet this presentation masks the dominance relation and juxtaposes the 1% chance of \$30 against the 1% chance of -\$10, which now drives the choice.

Tversky and Kahneman (1986) propose that the dominance relationship is intuitively 'more transparent' in the choice between options A and B. Hogarth and Reder (1986) note, however, that "Tversky and Kahneman do not specify the conditions under which people perceive problems as transparent or opaque" (p. S192); and neither did Savage (1954). This issue has apparently not been addressed in the subsequent literature: there is no general theory of 'transparency' of choice presentations.

We offer precise definitions of *presentations* or *frames* for choice under risk and over time, propose a property list for *transparent* frames, and show that these properties imply unique presentations of choice problems. We also define a *minimal* frame and identify it with many standard presentations of choice problems. We then apply a model of salience-based choice to derive behavioral predictions in minimal and transparent frames, and test for these predicted choice differences in a new experiment involving choice under risk. Finally, we extend the model to choice under ambiguity, where we derive the novel prediction that ambiguity aversion is frame-dependent. In Section 8 we discuss how this work relates to

³ Birnbaum and Navarrete (1998) and Luce (1998) refer to this combining of the probabilities of identical outcomes as "coalescing." Earlier, Starmer and Sugden (1993) had called the opposite operation "event-splitting."

recent literature on perception's role in choice under risk and over time (Birnbaum 2004; Birnbaum et al. 2008; Bordalo et al., 2012; Loomes 2010; Scholten and Read, 2010, and many others).

2. Presentations or Frames

Figure 1. Presentations or Frames for Decisions under Risk and over Time

				Choice I	rame	for Lot	teries			
	(x_1,y_1)	(p_1,q_1)	$(\mathbf{x}_2,\mathbf{y}_2)$	(p_{2},q_{2})		$(\mathbf{x}_{i},\mathbf{y}_{i})$	$(\mathbf{p_{i}},\mathbf{q_{i}})$		$(\mathbf{x}_{\mathbf{n}},\mathbf{y}_{\mathbf{n}})$	$(\mathbf{p}_{\mathbf{n}},\mathbf{q}_{\mathbf{n}})$
p	\mathbf{x}_1	\mathbf{p}_1	\mathbf{X}_2	\mathbf{p}_2	•••	Xi	\mathbf{p}_{i}	•••	X _n	$\mathbf{p_n}$
q	$\mathbf{y_1}$	\mathbf{q}_1	\mathbf{y}_2	\mathbf{q}_2	•••	$\mathbf{y}_{\mathbf{i}}$	$\mathbf{q_i}$	•••	$\mathbf{y_n}$	$\mathbf{q_n}$
			Che	oice Fran	ne for	Income	Streams	3		
	(x_1,y_1)	$(\mathbf{r}_1,\mathbf{t}_1)$	$(\mathbf{x}_2,\mathbf{y}_2)$	$(\mathbf{r}_{2},\mathbf{t}_{2})$		$(\mathbf{x}_{i},\mathbf{y}_{i})$	$(\mathbf{r}_{i},\mathbf{t}_{i})$		$(\mathbf{x}_{\mathbf{n}},\mathbf{y}_{\mathbf{n}})$	$(\mathbf{r}_{\mathbf{n}},\mathbf{t}_{\mathbf{n}})$
r	\mathbf{x}_1	\mathbf{r}_1	\mathbf{x}_2	\mathbf{r}_2	•••	Xi	$\mathbf{r}_{\mathbf{i}}$	•••	X _n	$\mathbf{r}_{\mathbf{n}}$
t	$\mathbf{y_1}$	$\mathbf{t_1}$	\mathbf{y}_2	$\mathbf{t_2}$	•••	$\mathbf{y_i}$	t _i		$\mathbf{y_n}$	t _n

2.1 Lotteries, Income Streams and their Frames

Frames present pairs of options as shown in Figure 1. Let X be a finite set of potential outcomes. A lottery is a mapping $p: X \to [0,1]$ such that $\sum_{x \in X} p(x) = 1$, and $\Delta(X)$ is the set of all such lotteries. Consider a pair of one-dimensional finite arrays \mathbf{p} and \mathbf{q} , representing lotteries p and q and offering a finite and equal number of outcomes denoted \mathbf{x}_i and $\mathbf{y}_i, i = 1, 2, ..., n$, where each \mathbf{x}_i occurs with probability \mathbf{p}_i and each \mathbf{y}_i occurs with probability \mathbf{q}_i : The top panel of Figure 1 illustrates the pair of arrays. We call $[\![\mathbf{p}, \mathbf{q}]\!]$ a frame or presentation of lottery pair $\{p, q\}$, and say that $[\![\mathbf{p}, \mathbf{q}]\!]$ presents $\{p, q\}$, if and only if $p(x) = \sum_{\{i \mid \mathbf{x}_i = x\}} \mathbf{p}_i$ and $q(y) = \sum_{\{i \mid \mathbf{y}_i = y\}} \mathbf{q}_i$: In words, all array probabilities of the same outcome must sum to the outcome's probability in the lottery presented by that array. Let supp(p) be the set of outcomes such that p(x) > 0, called the support of $p(\cdot)$, and let |supp(p)| denote the number of outcomes in a support. Note that p(x) > 0, they present.

For decisions over discrete time periods $i \in \{0, 1, 2, ..., T\}$, intertemporal *income streams* $r := (x_0, x_1, ..., x_T)$ and $t := (y_0, y_1, ..., y_T)$ assign finite sequences of outcomes to each period; denote the set of income streams by C. We also study choices from pairs $\{r, t\}$ of income streams presented by frame $[\![\mathbf{r}, \mathbf{t}]\!]$ as illustrated in the lower panel of Figure 1: Finite arrays \mathbf{r} and \mathbf{t} present income streams \mathbf{r} and \mathbf{t} , offering finite and equal numbers of outcomes $\mathbf{x_i}$ occurring in time periods $\mathbf{r_i}$ and outcomes $\mathbf{y_i}$ occurring in time periods $\mathbf{t_i}$. Bold font always denotes attributes in frames, while italic font always denotes lotteries, income streams, and the attributes in their supports.

2.2 Defining Minimal Frames and Transparent Frames

We provide here an intuitive treatment of minimal and transparent frames (our appendix gives formal definitions and uniqueness demonstrations). *Transparent frames* resemble Savage's (1954) state matrix

lottery presentations, but remove Savage's explicitly correlated payoffs. More precisely, as a theoretical matter, a transparent frame *contains no correlational information*—whether the actual probability-generating mechanism for the pair of lotteries implies correlated payoffs or not (as it does in Savage's state matrix presentation).⁴

Minimal frames are compact presentations of choices and, for choice under risk, formalize the 'prospect' presentation of lotteries pioneered by Kahneman and Tversky (1979). Minimal frames for choice under risk contain no redundancy: Minimal frames for choice under risk contain minimal redundancy: the same outcome only appears more than once in special circumstances.⁵ if one of the alternatives is degenerate or where lotteries have different support sizes. Similarly, in minimal frames for choice over time, any time period appears just once in presentations of income streams, and the frame contains the fewest columns necessary to present all of the non-zero payoffs in the income streams.

Building on the intuitive examples of Savage (1954) and Tversky and Kahneman (1986), we propose that a transparent frame for lotteries satisfies five formal properties:

- **1. Common Consequence Separation:** Identify the common consequences (the payoff-probability pairs that contain the same outcomes and corresponding probabilities) of the lotteries being compared and separate them from distinct consequences (the other payoff-probability pairs in the frame).
- **2. Monotonicity:** Order the outcomes of distinct consequences in decreasing order such that the ith best outcome in one lottery is in the same column as the ith best outcome in the other lottery.
- **3. Alignment:** Set the probabilities within each column equal to each other.
- **4. Completeness:** Ensure that the probabilities for each row in the frame sum to 1.
- **5. Relevance:** Ensure the probabilities in each column vector are positive.

of transparent presentation (e.g. one that makes correlational information explicit).

Our appendix provides an algorithm for constructing a unique transparent frame satisfying these five properties for any lottery pair (including those with different numbers of outcomes in their support).

Together, these properties have intuitive appeal and simplify comparison of alternatives by articulating relevant information and bringing it into attentional focus. Common Consequence Separation

this and grant that a proponent of an alternative normative or rational model might want a different formal definition

⁴ We think any compelling account of transparent presentation requires a normative basis. Under EU theory, payoff correlation (or lack of it) between alternatives in a pair is normatively irrelevant with completely specified outcomes: Therefore our definition of transparent presentation proceeds on that basis, not providing correlational information between alternatives. This does not mean that correlational information would not be provided to the decision maker in some form (say to a subject in an experiment, where a complete description of how probabilities are implemented to generate payoffs automatically includes it, as is the case in our experiment here); it just means the presentation itself does not provide such correlational information. Although we think most decision researchers regard EU (or SEU) as normative, prominent dissents remain (e.g. Allais 1953 and Ellsberg 1961). We acknowledge

⁵ The special circumstances are: (i) When one lottery is degenerate (see Section 4 below); and more generally (ii) When lotteries have different support sizes, the lottery with smaller support size must have one outcome repeated in order to complete the frame (see Appendix Section A.3 below).

isolates shared payoff-probability pairs, thus focusing attention on how the lotteries differ. Monotonicity ensures that the best payoffs in one lottery are compared to the best payoffs in another lottery. Because Alignment matches the probability dimensions of all payoff-probability pairs across a pair of alternatives, one need only compare payoffs and weight them. Completeness ensures that all outcomes in the support of a lottery are accounted for in the decision; and Relevance ensures that no irrelevant outcomes (those with probability zero) are considered. We think that, taken together, these properties will help a person focus on the tradeoffs needed to make a quality decision.

For decisions over time, we propose similar properties that a transparent frame should satisfy:

- **1. Common Consequence Separation:** Identify the common consequences (the payoff-time pairs that contain the same outcomes and the same corresponding delays) of the income streams being compared and separate them from distinct consequences.
- **2. Monotonicity:** Order the timing of distinct consequences in strictly increasing order such that the ith soonest period in one income stream is in the same column as the ith soonest period in the other.
- **3. Alignment:** Set the time periods within each column equal to each other.
- **4. Completeness:** Ensure that all time periods indexed by the income streams are included.
- **5. Relevance:** Ensure that only time periods indexed by the income streams are included.

These five properties also have intuitive appeal and justification. Common consequence separation puts attention on where the two alternatives differ. Monotonicity reflects the intuition that time has a natural forward direction and it may help one to consider time periods sequentially. The alignment property ensures that the time periods within each column of a frame are the same, standardizing outcomes within each column to have the same 'time value of money'. Alignment also enables a decision maker who is comparing two payoff-time pairs in a given pair of columns to focus on the differences in payoffs, rather than trading off both risk and time within those columns. Completeness ensures that all relevant time periods and payoffs are considered. The relevance property ensures that only relevant time periods and payoffs are considered. In particular, it encourages the decision maker to be forward looking as it does not display sunk costs (e.g., from previous income outcomes) that occurred prior to the dates indexed by the income streams. We show in the appendix that for any pair of income streams there is a unique frame satisfying these five properties.

3. Salience Weighted Utility over Presentations

Leland and Schneider's (2017) Salience Weighted Utility over Presentations (SWUP) is a simple decision model that operates on frames as we define them here.⁶ We note, however, that the theory of framing as developed here is independent of SWUP. In particular, other decision models may be applied

⁶ Leland and Schneider (2017) focus entirely on developing SWUP, the salience-based decision model for frames, but do not develop the predictions of this model for different frames (nor test these) as we do here.

to minimal and transparent frames and behavior can be experimentally tested in minimal and transparent frames without necessarily invoking SWUP. In both of these respects, the concepts of minimal and transparent frames are more general than SWUP. In our analysis, SWUP is one plausible frame-dependent decision model that we use to illustrate the predicted differences between minimal and transparent frames.

3.1 Salience Weighted Utility over Presentations for Choice under Risk

In a standard EU model of choice under risk, p is chosen over q if and only if (1) holds:

(1)
$$\sum_{x \in X} p(x)u(x) > \sum_{y \in X} q(y)u(y),$$

where u(x) is a utility function denoting payoffs to the decision maker from outcomes x. Leland and Schneider (2017) consider choices over frames as in the top panel of Figure 1, where (1) can be written equivalently as (2) (recall that bold font denotes outcomes and probabilities in a frame):

(2)
$$\sum_{i=1}^{n} \mathbf{p_i} u(\mathbf{x_i}) > \sum_{i=1}^{n} \mathbf{q_i} u(\mathbf{y_i}),$$

Inequality (1) pertains to choices over lotteries p and q (regardless of how they are framed), whereas (2) pertains to choices over a particular frame $[\![\mathbf{p},\mathbf{q}]\!]$ presenting lotteries p and q. Note that (1) and (2) provide an *alternative-based evaluation* - one lottery is strictly preferred to another, if and only if it yields a greater expected payoff to the decision maker.

Building on Leland and Sileo (1998), the alternative-based evaluation in (2) may be rewritten equivalently as an *attribute-based evaluation* such that \mathbf{p} is chose over \mathbf{q} if and only if (3) holds:

(3)
$$\sum_{i=1}^{n} [(\mathbf{p_i} - \mathbf{q_i})(u(\mathbf{x_i}) + u(\mathbf{y_i}))/2 + (u(\mathbf{x_i}) - u(\mathbf{y_i}))(\mathbf{p_i} + \mathbf{q_i})/2] > 0.$$

Note that (2) and (3) operate over frames rather than over lotteries directly. Leland and Schneider (2017) then allow for the possibility that the agent systematically overweights salient differences in probabilities and payoffs by introducing salience weights $\phi(\mathbf{p_i}, \mathbf{q_i})$ on probability differences and $\mu(\mathbf{x_i}, \mathbf{y_i})$ on payoff differences. This yields Leland and Schneider's Salience-Weighted Utility over Presentations (SWUP) model of choice under risk, in which \mathbf{p} is chosen over \mathbf{q} if and only if

(4)
$$\sum_{i=1}^{n} [\phi(\mathbf{p_i}, \mathbf{q_i})(\mathbf{p_i} - \mathbf{q_i})(u(\mathbf{x_i}) + u(\mathbf{y_i}))/2 + \mu(\mathbf{x_i}, \mathbf{y_i})(u(\mathbf{x_i}) - u(\mathbf{y_i}))(\mathbf{p_i} + \mathbf{q_i})/2] > 0.$$

3.2 Salience Weighted Utility for Choice over Time

One may also derive a SWUP model for framed choice over time (as in the lower panel of Figure 1). In the discounted utility (or DU) model, a person chooses income stream a over b if and only if (5) holds:

(5)
$$\sum_{i=0}^{T} \delta^{i} u(x_i) > \sum_{i=0}^{T} \delta^{i} u(y_i),$$

where δ is a constant discount factor. Via steps resembling (2), (3), and (4), Leland and Schneider (2017) derive (6) from (5), generalizing DU theory to allow for overweighting of salient differences in payoffs

⁷ All of our theoretical results hold when u(x) = x. However we leave open the possibility that u(x) is weakly concave for gains (as in standard risk-averse EU), or additionally weakly convex for losses and exhibiting loss aversion (as in Wakker and Tversky's 1993 nonstandard Sign-Dependent Expected Utility or SDEU).

and time delays. Placing salience weights $\theta(\mathbf{r_i}, \mathbf{t_i})$ on time differences and $\mu(\mathbf{x_i}, \mathbf{y_i})$ on payoff differences gives this salience-weighted evaluation in which \mathbf{r} is always chosen over \mathbf{t} if and only if

(6)
$$\sum_{i}^{m} [\theta(\mathbf{r_i}, \mathbf{t_i})(\delta^{\mathbf{r_i}} - \delta^{\mathbf{t_i}})(u(\mathbf{x_i}) + u(\mathbf{y_i}))/2 + \mu(\mathbf{x_i}, \mathbf{y_i})(u(\mathbf{x_i}) - u(\mathbf{y_i}))(\delta^{\mathbf{r_i}} + \delta^{\mathbf{t_i}})/2] > 0,$$
 whenever the frame $[\![\mathbf{r}, \mathbf{t}]\!]$ presents income streams r and t .

3.3 Salience Weighted Utility for Choice under Ambiguity

Leland and Schneider (2017) introduced SWUP for choices under risk and over time. Here we extend SWUP to the domain of uncertainty, where probabilities of some events are unknown. Suppose there is a finite set of possible states of nature $s \in \{1,2,...,S\}$, where a lottery is assigned to be played in each state. Denote uncertain prospects by f and g, where f assigns lottery f(s) to each state s and g assigns lottery g(s) to each state s. In the classic alternative-based evaluation, there is assumed to be a unique subjective probability distribution, π_s , over states (Anscombe and Aumann, 1963) such that f is preferred over g if and only if (7) holds (where $f_s(x)$ is the probability of outcome x in state s):

(7)
$$\sum_{s \in S} \sum_{x \in X} \pi_s[f_s(x)u(x)] > \sum_{s \in S} \sum_{y \in X} \pi_s[g_s(y)u(y)].$$

Let there be a frame for each state, where *is* indexes the ith attribute in the state *s* frame. Given two multidimensional arrays $\mathbf{f} = \{\mathbf{f^1}, ..., \mathbf{f^S}\}$ and $\mathbf{g} = \{\mathbf{g^1}, ..., \mathbf{g^S}\}$, we define a formula analogous to (5). \mathbf{f} is always chosen over \mathbf{g} if and only if (8) holds:

(8)
$$\sum_{s}^{s} \sum_{i}^{n(s)} \pi_{s}[\mathbf{f_{is}}u(\mathbf{x_{is}})] > \sum_{s}^{s} \sum_{i}^{n(s)} \pi_{s}[\mathbf{g_{is}}u(\mathbf{y_{is}})].$$

Next, we introduce the corresponding model of salience-weighted evaluation in which \mathbf{f} is always chosen over \mathbf{g} if and only if (9) holds:

(9)
$$\sum_{s}^{S} \sum_{i=1}^{n} \pi_{s} [\phi(\mathbf{f_{is}}, \mathbf{g_{is}})(\mathbf{f_{is}} - \mathbf{g_{is}})(u(\mathbf{x_{is}}) + u(\mathbf{y_{is}}))/2 + \mu(\mathbf{x_{is}}, \mathbf{y_{is}})(u(\mathbf{x_{is}}) - u(\mathbf{y_{is}}))(\mathbf{f_{is}} + \mathbf{g_{is}})/2] > 0.$$

We refer to agents who choose according to salience-based evaluation models (the representations 4, 6, and 9) as *focal thinkers* since such agents focus on salient differences in attributes. Such an agent chooses the alternative which 'looks better' with respect to that agent's perceptual system.

3.4 Properties of Salience Perceptions

The salience functions μ , ϕ and θ alone determine how the behavior of a focal thinker differs from a rational agent (who chooses according to formulas 1, 5 and 7). We assume that salience functions exhibit the two properties of the perceptual system in Definition 1, introduced by Bordalo et al. (2012).

Definition 1 (Salience Function): A salience function $\sigma(\mathbf{a_i}, \mathbf{b_i})$ is any continuous, non-negative and symmetric ($\sigma(\mathbf{a_i}, \mathbf{b_i}) = \sigma(\mathbf{b_i}, \mathbf{a_i})$) function of $\mathbf{a_i}$ and $\mathbf{b_i} \in \mathbb{R}$ that satisfies the following two properties:

- **1. Ordering:** If $[a_i', b_i'] \subset [a_i, b_i]$ then $\sigma(a_i', b_i') < \sigma(a_i, b_i)$.
- **2. Diminishing Absolute Sensitivity (DAS):** $\sigma(.)$ exhibits diminishing sensitivity if for any $\mathbf{a_i}$, $\mathbf{b_i}$, $\epsilon > 0$, $\sigma(\mathbf{a_i} + \epsilon, \mathbf{b_i} + \epsilon) < \sigma(\mathbf{a_i}, \mathbf{b_i})$.

In addition to properties 1 and 2 from Bordalo et al. (2012), we will allow for the possibility that a salience function satisfies a third property, increasing proportional sensitivity or IPS:

3. Increasing Proportional Sensitivity (IPS): $\sigma(.)$ exhibits increasing proportional sensitivity if for any $\mathbf{a_i} > 0$, $\mathbf{b_i} \ge 0$ and any $\alpha > 1$, $\sigma(\alpha \mathbf{a_i}, \alpha \mathbf{b_i}) > \sigma(\mathbf{a_i}, \mathbf{b_i})$.

There is a close relationship between DAS and IPS: DAS implies that for a fixed absolute difference, the perceptual system is more sensitive to larger ratios, while IPS implies that for a fixed ratio, the perceptual system is more sensitive to larger absolute differences. We will show that for focal thinkers, IPS for ϕ implies the general version of the Allais common ratio effect, and DAS for ϕ implies ambiguity aversion in Ellsberg's paradox in minimal frame presentations. Our approach thus provides a unified treatment of Allais-style and Ellsberg-style behavior and shows that they can be derived from basic properties of the probability salience function ϕ without requiring any parametric assumptions about the underlying salience functions or utility functions.

3.5 Computable Parameter-Free Salience Functions for Calculation and Estimation

Computable parameter-free salience functions are useful for both applied theory calculations and statistical estimation of salience-based models such as SWUP. One alternative is the parameter-free DAS salience function (Bordalo et al., 2013):

(10)
$$\sigma(\mathbf{a_i}, \mathbf{b_i}) = \frac{|\mathbf{a_i} - \mathbf{b_i}|}{|\mathbf{a_i}| + |\mathbf{b_i}|} \text{ whenever } \mathbf{a_i} > 0 \text{ or } \mathbf{b_i} > 0; \text{ and } \sigma(0,0) = 0.$$

Although simple and parameter-free, (10) does not satisfy IPS. For our estimations in Section 5.3 below, we prefer a salience function that satisfies IPS. Given behaviors such as the fourfold pattern of risk attitudes that follow from IPS (Leland and Schneider, 2017), along with the general form of the Allais common ratio effect, it seems desirable for a salience function to satisfy that property.

Therefore we introduce a new computable and parameter-free salience function satisfying both DAS and IPS. To motivate this, consider the two frames in Figure 2. Here, the salience function in (10) assigns

 $x_{4,}y_{4}$ x_1, y_1 p_{1},q_{1} x_2, y_2 p_2,q_2 x_3, y_3 p_3,q_3 p_{4},q_{4} 200 0 300 0.25 0.25 100 0.25 0.25 p 100 0.25 100 0.25 100 0.25 100 0.25 p_{1},q_{1} x_{2}, y_{2} p_2,q_2 x_{3}, y_{3} p_3,q_3 $x_{4,}y_{4} \\$ p_4, q_4 300 0.25 2000 0.25 1000 0.25 0 0.25 p 100 0.25 1000 0.25 1000 0.25 1000 0.25

Figure 2. Context-dependence and Scale-dependence of Salience Perception

the same salience value $\sigma(300,100) = 0.5$ to the comparison between 300 and 100 in both frames. But perhaps that comparison is more salient in the top frame in Figure 2 (where other payoffs are in hundreds of dollars, so that the \$200 difference seems larger in the top context) than in the bottom frame (where other payoffs are in thousands of dollars, so that the \$200 difference seems smaller in the bottom context).

The context-dependence and scale-dependence of salience perception illustrated in Figure 2 can be accommodated by including a function that depends on all outcomes (or all probabilities) in a frame. One plausible candidate is the Euclidean norm which can be viewed as taking the second moment of value deviations from 0 for all outcomes (or all probabilities) in a frame. Let (\mathbf{a}, \mathbf{b}) denote a vector of length 2n, formed by horizontally concatenating a pair of like dimension vectors in a frame (i.e., all outcomes in a frame, or all probabilities in a frame); and let $\|\mathbf{a}, \mathbf{b}\|$ denote the Euclidean norm of the vector (\mathbf{a}, \mathbf{b}) . A context-dependent parameter-free DAS-IPS salience function can be defined as in (11):

(11)
$$\sigma(\mathbf{a_i}, \mathbf{b_i} | \mathbf{a}, \mathbf{b}) = \frac{|\mathbf{a_i} - \mathbf{b_i}|}{|\mathbf{a_i}| + |\mathbf{b_i}| + |\mathbf{a_i}\mathbf{b}|}.$$

This satisfies DAS and IPS for any frame, and under (11) the salience of (300, 100) is greater in the top frame of Figure 2 than in the bottom frame. We will use this *DAS-IPS* salience function for SWUP estimations in Section 5.3 below.

4. Minimal and Transparent Frames for Choice under Risk

We now consider what SWUP implies when alternatives are presented in a minimal versus a transparent frame, beginning with choices under risk. For frames with degenerate lotteries (those yielding a single outcome with probability 1), it seems almost unavoidable that one compares each outcome in the non-degenerate lottery \mathbf{q} to the unique outcome in the degenerate lottery \mathbf{p} . So when one lottery in a pair is degenerate, we adopt the convention that monotone minimal frames appear as shown in Figure 3.

Figure 3. Choice Frame with a Degenerate Lottery p

	(x_1, y_1)	$(\mathbf{p_1}, \mathbf{q_1})$	(x_2, y_2)	$(\mathbf{p}_2, \mathbf{q}_2)$	•••	$(\mathbf{x_i}, \mathbf{y_i})$	$(\mathbf{p_i}, \mathbf{q_i})$	•••	$(\mathbf{x}_{\mathbf{n}}, \mathbf{y}_{\mathbf{n}})$	(p_n, q_n)
p	X	$\mathbf{q_1}$	X	$\mathbf{q_2}$	•••	X	$\mathbf{q_i}$	•••	X	$\mathbf{q_n}$
q	y ₁	q_1	y_2	\mathbf{q}_2	•••	y _i	$\mathbf{q_i}$	•••	y _n	$\mathbf{q}_{\mathbf{n}}$

4.1 The Stochastic Dominance Axiom in Minimal and Transparent Frames

If lottery p offers at least as good an outcome as q does at every probability increment—and a strictly better outcome at some probability increment—then p stochastically dominates q. Rational choice requires consistency with stochastic dominance. Consider the example from Section 1, shown in Figure 4.

Figure 4. Stochastic Dominance in Minimal and Transparent Frames

		Stochast	ic Domir	nance in	the Mini	mal Fra	me				
x_1, y_1 p_1, q_1 x_2, y_2 p_2, q_2 x_3, y_3 p_3, q_3 x_4, y_4 p_4, q_4											
	p' 0	0.90	45	0.07	-10	0.01	-15	0.02			
q' 0 0.90 45 0.06 30 0.01 -15 0.03											
V . V		ochastic			- -			V - V -	n - 0		
$\mathbf{p} = \frac{x_{1}, y_{1}}{45}$		-10	p_{2},q_{2} 0.01	45	0.06	$x_{4,}y_{4}$	0.90	-15	p _{5,} q 0.0		
-		-		+			1		0.0		
q 30	0.01	-15	0.01	45	0.06	0	0.90	-15			

All subjects chose the stochastically dominant lottery p in the transparent frame, but many chose the dominated lottery q' in the minimal frame. For transparent frames, a focal thinker satisfies stochastic dominance since any differences in (4) favor the stochastically dominant lottery.

4.2 The Independence Axiom in Minimal and Transparent Frames

The Allais common consequence paradox (Allais, 1953) involves choices like those in the left panel of Figure 5. A decision maker chooses between lottery q, offering \$2400 with certainty and a lottery p, offering a 33% chance of \$2500, a 66% chance of \$2400, and a 1% chance of \$0. The decision maker then chooses between lottery \tilde{q} offering a 34% chance of \$2400 and lottery \tilde{p} offering a 33% chance of \$2500. Kahneman and Tversky (1979) report that most subjects choose q over p and choose \tilde{p} over \tilde{q} . This preference pattern is inconsistent with EU which predicts preferences of either p and \tilde{p} or q and \tilde{q} .

In the minimal frame [p, q], the comparison of 2400 and 0 is more salient than that of 2500 and 2400, favoring q. However, in the minimal frame $[\tilde{p}, \tilde{q}]$, the comparison between 2400 and 0 is not cued. Instead, the decision maker compares the upside of winning 2500 instead of 2400 with the downside of forfeiting a 1% chance in the probability of winning. To the extent this \$100 difference is more salient

Figure 5. The Allais Paradox in Minimal and Transparent Frames

Allais Paradox in Minimal Frames

Allais Paradox in Transpare

	Allais	Parado	ox in M	linimal	Frame	es
	x_1,y_1	p_{1},q_{1}	x_{2}, y_{2}	p_{2}, q_{2}	x_{3}, y_{3}	p_3,q_3
p	2500	0.33	2400	0.66	0	0.01
l	2400	0.33	2400	0.66	2400	0.01
	V. V.	n. a.	Ya Va	n _a (l _a		
~	x_1, y_1	p_{1},q_{1}	x_2, y_2	p_2,q_2	1	
p	2500	0.33	0	0.67		
ñ	2400	0.34	0	0.66		

	Allais	Parad	ox in T	ranspa	rent Fi	rames
	$x_{1,}y_{1}$	p_{1},q_{1}	x_2, y_2	p_2,q_2	x_{3}, y_{3}	p_{3}, q_{3}
p	2500	0.33	0	0.01	2400	0.66
\mathbf{q}	2400	0.33	2400	0.01	2400	0.66
	x_1, y_1	p_1,q_1	x_2, y_2	p_{2},q_{2}	x_3,y_3	p_{3}, q_{3}
p	2500	0.33	0	0.01	0	0.66
q	2400	0.33	2400	0.01	0	0.66
-1						

than the 0.01 difference in probabilities, the decision maker chooses \tilde{p} over \tilde{q} . Now consider the transparent frames in Figure 5. Here, the components common to each decision (i.e., (\$2400, 0.66) in the choice between p and q and (\$0, 0.66) in the choice between \tilde{p} and \tilde{q}) are isolated and the decision in both cases depends on comparisons between 2500 and 2400 and between 2400 and 0.

Figure 6 displays a version of Allais' (1953) common ratio effect—another well-known EU violation. The minimal frames display a choice between lotteries p and q, offering an 80% chance of \$4000 versus \$3000 with certainty, and a choice between \tilde{p} and \tilde{q} , offering a 20% chance of \$4000 versus a 25% chance of \$3000. In this version, Kahneman and Tversky (1979) report that a majority of subjects chose q over p and chose \tilde{p} over \tilde{q} when the choices were presented in minimal frames. This response pattern violates EU which predicts choices of either p and \tilde{p} or q and \tilde{q} . In the transparent frames in Figure 6, the salient comparisons in both choices are between 3000 and 0 and between 4000 and 3000. Since the same comparisons are focal in both choices, one might predict more consistent behavior in transparent frames.

Figure 6. The Common Ratio Effect Presented in Minimal and Transparent Frames

	In M	inimal	Frame	es]	In Trai	nsparer	nt Fran	nes		
р q	x ₁ ,y ₁ 4000 3000	p ₁ ,q ₁ 0.80 0.80	0 3000	p ₂ ,q ₂ 0.20 0.20	p	x ₁ ,y ₁ 4000 3000	0.80 0.80	x ₂ ,y ₂ 0 3000	p ₂ ,q ₂ 0.20 0.20			
$\widetilde{\mathbf{p}}$ $\widetilde{\mathfrak{q}}$	4000 3000	p ₁ ,q ₁ 0.20 0.25	x ₂ ,y ₂ 0 0	0.80 0.75	$\widetilde{f q}$	x ₁ ,y ₁ 4000 3000	p ₁ ,q ₁ 0.20 0.20	x ₂ ,y ₂ 0 3000	p ₂ ,q ₂ 0.05 0.05	x ₃ ,y ₃ 0 0	p ₃ ,q ₃ 0.75 0.75	l

5. An Experiment on Framing and Decisions under Risk

Our experiment compares the predictions of three models – the leading normative model of decision making (EU), cumulative prospect theory (CPT) due to Tversky and Kahneman (1992), and the model of salience weighted utility over presentations (SWUP). Outcomes in our experiment involve only gains, so the predictions of CPT coincide with those of Quiggin's (1982) rank-dependent utility (or RDU). We use SWUP instead of the salience-based model in Bordalo et al. (2012) since our focus is on framing effects. The model in Bordalo et al. (2012) does not predict framing effects between minimal and transparent frames, but instead predicts choices are sensitive to correlations between lotteries.

Table 1 shows predictions made by EU, CPT/RDU, and SWUP. Under EU, choices satisfy stochastic dominance, independence and the property called frame invariance (different presentations of the same two lotteries will produce the same observed choices). While CPT/RDU permit common ratio and Allais' Paradox (common consequence) violations of EU, they rule out dominance violations and also imply frame invariance. SWUP predicts that none of the regularities are observed in transparent frames, but that all of them can be observed in minimal frames (so SWUP does not predict frame invariance either). For instance, even with linear utility, SWUP explains the classical versions of the common consequence effect and common ratio effect in Kahneman and Tversky (1979) and the violation of stochastic dominance in Tversky and Kahneman (1986) in minimal frames, but SWUP satisfies independence and stochastic dominance in transparent frames. Our experiment is designed to test these predictions.

Table 1. Predictions of Different Models for Minimal and Transparent Frames

	Common Ratio Effect		Allais	Paradox	Dominance Violation		
Model	Minimal	Transparent	Minimal	Transparent	Minimal	Transparent	
EU	*	*	*	*	*	*	
CPT/RDU	√	✓	✓	✓	*	*	
SWUP	✓	*	✓	*	✓	*	

Notes: ✓ indicates behavior is predicted; * indicates behavior is not predicted.

5.1 Experimental Procedure and Design

We seated 137 subjects⁸ at visually isolated computer terminals in lab cubicles. Each subject chose one lottery from a lottery pair (no indifference permitted) for 100 distinct pairs presented sequentially one at a time on a computer screen. After completing their 100 choices, each subject rolled a pair of ten-sided dice, randomly selecting one of their 100 chosen lotteries to count for payment: The subject then played out that lottery by selecting a numbered raffle ticket from an opaque bag, receiving a cash outcome to keep along with a promised flat \$7.00 for timely arrival and participation. Figure 7 shows minimal and transparent frame versions of one pair, as shown on subjects' computer screens (just one appearing on any screen). If a subject chose the 'red' lottery, she would draw a ticket from an opaque bag containing 100 red raffle tickets.⁹ If the number on the ticket was between 1 and 25, she received \$30. If the number was between 26 and 100, she received \$0. Computerized instructions (screen prints appear in the supplementary materials <u>SM</u>) explained this generally, using specific examples and follow-up tests (which subjects had to correctly answer before proceeding).

Figure 7. A Choice Pair in a Minimal Frame (top) and a Transparent Frame (bottom)

5	Select One	\$	Tickets	\$	Tickets	
	© Red	\$30	25	\$0	75	
	© Blue	\$40	20	\$0	80	
Select One	\$	Tickets	\$	Tickets	\$	Tickets
Red	\$40	20	\$0	5	\$ 0	75
Blue	\$30	20	\$30	5	\$0	75

Of the 100 lottery pairs, 40 are *test* pairs (shown in our <u>SM</u>) for hypothesis tests within and between transparent and minimal framed pairs, while 60 are *extra* pairs mostly designed to aid efficient estimation of structural model parameters. We present lottery pairs in blocked order (randomizing order within each block of ten pairs) to space related test pairs (also utilizing the extra pairs for this spacing purpose). Of the 40 test pairs, 18 test for common ratio effects in two groups CR.A and CR.B of 9 pairs each; 16 test for Allais' Paradox (common consequence effects) in four groups AP.A, AP.B, AP.C, and AP.D of 4 pairs each; and 6 test for dominance violations in three groups DV.A, DV.B and DV.C of 2 pairs each.

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⁸ The planned sample was 144 subjects (6 sessions of 24 each), but seven scheduled subjects failed to appear. Subjects were undergraduates students at a U.S. university in April and May of 2016.

⁹ A separate bag of blue tickets resolved "blue" lotteries: Use of two separate bags induces statistically independent lotteries. We do this for a specific purpose. Under both Regret Theory (Loomes and Sugden 1987a, p. 274) and the salience theory of Bordalo et al. (2012, p. 1259), dominance violations should not happen when lotteries are independent. So with two independent bags, any violations of dominance we observe in minimal frames—predicted by SWUP—are inconsistent with both Regret Theory and the salience model of Bordalo et al.

5.2 Results: Data Patterns and Conventional Hypothesis Tests

Experimental results in the two common ratio groups of pairs appear in the two panels of Figure 8. The inset in each panel is a prospect presentation of the "root pair" {S,R} generating each group of common ratio pairs. Because the safe lottery S is degenerate in each root pair, minimal and transparent presentations of the root pairs are identical: Hence in each panel, only one choice proportion is graphed above the common ratio of 1 (the root pair). However, minimal and transparent presentations differ for any pairs {S',R'} with common ratio less than 1 (the other pairs in each group), so two choice proportions are graphed above those common ratios. The solid (double) line connects minimal (transparent) frame observed risky lottery choice proportions within each group. We also show a Bayesian 90% confidence interval (based on the Jeffreys Prior as recommended by Brown, Cai and DasGupta 2001) for each proportion, and these illustrate two findings: Frame invariance dramatically fails in these common ratio groups; and relative to minimal framing, transparent framing strongly promotes choice of safe lotteries.

Under Conlisk's (1989) constant error model and EU, proportions of non-EU choice patterns $S \cup R'$ and $R \cup S'$ (involving the root pair and any other pair in a group) should be equivalent: Asterisks above each confidence interval in Figure 8 indicate rejection of this hypothesis (* at 5%, and ** at 1%) by a likelihood ratio test based on Conlisk's constant error model. Such rejections are universal in minimal frames and, moreover, observed minimal frame choice proportions display a characteristic crossing of the 0.5 Rubicon as the common ratio gets low enough. Neither result obtains in transparent frames: Risky choice proportions rise significantly at the lowest common ratios only, and never exceed 0.5—in keeping with simple strength-of-preference explanations (and so not clear evidence of an independence violation).

The four panels of Figure 9 show our results in the four Allais Paradox (common consequence) groups of pairs appear in the four panels of Figure 9. Each group has a pair $1 \{S_1, R_1\}$ and a pair 2

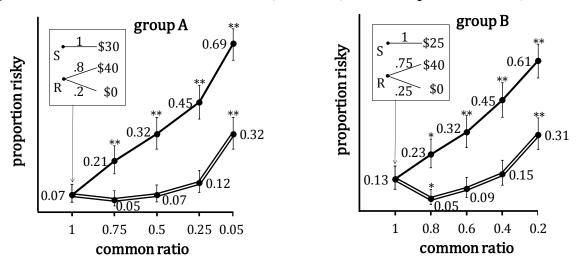


Figure 8: Common Ratio Effect in Minimal (Solid Line) and Transparent Frames (Double Line)



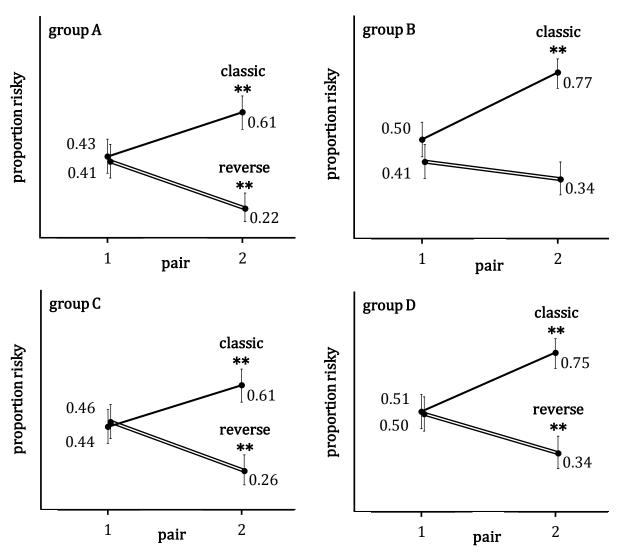
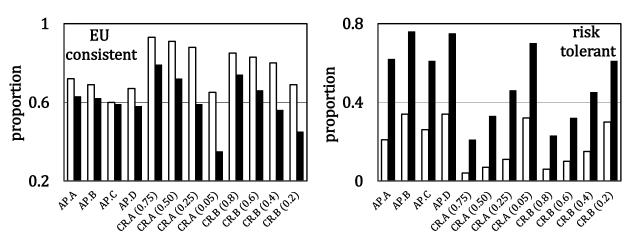


Figure 10. EU Consistency and Risk Tolerance in Minimal (black) and Transparent Frames (white)



{S₂, R₂} formed from pair 1 by changing a common consequence to zero. We presented these pairs in both minimal and transparent frames, so two choice proportions are graphed above each pair. As in Figure 8, the solid (double) line connects minimal (transparent) frame observed risky lottery choice proportions within each group, and we show the same type of Bayesian 90% confidence interval for each proportion. In pair 1 of each common consequence group, SWUP predicts no difference between these choices and the data bear this out. However, pair 2 choice proportions always differ strongly across the two types of frames: as in the common ratio groups, transparent framing strongly promotes the choice of safe lotteries relative to minimal framing.

In these groups, under Conlisk's (1989) constant error model and EU, proportions of non-EU choice patterns $S_1 \cup R_2$ and $R_1 \cup S_2$ should be equivalent. Asterisks above each confidence interval in Figure 9 indicate rejection of this hypothesis (* at 5%, and ** at 1%) by a likelihood ratio test. Figure 9 shows that this hypothesis is almost always rejected in both minimal and transparent frames; but there is a strong difference in the direction of the rejection across these frame types. As predicted by both CPT/RDU and SWUP, we observe the pattern $S_1 \cup R_2$ significantly more than the pattern $R_1 \cup S_2$ in minimal frames: This is the 'classic' Allais result (and so labeled in Figure 9). However, significant 'reverse' Allais results appear in three of the four groups in transparent frames ($S_1 \cup R_2$ occurs significantly *less* often than $R_1 \cup S_2$) and no theory we know of predicts this. SWUP predicts that no significant Allais pattern (classic or reverse) should appear in transparent frames—as observed in group B, but not in groups A, C, and D.

Choice patterns $S_1 \cup S_2$ and $R_1 \cup R_2$ (in each Allais Paradox group) and choice patterns $S \cup S'$ and $R \cup R'$ (in each common ratio group) are consistent with EU. The left panel of Figure 10 shows rates of EU consistency in transparent frames (the white bars) and minimal frames (the black bars). In all twelve pairs of pairs, EU consistency in transparent frames exceeds that in minimal frames, with nine of these twelve comparisons significant (at 10% or better) by likelihood ratio tests. The right panel of Figure 10 compares risk tolerance in minimal and transparent frames. ¹¹ Frame invariance (implied by EU and CPT/RDU) predicts that risk tolerance (R' or R_2 choice proportions) should not significantly differ between frames, while SWUP predicts greater risk tolerance in minimal frames. In all twelve comparisons (and significantly at 1% or better), risky choice proportions are greater in minimal than transparent frames: On average across these twelve pairs, risky choice proportions are 30.9 percentage points higher in minimal (than transparent) frames.

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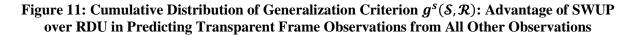
¹⁰ In Figure 5, one can see that the top two framings $[\![\mathbf{p}, \mathbf{q}]\!]$ of $\{p, q\}$ are identical *except* that the transparent frame isolates the common consequence column block at the right, while the monotone minimal frame places that column block at the center. Under SWUP such column block switching should have no effect on choice.

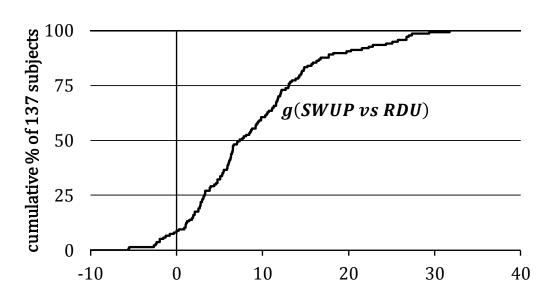
¹¹ Here we omit the common ratio "root pairs" $\{S,R\}$ since the minimal and transparent frame for these pairs are identical, and also omit Allais Paradox pairs $\{S_1, R_1\}$ since SWUP predicts no difference between the minimal and transparent framings of pair 1 in each common consequence group (see Figure 5 and fn. 9).

We expect that, as predicted by SWUP, dominance violations will be common in minimal frames (which is not predicted by EU or CPT/RDU) but very rare in transparent frames and our data bears this out. In transparent framings of the dominance violation pairs DV.A, DV.B and DV.C, dominance violations are just 3%, 2% and 3% of all choices, while in minimal framings dominance violations are 67%, 81% and 57% of all choices. This is a strong failure of frame invariance.

5.3 Results: Predicting Transparent Frame Behavior from All Other Behavior

The Generalization Criterion "g" discussed by Busemeyer and Wang (2000) provides a comparative prediction test of pairs of models. We let \mathcal{M} denote either of two models \mathcal{S} (for SWUP) and \mathcal{R} (for RDU/CPT). The generalization criterion answers the question: Which model $\mathcal{M} \in \{\mathcal{S}, \mathcal{R}\}$ best predicts observations in one subset of the data, given estimates obtained using a wholly different subset of the data? Here we let j=1,2,...,100 index pairs and s=1,2,...,137 index subjects; and we let the indicators c_j^s code observations. Let \mathcal{J} denote either of two mutually exclusive and exhaustive subsets of our 100 pairs j: \mathcal{T} (all pairs j transparently framed, 30 pairs in all) and its complement \mathcal{N} (all pairs j not transparently framed—mostly minimally framed—70 pairs in all). Let $c_{\mathcal{J}}^s$ denote the vectors of all observations in each subset $\mathcal{J} \in \{\mathcal{T}, \mathcal{N}\}$. Let $\mathcal{L}_{\mathcal{M}}^s(c_{\mathcal{J}}^s|\theta_{\mathcal{M}})$ be the log likelihood function of the observations $c_{\mathcal{J}}^s$, given model \mathcal{M} and its parameter vector $\theta_{\mathcal{M}}$, and let $\hat{\theta}_{\mathcal{M},\mathcal{N}}^s$ be the maximum likelihood





 $c_j^s = 1$ when subject s chooses the riskier lottery (or the stochastically dominant lottery in the 6 dominance violation pairs) in pair j; otherwise, $c_i^s = 0$.

¹³ As shown in our <u>SM</u>, 8 of our extra pairs are in a third type of frame (the "payoff-aligned frame") which will have special usefulness for identifying properties of salience functions in another study based on these data.

estimates of $\theta_{\mathcal{M}}$ for subject s, based only on the observation vector $c_{\mathcal{N}}^s$: Frequently one calls $\hat{\theta}_{\mathcal{M},\mathcal{N}}^s$ the "in-sample estimates" (a.k.a. "training data" or "calibration" estimates) for each model \mathcal{M} and each subject s. Our supplementary material \underline{SM} details the specification and estimation of the two models.

These in-sample estimates—wholly estimated without any transparent frame behavior—then become predictors of the log likelihood of the "out-of-sample" observations c_T^S in transparently framed pairs: This measure of prediction information is $\mathcal{L}_{\mathcal{M}}^S(c_T^S|\hat{\theta}_{\mathcal{M},\mathcal{N}}^S)$. One then compares these information measures for the pair of models: In our application, the generalization criterion for subject s is $g^s(\mathcal{S},\mathcal{R}) = 2[\mathcal{L}_{\mathcal{S}}^S(c_T^S|\hat{\theta}_{\mathcal{S},\mathcal{N}}^S) - \mathcal{L}_{\mathcal{R}}^S(c_T^S|\hat{\theta}_{\mathcal{R},\mathcal{N}}^S)]$. When $g^S(\mathcal{S},\mathcal{R})$ is positive, SWUP's out-of-sample predictions are better than those of RDU. Figure 11 shows the cumulative distribution of $g^S(\mathcal{S},\mathcal{R})$ across our 137 subjects and demonstrates that SWUP is the overwhelmingly superior model by the generalization criterion: $g^S(\mathcal{S},\mathcal{R})$ is positive for 125 of the 137 subjects.

6. Minimal and Transparent Frames for Choice over Time: Present Bias and Hidden Zero Effects

Consider the minimal frames in Figure 12. The stationarity axiom of DU theory implies that people should choose either **r** and **r'** or **t** and **t'**. However, in choices such as these, experiments show that people frequently choose **r** and **t'**, a finding termed *present bias* (Laibson, 1997). Present bias occurs in the minimal frames in Figure 12 since the comparison between receiving money today versus in one year is more salient than the comparison of \$75 versus \$100, but this monetary comparison is more salient than receiving payment in 10 or 11 years. However, in transparent frames, the focal comparisons in both choices are between \$75 and \$0 and between \$100 and \$0. Switching from minimal to transparent frames provides a formal explanation of the hidden zero effect (Magen et al., 2008; Radu et al., 2011; Read et al., 2017) in which behavior becomes more patient when the opportunity costs of income (such as receiving \$0 instead of \$100 in 1 year) are made salient. Transparent frames retain the second choice of **t'** over **r'** from minimal frames, but shift the first choice toward preferring **t** over **r** via the hidden zero effect. The prediction of more patient behavior in transparent frames is also consistent with the finding by Fisher and Rangel (2014) that shifting attention from focusing on time to focusing on money reduces impatience

Figure 12. Present Bias in Minimal and Transparent Frames

	In Minimal Frames			In '	Transpa	rent Fra	mes
	x_{1}, y_{1}	Years		$x_{1,}y_{1}$	Years	x_2,y_2	Years
r	75	0	r	75	0	0	1
t	100	1	t	0	0	100	1
	x_{1}, y_{1}	Years		x_{1},y_{1}	Years	x_2,y_2	Years
r'	75	10	r'	75	10	0	11
t'	100	11	t'	0	10	100	11

since transparent frames increase the salience of the money dimension relative to minimal frames. Transparent frames may thus serve to induce more patient behavior and more time consistent behavior.

7. Minimal and Transparent Frames for Choice under Ambiguity: Ellsberg's Paradox

The SWUP model in (9) predicts ambiguity aversion in minimal frames, but also predicts that transparent frames will reduce it. To illustrate this we refer to four "option pairs" Schneider, Leland, and Wilcox (2018) presented to 79 subjects. To understand these option pairs, first suppose that the top pair {A, B} in Figure 13 was selected for payout at the conclusion of a subject's session, ¹⁴ and the subject chose B from it. The subject then blindly draws one ticket from an opaque bag containing an unknown mixture of red and blue tickets: If she draws a red ticket, she plays a lottery with a 75% chance of winning \$25 and a 25% chance of winning nothing; but if she draws a blue ticket, she instead plays a lottery with a 25% chance of winning \$25 and a 75% chance of winning nothing. If she had chosen option A instead, she would play a lottery with a 50% chance of winning \$25 and a 50% chance of winning nothing regardless of the ticket color she draws. The pair {A', B'} is similar except that in option B' the "good" state is reversed. For these minimal frames, the experiment replicates Ellsberg's Paradox, finding that people do not assign well-defined subjective probabilities to states, but rather prefer alternatives with known probabilities over unknown probabilities—a preference pattern called *ambiguity aversion*.

Red Ticket Blue Ticket A \$25 0.50 **\$0** 0.50 \$25 0.50 **\$0** 0.50 В \$25 0.75 \$0 0.25 \$25 0.25 \$0 0.75 Red Ticket Blue Ticket $\mathbf{A'}$ \$25 0.50 0.50 \$25 0.50 **\$0** 0.50 **\$0** B'\$25 \$25 0.25 \$0 0.75 0.75 \$0 0.25

Figure 13. The Ellsberg Paradox in Minimal Frames

With a uniform prior over states¹⁵ and normalizing u(25) = 1 and u(0) = 0, SWUP implies that A is chosen over B in the minimal frame shown in the top panel of Figure 13 if inequality (12) holds:

(12)
$$0.5\phi(0.5,0.75)(-0.25) + 0.5\phi(0.5,0.25)(0.25) > 0.$$

By symmetry and DAS, $\phi(0.5,0.25) > \phi(0.5,0.75)$, so (12) holds for all salience and utility functions. Intuitively, the salient comparisons between A and B are between a 0.50 probability of winning and a 0.75 probability (in the 'red' state) and between probabilities of 0.50 and 0.25 (in the 'blue' state). Diminishing absolute sensitivity implies a focal thinker will be more sensitive to the latter comparison. Similarly, A' is chosen over B' in the bottom panel of Figure 13, yielding ambiguity aversion.

¹⁴ A subject chose from 60 pairs and, at the session's end, one of the 60 pairs was randomly chosen for payout.

¹⁵ In estimating a mean-dispersion model of ambiguity preference to explain their data, Schneider, Leland, and Wilcox (2018) estimate the subjective prior assigned to red and blue ticket states to be near uniform for the great majority of their 79 subjects.

Figure 14 shows Ellsberg-style choices in transparent frames. If the decision maker has a uniform prior over red and blue ticket states, then SWUP predicts ambiguity aversion in minimal frames, and ambiguity-neutrality (indifference between A and B) in transparent frames. Ambiguity neutrality in transparent frames follows from symmetry of μ , in which case (13) holds:

(13)
$$0.5\mu(0,25)(-6.25) + 0.5\mu(25,0)(6.25) = 0.$$

The transparent frame focuses attention on comparing \$0 and \$25 rather than probabilities. If the decision maker has a uniform prior over red and blue ticket states, then a focal thinker exhibits ambiguity aversion in minimal frames, and ambiguity-neutrality (indifference between A and B) in transparent frames. Schneider, Leland, and Wilcox (2018) find support for this directional difference in ambiguity attitudes between minimal and transparent frames (p = 0.0008 by a sign test against the null that the treatment effect has a zero location) and estimate that ambiguity premia in minimal frames are roughly double the ambiguity premia in transparent frames.

Red Ticket Blue Ticket 0.25 \$25 0.50 \$0 0.25 \$25 0.25 \$25 0.25 \$0 0.50 Α В \$25 0.25 \$25 \$0 0.25 \$25 0.50 0.25 \$0 0.25 \$0 0.50 Red Ticket Blue Ticket 0.25 0.50 0.25 \$25 Α′ \$25 \$25 0.25 \$0 \$0 0.50 \$0 0.25 \mathbf{B}' 0.25 0.25 \$25 0.25 \$0 \$25 \$25 0.50 \$0 0.50 0.25

Figure 14. The Ellsberg Paradox in Transparent Frames

8. Related Literature

In tests of stochastic dominance, presentation effects such as the Tversky and Kahneman (1986) example have also been observed in several studies by Birnbaum and his colleagues (e.g., Birnbaum and Navarette, 1998; Birnbaum, 1999; Birnbaum, 2004, Birnbaum et al., 2008). In tests of the independence axiom of EU, many scholars supported Allais' (1953) prediction of common consequence violations in minimal frames, but there is mixed evidence regarding the prediction that common consequence effects will not occur in transparent frames. Moskowitz (1974) examined common consequence choice pairs presented either in a minimal written form as in the introduction, as tree diagrams, or as transparent matrices like the one proposed by Savage and shown in the introduction. Contrary to our predictions, there were no differences in the proportion of consistent responses among subjects given written or matrix presentations, although both were higher than for tree diagrams. However, Keller (1985), Birnbaum (2008), Incekara-Hafalir and Stecher (2012), and Harman and Gonzalez (2015) all find that more transparent presentations reduce the occurrence of common consequence violations.

Allais (1953) first predicted a common ratio effect for choices presented in minimal frames. Harless (1992) observed that minimal presentations of common ratio choices violate independence while

transparent frames do not. He presented different groups of subjects with choices such as a choice between S:(\$8,000, 0.10; \$0, 0.90) and R:(\$20,000, 0.05; \$0, 0.95) and one between S':(\$8,000, 0.60; \$0, 0.40) and R':(\$20,000, 0.30; \$0, 0.70). One group received both choices in a transparent matrix format in the top of Figure 15. Another group was presented with choices in a minimal 'tickets' frame in the bottom of Figure 15. As predicted by SWUP, Harless found that deviations from EU were unsystematic in the transparent frame, but systematic (in Allais' predicted direction) in the minimal frame. Harless also examined how different juxtapositions of payoffs across choices would influence the incidence and direction of common ratio violations in matrix frames. Those results are also consistent with the predictions of SWUP. Keller (1985) also found that proportional matrix presentations reduce the occurrence of common ratio violations; yet both Loomes and Sugden (1987b) and Keller still report many common ratio violations (with state matrix and proportional matrix presentations, respectively). 17

Figure 15. Transparent Matrix and Minimal 'Tickets' Presentations used by Harless (1992)

		Tra	nsparent Tic	kets Mat	rix Presenta	ation					
	1-5	6-10	11-100		1-30	31-60	61-100				
S	\$8	\$8	\$0	S'	\$8	\$8	\$0				
R	\$20	\$0	\$0	R'	\$20	\$0	\$0				
	Minimal Tickets Presentation										
S	Ticket 1-10	drawn	Win \$8	S'	Ticket 1-60) drawn	Win \$8				
3	Ticket 11-1	00 drawn	Win \$0	3	Ticket 61-1	00 drawn	Win \$0				
R	Ticket 1-5 d	lrawn	Win \$20	R'	Ticket 1-30) drawn	Win \$20				
K	Ticket 6-10	0 drawn	Win \$0		Ticket 31-1	Win \$0					

Starmer and Sugden (1993) and Humphrey (1995) studied a presentation effect known as event splitting: Presenting an event split into many separate ones (for gains) makes that event more attractive. SWUP explains event splitting effects *when* splitting an event transforms a presentation from a minimal to a transparent frame, but *not* when the probabilities within each column vector of a frame are aligned in both presentations (so that salient payoff comparisons are the same in original and split frames).

Many domain specific choice models share features with SWUP and will accommodate some of the implications discussed above. Loomes (2010) perceived relative argument model (PRAM) of risky choice, for example, posits that agents make cross-lottery comparisons of probabilities and of payoffs that reflect diminishing absolute sensitivity and then chose based on which comparison most favors one alternative over the other. However, PRAM applies to choices involving at most three outcomes which

¹⁶ Loomes and Sudgen's (1982) Regret theory and Bordalo et alia's (2012) salience based model of risky choice also predict many of these effects but only for statistically independent lotteries. Leland (1998) shows the behavior obtains for statistically dependent lotteries as well.

¹⁷ Except for Loomes and Sugden (1987b), all the experiments in this paragraph examine hypothetical choices. Therefore we think our incentivized experiment adds substantially to this particular strand of the literature.

limits its applicability. Scholten and Read's (2010) tradeoff model of intertemporal choices, like SWUP, posits that decisions result not from alternative-based discounting but from trading off time differences against payoff differences across alternatives. However, it has not been extended to risky or ambiguous choice and leaves no room for discounting: SWUP incorporates both discounting and context-dependence. Birnbaum and Chavez's (1997) transfer-of-attention-exchange (TAX) is an early example of a model in which aspects of lotteries result in some attributes receiving more attention and disproportionate weight. TAX predicts the framing effects implied by SWUP for choices under risk but has not been extended to intertemporal choices nor those under uncertainty.

Loomes and Sudgen's (1982, 1987a) Regret theory and Bordalo et. alia's (2012) salience theory also posit comparative decision processes that overweight large payoff differences (either for their potential to induce regret or as a consequence of their perceived salience). While these models make predictions in the context of risky choice that are similar to SWUP, their predictions depend critically on whether the alternatives are assumed to be statistically dependent or independent, as it is the correlation between lotteries that determines which comparisons are made. Under SWUP, comparisons made are simply determined by how options are presented and what payoffs and probabilities (or time periods) are aligned in the presentation, and not at all on how probability-generating mechanisms are related across the two lotteries in that presentation. Because of this, the dominance violations we observed in our experiment (which induces statistically independent lotteries) are predicted by SWUP but ruled out by Regret theory and the salience theory of Bordalo et. alia.¹⁸

9. Summary

We now gather our results on framing effects for risk, time, and ambiguity in three propositions. Earlier in sections 4, 6 and 7 we illustrated the claims in Proposition 1 (see our SM for the calculations).

Proposition 1 (Minimal Frames and Behavioral Biases):

A focal thinker with linear utility u(x) = x and either the DAS (10) or DAS-IPS (11) salience function:

- (i) violates stochastic dominance in the minimal frame in Figure 4;
- (ii) exhibits the Allais paradox in the minimal frames in Figure 5;
- (iii) exhibits the common ratio effect in the minimal frames in Figure 6;
- (iv) exhibits present bias in the minimal frames in Figure 12 (for annual discount factor $\delta \in [0.41, 0.95]$); and
- (v) exhibits Ellsberg's paradox in the minimal frames in Figure 13.

For minimal frames, we prove more general results regarding two of the most robust and most well-known violations of EU theory: the Allais common ratio effect and the Ellsberg paradox. Although these paradoxes are two of the oldest violations of rational choice theory, there has been relatively little work

¹⁸ See Loomes and Sugden (1987a, p. 274) and Bordalo et alia (2012, p. 1259).

investigating the precise relationship between them. Proposition 2 is proved in our Appendix, for general versions of the Allais common ratio effect and Ellsberg's paradox also defined in the Appendix.

Proposition 2 (Allais, Ellsberg, and Salience Perception): For monotone minimal frames:

- (i) A focal thinker exhibits the general Allais common ratio effect if and only if ϕ satisfies increasing proportional sensitivity.
- (ii) Under a uniform prior, a focal thinker exhibits ambiguity aversion in Ellsberg's paradox if ϕ satisfies diminishing absolute sensitivity.

Proposition 2 is general and establishes that two basic properties of the perceptual system (greater sensitivity to larger absolute differences for a fixed ratio (IPS), and greater sensitivity to larger ratios for a fixed absolute difference (DAS)) directly imply the most robust violations of EU theory without any parametric assumptions regarding the form of the agent's salience functions or utility functions.

Without defining them precisely, Savage (1954) and Tversky and Kahneman (1986) argued that transparent presentations would reduce violations of rational choice theory. We formalized 'transparent presentation' of choice alternatives for risk, ambiguity, and time using a set of properties which imply unique transparent frames. With this done we can now state a theorem concerning transparent frames—converting a long-standing suggestion into a set of falsifiable statements. Consider four types of systematic violations of rational choice theory: violations of stochastic dominance, Allais paradox violations of EU theory, present biased violations of DU theory, and Ellsberg paradox violations of subjective expected utility (or SEU) theory. We show they should all vanish under our definition of transparent framing: This *Transparent Frame Theorem* is proved in our Appendix.

Proposition 3 (Transparent Frame Theorem): For transparent frames, a focal thinker will not exhibit the following violations of rational choice theory, even if the focal thinker exhibits them in minimal frames:

- (i) Violations of Stochastic Dominance;
- (ii) Allais Paradox violations of EU theory;
- (iii) Common Ratio violations of EU theory;
- (iv) Present Bias violations of DU theory; and
- (v) Ellsberg Paradox violations of SEU theory.

The transparent frame theorem is general in that it does not depend on the form of the decision maker's salience functions or on the form of the decision maker's utility function or discount factor or subjective beliefs, or on the particular parameter values used for the paradoxes, and it applies to choices across the domains of risk, time, and uncertainty.

For risk, transparent frames are similar to the state matrices employed by Savage (1954) (but implying neither correlation nor independence between payoffs) and to the 'canonical split form' of

Birnbaum and Schmidt's (2015) tree presentation of lotteries, but are distinct in that the canonical split form does not separate common consequences from distinct consequences. For decisions under risk, minimal frames are related to Birnbaum's (1999) tree presentation of lotteries in 'coalesced form:' In particular, a choice set in which all lotteries are in coalesced form generates a minimal frame. We are not aware of any previous attempt to formalize different presentation formats for income streams over time.

Our definitions of frames and our formal distinctions between minimal and transparent frames provide a unified foundation for analyzing choice presentations across three major domains of individual choice. Combine that foundation with a decision model that operates on frames (such as SWUP), and a formal logic of framing effects emerges. The same mathematical structure—and the same psychological intuition—explains a variety of the most robust and well-known behavioral biases across the domains of risk, uncertainty, and time, violating four of the most well-known axioms in rational choice theory (stochastic dominance, independence, stationarity, and the sure-thing principle). Focal thinkers will violate these axioms in minimal frames but satisfy them in transparent frames. Evidence from previous literature and from our own experiment suggests that biases are reduced, but not eliminated, when the presentation of choice alternatives is made transparent.

Appendix

Here, $\mathbf{v} \geq \mathbf{w}$ ($\mathbf{v} \approx \mathbf{w}$) means that for focal thinkers, option v 'looks strictly better than' ('looks indifferent to') option w when a frame $[\![\mathbf{v}, \mathbf{w}]\!]$ (of type specified in each proposition) presents the choice pair $\{v, w\}$.

A.1 Proof of Proposition 2 (Allais and Ellsberg in Minimal Frames)

We define the general form of the Allais common ratio effect below:

Definition 2 (General Allais Common Ratio Effect): Consider monotone minimal frames (i) and (ii) in Figure A.1. The *general Allais common ratio effect* holds for focal thinkers if for all $\mathbf{y} > \mathbf{x} > \mathbf{0}$, $1 \ge \mathbf{p_1} > \mathbf{q_1} > 0$ and $\alpha \in (0,1)$, $\mathbf{p} \approx \mathbf{q}$, implies $\mathbf{q}' \approx \mathbf{p}'$.

Figure A.1 Minimal Frame for the General Allais Common Ratio Effect

(i)	$(\mathbf{x}_1,\mathbf{y}_1)$	(p_1,q_1)	$(\mathbf{x}_2,\mathbf{y}_2)$	(p_{2},q_{2})	(ii)	$(\mathbf{x}_1,\mathbf{y}_1)$	(p_{1},q_{1})	(x_2,y_2)	(p_{2},q_{2})
р	X	p_1	0	$1 - p_1$	\mathbf{p}'	X	αp_1	0	$1 - \alpha p_1$
q	у	q_1	0	$1-q_1$	\mathbf{q}'	у	αq_1	0	$1-\alpha q_1$

Proof of Proposition 2 (i): Note that $\mathbf{p} \approx \mathbf{q}$ in Figure A.1 choice (i) if and only if

$$\mu(\mathbf{x}, \mathbf{y}) \left(u(\mathbf{y}) - u(\mathbf{x}) \right) \left[\frac{\mathbf{p}_1 + \mathbf{q}_1}{2} \right] = \phi(\mathbf{p}_1, \mathbf{q}_1) (\mathbf{p}_1 - \mathbf{q}_1) \left[\frac{u(\mathbf{y}) + u(\mathbf{x})}{2} \right].$$

Also note that $\mathbf{q}' > \mathbf{p}'$ if and only if $\mu(\mathbf{x}, \mathbf{y}) \left(u(\mathbf{y}) - u(\mathbf{x}) \right) \left[\frac{\mathbf{p_1} + \mathbf{q_1}}{2} \right] > \phi \left(\alpha \mathbf{p_1}, \alpha \mathbf{q_1} \right) \left(\mathbf{p_1} - \mathbf{q_1} \right) \left[\frac{u(\mathbf{y}) + u(\mathbf{x})}{2} \right]$.

By IPS, scaling αp_1 and αq_1 each by $\frac{1}{\alpha}$ leads to $\phi(\alpha p_1, \alpha q_1) < \phi(p_1, q_1)$ for all $\alpha \in (0,1)$. Letting $\mathbf{k} \equiv 1/\alpha$, the common ratio effect holds if and only if $(\alpha p_1, \alpha q_1) < \phi(\mathbf{k} \alpha p_1, \mathbf{k} \alpha q_1)$ for all $\mathbf{k} > 1$. \blacksquare Next, consider Ellsberg's (1961) two-color paradox. There are two urns. Urn 1 contains 50 red and 50 black balls. Urn 2 contains an unknown mixture of 100 red and black balls. A person is given two choices:

Choice 1: Choose between A and B

Choice 2: Choose between C and D

A. Win \$100 if red is drawn from Urn 1

C. Win \$100 if black is drawn from Urn 1

B. Win \$100 if red is drawn from Urn 2

D. Win \$100 if black is drawn from Urn 2

SEU requires choices of either A and D or B and C. However, Ellsberg found that most people choose A and C (options with objective probabilities) over B and D (options with ambiguous probabilities) thereby exhibiting ambiguity aversion. The minimal frame for these choices (for each state $s \in \{0, 1, ..., 100\}$, the actual number of red balls in Urn 2) is displayed in Figure A.2, where q(s) is the probability of drawing a red ball from Urn 2 in state s.

Figure A.2. Minimal Frame for the Two-Color Ellsberg Paradox

A(s)	\$100	0.5	\$0	0.5
A(s) B(s)	\$100	q(s)	\$0	1 - q(s)
C (s) D(s)	\$100	0.5	\$0	0.5
D(s)	\$100	1-q(s)	\$0	q(s)

Definition 3 (Ambiguity Aversion in Ellsberg's Paradox): For the frames in Figure A.2, a focal thinker exhibits *ambiguity aversion in Ellsberg's paradox* if A > B and C > D.

Proof of Proposition 2 (ii): This proof is for Ellsberg's two-color paradox. An analogous argument resolves Ellsberg's three-color paradox. Let s denote the number of red balls in Urn 2. Since the number of black balls is 100 - s, the state of the urn is fully characterized by s. For each state, the presentation for Choice 1 is given by Figure A.2, where $\mathbf{q}(s)$ is the probability of drawing a red ball from Urn 2 in state s. Normalize payoffs such that $u(\mathbf{100}) = 1$, and $u(\mathbf{0}) = 0$. For a focal thinker, A is chosen over B if and only if inequality (14) holds. Under a uniform prior, (14) becomes (15):

(14)
$$\sum_{s=1}^{m} \pi_s \left[\phi(\mathbf{0}.\mathbf{5}, \mathbf{q}(\mathbf{s})) (\mathbf{0}.\mathbf{5} - \mathbf{q}(\mathbf{s})) \right] > 0.$$

$$(15) \qquad \qquad \frac{1}{101} \left[\sum_{s=0}^{s=50} \varphi \left(\textbf{0.5}, \textbf{q(s)} \right) \left(\textbf{0.5} - \textbf{q(s)} \right) + \sum_{s=51}^{s=100} \varphi \left(\textbf{0.5}, \textbf{q(s)} \right) \left(\textbf{0.5} - \textbf{q(s)} \right) \right] > 0, \text{ which implies}$$

(16)
$$\sum_{s=0}^{s=50} \phi(\mathbf{0}.\mathbf{5}, \mathbf{q}(\mathbf{s})) (\mathbf{0}.\mathbf{5} - \mathbf{q}(\mathbf{s})) + \sum_{s=0}^{s=50} \phi(\mathbf{0}.\mathbf{5}, \mathbf{1} - \mathbf{q}(\mathbf{s})) (\mathbf{q}(\mathbf{s}) - \mathbf{0}.\mathbf{5}) > 0.$$

To see that (14) holds, note that for each $\mathbf{q}(s) \in [0,0.5)$ diminishing absolute sensitivity and symmetry of ϕ imply $\phi(0.5,1-\mathbf{q}(s)) < \phi(0.5,\mathbf{q}(s))$. In particular, by symmetry, $\phi(0.5,1-\mathbf{q}(s)) = \phi(0.5+0.5-\mathbf{q}(s),\mathbf{q}(s)+0.5-\mathbf{q}(s)) = \phi(0.5+\epsilon,\mathbf{q}(s)+\epsilon)$. Thus, by diminishing absolute

sensitivity, (16) holds, yielding a choice for the risky over the ambiguous urn. The argument follows analogously for the choice between C and D, resulting in ambiguity aversion. ■

A.2 Proof of Proposition 3 (Transparent Frame Theorem)

Proof of Proposition 3 (i): Lottery p is defined to stochastically dominate q if $P(x) \le Q(x)$ for all $x \in X$, with at least one strict inequality, where P(x) and Q(x) are the cumulative distribution functions corresponding to p and q, respectively. Whenever p stochastically dominates q, $\mathbf{x_i} \ge \mathbf{y_i}$ and $\mathbf{p_i} - \mathbf{q_i} = 0$ for all i in a transparent frame. Thus, the salience weights in (4) favor \mathbf{p} over \mathbf{q} in each binary comparison where the differences are not zero.

Proof of Proposition 3 (ii), (iii): We show that this result applies to both Allais paradoxes – the common consequence effect and the common ratio effect. The Allais common consequence choices in transparent frames are displayed in Figure A.3, where x > y > 0, q > p, and $q, p \in (0,1)$.

Figure A.3. The Allais Common Consequence Effect in Transparent Frames

p q	x y	p p	0 y	q – p q – p	y y	1 – q 1 – q
\mathbf{p}'	X	р	0	q – p	0	1 – q
\mathbf{q}'	у	р	y	q – p	0	1 – q

A focal thinker does not exhibit the common consequence effect if either $\mathbf{p} \geq \mathbf{q}$ and $\mathbf{p}' \geq \mathbf{q}'$ or $\mathbf{q} \geq \mathbf{p}$ and $\mathbf{q}' \geq \mathbf{p}'$. For focal thinkers, (4) implies that common consequences cancel, so $\mathbf{p} \geq \mathbf{q}$ iff $\mathbf{p}' \geq \mathbf{q}'$. \blacksquare The choices for the general form of the Allais common ratio effect are shown in transparent frames in Figure A.4, where $\mathbf{y} > \mathbf{x} > \mathbf{0}$, $\mathbf{1} \geq \mathbf{p} > \mathbf{q} > \mathbf{0}$, and $\mathbf{\alpha} \in (0,1)$.

Figure A.4. The Allais Common Ratio Effect in Transparent Frames

р	X	q	0	$\mathbf{p} - \mathbf{q}$	0	1 – p
q	y	q	y	$\mathbf{p} - \mathbf{q}$	0	1 – p
!			_	()		1
p'	X	αq	U	$\alpha(\mathbf{p}-\mathbf{q})$	U	$1-\alpha p$
\mathbf{q}'	\mathbf{y}	αq	y	$\alpha(\mathbf{p}-\mathbf{q})$	0	$1 - \alpha p$

A focal thinker does not exhibit the Allais common ratio paradox if either $p \ge q$ and $p' \ge q'$ or $q \ge p$ and $q' \ge p'$. For focal thinkers, (4) implies that the constant α factors out, so that $p \ge q$ iff $p' \ge q'$. \blacksquare **Proof of Proposition 3 (iv):** Choices that test present bias are shown in Figure A.5 in transparent frames.

Figure A.5. Present Bias in Transparent Frames

SS	X	0	0	Δ	SS'	X	t	0	$\mathbf{t} + \Delta$
LL	0	0	y	Δ	LL'	0	t	y	$\mathbf{t} + \Delta$

Present bias is absent if either $SS >_t LL$ and $SS' >_t LL'$ or $LL >_t SS$ and $LL' >_t SS'$. For a focal thinker, (6) gives both $SS >_t LL$ iff $\mu(\mathbf{x}, \mathbf{0})u(\mathbf{x}) > \mu(\mathbf{0}, \mathbf{y})u(\mathbf{y})\delta^{\Delta}$ and $SS' >_t LL'$ iff $\mu(\mathbf{x}, \mathbf{0})u(\mathbf{x})\delta^t > \mu(\mathbf{0}, \mathbf{y})u(\mathbf{y})\delta^{t+\Delta}$. Since δ^t can be factored out of the latter, $SS >_t LL$ iff $SS' >_t LL'$.

Proof of Proposition 3 (v): The transparent frames for Choices 1 and 2 in Ellsberg's two-color paradox are shown in Figure A.6, where state $s \in \{0,1,...,100\}$ indexes the number of red balls in Urn 2, and p(s) = |50 - s|/100. Note that when p(s) = 0.25, Figure A.6 resembles the frame in Figure 14. Ellsberg's paradox absent if either $\mathbf{A} \geq \mathbf{B}$ and $\mathbf{D} \geq \mathbf{C}$ or $\mathbf{B} \geq \mathbf{A}$ and $\mathbf{C} \geq \mathbf{D}$ or if there is indifferent in both choices. Without loss of generality, set u(100) = 1 and u(0) = 0. Denote the set of states favoring A by \overline{S} and the set of states favoring B by S. The SWUP evaluation for the choice between A and B is:

(17)
$$\sum_{s \in \overline{S}} \pi_s \, \mu(100,0) \big(p(s) \big) + \sum_{s \in \underline{S}} \pi_s \, \mu(0,100) \big(-p(s) \big)$$

States favoring A: $s \in \{0,1,...,50\}$ States favoring B: $s \in \{51,52,...,100\}$ A \$100 0.5 - p(s)0.5 \$100 0.5 p(s)\$0 0.5 - p(s)В \$100 0.5 - p(s)\$0 0.5 \$100 p(s)\$100 \$0 0.5 - p(s)States favoring C: $s \in \{51,52,...,100\}$ States favoring D: $s \in \{0,1,...,50\}$ C p(s)\$100 0.5 - p(s)0.5 p(s)\$100 0.5 \$0 0.5 - p(s)\$0 \$100 \$100 \$100

Figure A.6. The Ellsberg Paradox in Transparent Frames

where π_s is the subjective probability that the true state is s. All other differences within each column in the frame cancel. Under a uniform prior, the decision maker is indifferent between A and B (by symmetry of μ) in which case the evaluation in (17) equals zero. Moreover, even if the distribution is not uniform, (17) implies ambiguity neutrality since if (17) is positive, the decision maker would prefer A and D since the same set of states favor A and D. If (17) is negative, the decision maker prefers B and C. This argument extends analogously to Ellsberg's (1961) three-color paradox.

A.3 Minimal Frames

Definition 4 (Minimal Frame: Risk): For two non-degenerate lotteries $p, q \in \Delta(X)$, frame $[\![p, q]\!]$ (such as in the top panel of Figure 1) is *minimal* if $\forall i \neq j$, $\mathbf{x_i} \neq \mathbf{x_i}$ and $\mathbf{y_i} \neq \mathbf{y_i}$.

Definition 5 (Monotone Frame: Risk): For two non-degenerate lotteries $p,q \in \Delta(X)$, frame $[\![p,q]\!]$ (such as in the top panel of Figure 1) is *monotone* if $\mathbf{x_1} \geq \mathbf{x_2} \geq \cdots \geq \mathbf{x_n}$ and $\mathbf{y_1} \geq \mathbf{y_2} \geq \cdots \geq \mathbf{y_n}$. The following result is immediate.

Proposition 4 (Uniqueness of Monotone Minimal Frames: Risk): For any lotteries $p, q \in \Delta(X)$ with |supp(p)| = |supp(q)|, a monotone minimal frame $[\![\mathbf{p}, \mathbf{q}]\!]$ is unique up to the operation of rowswitching.

Proof: A frame that is minimal and monotonic is strictly monotonic. Hence, for two lotteries with the same support size, the ith best outcome of p is in the same column vector as the ith best outcome of q.

Minimal frames can also be defined for choices over time.

Definition 6 (Minimal Frame: Time): For income streams $r, t \in C$, frame $[\![\mathbf{r}, \mathbf{t}]\!]$ (such as in the bottom panel of Figure 1) is *minimal* if it has the smallest number of columns necessary to present $\{r, t\}$.

Definition 7 (Monotone Frame: Time)¹⁹: For income streams $r, t \in C$, frame [r, t] (such as in the bottom panel of Figure 1) is *monotone* if $\mathbf{r_1} < \mathbf{r_2} < \cdots < \mathbf{r_n}$ and $\mathbf{t_1} < \mathbf{t_2} < \cdots < \mathbf{t_n}$.

Let the support of an income stream, r, be the set of non-zero outcomes in r, and denote it by supp(r).

Proposition 5 (Uniqueness of Monotone Minimal Frames: Time): For any streams $r, t \in C$, with |supp(r)| = |supp(t)|, a monotone minimal frame $[\![\mathbf{r}, \mathbf{t}]\!]$ is unique up to the operation of row-switching. The proof of Proposition 5 is analogous to that of Proposition 4. Propositions 4 and 5 guarantee uniqueness of monotone minimal frames when both lotteries or both income streams have the same support size.

A.4 Transparent Frames

We next define transparent frames and show they are uniquely defined under general conditions (even if the lotteries or income streams have different support sizes). Given two lotteries, p and q, a pair (x, r), consisting of an outcome x with probability r, is a common consequence if $x \in supp(p) \cap supp(q)$ and $p(x) \ge r$, $q(x) \ge r$. All other pairs of outcomes and corresponding probabilities are distinct consequences. That is, a common consequence between two lotteries is one with the same outcome occurring with the same probability in both lotteries. We say (x,r) is a maximal common consequence if (x,r) is a common consequence for which p(x) = r or q(x) = r (or both). Our definition of transparent frames is constructive: it specifies the construction of a unique transparent frame for any pair of lotteries.

Definition 8 (Transparent Frame: Risk): A *transparent frame* for lotteries $[\![\mathbf{p}, \mathbf{q}]\!]$ is the special case of the frame in the top panel of Figure 1 that has the following properties:

- (1) Presentation of outcome-probability pairs (Common Consequence Separation): All (maximal) common consequences are separated from all distinct consequences such that all maximal common consequences are adjacent, and all distinct consequences are adjacent, as shown in Figure A.7.²⁰
- (2) Presentation of outcomes (monotonicity): Outcomes are ordered such that

$$x_n \ge \cdots \ge x_1$$
; $y_n \ge \cdots \ge y_1$; $z_k > \cdots > z_1$.

(3) Presentation of probabilities (alignment): Probabilities are presented so that $\forall i \in [k+1, k+n]$:

(i)
$$\mathbf{p_i} = \mathbf{q_i}$$
; and (ii) Given $\mathbf{x_i} = x$ and $\mathbf{y_i} = y$, $\mathbf{p_i} = min(p(x) - \sum_{j < i: \mathbf{x_j} = x} \mathbf{p_j}, q(y) - \sum_{j < i: \mathbf{y_j} = y} \mathbf{q_j})$.

(4) Relevance: $p_i > 0$ for all $i \in [1, k + n]$.

¹⁹ For our results, it does not matter whether monotone frames of income streams are monotonic in outcomes or monotonic in time periods, nor does it matter whether monotonic frames are increasing or decreasing.

In Figure A.7, there are k common consequences (with corresponding outcomes $z_1,...,z_k$) where $k \ge 0$. The remaining pairs of payoff column vectors and corresponding probability column vectors are distinct consequences.

Figure A.7. Transparent Frame of Lotteries p and q

						p_{k+1}				
q	y _n	p_{k+n}	 y_{k+i}	p_{k+i}	 y ₁	p_{k+1}	$\mathbf{z}_{\mathbf{k}}$	$\mathbf{p_k}$	 \mathbf{z}_1	$\mathbf{p_1}$

In (3-ii), p(x) is the overall probability of outcome x in lottery p, and $\mathbf{p_j}$ is the probability of outcome $\mathbf{x_j}$ in the jth payoff column vector in the frame. Note that $\sum_{j < i: \mathbf{x_j} = x} \mathbf{p_j}$ is the cumulative probability of outcome x that is summed over the preceding columns in the frame. The expression for $\mathbf{p_i}$ ensures that either the *entire* remaining probability mass for outcome x or for outcome y will be completely used in column \mathbf{i} in the frame. The algorithm for computing these probabilities (property (3-ii)) thus ensures *compactness* (the frame has the fewest cells subject to satisfying properties (1), (2), (3-i), and (4)) and it ensures *completeness* – the probabilities in each row vector of the frame sum to 1. The algorithm in (3-ii) also ensures *uniqueness* (the frame is uniquely defined even for lotteries with different support sizes), and it is constructive, by specifying how to generate the frame given any pair of lotteries.

Proposition 6 (Uniqueness of Transparent Frames: Risk): For any lotteries $p, q \in \Delta(X)$ with finite support, there is a unique transparent frame $[\![p,q]\!]$ up to the operation of row-switching.

Proof of Proposition 6 (Uniqueness of Transparent Frames: Risk): By monotonicity, the display of the k maximal common consequences is strictly monotonic, and thereby unique. For distinct consequences, monotonicity, in conjunction with the algorithm for $\mathbf{p_i}$, (alignment properties (3-i) and (3-ii)) uniquely determines each subsequent cell in the frame and so the frame is unique (There is only one possible specification implied by $\mathbf{p_i}$ for each probability and corresponding payoff column vector for each cell $i \in (k+1,k+n)$). \blacksquare^{21}

Definition 9 (Transparent Frame: Time): A *transparent frame* $[\![\mathbf{r}, \mathbf{t}]\!]$ for income streams $r, t \in \mathcal{C}$ is the special case of the frame in the bottom panel of Figure 1 that satisfies the following properties.

- (1) Common Consequence Separation: Common consequences are separated from distinct consequences such that common consequences are adjacent and distinct consequences are adjacent as in Figure A.8²².
- (2) Alignment: $t_i = r_i$ for all i.
- (3) Monotonicity: Time periods are ordered monotonically such that $t_1 < \cdots < t_k$; $t_{k+1} < \cdots < t_n$.
- (4) Completeness and Relevance: The first and last period in $[\![\mathbf{r}, \mathbf{t}]\!]$ are the same as in r and t and $\mathbf{n} = T$, where \mathbf{n} is the number of periods in the frame and T is the time horizon of the income streams.

²¹ Note that in the cases of transparent risk frames, both alignment properties (3-i) and (3-ii) are used to confer uniqueness. For choices over time, we only require an alignment property similar to (3-i).

²² In Figure A.8, there are m common consequences (with corresponding outcomes $z_1,...,z_m$) where $m \ge 0$. The remaining pairs of payoff column vectors and corresponding time column vectors are distinct consequences.

Figure A.8. Transparent Frame of Income streams r and t

	- 1	-		_	-	_	-	1		
X ₁	$\mathbf{r_1}$		$\mathbf{x_i}$	$\mathbf{r_i}$	 $\mathbf{x_k}$	r_k	$\mathbf{z_1}$	r_{k+1}	 Zm	rn
V ₁	t ₁		Vi	t_i	 Vk	tk	Z ₁	t_{k+1}	 Zm	tn
										$egin{array}{cccccccccccccccccccccccccccccccccccc$

Proposition 7 (Uniqueness of Transparent Frames: Time): For any income streams $r, t \in C$ over a horizon of T periods there is a unique transparent frame $[\![\mathbf{r}, \mathbf{t}]\!]$ up to the operation of row-switching.

Proof of Proposition 7 (Uniqueness of Transparent Frames: Time): By completeness and relevance, all (and only) time periods that index the income streams are displayed in the frame. By alignment, the i^{th} period in stream r in the frame is in the same column vector as the i^{th} period in stream t. By monotonicity, the ordering of the periods in the frame is uniquely determined and the frame is unique.

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