# A result on large induced subgraphs with prescribed residues in bipartite graphs 

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#### Abstract

It was proved by Scott that for every $k \geqslant 2$, there exists a constant $c(k)>0$ such that for every bipartite $n$-vertex graph $G$ without isolated vertices, there exists an induced subgraph $H$ of order at least $c(k) n$ such that $\operatorname{deg}_{H}(v) \equiv 1(\bmod k)$ for each $v \in H$. Scott conjectured that $c(k)=\Omega(1 / k)$, which would be tight up to the multiplicative constant. We confirm this conjecture.


Mathematics Subject Classifications: 05C07, 05C35

## 1 Introduction

Given a graph $G$ and integers $q>r \geqslant 0$, we define $f(G, r, q)$ to be the maximum order of an induced subgraph $H$ of $G$ where $\operatorname{deg}_{H}(v) \equiv r(\bmod q)$ for all $v \in H$ (or if no such $H$ exists, we set $f(G, r, q)=0)$.

There are many questions and conjectures concerning the behavior of $f(G, r, q)$ for various $G, r, q$. An old unpublished result of Gallai in this area is that ${ }^{1} f(G, 0,2) \geqslant n / 2$ for every $n$-vertex graph (see [7, Excercise 5.17] for a proof). Further questions about the behavior of $f$ received attention around 20-30 years ago (see e.g., [2, 3, 9, 10]). And more recently, this topic has had a renaissance (see e.g., $[1,5,6,8]$ ).

This note will focus on an old result of Scott. For positive integer $k$, we define $c(k)$ to be $\inf _{G}\{f(G, 1, k) /|G|\}$ where $G$ ranges over all bipartite graphs with $\delta(G) \geqslant 1$. The following was proved by Scott:

[^0]Theorem 1 ([10, Lemma 8]). Let $k \geqslant 2$. Then

$$
1 /\left(2^{k}+k+1\right) \leqslant c(k) \leqslant 1 / k .
$$

Scott observed that a slightly more careful argument could further show that $c(k)=$ $\Omega\left(\frac{1}{k^{2} \log k}\right)$.

In this note we give an improved lower bound to $c(k)$ which is optimal up to the (implied) multiplicative constant.

Theorem 2. Let $k \geqslant 2$. Then $c(k)=\Omega(1 / k)$.
This is done by taking the improved argument suggested by Scott, and then applying a dyadic pigeonhole argument which was previously overlooked.

## 2 Proof of Theorem 2

We will need the following result on the mixing time of random walks modulo $k$.
Lemma 3. Let $X_{i}$ be i.i.d. random variables that sample $\{0,1\}$ uniformly at random. If $N \geqslant k^{3}$, then ${ }^{2} \mathbb{P}\left(\sum_{i=1}^{N} X_{i} \equiv 1(\bmod k)\right) \geqslant\left(1-o_{k}(1)\right) / k$.
Lemma 3 is a mild corollary of several known results, and we note $k^{3}$ could be replaced with $k^{2} \log k$ (or any function which is $\omega\left(k^{2}\right)$ ).

For convenience, a fully elementary proof of Lemma 3 is provided in Appendix A.
In [10], when Scott outlined how to prove $c(k) \geqslant \Omega\left(\frac{1}{k^{2} \log k}\right)$, he noted that Lemma 3 (the key to the improvement) can be derived by slightly modifying the argument in [4, Theorem 2 of Chapter 3]. These appropriate modifications now appear in [5]. Namely, the interested reader can confirm that Lemma 3 follows from the proof ${ }^{3}$ of [5, Lemma 2.3]. Both of these proofs rely on discrete Fourier Analysis.

We now proceed to the main proof.
Proof of Theorem 2. Let $G$ be an $n$-vertex bipartite graph with $\delta(G) \geqslant 1$, and let $V_{1}, V_{2}$ bipartition $G$ with $\left|V_{1}\right| \geqslant\left|V_{2}\right|$. We shall write $c_{1}, c_{2}$ to denote small positive quantities which will be determined later (it would suffice to take $c_{1}=1 / 4, c_{2}=1 / 2$, but for clarity and a slightly better constant we will only consider their values at the end of the proof and shall have them depend slightly on $k$ ). Our proof splits into three cases.

We take $W_{1} \subset V_{2}$ to be a minimal set satisfying $\left|N(v) \cap W_{1}\right|>0$ for all $v \in V_{1}$ (i.e., $W_{1}$ is a minimal dominating set of $V_{1}$ ). By minimality of $W_{1}$, for each $w \in W_{1}$ there must exist $v_{w} \in V_{1}$ where $N\left(v_{w}\right) \cap W_{1}=\{w\}$. Let $S_{1}=\left\{v_{w}: w \in W_{1}\right\}$. We conclude that $W_{1} \cup S_{1}$ induces a matching in $G$, proving that $f(G, 1, k) \geqslant 2\left|W_{1}\right|$.

Hence, we will be done if $\left|W_{1}\right| \geqslant c_{1}\left|V_{1}\right| / k$ (this is "Case 1"). So we continue assuming $\left|W_{1}\right|<c_{1}\left|V_{1}\right| / k$.

[^1]For $2 \leqslant i \leqslant k-1$, we inductively create sets $W_{i}, S_{i}$. We take $W_{i} \subset W_{i-1}$ to be a minimal dominating set of $V_{1} \backslash\left(\bigcup_{j=1}^{i-1} S_{j}\right)$. And like in the above, we take $S_{i} \subset V_{1} \backslash\left(\bigcup_{j=1}^{i-1} S_{j}\right)$ so that $W_{i} \cup S_{i}$ induces a matching in $G$.

Let $T=V_{1} \backslash\left(\bigcup_{i=1}^{k-1} S_{i}\right)$. We have

$$
\begin{aligned}
|T| & =\left|V_{1}\right|-\sum_{i=1}^{k-1}\left|S_{i}\right| \\
& =\left|V_{1}\right|-\sum_{i=1}^{k-1}\left|W_{i}\right| \\
& \geqslant\left|V_{1}\right|-(k-1)\left|W_{1}\right| \\
& \geqslant\left(1-c_{1}\right)\left|V_{1}\right| .
\end{aligned}
$$

Next, let $T^{*}=\left\{v \in T:\left|N(v) \cap W_{k-1}\right| \geqslant k^{3}\right\}$. Supposing that $\left|T^{*}\right| \geqslant c_{2}\left|V_{1}\right|$ (this is "Case 2"), we will deduce that $f(G, 1, k) \geqslant\left(c_{2}-o_{k}(1)\right)\left|V_{1}\right| / k$.

Indeed, let $U \subset W_{k-1}$ be a random subset where each element is included (independently) with probability $1 / 2$. We set $T_{U}=\{v \in T:|N(v) \cap U| \equiv 1(\bmod k)\}$. By Lemma 3, we have that $\mathbb{P}\left(v \in T_{U}\right) \geqslant\left(1-o_{k}(1)\right) / k$ for each $v \in T^{*}$. Thus by linearity of expectation we may fix some $U \subset W_{k-1}$ where $\left|T_{U}\right| \geqslant\left|T^{*}\right|\left(1-o_{k}(1)\right) / k \geqslant\left(c_{2}-o_{k}(1)\right)\left|V_{1}\right| / k$. Next choosing $S \subset \bigcup_{i=1}^{k-1} S_{i}$ so that $\left|N(u) \cap\left(T_{U} \cup S\right)\right| \equiv 1(\bmod k)$ for each $u \in U$, we have that $S \cup U \cup T_{U}$ induces a subgraph in $G$ demonstrating that $f(G, 1, k) \geqslant\left|S \cup U \cup T_{U}\right| \geqslant$ $\left|T_{U}\right| \geqslant\left(c_{2}-o_{k}(1)\right)\left|V_{1}\right| / k$.

Otherwise, we must have that $T \backslash T^{*}$, the set of $v \in T$ where $\left|N(v) \cap W_{k-1}\right|<k^{3}$, has $>\left(1-c_{1}-c_{2}\right)\left|V_{1}\right|$ elements (this is "Case 3"). By dyadic pigeonhole, there exists some $0 \leqslant p \leqslant \log \left(k^{3}\right)=O(\log k)$ so that

$$
\begin{aligned}
\left|\left\{v \in T: 2^{p} \leqslant\left|N(v) \cap W_{k-1}\right|<2^{p+1}\right\}\right| & \geqslant\left|T \backslash T^{*}\right| / O(\log k) \\
& \geqslant\left(1-c_{1}-c_{2}\right)\left|V_{1}\right| / O(\log k)
\end{aligned}
$$

Take $T^{\prime}=\left\{v \in T: 2^{p} \leqslant\left|N(v) \cap W_{k-1}\right|<2^{p+1}\right\}$ to be this large set.
We let $U \subset W_{k-1}$ be a random subset so that each element is included (independently) with probability $1 / 2^{p}$. Defining $T_{U}$ as before, some casework ${ }^{4}$ shows $\mathbb{P}\left(v \in T_{U}\right) \geqslant e^{-2}$ for each $v \in T^{\prime}$. Hence, by linearity of expectation, we may fix $U$ so that $\left|T_{U}\right| \geqslant e^{-2}\left|T^{\prime}\right|$. As above we may find $S \subset \bigcup_{i=1}^{k-1} S_{i}$ so that $S \cup U \cup T_{U}$ demonstrates that $f(G, 1, k) \geqslant$ $\left|S \cup U \cup T_{U}\right| \geqslant e^{-2}\left(1-c_{1}-c_{2}\right)\left|V_{1}\right| / O(\log k)$.

Now fix any sufficiently small $\epsilon>0$. Letting $c_{1}=1 / 3-\epsilon / 2, c_{2}=2 / 3-\epsilon$, we get that each of the first two cases imply that $f(G, 1, k) \geqslant\left(2 / 3-\epsilon-o_{k}(1)\right)\left|V_{1}\right| / k \geqslant(1 / 3-\epsilon-$ $\left.o_{k}(1)\right) n / k$ (since $\left.\left|V_{1}\right| \geqslant\left|V_{2}\right|\right)$. Meanwhile with $\epsilon$ fixed, the third case implies $f(G, 1, k)=$ $\Omega_{\epsilon}(n / \log k)$. Taking $\epsilon \downarrow 0$ as $k \rightarrow \infty$ we have that $f(G, 1, k) \geqslant\left(1 / 3-o_{k}(1)\right) n / k$.

[^2]As a closing remark, we note it is still open whether $c(k)=1 / k$ for all $k$ (as noted in [10], considering $K_{k, k}$ demonstrates that $\left.c(k) \leqslant 1 / k\right)$. Even for $k=2$, the best known bounds are $1 / 4 \leqslant c(2) \leqslant 1 / 2$, with the lower bound coming from [9, Theorem 2].

## A An elementary proof of our lemma

Throughout, we let $\operatorname{Bin}(N, 1 / 2)$ denote the sum of $N$ i.i.d. random variables that uniformly sample $\{0,1\}$. For integers $N, i, k, r$, we let

$$
\begin{aligned}
& p_{N, i}:=\mathbb{P}(\operatorname{Bin}(N, 1 / 2)=i)=\binom{N}{i} 2^{-N}, \\
& P_{N, k, r}:=\mathbb{P}(\operatorname{Bin}(N, 1 / 2) \equiv r \quad(\bmod k)) .
\end{aligned}
$$

Also, for integer $N$ we let $m_{N}:=\max _{i}\left\{p_{N, i}\right\}=\binom{N}{\lfloor N / 2\rfloor} 2^{-N}$. A standard application of Stirling's formula gives $m_{N} \leqslant 1 / \sqrt{N}$ for sufficiently large $N$ (in fact $m_{N} \sqrt{N} \rightarrow \sqrt{2 / \pi}$ ).

We will first show the following.
Proposition 4. For all $N, k, r, r^{\prime},\left|P_{N, k, r}-P_{N, k, r^{\prime}}\right| \leqslant 2 m_{N}$.
Proof. First, we note that $|i-N / 2|<\left|i^{\prime}-N / 2\right|$ implies $p_{N, i}>p_{N, i^{\prime}}$ (since $\binom{N}{x}$ is a concave function of $x$ on the domain $[0, N]$, which is symmetric and achieves it maximum at $x=N / 2$ ). Hence, writing $I(E)$ to denote the indicator of an event $E$, we get

$$
\begin{aligned}
\left|P_{N, k, r}-P_{N, k, r^{\prime}}\right| & \leqslant\left|\sum_{i=1}^{N / 2} p_{N, i}\left(I(i \equiv r \quad(\bmod k))-I\left(i \equiv r^{\prime} \quad(\bmod k)\right)\right)\right| \\
& +\left|\sum_{i=1}^{N / 2} p_{N, N-i}\left(I(N-i \equiv r \quad(\bmod k))-I\left(N-i \equiv r^{\prime} \quad(\bmod k)\right)\right)\right| \\
& \leqslant 2 m_{N}
\end{aligned}
$$

(here the last step follows from the fact that the non-zero summands of each sum alternate in sign, are increasing in absolute value, and are bounded in absolute value by $m_{N}$ ).

We can now establish the desired lemma.
Proof of Lemma 3. For integers $N, k$, let $M_{N, k}=\min _{r}\left\{P_{N, k, r}\right\}$.
Obviously, we always have $P_{N, k, 1} \geqslant M_{N, k}$. So it suffices to show that for $N \geqslant k^{3}$, that $M_{N, k} \geqslant\left(1-o_{k}(1)\right) / k$.

By Proposition 4, we see that

$$
1=\sum_{r=1}^{k} P_{N, k, r} \leqslant k\left(M_{N, k}+2 m_{N}\right),
$$

hence

$$
M_{N, k} \leqslant\left(1-2 k m_{N}\right) / k .
$$

Assuming $N \geqslant k^{3}$ with $k$ sufficiently large, we have

$$
M_{N, k} \leqslant(1-2 / \sqrt{k}) / k=\left(1-o_{k}(1)\right) / k
$$

as desired.

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[^0]:    ${ }^{1}$ Actually what Gallai proved was slightly stronger. He showed that for each graph $G$, we can partition $V(G)$ into two parts $A, B$ so that $\operatorname{deg}_{G[A]}(v) \equiv 0(\bmod 2)\left(\right.$ respectively $\left.\operatorname{deg}_{G[B]}(v) \equiv 0(\bmod 2)\right)$ for each $v \in A$ (respectively $v \in B$ ).

[^1]:    ${ }^{2}$ Throughout this paper, we write $o_{k}(1)$ to denote quantities that tend to zero as $k \rightarrow \infty$.
    ${ }^{3}$ In [5], the statement of their lemma hides some constants which are necessary to verify our statement of Lemma 3.

[^2]:    ${ }^{4}$ If $p=0$, then $U=W_{k-1}$ and this probability is one. Otherwise this probability is at least $\binom{\left|N(v) \cap W_{k-1}\right|}{1}\left(1-2^{-p}\right)^{\left|N(v) \cap W_{k-1}\right|} 2^{-p} \geqslant\left(1-2^{-p}\right)^{2^{p+1}-1} \geqslant e^{-2}$.

